Identification, ARMA Models, and Wold Decomposition

TA Session 2: Reduced-Form Time Series Analysis

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Today's Roadmap

Hour 1: Theory (25-30 minutes per topic)

1. **Identification in Time Series**
   - ▶ Global vs. local identification
   - ▶ Observational equivalence and ACGF
   - ▶ Causality & invertibility

2. **Reduced-Form Models: AR, MA, ARMA**
   - ▶ Definitions, stationarity, and invertibility
   - ▶ Observational equivalence in MA processes

3. **Wold's Decomposition Theorem**

Hour 2: Exercises

- ▶ Tim's Practice Problems (Ex. 2, 3)
- ▶ Medeiros ARMA Exercises (1, 2, 3, 6) + Vector Processes (Ex. 4)

Part I: Identification in Time Series

The Challenge of Learning from Data

Always Remember: Identification is a non-testable assumption!

The Identification Problem: Setup

Setup: We have a parametric model $\mathcal{M}(\theta)$ with structural parameters $\theta \in \Theta$.

The model generates a family of distributions for observed data:

$$\mathbb{D}(\mathbf{Y}^T \mid \theta), \quad \text{where } \mathbf{Y}^T = \{y_t\}_{t=1}^{T}$$

The Fundamental Question:

Can we uniquely determine θ from the distribution of the data?

Why This Matters

If two different values of θ generate the **same distribution** of data, then we **cannot distinguish** between them using any amount of data!

Global and Local Identification

Definition 1 (Global Identification)

$\theta_0 \in \Theta$ is **globally identified** if for all $\theta \in \Theta$:

$$\mathbb{D}(\mathbf{Y}^T \mid \theta) = \mathbb{D}(\mathbf{Y}^T \mid \theta_0) \implies \theta = \theta_0$$

Definition 2 (Local Identification)

$\theta_0 \in \Theta$ is **locally identified** if there exists a neighborhood U_{θ_0} such that for all $\theta \in U_{\theta_0}$:

$$\mathbb{D}(\mathbf{Y}^T \mid \theta) = \mathbb{D}(\mathbf{Y}^T \mid \theta_0) \implies \theta = \theta_0$$

Intuition: Global $\Rightarrow$ unique in entire space; Local $\Rightarrow$ unique in a neighborhood.

Linear Gaussian Models: Why They Matter for Time Series

For **linear Gaussian models**: $\mathbf{Y}^T \sim \mathcal{N}(\mu(\theta), \Sigma(\theta))$

The mean and variance **fully characterize** the distribution.

Connection to Time Series

Most time series models we study are linear Gaussian!

- ▶ AR, MA, ARMA processes with Gaussian innovations
- ▶ VARs and linearized DSGE models
- ▶ Wold decomposition assumes covariance stationarity

$\Rightarrow$ **Second-moment properties** (autocovariances) become the key object for identification!

Implication: If two parameter values yield the same autocovariance structure, they are **observationally equivalent**.

Observational Equivalence

Definition 3 (Observational Equivalence)

Two models $\mathcal{M}(\theta_1)$ and $\mathcal{M}(\theta_2)$ are **observationally equivalent** if:

$$\mathbb{D}(\mathbf{Y}^T \mid \theta_1) = \mathbb{D}(\mathbf{Y}^T \mid \theta_2)$$

In Time Series: For covariance-stationary Gaussian processes, observational equivalence reduces to having the **same second-moment properties**.

The Identification Challenge

Different structural models/parameters can generate **identical autocovariance functions**.

$\Rightarrow$ Data alone cannot distinguish between them!

Next: How do we formalize "same second-moment properties"?

The Autocovariance Generating Function (ACGF)

Tool for comparing second-moment properties:

Definition 4 (ACGF)

For a covariance-stationary process with autocovariances $\Gamma(j)$:

$$\Omega(z) = \sum_{j=-\infty}^{\infty} \Gamma(j) z^j$$

Key Insight: The Identification Challenge in Macroeconometrics

Two time series models are observationally equivalent if and only if they have the **same ACGF**:

$$\Omega(z; \theta_1) = \Omega(z; \theta_2) \quad \forall z$$

This is the root of why structural interpretation requires **additional identifying assumptions**!

Preview: Types of Indeterminacy

Observational equivalence can lead to **indeterminacy** — multiple structural interpretations consistent with the same data.

Three main sources (covered in detail in the VAR section):

1. **Static Indeterminacy**: Orthogonal rotations of shocks
2. **Dynamic Indeterminacy**: Non-fundamental representations
3. **Size Indeterminacy**: More shocks than observables

Example: MA(1) "Root Flipping"

Consider $y_t = u_t + \theta u_{t-1}$ with $|\theta| < 1$.
The process with $\theta^* = 1/\theta$ and rescaled variance has **identical autocovariances**!
Both: $\rho_1 = \frac{\theta}{1+\theta^2}$ (same for θ and $1/\theta$)

We'll develop identification strategies when we cover Structural VARs.

Causality and Invertibility: General Definitions

Let $\{Y_t\}$ be a stationary process and $\{Z_t\}$ an input/shock process.

Definition 5 (Causality)

$\{Y_t\}$ is **causal with respect to** $\{Z_t\}$ if Y_t is $\mathcal{F}_t^Z$-measurable for all t.

Definition 6 (Invertibility)

$\{Y_t\}$ is **invertible with respect to** $\{Z_t\}$ if Z_t is $\mathcal{F}_t^Y$-measurable for all t.

Interpretations:

- ▶ Y_t depends only on **current and past** values of Z — no future inputs needed.
- ▶ The shocks Z_t can be **recovered from current and past observables** Y_t.

VARMA: Causality and Invertibility

Definition 7

A VARMA process $\{y_t\}$ is **causal** w.r.t. $\{z_t\}$ if $y_t \in \text{span}(z_\tau, -\infty < \tau \leq t)$

It is **invertible** w.r.t. $\{z_t\}$ if $z_t \in \text{span}(y_\tau, -\infty < \tau \leq t)$

Sufficient conditions for $\Phi(L)y_t = \Theta(L)z_t$:

- **Causality**: $\Phi(L)$ has 1-sided inverse $\Rightarrow$ VMA(∞): $y_t = \Psi(L)z_t$
- **Invertibility**: $\Theta(L)$ has 1-sided inverse $\Rightarrow$ VAR(∞): $\Pi(L)y_t = z_t$

Key takeaways:

- Structural macro models yield VMA representations $\Rightarrow$ always causal
- Invertibility is **far from guaranteed!** (e.g., 5 shocks but 2 observables)

Why Invertibility Matters

For estimation and structural analysis:

If the process is invertible:
- ▶ Innovations z_t can be recovered from data
- ▶ We can estimate VAR and back out the shocks
- ▶ IRFs and variance decompositions become meaningful

If NOT invertible:
- ▶ VAR residuals are **not** the structural shocks
- ▶ Standard SVAR methods may give misleading results
- ▶ Need alternative identification strategies

Necessary Condition

For invertibility: $n_{\text{shocks}} \leq n_{\text{observables}}$
But this is **not sufficient**! (Forward-looking behavior can break invertibility)

Part I: Summary

Concept	Key Point
Global Identification	Unique θ in entire parameter space
Local Identification	Unique in neighborhood
Linear Gaussian	Second moments fully characterize distribution
Observational Equiv.	Same ACGF $\Rightarrow$ indistinguishable
Causality	Y_t is $\mathcal{F}_t^Z$-measurable (depends on current/past Z's)
Invertibility	Z_t is $\mathcal{F}_t^Y$-measurable (recoverable from current/past Y's)

Key Insight: In linear Gaussian time series, identification hinges on second-moment properties. Observational equivalence $\Rightarrow$ **indeterminacy**, which requires additional assumptions to resolve (covered in VAR section).

