

NONPARAMETRICS

-APPLIED MULTIVARIATE ANALYSIS & STATISTICAL LEARNING-

Lecturer: Darren Homrighausen, PhD

Preamble:

- Nonparametric regression provides a flexible fit, but can only be used in low dimensions
- Additive models are a relaxation of nonparametric models that can be fit in much higher dimensions

FROM LINEAR TO NONLINEAR MODELS

GOAL: Develop a prediction function $\hat{f} : \mathbb{R}^p \rightarrow \mathbb{R}$ for predicting Y given an X

Commonly, $f(X) = X^\top \beta$

(e.g. Constrained linear regression)

This greatly simplifies algorithms, while not sacrificing too much flexibility

However, sometimes directly modeling the nonlinearity is more natural

PREDICTION VIA LOCAL AVERAGING

The fundamental quantities of interest we have been modeling are the **Bayes' rules**

$$\mathbb{E}[Y|X] \quad \text{or} \quad \arg \max_g \mathbb{P}(Y = g|X)$$

We know how to estimate expectations:

→ If Y_1, Y_2, \dots, Y_n all have expectation μ , then

$$\hat{\mu} = \frac{1}{n} \sum_{i=1}^n Y_i$$

is an intuitive estimator of μ
(and a reasonable prediction of a new Y)

PREDICTION VIA LOCAL AVERAGING

Similarly, we can estimate $\mathbb{E}[Y|X_*]$ with \mathcal{D} :

$$\hat{f}(X_*) = \frac{1}{n_*} \sum_{i=1}^n Y_i \mathbf{1}(X_i = X_*)$$

where $n_* = \sum_{i=1}^n \mathbf{1}(X_i = X_*)$ is the number of training observations where $X_i = X_*$

IN WORDS: we are taking an average of all the observations Y_i such that $X_i = X_*$

(Of course, this is all conditional expectation really is)

There is a big problem: There generally aren't any X_i at X_* !

PREDICTION VIA LOCAL AVERAGING

Suppose we relax the constraint $X_i = X_*$ a bit and include points that are **close enough** instead

Define a generic function “neighborhood” which takes in a value X_* and reports:

$$\text{neighborhood}(X_*) = \text{all } X_i \text{ in the “neighborhood” of } X_*$$

To estimate $\mathbb{E}Y|X_*$, average the **nearby** training observations

$$\hat{f}(X_*) = \frac{1}{n_*} \sum_{i=1}^n Y_i \mathbf{1}(X_i \in \text{neighborhood}(X_*))$$

where $n_* = |\text{neighborhood}(X_*)|$ is the number of training observations where X_i is in the “neighborhood” of X_*

PREDICTION VIA LOCAL AVERAGING

We need to define $\text{neighborhood}(X_*)$

(If it hasn't become clear, yet, this is a key aspect characterizing different methods)

EXAMPLE: Let's define “neighborhood” in terms of having Euclidean norm less than or equal to some threshold t

$$\mathbf{1}(X_i \in \text{neighborhood}(X_*)) = \mathbf{1}(\|X_i - X_*\|_2 \leq t)$$

Here...

- ... the Euclidean norm quantifies **distance**
- ... t quantifies **close**

(In fact, it is a tuning parameter)

PREDICTION VIA LOCAL AVERAGING: $X \in \mathbb{R}$

$$\hat{f}(X_*) = \frac{1}{n_*} \sum_{i=1}^n Y_i \mathbf{1}(\|X_i - X_*\|_2 \leq t), \quad \text{and} \quad X_* = 0.25$$

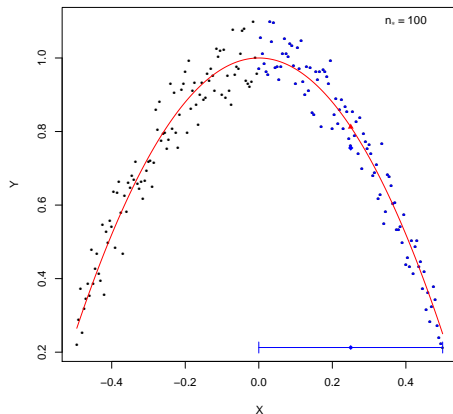


FIGURE: $t = 0.25$

PREDICTION VIA LOCAL AVERAGING: $X \in \mathbb{R}$

$$\hat{f}(X_*) = \frac{1}{n_*} \sum_{i=1}^n Y_i \mathbf{1}(\|X_i - X_*\|_2 \leq t), \quad \text{and} \quad X_* = 0.25$$

