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## Ex1

dataset <- read.table('wine.txt', header = T) # se serve aggiungi il nome
delle colonne col.names=c(' ',..)
dataset

load("mcshapiro.test.RData")

group1 <- factor(dataset$color) # qui species è il primo grouping
group2 <- factor(dataset$region) # qui gender è il secondo grouping

x <- dataset[,1] # salvo la variabile numerica in x

g <- length(levels(group1))
b <- length(levels(group2))
n <- length(x) / (g*b)

# Check hyp
group_12 <- group1
levels(group_12) <- c('R_P', 'R_T', 'R_V', 'W_P', 'W_T', 'W_V') #variable which
divides the g*b groups (only for plot)
group_12[group1=='red' & group2=='Piemonte'] <- 'R_P'
group_12[group1=='red' & group2=='Toscana'] <- 'R_T'
group_12[group1=='red' & group2=='Veneto'] <- 'R_V'
group_12[group1=='white' & group2=='Piemonte'] <- 'W_P'
group_12[group1=='white' & group2=='Toscana'] <- 'W_T'
group_12[group1=='white' & group2=='Veneto'] <- 'W_V'

# 1) normality (multivariate) in each group (4 test)
load("mcshapiro.test.RData")
Ps <- c(shapiro.test(x[ group_12==levels(group_12)[1]])$p,
        shapiro.test(x[ group_12==levels(group_12)[2]])$p,
        shapiro.test(x[ group_12==levels(group_12)[3]])$p,
        shapiro.test(x[ group_12==levels(group_12)[4]])$p,
        shapiro.test(x[ group_12==levels(group_12)[5]])$p,
        shapiro.test(x[ group_12==levels(group_12)[6]])$p)

Ps
# 0.2369077 0.4612391 0.9702243 0.6525198 0.5279147 0.6612603
# ok

# 2) homogeneity of the covariance (qualitatively)
S1 <- var(x[ group_12==levels(group_12)[1]])
S2 <- var(x[ group_12==levels(group_12)[2]])
S3 <- var(x[ group_12==levels(group_12)[3]])
S4 <- var(x[ group_12==levels(group_12)[4]])
S5 <- var(x[ group_12==levels(group_12)[5]])
S6 <- var(x[ group_12==levels(group_12)[6]])

# test of homogeneity of variances (1D)
# H0: sigma.1 = .. = sigma.g
# H1: there exist i,j s.t. sigma.i!=sigma.j

bartlett.test(x, group_12)

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# ok

### Model with interaction (complete model):
###  $x_{ijk} = \mu + \tau_i + \beta_j + \gamma_{ij} + \epsilon_{ijk}$ ;  $\epsilon_{ijk} \sim N(0, \sigma^2)$ ,
###  $i=1,2$  (effect Fact1),  $j=1,2$  (effect Fact2)

fit.aov2.int <- aov(x ~ group1 + group2 + group1:group2) # aov(y ~ factor1 +
factor2 + interaction)
summary.aov(fit.aov2.int)

# b)
# reduce
fit.aov2.ad <- aov(x ~ group1 + group2)
summary.aov(fit.aov2.ad)

fit.aov1 <- aov(x ~ group1)
summary.aov(fit.aov1)

# Estimate the great mean  $\mu$ :
m <- mean(x)

# Estimate  $\tau_i$ ,  $\beta_j$ :
tau1 <- mean(x[group1=='red']) - m
tau2 <- mean(x[group1=='white']) - m # devo avere g righe

# variance:
DF <- fit.aov1$df # fit è il modello finale

Spooled <- sum(fit.aov1$res^2)/DF
Spooled # estimate for the variance

# BonfCI
DF <- fit.aov1$df # fit è il modello finale

Spooled <- sum(fit.aov1$res^2)/DF # SS_res= Spooled * DF= sum(eps_hat_i) =
sum(fit$res^2) ~  $\sigma^2 \chi^2(n-g)$ 
means <- tapply(x, group1, mean)
names(means) <- levels(group1)
means

n_group = length(x)/2 # qui sto assumendo che i livelli di ogni gruppo
abbiano stessa numerosità
alpha <- 0.01
k <- 2+1 # g Conf Int for the means and 1 for the variance
# se voglio solo Bonf per le medie o per var ricordati di cambiare il k !

# l'intervallo per le g medie lo trovo dividendo una normale per la radice
della Chi/DF -> t

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BF_means      <- rbind(cbind(means - sqrt(Spooled / n_group) * qt(1 - alpha /
(2*k), DF),
                           means + sqrt(Spooled / n_group) * qt(1 - alpha /
(2*k), DF)))

# l'intervallo per sigma^2 lo trovo dalla distribuzione di SS_res ~ sigma^2
Chi(n-g)
# => chi^2_alpha/2 < SS_res / sigma^2 < chi^2_(1-alpha/2) con SS_res= Spooled
* DF
BF_var = c(Spooled * DF / qchisq(1 - alpha / (2*k), DF),
           Spooled * DF / qchisq(alpha / (2*k), DF))

BF = c(BF_means, BF_var)
BF

BF_means
BF_var

```