Part II: Reduced-Form Models

AR, MA, and ARMA Processes

The AR(p) Process: Definition

Definition 8 (Autoregressive Process of Order p)

A stochastic process $\{Y_t\}$ is AR(p) if:

$$Y_t = \mathbb{E}[Y_t \mid \mathcal{F}_{t-1}] + u_t = \phi_0 + \phi_1 y_{t-1} + \cdots + \phi_p y_{t-p} + u_t$$

Equivalently: $\Phi_p(L)y_t = \phi_0 + u_t$, where:

$$\Phi_p(L) = 1 - \phi_1 L - \phi_2 L^2 - \cdots - \phi_p L^p$$

and u_t is white noise: $\mathbb{E}[u_t \mid \mathcal{F}_{t-1}] = 0$, $\mathbb{E}[u_t^2 \mid \mathcal{F}_{t-1}] = \sigma^2$.

Interpretation: Current value = conditional expectation given past + innovation.

Note: AR processes are **always invertible** (by construction, $u_t = y_t - \mathbb{E}[y_t \mid \mathcal{F}_{t-1}]$).

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The Companion Form

Idea: Cast AR(p) as a VAR(1) by stacking variables.

Definition 9 (Companion Form)

$$\underbrace{\begin{bmatrix} Y_t \\ Y_{t-1} \\ \vdots \\ Y_{t-p+1} \end{bmatrix}}_{\mathbf{X}_t} = \begin{bmatrix} \phi_0 \\ 0 \\ \vdots \\ 0 \end{bmatrix} + \underbrace{\begin{bmatrix} \phi_1 & \phi_2 & \cdots & \phi_p \\ 1 & 0 & \cdots & 0 \\ \vdots & \ddots & & \vdots \\ 0 & \cdots & 1 & 0 \end{bmatrix}}_{F} \underbrace{\begin{bmatrix} Y_{t-1} \\ Y_{t-2} \\ \vdots \\ Y_{t-p} \end{bmatrix}}_{\mathbf{X}_{t-1}} + \begin{bmatrix} u_t \\ 0 \\ \vdots \\ 0 \end{bmatrix}$$

Compactly: $\mathbf{X}_t = c + F\mathbf{X}_{t-1} + v_t$

Why useful?
▶ $\{\mathbf{X}_t\}$ is Markovian even though $\{Y_t\}$ is not
▶ Stationarity analysis reduces to eigenvalue analysis of F

AR(p): Stationarity Conditions

Theorem 10 (Asymptotic Stationarity of AR(p))

*An AR(p) process is **asymptotically stationary** if the eigenvalues of the companion matrix F are all **inside the unit circle**.*

Sketch.

Iterating backward: $\mathbf{X}_t = F^t \mathbf{X}_0 + \sum_{i=0}^{t-1} F^i v_{t-i}$. For this not to explode, need $F^t \to 0$, which requires $|\lambda_i| < 1$ for all eigenvalues. $\qquad \square$

Corollary 11 (Characteristic Polynomial Version)

*AR(p) is stationary iff roots of $\Phi_p(z) = 1 - \phi_1 z - \cdots - \phi_p z^p = 0$ lie **outside the unit circle**.*

Connection: Eigenvalues of F are reciprocals of polynomial roots.

AR(p): Moments (Yule-Walker Equations)

For a weakly-stationary AR(p) process:

Expected value:

$$\mathbb{E}[y_t] = \frac{\phi_0}{1 - \sum_{i=1}^{p} \phi_i}$$

Autocovariance (Yule-Walker Equations):

$$\gamma_k = \sum_{i=1}^{p} \phi_i \gamma_{k-i}, \quad k \geq 1$$

Variance:

$$\gamma_0 = \sigma^2 + \sum_{i=1}^{p} \phi_i \gamma_i$$

Key Pattern

ACF **decays geometrically** — signature of AR processes!

AR(1) Example: Monetary Policy and Inflation

Recall from Lecture 1: The simple monetary transmission model:

$$y_t = -\gamma(i_t - \mathbb{E}_t[\pi_{t+1}]) + u_{1t} \quad \text{(IS curve)}$$
$$\pi_t = \lambda y_t + u_{2t} \quad \text{(Phillips curve)}$$
$$i_t = \rho \pi_t + i^* \quad \text{(Taylor rule)}$$

Reduced form for inflation:

$$\pi_t = \phi_0 + \phi_1 \pi_{t-1} + v_t$$

where $\phi_1 = 1 + \lambda\gamma(\rho - 1)$ (function of structural parameters!)

Stability requires: $|\phi_1| < 1 \Leftrightarrow$ **Taylor Principle**: $\rho > 1$

Key insight: The reduced-form AR coefficient encodes structural policy parameters!

AR(1) Inflation Example: Behavior by ρ

Cases (with $\lambda = 0.5$, $\gamma = -0.5$):

Taylor Rule ρ	Reduced-form ϕ_1	Behavior
$\rho = 0$	$\phi_1 > 1$	Explosive inflation
$0 < \rho < 1$	$\phi_1 > 1$	Explosive inflation
$\rho = 1$	$\phi_1 = 1$	Random walk
$1 < \rho < 5$	$0 < \phi_1 < 1$	Stationary, persistent
$5 < \rho < 9$	$-1 < \phi_1 < 0$	Stationary, oscillating
$\rho > 9$	$\phi_1 < -1$	Explosive (over-reaction)

ACF: $\rho_k = \phi_1^k$ — geometric decay toward zero.

See figures in Appendix → Figures

The MA(q) Process: Definition

Definition 12 (Moving Average Process of Order q)

A process $\{y_t\}$ is MA(q) if:

$$y_t = \theta_0 + u_t + \theta_1 u_{t-1} + \cdots + \theta_q u_{t-q} = \theta_0 + \theta_q(L)u_t$$

where $\theta_q(L) = 1 + \theta_1 L + \cdots + \theta_q L^q$ and $u_t \sim WN(0, \sigma^2)$.

Key difference from AR: Current value is a weighted sum of current and **past innovations** (not past values of y).

Proposition 1

*All finite-order MA(q) processes are **covariance-stationary** and **mean-ergodic**.*

Intuition: Finite sum of stationary (white noise) terms is always stationary.

MA(q): Moments

Expected value: $\mathbb{E}[y_t] = \theta_0$

Variance:

$$\gamma_0 = \sigma^2 \left(1 + \sum_{j=1}^{q} \theta_j^2\right)$$

Autocovariance:

$$\gamma_k = \begin{cases} \left(\theta_k + \sum_{j=1}^{q-k} \theta_j \theta_{j+k}\right)\sigma^2 & \text{if } k \leq q \\ 0 & \text{if } k > q \end{cases}$$

Key Pattern

ACF **cuts off** after lag q — signature of MA processes!
(Contrast with AR: geometric decay)

MA(q): Invertibility

Definition 13 (Invertibility)

An MA(q) process is **invertible** if u_t lies in the space spanned by current and past observables: $u_t \in \text{span}(y_\tau, -\infty < \tau \leq t)$.