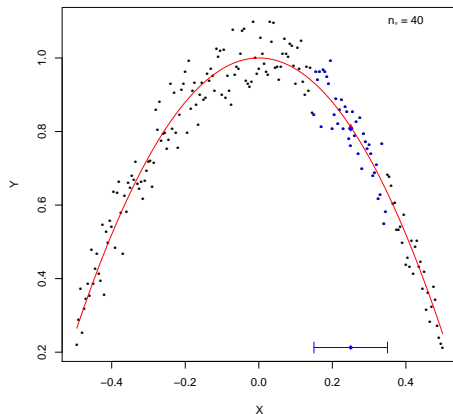


FIGURE: $t = 0.1$

PREDICTION VIA LOCAL AVERAGING: $X \in \mathbb{R}$

$$\hat{f}(X_*) = \frac{1}{n_*} \sum_{i=1}^n Y_i \mathbf{1}(\|X_i - X_*\|_2 \leq t), \quad \text{and} \quad X_* = 0.25$$

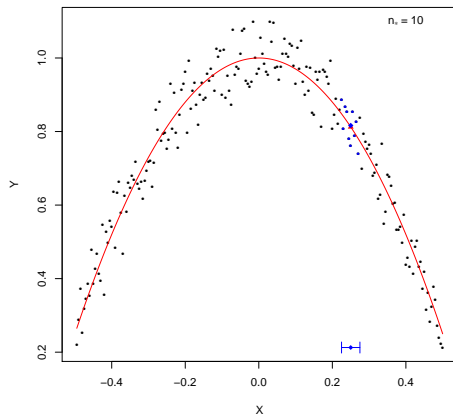


FIGURE: $t = 0.025$

PREDICTION VIA LOCAL AVERAGING: $X \in \mathbb{R}$

$$\hat{f}(X_*) = \frac{1}{n_*} \sum_{i=1}^n Y_i \mathbf{1}(\|X_i - X_*\|_2 \leq t), \quad \text{and} \quad X_* = 0.25$$

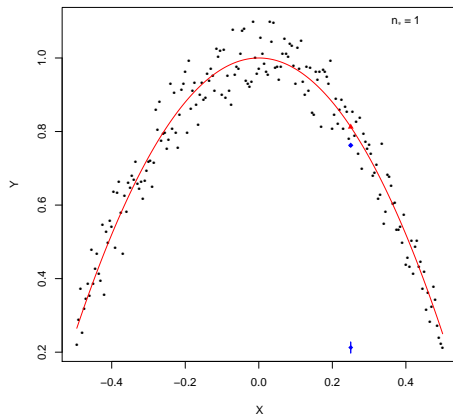


FIGURE: $t = 0.0001$

PREDICTION VIA LOCAL AVERAGING: $X \in \mathbb{R}$

$$\hat{f}(X_*) = \frac{1}{n_*} \sum_{i=1}^n Y_i \mathbf{1}(\|X_i - X_*\|_2 \leq t)$$

