HW 4: Fitting Accident Data with regression line

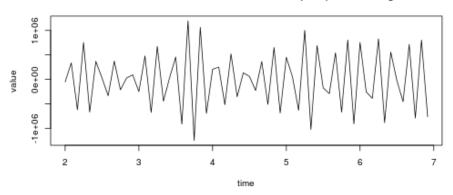
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1. Read in acci.txt file from the course web site. Plot the time series. Take difference with lag 12. Plot the differenced data. Call it D2.

```
acci <- read.table("acci.txt", header = TRUE)</pre>
D1 <- ts(acci, frequency = 12)
D2 \leftarrow diff(D1, d = 12)
##
   Augmented Dickey-Fuller Test
##
## data: D2
## Dickey-Fuller = -19.95, Lag order = 3, p-value = 0.01
## alternative hypothesis: stationary
pp.test(D2)
   Phillips-Perron Unit Root Test
##
## data: D2
## Dickey-Fuller Z(alpha) = -105.5, Truncation lag parameter = 3,
## p-value = 0.01
## alternative hypothesis: stationary
kpss.test(D2)
##
##
   KPSS Test for Level Stationarity
## data: D2
## KPSS Level = 0.0483, Truncation lag parameter = 1, p-value = 0.1
layout(matrix(c(1, 1, 2, 3), 2, 2, byrow = TRUE))
```

ACCI Time Series with Year-over-Year (YOY) differencing



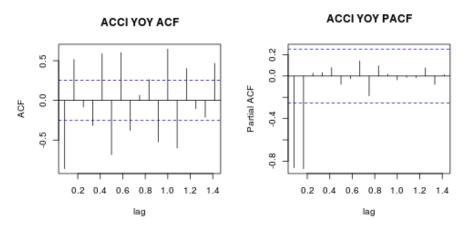


Figure 1: plot of chunk q1

2. Fit D2 with linear trend using OLS. Is the slope significant? How did you determine? Can the output from the screen be trusted? Based on the RSE, adjusted ρ^2 and p-value we fail to reject H_0 and cannot regard the slope of the OLS fit of the year-over-year differences as significant. This indicates that the values obtained in D2 are dependent upon time and should probably not be trusted. When we plot the forecast of the residuals we see a repeating pattern in the data and the forecast as well.

```
time <- 1:length(D2)</pre>
(err_mod <- summary(lm_D2 <- lm(D2 ~ time)))</pre>
##
## Call:
## lm(formula = D2 ~ time)
## Residuals:
##
        Min
                  10
                       Median
                                     30
                                              Max
## -1245097 -504589
                         21005
                                 486720 1194118
##
## Coefficients:
##
               Estimate Std. Error t value Pr(>|t|)
## (Intercept)
                  14854
                             159786
                                        0.09
                                                 0.93
## time
                    -663
                               4556
                                       -0.15
                                                 0.88
##
## Residual standard error: 611000 on 58 degrees of freedom
## Multiple R-squared: 0.000365,
                                     Adjusted R-squared:
## F-statistic: 0.0212 on 1 and 58 DF, p-value: 0.885
resid <- lm_D2$residuals
(arima_mdl \leftarrow auto.arima(ts(resid, start = c(2, 1), freq = 12)))
## Series: ts(resid, start = c(2, 1), freq = 12)
## ARIMA(4,0,0)(0,1,1)[12] with drift
##
## Coefficients:
##
                                                    drift
            ar1
                     ar2
                             ar3
                                      ar4
                                             sma1
         -2.965
                 -3.879
                          -2.629
                                  -0.779
                                          -0.397
                                                   659.77
##
## s.e.
          0.087
                  0.204
                           0.202
                                   0.083
                                            0.224
                                                    28.18
## sigma^2 estimated as 1.06e+09: log likelihood=-425.1
## AIC=864.2
               AICc=867
                           BIC=877.4
adf.test(resid)
```

```
## alternative hypothesis: stationary
pp.test(resid)
##
##
   Phillips-Perron Unit Root Test
##
## data: resid
## Dickey-Fuller Z(alpha) = -105.5, Truncation lag parameter = 3,
## p-value = 0.01
## alternative hypothesis: stationary
layout(matrix(c(1, 2, 3), 3, 1, byrow = TRUE))
plot(D2)
abline(lm_D2)
plot(x = 1:length(resid), y = resid, type = "1", main = "Fits vs. Residuals",
    ylab = "residuals")
plot(forecast(arima_mdl))
layout(matrix(c(1, 2, 3, 4), 2, 2, byrow = TRUE))
plot(lm_D2)
3. Fit residuals from (2) with seasonal ARIMA with d=0, D=0.
(i.e. ARIMA(p,0,q)x(P,0,Q)12). Is the seasonal part necessary?
Based on the results above, the standard error for the seasonal component is
significant and therefore should be included in the forecast.
(arima00 \leftarrow auto.arima(ts(lm_D2$residuals, start = c(2, 1), freq = 12), d = 0,
    D = 0))
## Series: ts(lm_D2\$residuals, start = c(2, 1), freq = 12)
## ARIMA(3,0,0)(1,0,0)[12] with zero mean
##
## Coefficients:
```