Equivalently, MA(q) can be written as AR(∞):

$$u_t = \theta_q(L)^{-1}(y_t - \theta_0) = \sum_{j=0}^{\infty} \psi_j(y_{t-j} - \theta_0)$$

Theorem 14 (Invertibility Condition)

*MA(q) is invertible iff roots of $\theta_q(z) = 1 + \theta_1 z + \cdots + \theta_q z^q = 0$ lie **outside the unit circle**.*

Practical importance: Invertibility allows expressing u_t as function of observables — essential for estimation!

MA(1): Observational Equivalence Example

Key fact: For any MA process, there exists a family of observationally equivalent alternatives with the same mean and autocovariances.

Example: MA(1) with $y_t = u_t + \theta u_{t-1}$, $u_t \sim (0, \sigma^2)$.

Autocorrelation at lag 1:

$$\rho_1 = \frac{\theta}{1 + \theta^2}$$

Observation: Both θ and $1/\theta$ give the **same** ρ_1!

The **invertible** representation ($|\theta| < 1$) is obs. equivalent to a **non-invertible** one:

- ▶ Original: θ, σ^2
- ▶ Alternative: $\theta^* = 1/\theta$, $\sigma^{*2} = \theta^2 \sigma^2$

Resolution

We **cannot test** statistically whether a process is invertible!
We impose invertibility as an *a priori* identifying assumption.

MA Observational Equivalence: ACGF Perspective

Tool: The Autocovariance Generating Function.

Proposition 2

For MA(q): $y_t - \theta_0 = \theta_q(L)\varepsilon_t$, the ACGF is:

$$\Gamma_Y(z) = \theta_q(z)\theta_q(z^{-1})\sigma_\varepsilon^2$$

Two MA processes are observationally equivalent if they have:
1. Same mean: $\theta_0^X = \theta_0^Y$
2. Same ACGF: $\Gamma_X(z) = \Gamma_Y(z)$

For MA(1): Matching $\Gamma(z) = (1 + \theta z)(1 + \theta z^{-1})\sigma^2$ allows the "flip":

$$\theta \leftrightarrow 1/\theta, \quad \sigma^2 \leftrightarrow \theta^2\sigma^2$$

The ARMA(p, q) Process

Definition 15 (ARMA Process)

A process $\{Y_t\}$ is ARMA(p, q) if:

$$Y_t = \phi_0 + \sum_{i=1}^{p} \phi_i Y_{t-i} + u_t + \sum_{j=1}^{q} \theta_j u_{t-j}$$

Or: $\Phi_p(L)Y_t = \phi_0 + \Theta_q(L)u_t$

Proposition 3 (Stationarity and Invertibility)

▶ **Stationarity**: *Roots of $\Phi_p(z) = 0$ outside unit circle (AR part)*
▶ **Invertibility**: *Roots of $\Theta_q(z) = 0$ outside unit circle (MA part)*

ACGF:

$$\Gamma(z) = \frac{\Theta_q(z)\Theta_q(z^{-1})}{\Phi_p(z)\Phi_p(z^{-1})}\sigma_\varepsilon^2$$

ARMA: Common Factor Problem

Identification issue: If AR and MA polynomials share a common root, the model is **not identified**.

Example: ARMA(1,1) with $(1 - \phi L)y_t = (1 + \theta L)u_t$
If $\phi = -\theta$ (common factor):

$$(1 - \phi L)y_t = (1 - \phi L)u_t \implies y_t = u_t$$

The ARMA(1,1) is **observationally equivalent to white noise**!

Solution

For a **canonical** (identified) ARMA representation, require:
$\Phi_p(z)$ and $\Theta_q(z)$ have **no common roots**.

ARMA: Equivalent Representations

A stationary and invertible ARMA(p, q) admits:

1. MA(∞) Representation:

$$y_t = \mu + \psi(L)u_t = \mu + \sum_{j=0}^{\infty} \psi_j u_{t-j}$$

where $\psi(L) = \Phi_p(L)^{-1}\Theta_q(L)$

2. AR(∞) Representation:

$$\pi(L)(y_t - \mu) = u_t$$

where $\pi(L) = \Theta_q(L)^{-1}\Phi_p(L)$

Key Insight

ARMA provides a **parsimonious parametric approximation** to MA(∞)/AR(∞) — crucial for Wold decomposition!

Summary: AR vs MA vs ARMA

Property	$\mathbf{AR}(p)$	$\mathbf{MA}(q)$	$\mathbf{ARMA}(p,q)$
Stationarity	Roots of $\Phi(z)$ outside UC	Always stationary	Roots of $\Phi(z)$ outside UC
Invertibility	Always invertible	Roots of $\Theta(z)$ outside UC	Roots of $\Theta(z)$ outside UC
ACF pattern	Geometric decay	Cuts off after lag q	Geometric decay
PACF pattern	Cuts off after lag p	Geometric decay	Geometric decay

UC = Unit Circle; **PACF** = Partial Autocorrelation Function

Part III: Wold's Decomposition Theorem

The Fundamental Representation of Stationary Processes

Wold's Representation Theorem

Theorem 16 (Wold, 1938)

*Let $\{y_t\}$ be an n-dimensional **covariance-stationary** time series. There exists an $n \times n$ lag polynomial $\Psi(L)$ and $u_t \sim WN(0, \Sigma)$ such that:*

$$y_t = \Psi(L)u_t + d_t, \quad \Psi(L) = I + \sum_{\ell=1}^{\infty} \Psi_\ell L^\ell$$

where:

1. *$\Psi(L)$ is **square-summable***
2. *$u_t = y_t - \mathbb{E}^*[y_t \mid \{y_\tau\}_{-\infty < \tau \leq t-1}]$ (one-step-ahead forecast errors)*
3. *$\{y_t\}$ is **invertible** with respect to u_t*
4. *$\{d_t\}$ is a **purely deterministic** process*

In words: Any covariance-stationary time series = VMA(∞) + deterministic

Why Wold Matters: The Key Insight

The Wold decomposition does something **very simple**: it splits a process $\{y_t\}$ into:

1. **One-step-ahead prediction errors** u_t
2. **A perfectly predictable residual** d_t

Three Equivalent Representations

For second-order properties, we can **freely map** between:

1. **Autocovariance function**: $\{\Gamma_k\}_{k=0}^{\infty}$
2. **Spectral density**: $s_y(\omega)$
3. **Wold decomposition**: $\{\Psi_j, \Sigma\}$

Mapping: $\Psi_\ell = \mathrm{Cov}(y_t, u_{t-\ell})\Sigma^{-1}$

Key: The Wold decomposition is **identifiable from data**!

Wold: Practical Implications

Problem: MA(∞) has infinitely many parameters — cannot estimate!

Solution: ARMA(p, q) provides a parsimonious approximation:

$$\Phi_p(L)y_t = \Theta_q(L)u_t \quad \Leftrightarrow \quad y_t = \frac{\Theta_q(L)}{\Phi_p(L)}u_t$$

The ratio $\psi(L) = \Theta_q(L)/\Phi_p(L)$ approximates the Wold coefficients $\{\Psi_i\}$.

Why This Matters

- ▶ We can **estimate** finite-parameter ARMA/VAR models
- ▶ These approximate the infinite-parameter Wold representation
- ▶ From VAR, we can also get VAR(∞) representation: $A(L)y_t = u_t$

Wold vs. Structural: A Critical Warning

Wold Innovations $\neq$ Structural Shocks!

The Wold coefficients $\{\Psi_j\}$ are **reduced-form** objects.
Nothing guarantees they are interesting — they are just coefficients on reduced-form prediction errors!

