

Conditional PDF, given an event

$$p_X(x) = \mathbf{P}(X = x)$$

$$f_X(x) \cdot \delta \approx \mathbf{P}(x \leq X \leq x + \delta)$$

$$p_{X|A}(x) = \mathbf{P}(X = x \mid A)$$

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$$\mathbf{P}(X \in B) = \sum_{x \in B} p_X(x)$$

$$\mathbf{P}(X \in B) = \int_B f_X(x) dx$$

$$\mathbf{P}(X \in B \mid A) = \sum_{x \in B} p_{X|A}(x)$$

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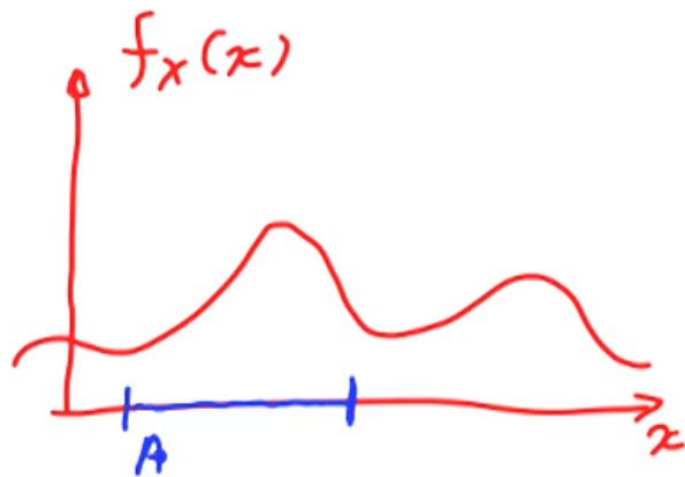
$$\mathbf{P}(X \in B \mid A) = \sum_{x \in B} p_{X|A}(x)$$

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$$\sum_x p_{X|A}(x) = 1$$

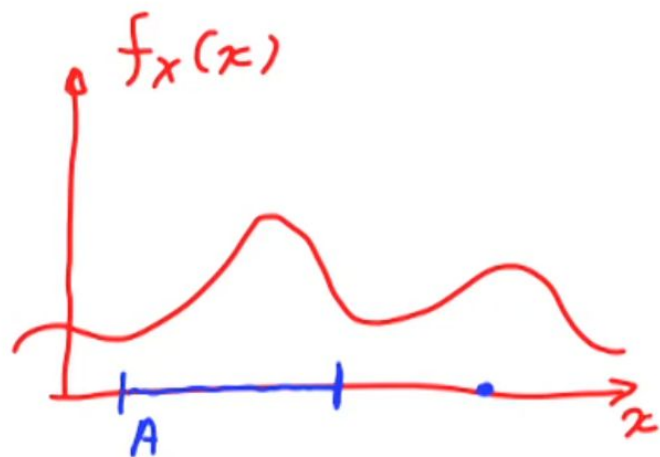
$$\int f_{X|A}(x) dx = 1$$

Conditional PDF of X , given that $X \in A$



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$$\mathbf{P}(x \leq X \leq x + \delta \mid X \in A) \approx f_{X|X \in A}(x) \cdot \delta$$

