Bayesian Statistics Homework 1

Gelman et al., *A Weakly Informative Default Prior Distribution For Logistic and Other Regression Models* (2008)
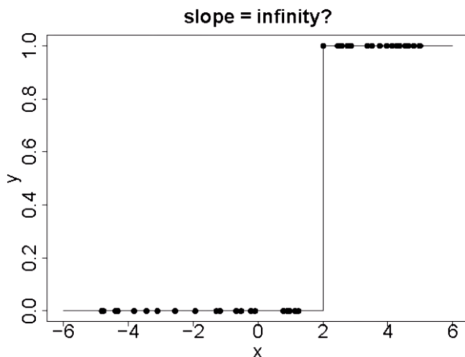
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Logistic regression: we have some problems...

In some cases, Maximum Likelihood estimates of the coefficients can lead to unstable results.

slope = infinity?

The need of (weak) prior information

Adding a prior distribution over regression's coefficients can regularize the estimates.

Which kind of prior do we use?

- Non-informative prior $\longrightarrow$ low regularization effect
- Fully informative prior $\longrightarrow$ application-specific

The aim is to provide *weakly informative priors*

- minimal information, valid for any context
- able to give regularized coefficient estimates

General assumptions for weakly-informative priors

Which kind of minimal information is valid for any generic model?

"For logistic regression, a change of 5 in the logistic scale moves a probability from 0.01 to 0.5, or from 0.5 to 0.99"

These kind of situations are rare to encounter, thus it is expected that $\pi(\beta_i)$ is lower for values outside $[-5, 5]$.

General assumptions for weakly-informative priors

In order to have a default prior that could be used in many different contexts, it has to be defined over a common interpretable scale.

It is required to *standardize* the input variables:

- For binary varibales, inputs must have 0 mean and they have to differ 1 in their lower and upper values.
- Other variables must have 0 mean and standard deviation of 0.5.

General assumptions for weakly-informative priors

Which probability distribution can be used for this prior?

- prior independence: $\pi(\beta) = \pi(\beta_0) \cdot \pi(\beta_1) \cdot ... \cdot \pi(\beta_J)$
- t-Student family: flat-tailed distributions, easy and stable computations.

Typical choices can be:

- for the coefficient terms $\beta_1, ..., \beta_J$:
  $\pi(\beta_i) = t_{\nu=1}(\theta = 0, s = 2.5)$

- for the costant term β_0:
  $\pi(\beta_0) = t_{\nu=1}(\theta = 0, s = 10)$

A tool to use weakly-informative priors in R

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bayesglm
```

- **Goal**: get β_{MAP} and $V_{\beta_{MAP}}$ estimates.
- **Computation**: Considering the prior distribution for each coefficient as:

$$\pi(\beta_i) = \int d\sigma_i \pi(\beta_i|\sigma_i)\pi(\sigma_i) = \int d\sigma_i \mathcal{N}(\theta_i, \sigma_i^2)\text{inv-}\chi^2(\nu_i^2, s_i^2)$$

estimates of β_{MAP} and $V_{\beta_{MAP}}$ are obtained in a iterative way:

- for a fixed σ, approximate the priors as $\pi(\beta_i) \approx \mathcal{N}(\theta_i, \sigma_i^2)$ and get $\hat{\beta}$ and $\hat{V}_{\beta_{MAP}}$ by using weighted least squares.
- determine the expected value of the log-posterior density $\log p(\beta, \sigma|X, y)$ and maximize it with respect to σ in order to get to get the estimate $\hat{\sigma}$.
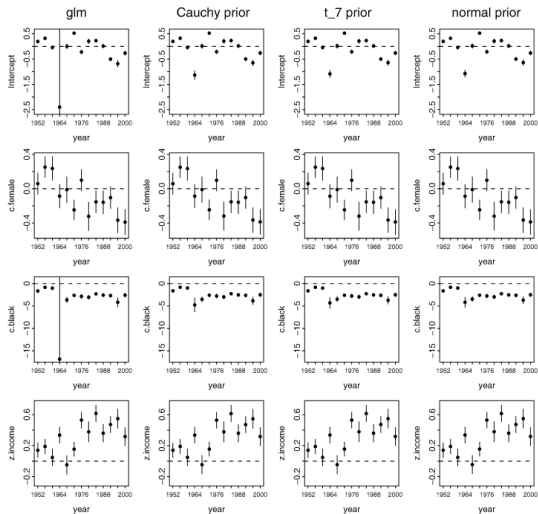
FIG. 2. *The left column shows the estimated coefficients (± 1 standard error) for a logistic regression predicting the probability of a Republican vote for president given sex, race, and income, as fit separately to data from the National Election Study for each election 1952 through 2000. [The binary inputs* `female` *and* `black` *have been centered to have means of zero, and the numerical variable* `income` *(originally on a 1–5 scale) has been centered and then rescaled by dividing by two standard deviations.]*

Dose, x_i (log g/ml)	Number of animals, n_i	Number of deaths, y_i
−0.86	5	0
−0.30	5	1
−0.05	5	3
0.73	5	5

```
# from glm:
            coef.est coef.se
(Intercept) -0.1      0.7
z.x         10.2      6.4
  n = 4, k = 2
  residual deviance = 0.1, null deviance = 15.8 (difference = 15.7)