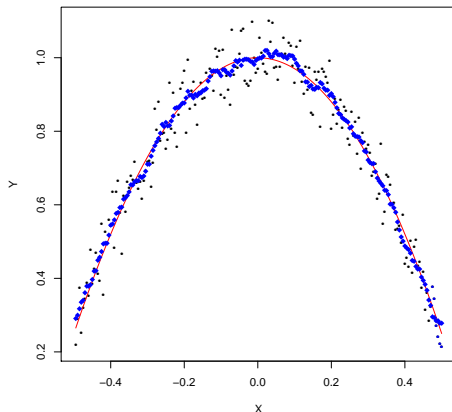


FIGURE: Over a grid of X_* 's, with $t = 0.025$

PREDICTION VIA LOCAL AVERAGING

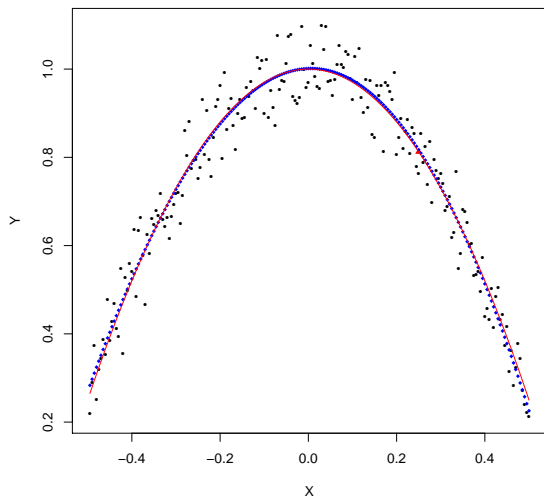


FIGURE: Using Regression Splines (ISL 7.4)

Nonparametric regression: Splines

POLYNOMIALS: REGRESSION

EXAMPLE: Fitting polynomials

Suppose we want to consider the following 4-dimensional feature transformation of a original 2-dimensional feature vector $X = (x_1, x_2)^\top \in \mathbb{R}^2$:

$$\Phi(X) = (1, \phi_1(X), \phi_2(X), \phi_3(X)) = (1, x_1, x_2, x_2^2)$$

For regression, we can model

$$\begin{aligned} f_*(X) &= \beta^\top \Phi(X) \\ &= \beta_0 + \beta_1 \text{balance} + \beta_2 \text{income} + \beta_3 \text{income}^2 \end{aligned}$$

See ISL 7.1 for details on fitting polynomials

POLYNOMIALS: CLASSIFICATION

EXAMPLE: Fitting polynomials

Suppose we want to consider the following 4-dimensional feature transformation of a original 2-dimensional feature vector $X = (x_1, x_2)^\top \in \mathbb{R}^2$:

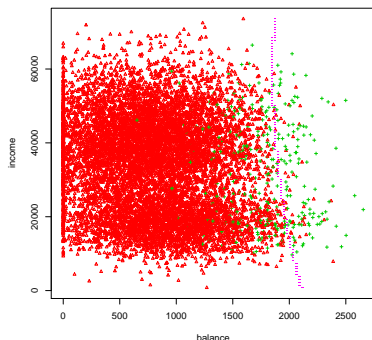
$$\Phi(X) = (1, \phi_1(X), \phi_2(X), \phi_3(X)) = (1, x_1, x_2, x_2^2)$$

Now, we can model

$$\begin{aligned} \log \left(\frac{p}{1-p} \right) &= \beta^\top \Phi(X) \\ &= \beta_0 + \beta_1 \text{balance} + \beta_2 \text{income} + \beta_3 \text{income}^2 \end{aligned}$$

POLYNOMIALS: CLASSIFICATION

```
out.glm = glm(default~balance + income +  
              I(income^2),family='binomial')
```



CONCLUSION: Linear rules in a transformed space can have nonlinear decisions in original features

BASIS FUNCTIONS

For nonparametric models the functions ϕ_k are **basis functions**

A general model would be

$$f_*(X) = \beta^\top \Phi(X)$$

Which we can fit via

$$\hat{\beta} = \underset{\beta}{\operatorname{argmin}} ||Y - \Phi\beta||_2^2,$$

where $\Phi = [\phi_k(X_i)] \in \mathbb{R}^{n \times K}$,

- k indexes the columns of Φ : $k = 1, \dots, K$
- i indexes the rows of Φ : $i = 1, \dots, n$

See ISL 7.3. Note that ISL uses b_k for basis. We use ϕ_k to connect to feature transformations

SPLINES

Choosing different ϕ_k , leads to different fits

A classic choice is (cubic) **splines**: $\phi_k(X) = (X - \xi)_+^3$

(Along with $\phi_k = X^k$ for $k = 0, 1, 2, 3$)

The ξ are **knots**. The knots are typically on a uniform grid

Splines are a basis that provides a very **smooth** representation of the regression function

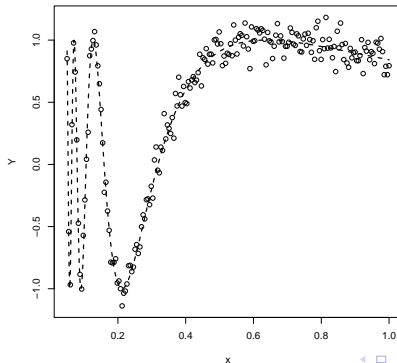
(Cubic splines do have a discontinuity in the third derivative)

But, again, we can view them as a way of defining a **distance** and **close**

Let's turn to the **DOPPLER FUNCTION** for an example..