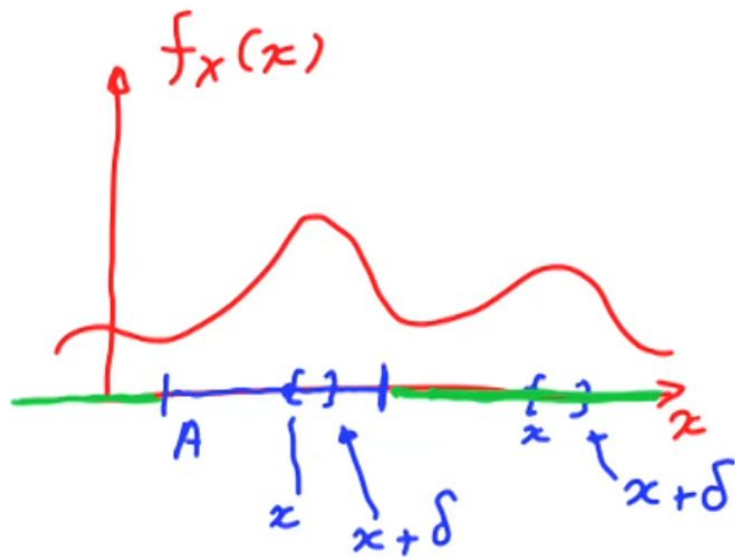


Conditional PDF of X , given that $X \in A$

$$P(x \leq X \leq x + \delta \mid X \in A) \approx f_{X|X \in A}(x) \cdot \delta$$

$$= \frac{P(x \leq X \leq x + \delta, X \in A)}{P(A)}$$

$$= \frac{P(x \leq X \leq x + \delta)}{P(A)} \approx \frac{f_X(x) \delta}{P(A)}$$



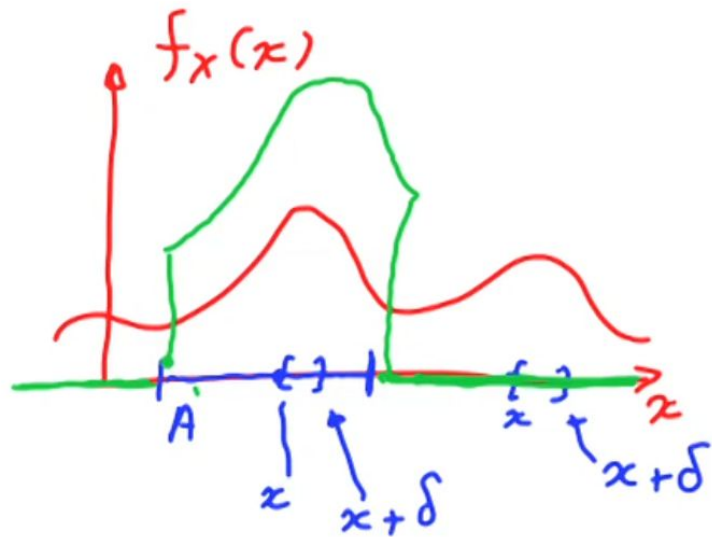
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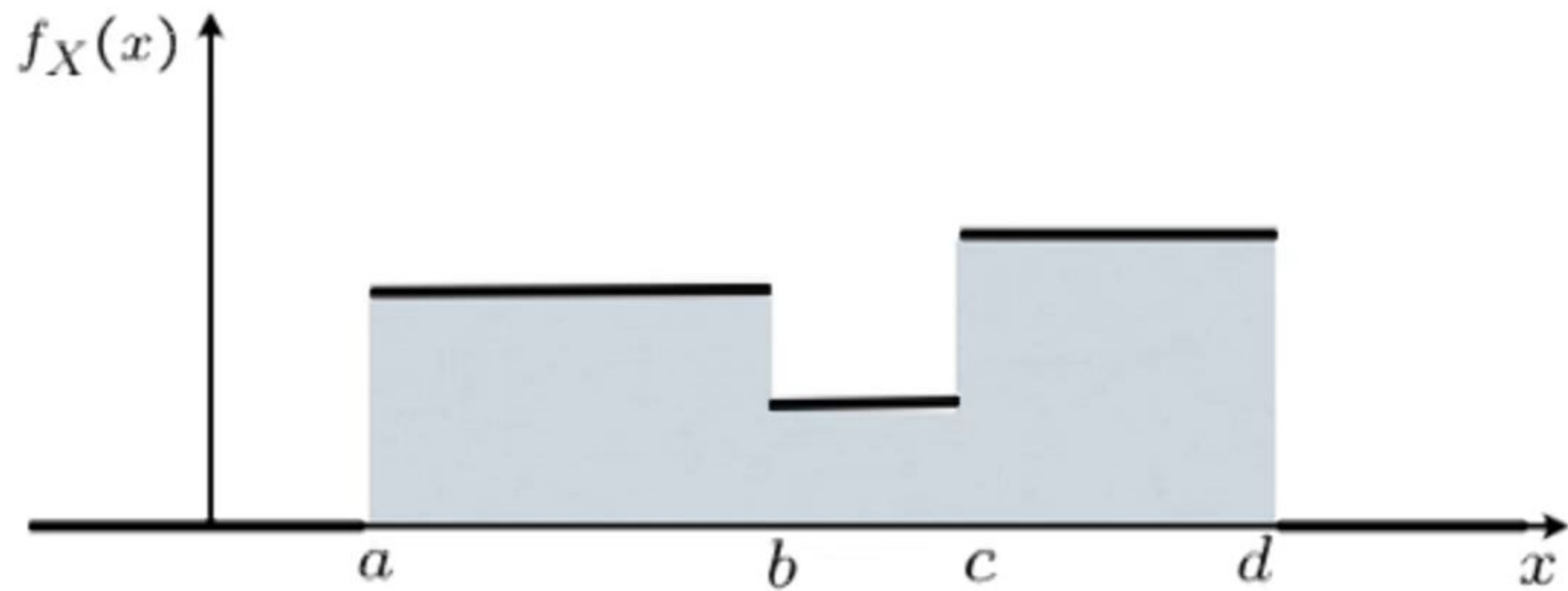
$$= \frac{P(x \leq X \leq x + \delta, X \in A)}{P(A)}$$