Structural VMA (what we want):

$$y_t = \sum_{\ell=0}^{\infty} \Theta_\ell \varepsilon_{t-\ell}, \quad \varepsilon_t = \text{structural shocks}$$

Wold representation (what we estimate):

$$y_t = \sum_{j=0}^{\infty} \Psi_j u_{t-j}, \quad u_t = \text{prediction errors}$$

The gap: Identifying Θ_ℓ from Ψ_j requires **additional assumptions**!
$\Rightarrow$ This is the **identification problem** we'll address in Structural VARs.

Summary

Identification:

- ▶ Global vs. local: uniqueness in entire space vs. neighborhood
- ▶ Linear Gaussian $\Rightarrow$ second moments matter
- ▶ Observational equivalence: same ACGF $\Rightarrow$ indistinguishable
- ▶ Causality: y_t from current/past z's; Invertibility: z_t from y's

Reduced-Form Models:

- ▶ AR(p): Stationary if roots outside UC; ACF decays geometrically
- ▶ MA(q): Always stationary; invertible if roots outside UC; ACF cuts off
- ▶ ARMA: Combines both; common factor $\Rightarrow$ non-identification

Wold's Theorem:

- ▶ Any stationary process $= \text{VMA}(\infty) + $ deterministic
- ▶ Wold innovations $=$ prediction errors (**not** structural shocks!)
- ▶ Reduced-form, identifiable; structural interpretation needs more

Part IV: Exercises

Problem Statement

[Exercise 2 from practice_problems_week2.pdf]

Problem to be inserted from the problem set.

Solution Approach:

1. Step 1: [To be completed]
2. Step 2: [To be completed]
3. Step 3: [To be completed]

Exercise 2: Solution

Detailed Solution:

[Solution steps to be added once problem statement is available]

Key Takeaway

[Main insight from this exercise]

Tim's Practice Problems: Exercise 3

Problem Statement

[Exercise 3 from practice_problems_week2.pdf]

Problem to be inserted from the problem set.

Solution Approach:

1. Step 1: [To be completed]
2. Step 2: [To be completed]
3. Step 3: [To be completed]

Exercise 3: Solution

Detailed Solution:

[Solution steps to be added once problem statement is available]

Key Takeaway

[Main insight from this exercise]

Medeiros ARMA: Exercise 1 — Statement

Problem

Let y_t, $t = 1, \ldots, T$, be a time series described by an MA(1) process:

$$y_t = \mu + \varepsilon_t + \theta\varepsilon_{t-1}, \quad \varepsilon_t \sim \mathrm{IID}(0, \sigma_\varepsilon^2)$$

Suppose that y_t is observed with **additive measurement noise**, i.e., $x_t = y_t + u_t$, where $u_t \sim \mathrm{IID}(0, \sigma_u^2)$ and $\mathbb{E}(\varepsilon_t u_s) = 0$ for all t, s.

Compute:

(a) The mean and variance of x_t.

(b) The autocovariances of x_t.

(c) Which ARMA(p, q) process best describes x_t?

Medeiros ARMA: Exercise 1 — Solution (a)

(a) Mean and Variance of x_t:

Since $x_t = y_t + u_t = \mu + \varepsilon_t + \theta\varepsilon_{t-1} + u_t$:

Mean:

$$\mathbb{E}(x_t) = \mathbb{E}(\mu + \varepsilon_t + \theta\varepsilon_{t-1} + u_t) = \mu$$

Variance: Since ε_t, ε_{t-1}, and u_t are mutually uncorrelated:

$$\gamma_0 = \text{Var}(x_t) = \text{Var}(\varepsilon_t) + \theta^2\text{Var}(\varepsilon_{t-1}) + \text{Var}(u_t)$$
$$= \sigma_\varepsilon^2 + \theta^2\sigma_\varepsilon^2 + \sigma_u^2 = \boxed{(1 + \theta^2)\sigma_\varepsilon^2 + \sigma_u^2}$$

(b) Autocovariances of x_t:

For $k \geq 1$:

$$\gamma_k = \text{Cov}(x_t, x_{t-k}) = \text{Cov}(\varepsilon_t + \theta\varepsilon_{t-1} + u_t, \varepsilon_{t-k} + \theta\varepsilon_{t-k-1} + u_{t-k})$$

For $k = 1$: The only surviving term is $\text{Cov}(\theta\varepsilon_{t-1}, \varepsilon_{t-1})$:

$$\gamma_1 = \theta \cdot \sigma_\varepsilon^2 = \boxed{\theta\sigma_\varepsilon^2}$$

For $k \geq 2$: There are no overlapping terms:

$$\gamma_k = \boxed{0}, \quad k \geq 2$$

Result

The autocovariance structure is: $\gamma_0 = (1 + \theta^2)\sigma_\varepsilon^2 + \sigma_u^2$, $\gamma_1 = \theta\sigma_\varepsilon^2$, $\gamma_k = 0$ for $k \geq 2$.

(c) Which ARMA(p, q) describes x_t?

The autocovariance "cuts off" after lag $1 \Rightarrow$ **MA(1)** behavior!

Therefore, x_t can be written as:

$$x_t = \mu + \eta_t + \tilde{\theta}\eta_{t-1}, \quad \eta_t \sim \text{WN}(0, \sigma_\eta^2)$$

Matching parameters: We need:

$$\gamma_0^{\text{MA(1)}} = (1 + \tilde{\theta}^2)\sigma_\eta^2 = (1 + \theta^2)\sigma_\varepsilon^2 + \sigma_u^2$$

$$\gamma_1^{\text{MA(1)}} = \tilde{\theta}\sigma_\eta^2 = \theta\sigma_\varepsilon^2$$

Conclusion

x_t is an **MA(1)** process, but with parameters different from the original MA(1)!
Measurement noise "hides" the true structural parameters $(\theta, \sigma_\varepsilon^2)$.

Problem

Let $\{y_t\}$ and $\{z_t\}$, $t = 1, \ldots, T$, be two stochastic processes:

$$y_t = \phi_1 y_{t-1} + \varepsilon_t, \quad z_t = \phi_2 z_{t-1} + u_t$$

where $\varepsilon_t \sim \text{IID}(0, \sigma_\varepsilon^2)$, $u_t \sim \text{IID}(0, \sigma_u^2)$, and $\mathbb{E}(\varepsilon_t u_s) = 0$ for all t, s.

Show that:

(a) If $\phi_1 = \phi_2$, then $x_t = y_t + z_t$ is an AR(1) process.

(b) If $\phi_1 \neq \phi_2$, then $x_t = y_t + z_t$ is an ARMA(2, q) process with $q \leq 1$.

(a) Case $\phi_1 = \phi_2 = \phi$:

Adding the two equations:

$$y_t + z_t = \phi(y_{t-1} + z_{t-1}) + \varepsilon_t + u_t$$
$$x_t = \phi x_{t-1} + w_t$$

where $w_t = \varepsilon_t + u_t$.

Checking that w_t is white noise:

- $\mathbb{E}(w_t) = 0$
- $\mathrm{Var}(w_t) = \sigma_\varepsilon^2 + \sigma_u^2$
- $\mathrm{Cov}(w_t, w_{t-k}) = 0$ for $k \neq 0$ (since ε and u are serially independent and mutually uncorrelated)

Conclusion

x_t is an **AR(1)** process with parameter ϕ and innovation $w_t \sim \mathrm{WN}(0, \sigma_\varepsilon^2 + \sigma_u^2)$.