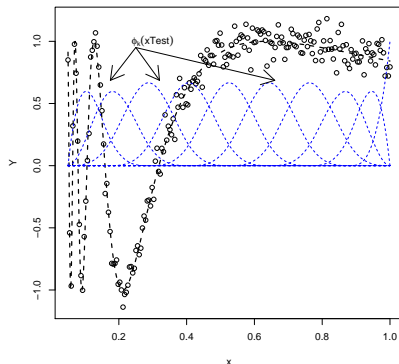
NONPARAMETRIC REGRESSION: EXAMPLE

```
x = seq(.05,1,length=200)
Y = sin(1/x) + rnorm(100,0,.1)
plot(x,Y)
xTest = seq(.05,1,length=1000)
lines(xTest,sin(1/xTest),col='black',lwd=2,lty=2)
```



NONPARAMETRIC REGRESSION: EXAMPLE

```
require(splines)
Phi = bs(x,df=10)
matlines(x=x,Phi,lty=2,type='l',col='blue')
```



NONPARAMETRIC REGRESSION: EXAMPLE

```
Phi = bs(x,df=K)
```

```
Yhat = predict(lm(Y~.,data=Phi))
```

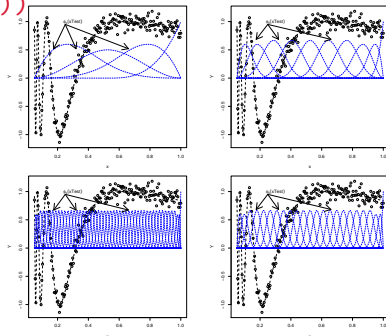
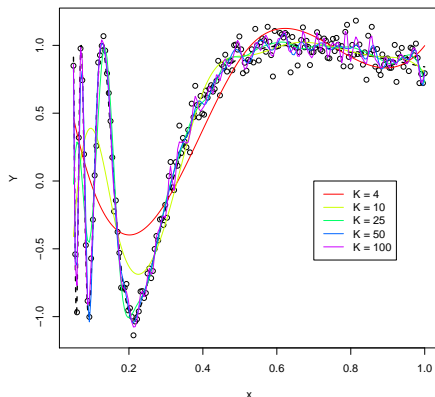


FIGURE: $K = 4, 10, 24, 50$,
clockwise from top left

SMOOTHING SPLINES

What we've just defined are known as **regression splines**

It required choosing the number of basis elements ϕ_k to control smoothness

Instead, we can dramatically over-parameterize: choose $K = n$

Then, we have to add regularization via a ridge-type penalty:

$$\min ||Y - \Phi\beta||_2^2 + \lambda||\nabla^2\beta||_2^2$$

The λ is typically chosen via GCV

(Don't worry overmuch about the form of the penalty term, just the 'type')

FROM LINEAR TO NONLINEAR MODELS

QUESTION: Why don't we always fit such a flexible model?

ANSWER: This works great if p is small (though it is easier to interpret linear regression output)

(and the specification of **distance** and **close** is good)

However, as p gets large

- **nothing** is nearby
- **all** points are on the boundary

(Hence, predictions are generally extrapolations)

These aspects make up (part) of the **curse of dimensionality**

CURSE OF DIMENSIONALITY

Fix the dimension p

(Assume p is even to ignore unimportant digressions)

Let S be a hypersphere with radius r

Let C be a hypercube with side length $2r$

Then, the volume of S and C are, respectively

$$V_S = \frac{r^p \pi^{p/2}}{(p/2)!} \text{ and } V_C = (2r)^p$$

(Interesting observation: this means for $r < 1/2$ the volume of the hypercube goes to 0, but the diagonal length is always $\propto \sqrt{p}$. Hence, the hypercube gets quite 'spiky' and is actually horribly jagged. Regardless of radius, the hypersphere's volume goes to zero quickly.)