$$= \frac{P(x \leq X \leq x + \delta)}{P(A)} \approx \frac{f_X(x) \delta}{P(A)}$$

$$f_{X|X \in A}(x) = \begin{cases} 0, & \text{if } x \notin A \\ \frac{f_X(x)}{P(A)}, & \text{if } x \in A \end{cases}$$

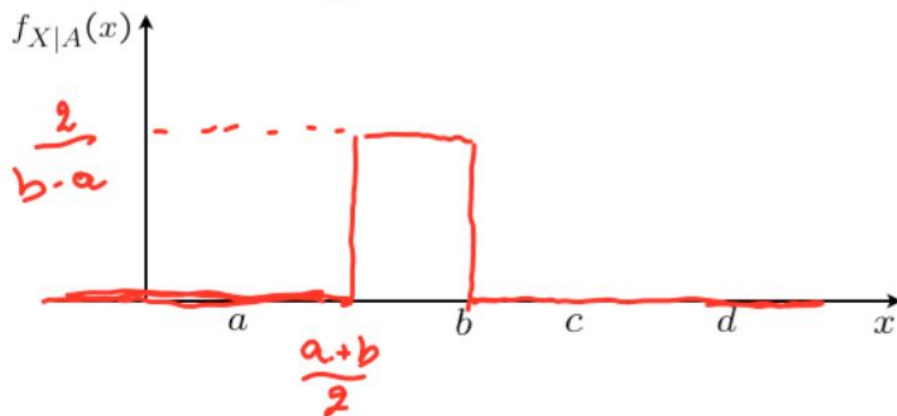
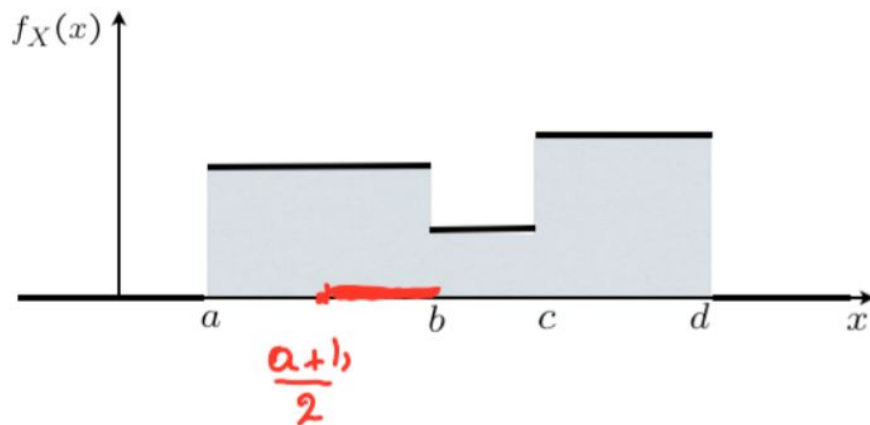