(b) Case $\phi_1 \neq \phi_2$:

Strategy: Write in lag-operator notation and combine.

The individual processes are:

$$(1 - \phi_1 L)y_t = \varepsilon_t, \quad (1 - \phi_2 L)z_t = u_t$$

To eliminate y_t and z_t separately, apply cross-filters:

$$(1 - \phi_2 L)(1 - \phi_1 L)y_t = (1 - \phi_2 L)\varepsilon_t$$
$$(1 - \phi_1 L)(1 - \phi_2 L)z_t = (1 - \phi_1 L)u_t$$

Summing:

$$(1 - \phi_1 L)(1 - \phi_2 L)(y_t + z_t) = (1 - \phi_2 L)\varepsilon_t + (1 - \phi_1 L)u_t$$

Expanding:

$$[1 - (\phi_1 + \phi_2)L + \phi_1\phi_2 L^2]x_t = \varepsilon_t - \phi_2\varepsilon_{t-1} + u_t - \phi_1 u_{t-1}$$

Right-hand side: Define $\eta_t = \varepsilon_t + u_t - \phi_2\varepsilon_{t-1} - \phi_1 u_{t-1}$.

Is this an MA(1) process in general? Compute autocovariances of η_t:

- $\gamma_0^\eta = (1 + \phi_2^2)\sigma_\varepsilon^2 + (1 + \phi_1^2)\sigma_u^2$
- $\gamma_1^\eta = -\phi_2\sigma_\varepsilon^2 - \phi_1\sigma_u^2$
- $\gamma_k^\eta = 0$ for $k \geq 2$

Conclusion

x_t follows an **ARMA(2,1)** process with:

- AR part: $(1 - \phi_1 L)(1 - \phi_2 L)$ with roots $1/\phi_1$ and $1/\phi_2$
- MA part: order $q \leq 1$ (it can be 0 if $\gamma_1^\eta = 0$)

Medeiros ARMA: Exercise 3 — Statement

Problem

Consider two independent AR(1) processes x_t and y_t defined by:

$$(1 - \rho_1 L)x_t = u_t, \ |\rho_1| < 1 \quad \text{and} \quad (1 - \rho_2 L)y_t = v_t, \ |\rho_2| < 1$$

where $\mathbb{E}(u_t) = \mathbb{E}(v_t) = 0$, $\mathbb{E}(u_t v_\tau) = 0$ for all t, τ,
$\mathbb{E}(u_t u_\tau) = \sigma_u^2$ if $t = \tau$ (0 otherwise), $\mathbb{E}(v_t v_\tau) = \sigma_v^2$ if $t = \tau$.
Let $z_t = x_t + y_t$.

Tasks:
(a) Compute $\mathbb{E}(x_{t+1}|\mathcal{F}_t)$, $\mathbb{E}(y_{t+1}|\mathcal{F}_t)$, $\mathbb{E}(z_{t+1}|\mathcal{F}_t)$, and the forecast error variance V_D.
(b) Show that z_t is ARMA(2,1) and find θ.
(c) Conditions for z_t to be AR(2)? Show that if ρ_1, ρ_2 have the same sign, then it cannot be AR(2).
(d) Show that $\theta \neq +1$

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(a) Conditional forecasts:

Since x_t and y_t are AR(1):

$$\mathbb{E}(x_{t+1}|\mathcal{F}_t) = \rho_1 x_t, \quad \mathbb{E}(y_{t+1}|\mathcal{F}_t) = \rho_2 y_t$$

Para $z_t = x_t + y_t$:

$$\mathbb{E}(z_{t+1}|\mathcal{F}_t) = \rho_1 x_t + \rho_2 y_t$$

Forecast error $e_{t+1} = z_{t+1} - \mathbb{E}(z_{t+1}|\mathcal{F}_t)$:

$$e_{t+1} = (x_{t+1} + y_{t+1}) - (\rho_1 x_t + \rho_2 y_t) = u_{t+1} + v_{t+1}$$

Error variance V_D:

$$V_D = \text{Var}(e_{t+1}) = \text{Var}(u_{t+1}) + \text{Var}(v_{t+1}) = \boxed{\sigma_u^2 + \sigma_v^2}$$

(b) Showing that z_t is ARMA(2,1):

From Exercise 2, we know that when $\rho_1 \neq \rho_2$:

$$[1 - (\rho_1 + \rho_2)L + \rho_1\rho_2 L^2]z_t = (1 + \theta L)\varepsilon_t$$

where ε_t is white noise with variance σ_ε^2.

Identification via autocovariances: For an MA(1), $\eta_t = (1 + \theta L)\varepsilon_t$:

$$\gamma_0^\eta = (1 + \theta^2)\sigma_\varepsilon^2, \quad \gamma_1^\eta = \theta\sigma_\varepsilon^2$$

From the previous exercise, the right-hand side has:

$$\gamma_0^\eta = (1 + \rho_2^2)\sigma_u^2 + (1 + \rho_1^2)\sigma_v^2$$

$$\gamma_1^\eta = -\rho_2\sigma_u^2 - \rho_1\sigma_v^2$$

Solving for θ:

The lag-1 autocorrelation of an MA(1) is:

$$\rho_1^{\text{MA}} = \frac{\gamma_1}{\gamma_0} = \frac{\theta}{1 + \theta^2}$$

Equating:

$$\frac{\theta}{1 + \theta^2} = \frac{-\rho_2 \sigma_u^2 - \rho_1 \sigma_v^2}{(1 + \rho_2^2)\sigma_u^2 + (1 + \rho_1^2)\sigma_v^2}$$

This is a quadratic equation in θ:

$$\theta^2 \cdot (\text{numerador}) - \theta \cdot (\text{denominador}) + (\text{numerador}) = 0$$

Two roots!

There are two solutions, θ and $1/\theta$, corresponding to observationally equivalent representations (invertible vs. non-invertible).

(c) Conditions for z_t to be AR(2):

z_t is AR(2) if the MA part disappears, i.e., $\theta = 0$, which requires $\gamma_1^\eta = 0$:

$$-\rho_2 \sigma_u^2 - \rho_1 \sigma_v^2 = 0 \quad \Rightarrow \quad \boxed{\rho_1 \sigma_v^2 = -\rho_2 \sigma_u^2}$$

If ρ_1 and ρ_2 have the same sign:

The left-hand side $\rho_1 \sigma_v^2$ and the right-hand side $-\rho_2 \sigma_u^2$ have opposite signs (since $\sigma_u^2, \sigma_v^2 > 0$).

Impossibility

If $\text{sgn}(\rho_1) = \text{sgn}(\rho_2)$, then the condition $\rho_1 \sigma_v^2 = -\rho_2 \sigma_u^2$ can **never** be satisfied. Therefore, z_t **cannot be AR(2)** when ρ_1 and ρ_2 have the same sign.

(d) Showing that $\theta \neq \pm 1$:

Suppose $\theta = 1$. Then $\rho_1^{\mathrm{MA}} = \frac{1}{1+1} = \frac{1}{2}$.

For this to occur, we would need:

$$\frac{-\rho_2 \sigma_u^2 - \rho_1 \sigma_v^2}{(1 + \rho_2^2)\sigma_u^2 + (1 + \rho_1^2)\sigma_v^2} = \frac{1}{2}$$

Since $|\rho_1|, |\rho_2| < 1$, the denominator $(1 + \rho_2^2)\sigma_u^2 + (1 + \rho_1^2)\sigma_v^2 > 2\sigma_u^2 + 2\sigma_v^2$ (using $\rho^2 < 1$).