CURSE OF DIMENSIONALITY

Hence, the ratio of the volumes of a circumscribed hypersphere by a hypercube is

$$\frac{V_C}{V_S} = \frac{(2r)^p \cdot (p/2)!}{r^p \pi^{p/2}} = \frac{2^p \cdot (p/2)!}{\pi^{p/2}} = \left(\frac{4}{\pi}\right)^d d!$$

where $d = p/2$

OBSERVATION: This ratio of volumes is increasing **really** fast. This means that all of the volume of a hypercube is near the corners. Also, this is independent of the radius.

Additive models

(ISL 7.7, 7.8.3)

ADDITIVE MODELS

The linear model looks like:

$$f(X) = \beta^\top X = \sum_{j=1}^p \beta_j x_j$$

The general nonparametric model looks like:

$$f(X) = \beta^\top \Phi(X) = \sum_{j=1}^K \beta_j \phi_j(X)$$

COMBINE THEM:

$$f(X) = f_1(x_1) + \cdots + f_p(x_p) = \sum_{j=1}^p f_j(x_j)$$

Estimation of such a procedure is not much more complicated than a fully linear model (as all inputs enter separately)

ADDITIVE MODELS (FOR REGRESSION)

Additive models are usually phrased using the **population level** expectation

(These get replaced with empirical versions)

The update is a Gauss-Seidel-type update

(The Gauss-Seidel method is an iterative scheme for solving linear, square systems)

This is for $j = 1, \dots, p, 1, \dots, p, 1 \dots$:

$$f_j(x_j) \leftarrow \mathbb{E} \left[Y - \sum_{k \neq j} f_k(x_k) | x_j \right]$$

Under fairly general conditions, this converges to $\mathbb{E}[Y|X]$

ADDITIVE MODELS (FOR REGRESSION)

Backfitting for additive models is roughly as follows:

Choose a univariate nonparametric smoother \mathcal{S} and form all univariate fits \hat{f}_j on each x_j marginally

(Common choice: splines)

Iterate over j until convergence:

1. Define the residuals $R_i = Y_i - \sum_{k \neq j} \hat{f}_k(X_{ik})$
2. Smooth the residuals $\hat{f}_j = \mathcal{S}(R)$
3. Center $\hat{f}_j \leftarrow \hat{f}_j - n^{-1} \sum_{i=1}^n \hat{f}_j(X_{ij})$

Report

$$\hat{f}(X) = \bar{Y} + \hat{f}_1(x_1) + \cdots + \hat{f}_p(x_p)$$

FITTING ADDITIVE MODELS R

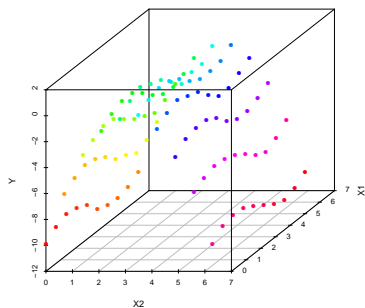
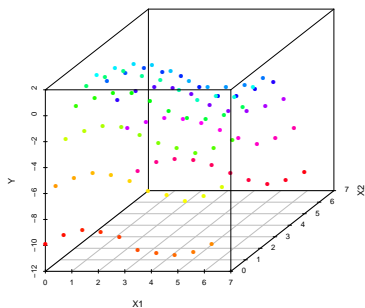
```
library(gam)
x  = seq(0,2*pi,length=10)
xx = expand.grid(x,x)
x1 = xx[,1]
x2 = xx[,2]

Y = sin(xx[,1]) - (xx[,2] - pi)^2 + rnorm(nrow(xx),0,.1)
sim = data.frame(x1=x1,x2=x2,Y=Y)

out = gam(Y~s(x1,3)+s(x2,3),data=sim)
```

ADDITIVE MODELS: SIMULATION

Plotting the simulated data: Y as a function of x_1, x_2 :