Example



Example

$$A: \frac{a+b}{2} \leq X \leq b$$



$$\mathbf{E}[X \mid A] = \frac{1}{2} \cdot \frac{a+b}{2} + \frac{1}{2} b$$

$$= \frac{1}{4} a + \frac{3}{4} b$$

$$\mathbf{E}[X^2 \mid A] =$$

$$\int_{\frac{a+b}{2}}^b \frac{2}{b-a} \cdot x^2 dx$$

Memorylessness of the exponential PDF

- Do you prefer a used or a new “exponential” light bulb?
- Bulb lifetime T : $\text{exponential}(\lambda)$

$$\mathbf{P}(T > x) = e^{-\lambda x}, \text{ for } x \geq 0$$

- we are told that $T > t$
- r.v. X : remaining lifetime

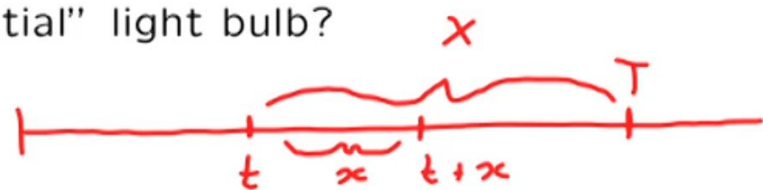
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– r.v. X : remaining lifetime $= T - t$



$$P(X > x | T > t) =$$

Memorylessness of the exponential PDF

- Do you prefer a used or a new “exponential” light bulb? **Probabilistically identical!**
- Bulb lifetime T : $\text{exponential}(\lambda)$



$$P(T > x) = e^{-\lambda x}, \text{ for } x \geq 0$$

– we are told that $T > t$

– r.v. X : remaining lifetime $= T - t$

$$P(X > x | T > t) = e^{-\lambda x}, \text{ for } x \geq 0$$

$$= \frac{P(T - t > x, T > t)}{P(T > t)} = \frac{P(T > t + x, T > t)}{P(T > t)} = \frac{P(T > t + x)}{P(T > t)}$$

$$= \frac{e^{-\lambda(t+x)}}{e^{-\lambda t}} = e^{-\lambda x}$$

Memorylessness of the exponential PDF

$$f_T(x) = \lambda e^{-\lambda x}, \quad \text{for } x \geq 0$$

$$\mathbf{P}(0 \leq T \leq \delta)$$

$$\mathbf{P}(t \leq T \leq t + \delta \mid T > t)$$

similar to an independent coin flip,
every δ time steps,
with $\mathbf{P}(\text{success}) \approx \lambda \delta$

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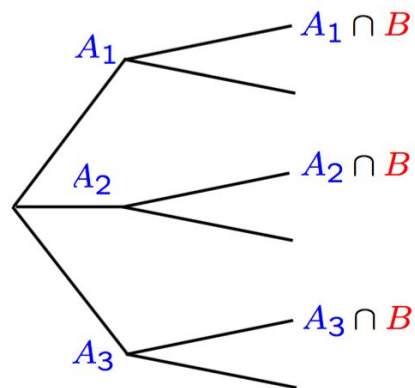
$$\mathbf{P}(0 \leq T \leq \delta) \approx f_T(0) \cdot \delta = \lambda \delta$$