The numerator has absolute value $|\rho_2 \sigma_u^2 + \rho_1 \sigma_v^2| < \sigma_u^2 + \sigma_v^2$.

Therefore, the absolute value of the ratio is:

$$\left| \rho_1^{\mathrm{MA}} \right| < \frac{\sigma_u^2 + \sigma_v^2}{2(\sigma_u^2 + \sigma_v^2)} = \frac{1}{2}$$

Conclusion

Since $|\rho^{\mathrm{MA}}| < 1/2$, and $\theta = \pm 1$ would imply $|\rho^{\mathrm{MA}}| = 1/2$, we have $\boxed{\theta \neq \pm 1}$.

(e) Forecasting using the ARMA(2,1):

Do modelo $[1 - (\rho_1 + \rho_2)L + \rho_1\rho_2 L^2]z_t = (1 + \theta L)\varepsilon_t$:

$$z_t = (\rho_1 + \rho_2)z_{t-1} - \rho_1\rho_2 z_{t-2} + \varepsilon_t + \theta\varepsilon_{t-1}$$

Forecast:

$$\mathbb{E}(z_{T+1}|\mathcal{F}_T) = (\rho_1 + \rho_2)z_T - \rho_1\rho_2 z_{T-1} + \theta\varepsilon_T$$

Error: $e_{T+1} = z_{T+1} - \mathbb{E}(z_{T+1}|\mathcal{F}_T) = \varepsilon_{T+1}$

Error variance V_A:

$$V_A = \text{Var}(\varepsilon_{T+1}) = \sigma_\varepsilon^2$$

Comparing $V_A - V_D$: $V_A = \sigma_\varepsilon^2$ and $V_D = \sigma_u^2 + \sigma_v^2$.

Since $(1 + \theta^2)\sigma_\varepsilon^2 = (1 + \rho_2^2)\sigma_u^2 + (1 + \rho_1^2)\sigma_v^2$, and $\theta^2 < 1$, $\rho_i^2 < 1$, we have $V_A \approx V_D$ (depending on the specific parameter values).

Partial Adjustment Model

The desired level y_t^* is related to x_t by: $y_t^* = \beta x_t$.
x_t is covariance-stationary. A fraction γ of the disequilibrium is removed through partial adjustment:

$$y_t - y_{t-1} = \gamma(y_t^* - y_{t-1}) + e_t, \quad e_t \sim N(0, \sigma_e^2),\ 0 < \gamma < 1$$

Compare with the adaptive expectations model:

$$y_t = \beta x_{t+1}^e + v_t, \quad v_t \sim N(0, \sigma_v^2)$$

$$x_{t+1}^e - x_t^e = \lambda(x_t - x_t^e), \quad 0 < \lambda < 1$$

Tasks: (a) Compare the models. (b) Show that OLS of y_t on y_{t-1} and x_t is consistent for the model (1)-(2), but inconsistent for (3)-(4).

(a) Comparing the two models:

Partial Adjustment Model: Substituting $y_t^* = \beta x_t$:

$$y_t - y_{t-1} = \gamma(\beta x_t - y_{t-1}) + e_t$$
$$y_t = \gamma\beta x_t + (1 - \gamma)y_{t-1} + e_t$$

Adaptive Expectations Model: Solving for x_{t+1}^e:

$$x_{t+1}^e = \lambda x_t + (1 - \lambda)x_t^e = \lambda \sum_{j=0}^{\infty}(1 - \lambda)^j x_{t-j}$$

Substituting into $y_t = \beta x_{t+1}^e + v_t$:

$$y_t = \beta\lambda \sum_{j=0}^{\infty}(1 - \lambda)^j x_{t-j} + v_t$$

Using the Koyck transformation: $y_t = \beta\lambda x_t + (1 - \lambda)y_{t-1} + v_t - (1 - \lambda)v_{t-1}$

Reduced-form representation of both models:

	Partial Adjustment	Adaptive Expectations
Form	$y_t = \alpha x_t + \phi y_{t-1} + e_t$	$y_t = \alpha x_t + \phi y_{t-1} + \eta_t$
x_t coeff.	$\alpha = \gamma\beta$	$\alpha = \beta\lambda$
y_{t-1} coeff.	$\phi = 1 - \gamma$	$\phi = 1 - \lambda$
Error	e_t (WN)	$\eta_t = v_t - (1-\lambda)v_{t-1}$ (MA(1)!)

Key Difference

▶ In partial adjustment: e_t is white noise ⇒ exogenous

▶ In adaptive expectations: η_t is MA(1) ⇒ correlated with y_{t-1}!

(b) OLS consistency:

Consider the regression: $y_t = \alpha x_t + \phi y_{t-1} + u_t$

Partial Adjustment Model: $u_t = e_t$ is white noise.

- ▶ $\text{Cov}(y_{t-1}, e_t) = 0$ (error is uncorrelated with past variables)
- ▶ $\text{Cov}(x_t, e_t) = 0$ (exogeneidade estrita)
- ▶ $\Rightarrow$ OLS is **consistent** ✓

Adaptive Expectations Model: $u_t = \eta_t = v_t - (1 - \lambda)v_{t-1}$

- ▶ y_{t-1} depends on v_{t-1} (through the recursive equation)
- ▶ $\text{Cov}(y_{t-1}, \eta_t) = \text{Cov}(y_{t-1}, -(1 - \lambda)v_{t-1}) \neq 0$
- ▶ $\Rightarrow$ OLS is **inconsistent** ×

Intuition

The regressor y_{t-1} is correlated with the error η_t due to the MA(1) component.

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(c) How to check which model is appropriate?

Strategy: Test for autocorrelation in the residuals.

1. **Estimate by OLS**: $y_t = \hat{\alpha} x_t + \hat{\phi} y_{t-1} + \hat{u}_t$

2. **Examine the residuals $\hat{u}_t$**:
   - ▶ If $\hat{u}_t$ is white noise $\Rightarrow$ Partial Adjustment Model
   - ▶ If $\hat{u}_t$ shows autocorrelation (MA(1)) $\Rightarrow$ Adaptive Expectations

3. **Formal tests**:
   - ▶ Durbin–Watson test (for AR(1) in the residuals)
   - ▶ Breusch–Godfrey test (for higher-order autocorrelation)
   - ▶ Ljung–Box test on the residual ACF

Caution

If the correct model is adaptive expectations, use appropriate estimation methods (GMM, NLLS, or MLE) instead of OLS.

Macroeconomic Model

$$
\begin{aligned}
(1): \quad & \pi_t = \lambda y_t + \pi_t^e + u_{1t}, \quad 0 < \lambda < 1 \\
(2): \quad & y_t = \gamma(i_{t-1} - \pi_t^e) + u_{2,t}, \quad -1 < \gamma < 0 \\
(3): \quad & \pi_t^e = \pi_{t-1} \\
(4): \quad & i_t = i^* + \rho(\pi_t - \pi^*), \quad \rho \geq 0
\end{aligned}
$$

where π_t = inflation, y_t = output gap, i_t = nominal interest rate,
$(u_{1t}, u_{2t})' \sim N(0, \Sigma)$.

Interpretation: (1) Phillips curve, (2) IS, (3) adaptive expectations, (4) Taylor rule.

Tasks:

(a) Show that π_t follows an $\mathrm{ARMA}(p, q)$. Relate the coefficients to the structural parameters.