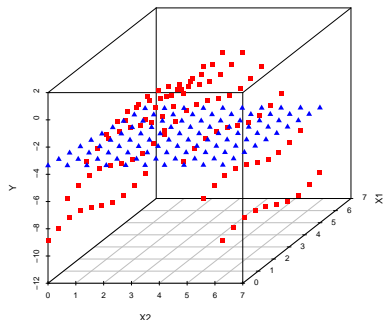
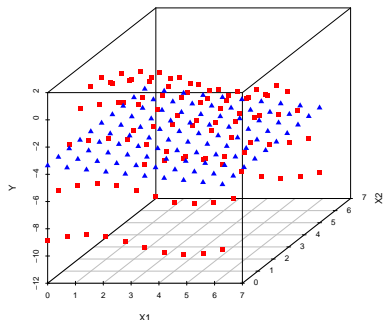
ADDITIVE MODELS: SIMULATION RESULTS

If we fit the (multiple) linear regression:

$$Y = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \epsilon,$$

we miss a lot of the structure (blue triangles)

The smoother GAM is able to capture it, however:



(These are the fitted values only. Red squares: GAM, Blue triangles: linear regression)

DETOUR: PLOTTING 3D IN R

```
out = scatterplot3d(X1,X2,Y,pch=16,type='n')
xyz = out$xyz.convert(X1,X2,out.pred)
points(xyz,col='red',pch=15)
xyz = out$xyz.convert(X1,X2,out.pred.lm)
points(xyz,col='blue',pch=17)
```

ADDITIVE MODELS (FOR CLASSIFICATION)

As squared error loss isn't quite right for classification, **additive logistic regression** is a popular approach

Suppose $Y \in \{-1, 1\}$

$$\log \left(\frac{\mathbb{P}(Y = 1|X)}{\mathbb{P}(Y = -1|X)} \right) = \sum_{j=1}^p f_j(x_j) = f(X)$$

This gets inverted in the usual way to acquire a probability estimate

$$\pi(X) = \mathbb{P}(Y = 1|X) = \frac{e^{f(X)}}{1 + e^{f(X)}}$$

$(f_j(x_j) = \beta_j x_j$ and $f(X) = X^\top \beta$ gives us (linear) logistic regression)

These models are usually fit by numerically maximizing the binomial likelihood, and hence enjoy all the asymptotic optimality prospects of MLEs

ADDITIVE MODELS (FOR CLASSIFICATION)

EXAMPLE: In **R**, this can be fit with the package **gam**

In the **gam** package there is a dataset **kyphosis**

This dataset examines a disorder of the spine

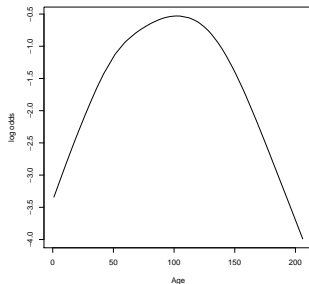
Let's look at two possible features **Age** and **Number**

(**Number** refers to the number of vertebrae that were involved in a surgery)

ADDITIVE MODELS (FOR CLASSIFICATION)

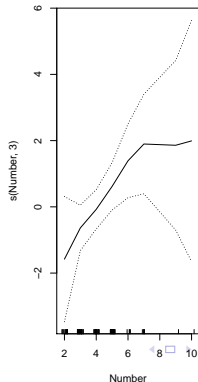
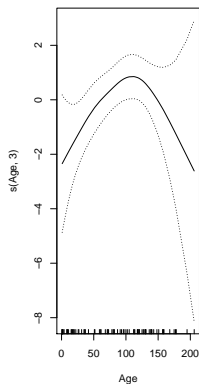
```
library(gam)
data(kyphosis)

out = gam(Kyphosis~s(Age,3),family=binomial,data=kyphosis)
out.pred = predict(out)
plot(sort(kyphosis$Age),out.pred[order(kyphosis$Age)],
     type='l',xlab='Age',ylab='log odds')
```



ADDITIVE MODELS (FOR CLASSIFICATION)

```
out = gam(Kyphosis ~ s(Age,3) + s(Number,3),  
          family = binomial, data=kyphosis)  
par(mfrow=c(1,2))  
plot(out,se=T)
```



Postamble:

- Nonparametric regression provides a flexible fit, but can only be used in low dimensions
(Curse of dimensionality)
- Additive models are a relaxation of nonparametric models that can be fit in much higher dimensions