$$\mathbf{P}(t \leq T \leq t + \delta \mid T > t) = \lambda \delta$$

similar to an independent coin flip,
every δ time steps,
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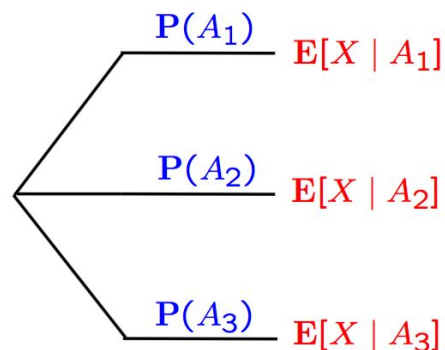
Total probability and expectation theorems



$$\mathbf{P}(B) = \mathbf{P}(A_1)\mathbf{P}(B | A_1) + \cdots + \mathbf{P}(A_n)\mathbf{P}(B | A_n)$$

$$p_X(x) = \mathbf{P}(A_1)p_{X|A_1}(x) + \cdots + \mathbf{P}(A_n)p_{X|A_n}(x)$$

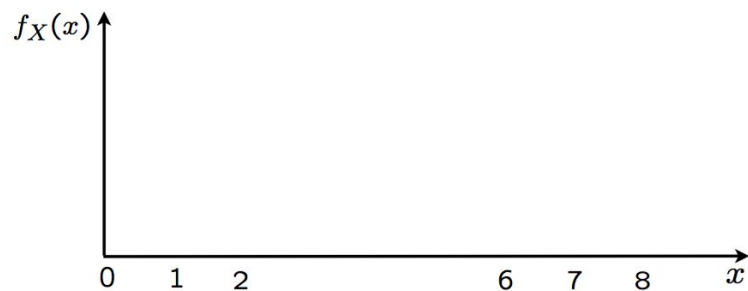
$$f_X(x) = \mathbf{P}(A_1)f_{X|A_1}(x) + \cdots + \mathbf{P}(A_n)f_{X|A_n}(x)$$



$$\mathbf{E}[X] = \mathbf{P}(A_1)\mathbf{E}[X | A_1] + \cdots + \mathbf{P}(A_n)\mathbf{E}[X | A_n]$$

Example

- Bill goes to the supermarket shortly, with probability $1/3$, at a time uniformly distributed between 0 and 2 hours from now; or with probability $2/3$, later in the day at a time uniformly distributed between 6 and 8 hours from now



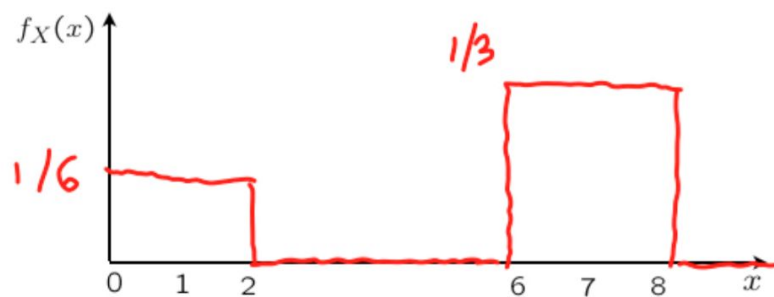
$$f_X(x) = P(A_1)f_{X|A_1}(x) + \cdots + P(A_n)f_{X|A_n}(x)$$

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- Bill goes to the supermarket shortly, with probability $1/3$, at a time uniformly distributed between 0 and 2 hours from now; or with probability $2/3$, later in the day at a time uniformly distributed between 6 and 8 hours from now

$$P(A_1) = \frac{1}{3} \quad f_{X|A_1} \sim \text{unif}[0, 2] \quad P(A_2) = \frac{2}{3} \quad f_{X|A_2} \sim U[6, 8]$$