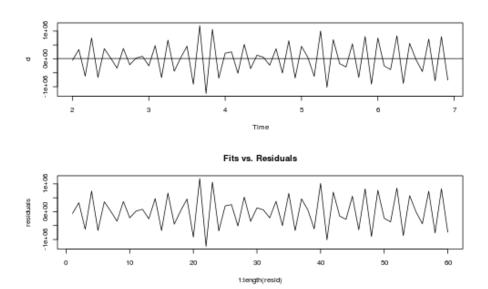
##

##

data: resid

Augmented Dickey-Fuller Test

Dickey-Fuller = -19.95, Lag order = 3, p-value = 0.01



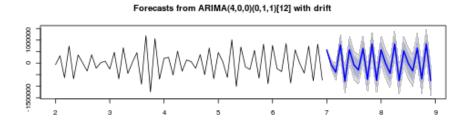


Figure 2: plot of chunk q2

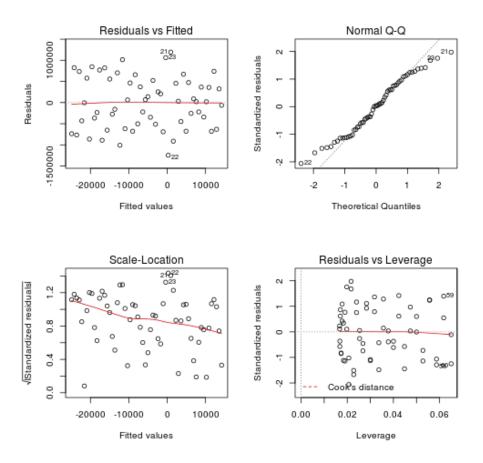


Figure 3: plot of chunk q2

```
##
            ar1
                                    sar1
##
                                  0.782
         -2.168
                  -1.869
                          -0.623
                                  0.081
          0.096
                           0.100
##
##
## sigma^2 estimated as 4.82e+09:
                                    log likelihood=-762.4
## AIC=1535
              AICc=1536
                           BIC=1545
```

4. Using the best model from #3, predict twelve months ahead in D2.

```
plot(forecast(arima00, h = 12))
```

Forecasts from ARIMA(3,0,0)(1,0,0)[12] with zero mean

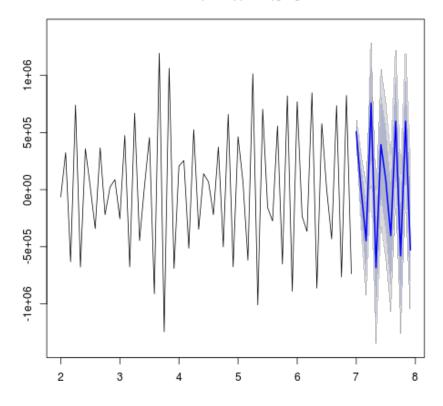


Figure 4: plot of chunk q4

5. Using the prediction from #4, predict twelve months ahead in D1 (original TS).

```
(q5D1 <- Arima(D1, order = c(3, 0, 0), seasonal = list(order = c(1, 0, 0), period = 12)))
## Series: D1
## ARIMA(3,0,0)(1,0,0)[12] with non-zero mean
##
## Coefficients:
## ar1 ar2 ar3 sar1 intercept
## 0.644 0.109 0.050 0.876 9317.0
## s.e. 0.124 0.140 0.126 0.049 915.4
##
## sigma^2 estimated as 121442: log likelihood=-532.9
## AIC=1078 AICc=1079 BIC=1092</pre>

q5 <- forecast(q5D1)
plot(q5)</pre>
```