(b) Consider $\rho = 0, 1, 2$. What can be said about inflation and monetary policy? Simulate 200 observations in MATLAB.

(c) Compute $\hat{\pi}_{t+j|t} = \mathbb{E}[\pi_{t+j}|\mathcal{F}_t]$ and $\mathrm{Var}(\pi_{t+j} - \hat{\pi}_{t+j|t})$.

(d) Let $\mathbf{x}_t = (\pi_t, y_t)'$. Find $\mathbb{D}(\mathbf{x}_t|\mathbf{x}_{t-1})$ and $\mathbb{E}(\pi_t|y_t, \mathbf{x}_{t-1})$.

(e) What would change if $\pi_t^e = \alpha\pi_{t-1} + (1 - \alpha)\pi_{t-1}^e$?

(a) Deriving the ARMA process for π_t:

Step 1: Substitute (3) and (4) into (2):

$$y_t = \gamma(i_{t-1} - \pi_{t-1}) + u_{2t}$$
$$= \gamma(i^* + \rho(\pi_{t-1} - \pi^*) - \pi_{t-1}) + u_{2t}$$
$$= \gamma(i^* - \rho\pi^*) + \gamma(\rho - 1)\pi_{t-1} + u_{2t}$$

Step 2: Substitute into (1):

$$\pi_t = \lambda[\gamma(i^* - \rho\pi^*) + \gamma(\rho - 1)\pi_{t-1} + u_{2t}] + \pi_{t-1} + u_{1t}$$
$$= c + [1 + \lambda\gamma(\rho - 1)]\pi_{t-1} + u_{1t} + \lambda u_{2t}$$

where $c = \lambda\gamma(i^* - \rho\pi^*)$ is a constant.

Result: π_t follows an **AR(1)** process:

$$\boxed{\pi_t = c + \phi \pi_{t-1} + \varepsilon_t}$$

where:

- $\phi = 1 + \lambda\gamma(\rho - 1)$
- $\varepsilon_t = u_{1t} + \lambda u_{2t}$ (a linear combination of the shocks)
- $\text{Var}(\varepsilon_t) = \sigma_{11}^2 + 2\lambda\sigma_{12} + \lambda^2\sigma_{22}^2$

Link to Structural Parameters

- $\phi < 1$ (stationarity) $\Leftrightarrow \lambda\gamma(\rho - 1) < 0$
- Since $\lambda > 0$ and $\gamma < 0$: $\phi < 1 \Leftrightarrow \rho > 1$
- This is the **Taylor principle**: $\rho > 1$ stabilizes inflation!

(b) Analysis for different values of ρ:

Suppose $\lambda = 0.5$, $\gamma = -0.5$. Then $\phi = 1 + 0.5 \cdot (-0.5) \cdot (\rho - 1) = 1 - 0.25(\rho - 1)$.

ρ	ϕ	Stationary?	Interpretation
0	1.25	No	Explosive (violates stability/unit-root condition)
1	1.00	No	Random walk
2	0.75	Yes	Stationary inflation

Economic Interpretation

- ▶ $\rho = 0$: the central bank does not respond to inflation $\Rightarrow$ explosive inflation
- ▶ $\rho = 1$: the central bank responds one-for-one $\Rightarrow$ constant real rate, inflation is I(1)
- ▶ $\rho > 1$: the central bank responds more than one-for-one $\Rightarrow$ stationary inflation

(c) Forecasts and forecast error variance:

For the AR(1) $\pi_t = c + \phi\pi_{t-1} + \varepsilon_t$ (assuming $|\phi| < 1$):

j-**step-ahead forecast**:

$$\hat{\pi}_{t+j|t} = \mu(1 - \phi^j) + \phi^j \pi_t$$

where $\mu = c/(1 - \phi)$ is the unconditional mean.

Forecast error variance:

$$\text{Var}(\pi_{t+j} - \hat{\pi}_{t+j|t}) = \sigma_\varepsilon^2 \sum_{i=0}^{j-1} \phi^{2i} = \sigma_\varepsilon^2 \cdot \frac{1 - \phi^{2j}}{1 - \phi^2}$$

Limit: As $j \to \infty$:

$$\lim_{j \to \infty} \text{Var}(\pi_{t+j} - \hat{\pi}_{t+j|t}) = \frac{\sigma_\varepsilon^2}{1 - \phi^2} = \text{Var}(\pi_t)$$

(d) Conditional distribution of $\mathbf{x}_t = (\pi_t, y_t)'$:

From the model equations:

$$\pi_t = c_\pi + \phi\pi_{t-1} + \varepsilon_t$$
$$y_t = c_y + \gamma(\rho - 1)\pi_{t-1} + u_{2t}$$

Therefore, $\mathbf{x}_t|\mathbf{x}_{t-1} \sim N(\boldsymbol{\mu}_t, \boldsymbol{\Sigma})$, where:

$$\boldsymbol{\mu}_t = \begin{pmatrix} c_\pi + \phi\pi_{t-1} \\ c_y + \gamma(\rho - 1)\pi_{t-1} \end{pmatrix}, \quad \boldsymbol{\Sigma} = \begin{pmatrix} \sigma_\varepsilon^2 & \lambda\sigma_{22}^2 + \sigma_{12} \\ \lambda\sigma_{22}^2 + \sigma_{12} & \sigma_{22}^2 \end{pmatrix}$$

Conditional expectation $\mathbb{E}(\pi_t|y_t, \mathbf{x}_{t-1})$:

Using the conditional normal formula:

$$\mathbb{E}(\pi_t|y_t, \mathbf{x}_{t-1}) = \mathbb{E}(\pi_t|\mathbf{x}_{t-1}) + \frac{\text{Cov}(\pi_t, y_t)}{\text{Var}(y_t)}[y_t - \mathbb{E}(y_t|\mathbf{x}_{t-1})]$$

(e) Generalized adaptive expectations: $\pi_t^e = \alpha\pi_{t-1} + (1-\alpha)\pi_{t-1}^e$

Solving recursively:

$$\pi_t^e = \alpha\sum_{j=0}^{\infty}(1-\alpha)^j\pi_{t-1-j}$$

Effect on the model: inflation now depends on the entire past history, not only on π_{t-1}.

Using the Koyck transformation:

$$\pi_t^e = \alpha\pi_{t-1} + (1-\alpha)\pi_{t-1}^e$$

The process for π_t becomes an **ARMA(2,1)** instead of AR(1):

$$\pi_t = c + \phi_1\pi_{t-1} + \phi_2\pi_{t-2} + \varepsilon_t + \theta\varepsilon_{t-1}$$

where the coefficients depend on $\alpha, \lambda, \gamma, \rho$ in a more complex way.

Exercise Session: Summary

Key Skills Practiced:

1. Checking stationarity conditions (roots outside unit circle)
2. Computing autocovariances and ACF for ARMA processes
3. Converting between AR and MA representations
4. Identifying common factors and parameter redundancy
5. Working with vector processes

For Next Session

Review these exercises and attempt additional problems from the appendix for extra practice.

Appendix

Detailed Material and Extended Examples

Checking Local Identification: The Information Matrix

Idea: Use the Kullback-Leibler divergence between distributions:

$$\Delta_{KL}(\theta \mid \theta_0) = - \int \log \left(\frac{\mathbb{D}(\mathbf{Y}^T \mid \theta)}{\mathbb{D}(\mathbf{Y}^T \mid \theta_0)} \right) \mathbb{D}(\mathbf{Y}^T \mid \theta_0) \, d\mathbf{Y}^T$$

By Jensen's inequality: $\Delta_{KL}(\theta \mid \theta_0) \geq 0$, with equality iff distributions are equal.