$$f_X(x) = P(A_1)f_{X|A_1}(x) + \cdots + P(A_n)f_{X|A_n}(x)$$

$$\bullet \quad E[X] = P(A_1)E[X | A_1] + \cdots + P(A_n)E[X | A_n]$$

$$\frac{1}{3} \cdot 1 + \frac{2}{3} \cdot 7$$

Jointly continuous r.v.'s and joint PDFs

$$p_X(x) \quad f_X(x)$$

$$p_{X,Y}(x, y) \quad f_{X,Y}(x, y)$$

$$p_{X,Y}(x, y) = \mathbf{P}(X = x \text{ and } Y = y) \geq 0$$

$$f_{X,Y}(x, y) \geq 0$$

$$\mathbf{P}((X, Y) \in B) = \sum_{(x,y) \in B} p_{X,Y}(x, y)$$

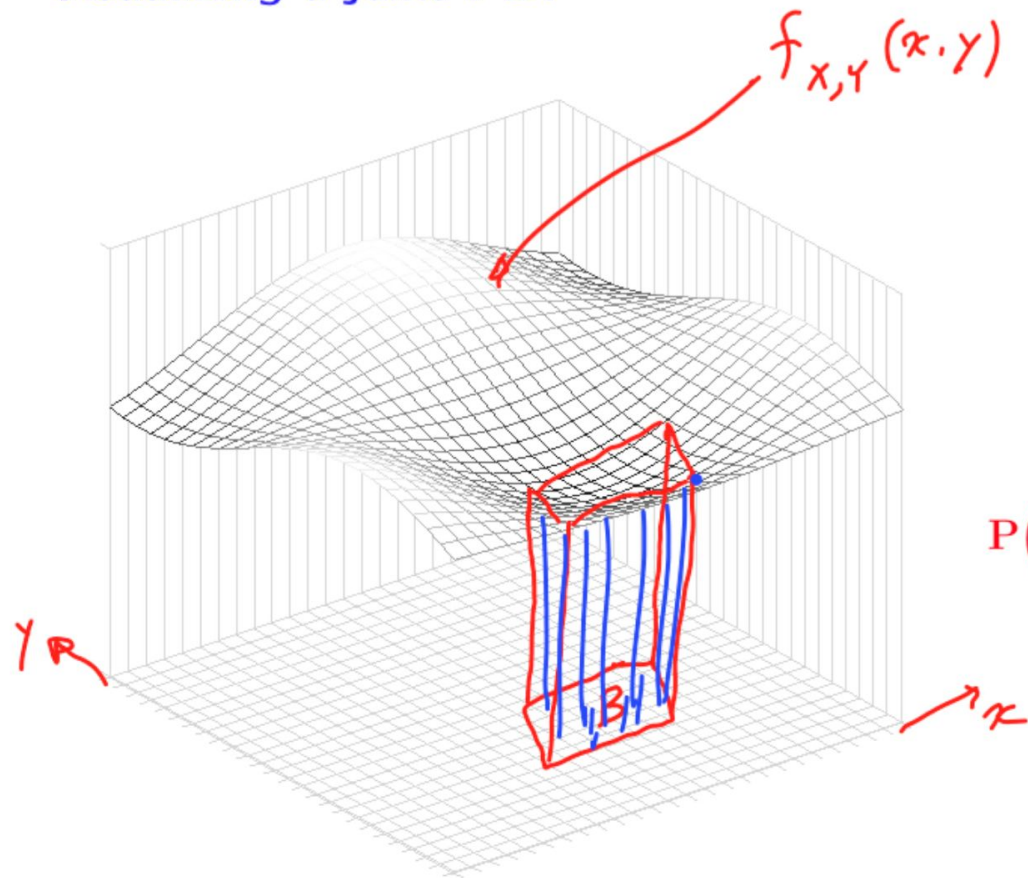
$$\mathbf{P}((X, Y) \in B) = \int \int_{(x,y) \in B} f_{X,Y}(x, y) dx dy$$

$$\sum_x \sum_y p_{X,Y}(x, y) = 1$$

$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{X,Y}(x, y) dx dy = 1$$

Definition: Two random variables are **jointly continuous** if they can be described by a joint PDF