# from bayesglm (Cauchy priors, scale 10 for const and 2.5 for other coef):
            coef.est coef.se
(Intercept) -0.2      0.6
z.x          5.4      2.2
  n = 4, k = 2
  residual deviance = 1.1, null deviance = 15.8 (difference = 14.7)
```
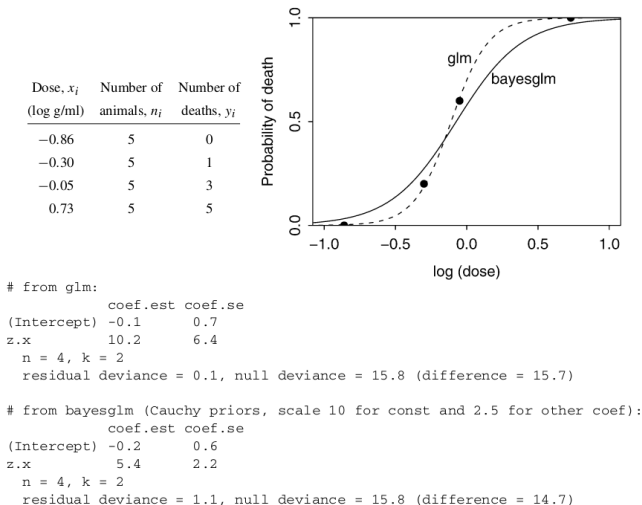
FIG. 3. *Data from a bioassay experiment, from Racine et al.* (1986), *and estimates from classical maximum likelihood and Bayesian logistic regression with the recommended default prior distribution. In addition to graphing the fitted curves (at top right), we show raw computer output to illustrate how our approach would be used in routine practice.*

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```
# from glm:
                    coef.est coef.sd                   coef.est coef.sd
(Intercept)           0.07    1.41   h39b.W1            -0.10 0.03
age.W1                0.02    0.02   pcs.W1             -0.01 0.01
mcs37.W1             -0.01    0.32   nonhaartcombo.W1  -20.99 888.74
unstabl.W1           -0.09    0.37   b05.W1             -0.07 0.12
ethnic.W3            -0.14    0.23   h39b.W2             0.02 0.03
age.W2                0.02    0.02   pcs.W2             -0.01 0.02
mcs37.W2              0.26    0.31   haart.W2            1.80 0.30
nonhaartcombo.W2      1.33    0.44   unstabl.W2          0.27 0.42
b05.W2                0.03    0.12   h39b.W3             0.00 0.03
age.W3               -0.01    0.02   pcs.W3              0.01 0.01
mcs37.W3             -0.04    0.32   haart.W3            0.60 0.31
nonhaartcombo.W3      0.44    0.42   unstabl.W3         -0.92 0.40
b05.W3               -0.11    0.11


# from bayesglm (Cauchy priors, scale 10 for const
                    and 2.5 for other coefs):
                    coef.est coef.sd                   coef.est coef.sd
(Intercept)          -0.84    1.15   h39b.W1            -0.08 0.03
age.W1                0.01    0.02   pcs.W1             -0.01 0.01
mcs37.W1             -0.10    0.31   nonhaartcombo.W1   -6.74 1.22
unstabl.W1           -0.06    0.36   b05.W1              0.02 0.12
ethnic.W3             0.18    0.21   h39b.W2             0.01 0.03
age.W2                0.03    0.02   pcs.W2             -0.02 0.02
mcs37.W2              0.19    0.31   haart.W2            1.50 0.29
nonhaartcombo.W2      0.81    0.42   unstabl.W2          0.29 0.41
b05.W2                0.11    0.12   h39b.W3            -0.01 0.03
age.W3               -0.02    0.02   pcs.W3              0.01 0.01
mcs37.W3              0.05    0.32   haart.W3            1.02 0.29
nonhaartcombo.W3      0.64    0.40   unstabl.W3         -0.52 0.39
b05.W3               -0.15    0.13
```

FIG. 4. *A logistic regression fit for missing-data imputation using maximum likelihood (top) and Bayesian inference with default prior distribution (bottom). The classical fit resulted in an error message indicating separation; in contrast, the Bayes fit (using independent Cauchy prior distributions with mean 0 and scale 10 for the intercept and 2.5 for the other coefficients) produced stable estimates. We would not usually summarize results using this sort of table, however, this gives a sense of how the fitted models look in routine data analysis.*
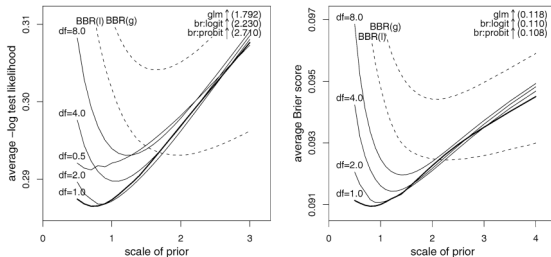
FIG. 6.   *Mean logarithmic score (left plot) and Brier score (right plot), in fivefold cross-validation averaging over the data sets in the UCI corpus, for different independent prior distributions for logistic regression coefficients. Higher value on the y axis indicates a larger error. Each line represents a different degrees-of-freedom parameter for the Student-t prior family. BBR(l) indicates the Laplace prior with the BBR algorithm of Genkin, Lewis, and Madigan (2007), and BBR(g) represents the Gaussian prior. The Cauchy prior distribution with scale 0.75 performs best, while the performances of `glm` and `brglm` (shown in the upper-right corner) are so bad that we could not capture them on our scale. The scale axis corresponds to the square root of variance for the normal and the Laplace distributions.*