6. Using your model from In-class Ex2-#2, (ARIMA(p,1,q)x(P,1,Q)12 model), predict twelve months ahead in D1.

```
(q6D1 <- auto.arima(D1, d = 1, D = 1))

## Series: D1
## ARIMA(0,1,1)(0,1,1)[12]
##
## Coefficients:
## ma1 sma1
## -0.426 -0.558
## s.e. 0.123 0.179
##
## sigma^2 estimated as 99480: log likelihood=-425.5
## AIC=857.1 AICc=857.5 BIC=863.3</pre>
q6 <- forecast(q6D1)
plot(q6)
```

Forecasts from ARIMA(3,0,0)(1,0,0)[12] with non-zero mean

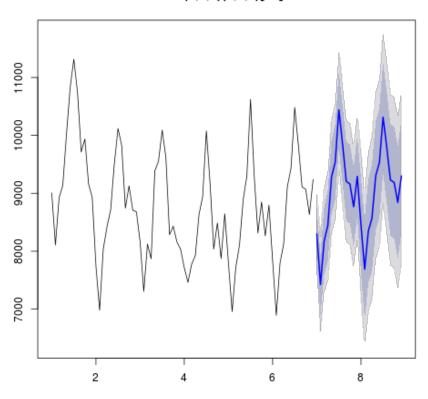


Figure 5: plot of chunk q5

Forecasts from ARIMA(0,1,1)(0,1,1)[12]

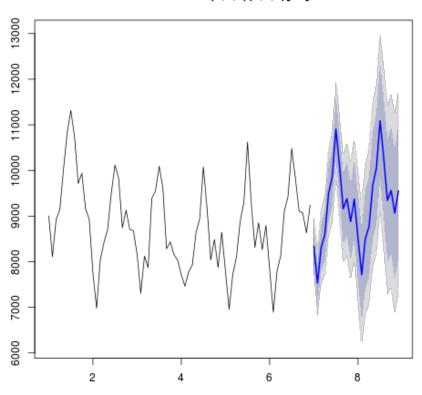


Figure 6: plot of chunk q6

7. Using your model from In-class Ex2-#3, (ARIMA(p,0,q)x(P,1,Q)12 model), predict twelve months ahead in D1.

```
(q7D1 \leftarrow auto.arima(D1, d = 0, D = 1))
## Series: D1
## ARIMA(2,0,0)(1,1,0)[12] with drift
## Coefficients:
##
          ar1 ar2
                      sar1
                              drift
##
        0.592 0.258 -0.350 -13.71
## s.e. 0.138 0.143 0.142
                               18.56
## sigma^2 estimated as 138798: log likelihood=-348
## AIC=705.9
              AICc=707
                        BIC=716.4
q7 <- forecast(q7D1)
plot(q7)
```

8. Compare your prediction from #5, #6, and #7. Plot the one-month predictions on the sampe plot. Which one do you like the best? Based on the results from questions 5, 6 & 7 I believe the model fit in question #7 to be the best based on its AIC, AICc & BIC statistics being the smallest for the three questions and for the standard errors it produces being among the smallest of the group.

```
plot(D1, xlim = c(1, 8))
points(x = seq(7, 8, length.out = 12), y = q7$mean[1:12])
points(x = seq(7, 8, length.out = 12), y = q6$mean[1:12], col = "red")
points(x = seq(7, 8, length.out = 12), y = q5$mean[1:12], col = "blue")
```

Forecasts from ARIMA(2,0,0)(1,1,0)[12] with drift

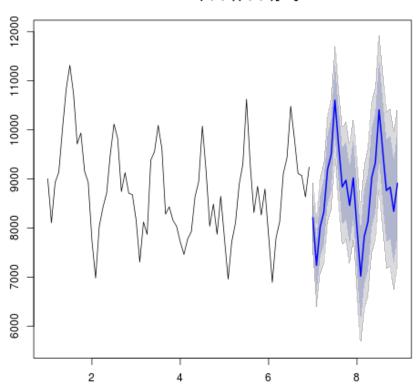


Figure 7: plot of chunk q7

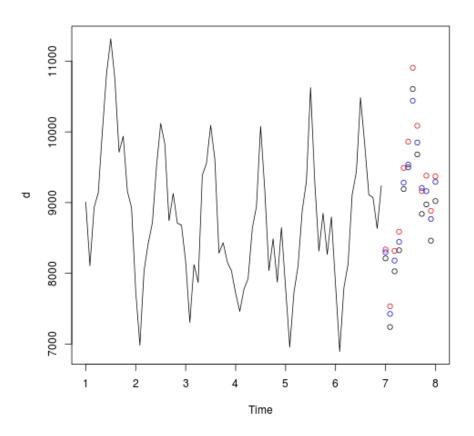


Figure 8: plot of chunk q8 $\,$