Visualizing a joint PDF



$$P((X,Y) \in B) = \int \int_{(x,y) \in B} f_{X,Y}(x,y) dx dy$$

On joint PDFs

$$\mathbf{P}\big((X, Y) \in B\big) = \int \int_{(x,y) \in B} f_{X,Y}(x, y) \, dx \, dy$$

$$\mathbf{P}(a \leq X \leq b, c \leq Y \leq d) = \int_c^d \int_a^b f_{X,Y}(x, y) \, dx \, dy$$

$$\mathbf{P}(a \leq X \leq a + \delta, c \leq Y \leq c + \delta) \approx f_{X,Y}(a, c) \cdot \delta^2$$

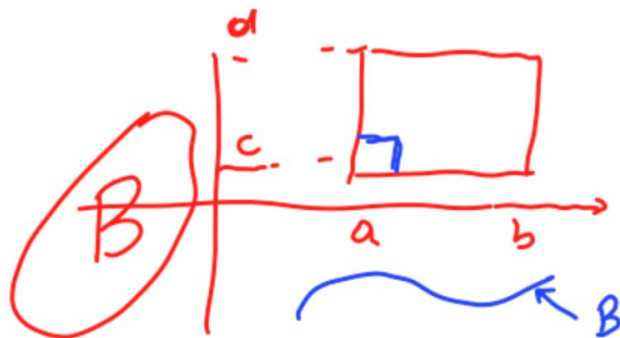
$f_{X,Y}(x, y)$: probability per unit area

$$\text{area}(B) = 0 \quad \Rightarrow \quad \mathbf{P}\big((X, Y) \in B\big) = 0$$

On joint PDFs

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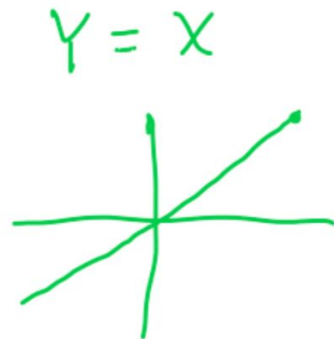
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$f_{X,Y}(x, y)$: probability per unit area