Second-order condition: The Hessian at θ_0 is the **Fisher Information Matrix**:

$$\mathcal{I}(\theta_0) = -\mathbb{E}_{\theta_0} \left[\nabla_{\theta^2} \log \mathbb{D}(\mathbf{Y}^T \mid \theta_0) \right]$$

Theorem 17 (Local Identification Condition)

*If $\mathcal{I}(\theta_0)$ is **positive definite** (full rank), then θ_0 is locally identified.*

Figure 1: Sensitivity of ϕ_1 to Taylor Rule coefficient ρ

[Figure: param_combination.png from metrics bible]

Under the simple monetary policy model with $\lambda = 0.5$, $\gamma = -0.5$:

▶ Stability requires $1 < \rho < 9$

▶ Taylor principle ($\rho > 1$) ensures stationary inflation

← Back

AR(1) Inflation: ACF Patterns

Figure 2: ACF under different Taylor Rule calibrations

[Figure: acf_pi_1.png, acf_pi_2.png from metrics bible]

Key observations:

- ▶ $\rho \in \{2, 3, 4\}$: Persistent, non-oscillatory decay
- ▶ $\rho \in \{6, 7, 8\}$: Faster decay, potentially oscillating

A3: Detailed Proofs

Proof: MA(q) Stationarity

Claim: All finite-order MA(q) processes are covariance-stationary.

Proof:

Mean:

$$\mathbb{E}[y_t] = \mathbb{E}\left[u_t + \sum_{j=1}^{q} \theta_j u_{t-j} + \theta_0\right] = \theta_0$$

Autocovariance:

$$\Gamma_k = \mathbb{E}[(y_t - \mu)(y_{t-k} - \mu)]$$
$$= \mathbb{E}\left[\left(\sum_{j=0}^{q} \theta_j u_{t-j}\right)\left(\sum_{m=0}^{q} \theta_m u_{t-k-m}\right)\right]$$

Since $\mathbb{E}[u_s u_r] = \sigma^2 \mathbf{1}_{s=r}$:

$$\Gamma_k = \begin{cases} \left(\theta_k + \sum_{j=1}^{q-k} \theta_j \theta_{j+k}\right)\sigma^2 & k \leq q \\ 0 & k > q \end{cases}$$

Proof: Invertibility Condition for MA(q)

Claim: MA(q) is invertible iff roots of $\theta_q(z) = 0$ lie outside unit circle.

Proof: Start with $y_t = \theta_0 + \theta_q(L)u_t$. We need:

$$u_t = \theta_q(L)^{-1}(y_t - \theta_0)$$

Factor the polynomial: $\theta_q(L) = (1 - \lambda_1 L)(1 - \lambda_2 L) \cdots (1 - \lambda_q L)$
where λ_i are reciprocals of the roots.

For each factor: $(1 - \lambda_i L)^{-1} = \sum_{k=0}^{\infty} (\lambda_i L)^k$
This converges iff $|\lambda_i| < 1$, i.e., roots $|1/\lambda_i| > 1$ (outside UC).

If convergent:

$$\theta_q(L)^{-1} = \sum_{j=0}^{\infty} \psi_j L^j \quad \Rightarrow \quad u_t = \sum_{j=0}^{\infty} \psi_j (y_{t-j} - \theta_0) \in \text{span}\{y_t, y_{t-1}, \ldots\}$$

$\square$

MA(∞) Representation of AR(p)

Claim: A stationary AR(p) can be written as MA(∞).

Proof (AR(1) case): Let $y_t = \phi_1 y_{t-1} + u_t$ with $|\phi_1| < 1$.
Iterating backward (assuming process started infinitely far in the past):

$$
\begin{aligned}
y_t &= \phi_1 y_{t-1} + u_t \\
&= \phi_1(\phi_1 y_{t-2} + u_{t-1}) + u_t \\
&= \cdots \\
&= \sum_{i=0}^{\infty} \phi_1^i u_{t-i} + \lim_{n \to \infty} \phi_1^n y_{t-n}
\end{aligned}
$$

Since $|\phi_1| < 1$: $\phi_1^n \to 0$ and $\sum |\phi_1|^i < \infty$.
Therefore: $y_t = \sum_{i=0}^{\infty} \psi_i u_{t-i}$ with $\psi_i = \phi_1^i$.
General case: Use companion form and iterate. $\square$

A4: Extended Examples

Example: Observational Equivalence in MA(2)

Consider MA(2): $y_t = u_t + \theta_1 u_{t-1} + \theta_2 u_{t-2}$

Factor: $\theta_2(z) = 1 + \theta_1 z + \theta_2 z^2 = (1 - \lambda_1 z)(1 - \lambda_2 z)$

where $\theta_1 = -(\lambda_1 + \lambda_2)$, $\theta_2 = \lambda_1 \lambda_2$

Invertible representation: Roots outside UC $\Rightarrow |\lambda_i| < 1$

Non-invertible alternative: "Flip" each root:

$$\lambda_i^* = 1/\lambda_i \quad \Rightarrow \quad \theta_1^* = \theta_1/\theta_2, \quad \theta_2^* = 1/\theta_2$$

Match ACGF: Scale variance by $\sigma^{*2} = \sigma^2 \theta_2^2$

Result: These have **identical** autocovariance structure but different impulse responses!

Example: NK Model Invertibility (Wolf, 2022)

Three-equation NK model with demand, supply, and monetary shocks:

$$\begin{pmatrix} y_t \\ \pi_t \\ i_t \end{pmatrix} = \frac{1}{1 + \phi_\pi \kappa} \begin{pmatrix} \sigma^d & \phi_\pi \sigma^s & -\sigma^m \\ \kappa \sigma^d & -\sigma^s & -\kappa \sigma^m \\ \phi_\pi \kappa \sigma^d & -\phi_\pi \sigma^s & \sigma^m \end{pmatrix} \begin{pmatrix} \varepsilon_t^d \\ \varepsilon_t^s \\ \varepsilon_t^m \end{pmatrix}$$

With all 3 observables: Impact matrix Θ is invertible for standard parameters.

With only y_t and π_t:

- ▶ Can back out ε_t^s (supply shock)
- ▶ **Cannot** disentangle ε_t^d and ε_t^m!

Lesson

Even in well-specified models, fewer observables than shocks $\Rightarrow$ non-invertibility!

A5: Reference Material

Key Formulas: Quick Reference

AR(p):
- Process: $\Phi_p(L)y_t = \phi_0 + u_t$
- Stationarity: Roots of $\Phi_p(z) = 0$ outside UC
- Mean: $\mu = \phi_0/(1 - \sum \phi_i)$
- Yule-Walker: $\gamma_k = \sum_{i=1}^{p} \phi_i \gamma_{k-i}$

MA(q):
- Process: $y_t = \theta_0 + \Theta_q(L)u_t$
- Always stationary
- Invertibility: Roots of $\Theta_q(z) = 0$ outside UC
- $\gamma_k = 0$ for $k > q$ (cutoff)

ARMA(p, q):
- Process: $\Phi_p(L)y_t = \phi_0 + \Theta_q(L)u_t$
- ACGF: $\Gamma(z) = \frac{\Theta_q(z)\Theta_q(z^{-1})}{\Phi_p(z)\Phi_p(z^{-1})}\sigma^2$

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