$$\text{area}(B) = 0 \Rightarrow \mathbf{P}((X, Y) \in B) = 0$$



From the joint to the marginals

$$p_X(x) = \sum_y p_{X,Y}(x, y)$$

$$f_X(x) = \int f_{X,Y}(x, y) dy$$

$$p_Y(y) = \sum_x p_{X,Y}(x, y)$$

$$f_Y(y) = \int f_{X,Y}(x, y) dx$$

Joint, Marginal and Conditional Densities

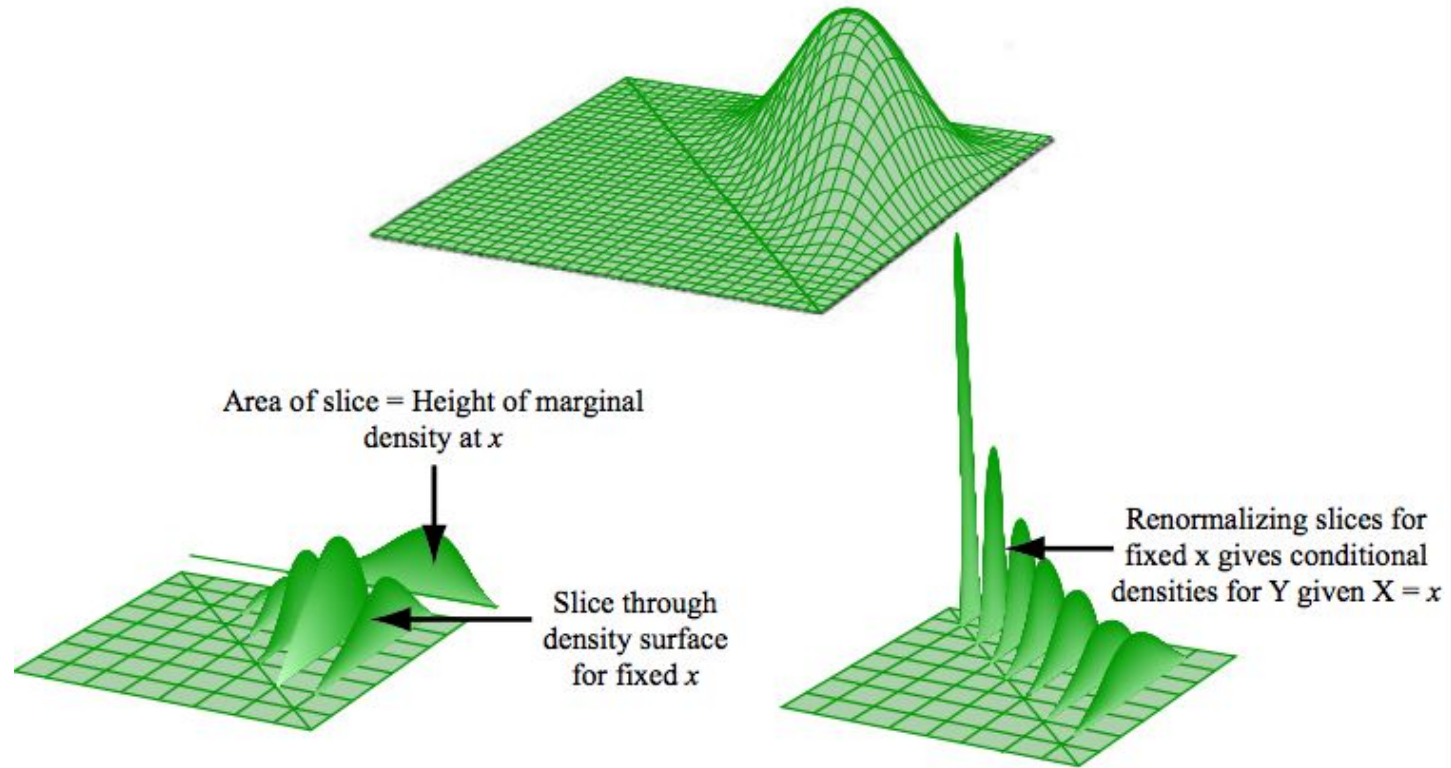


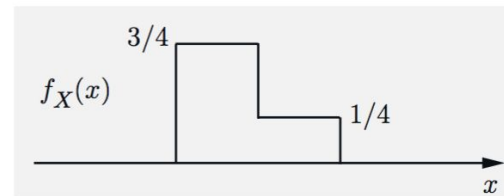
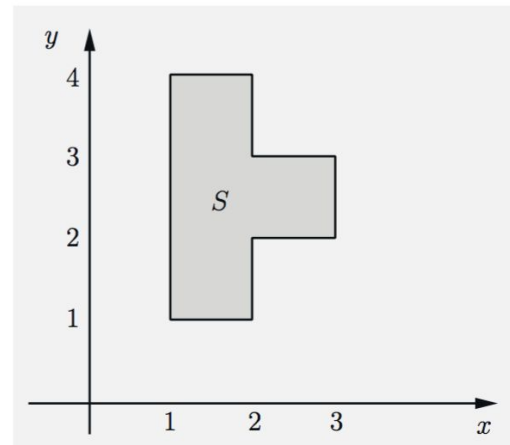
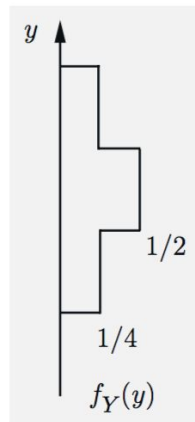
Image adapted from
Probability, by J. Pittman, 1999.

Uniform joint PDF on a set S

$$f_{X,Y}(x,y) = \begin{cases} \frac{1}{\text{area of } S}, & \text{if } (x,y) \in S, \\ 0, & \text{otherwise.} \end{cases}$$

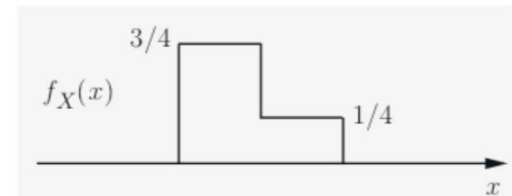
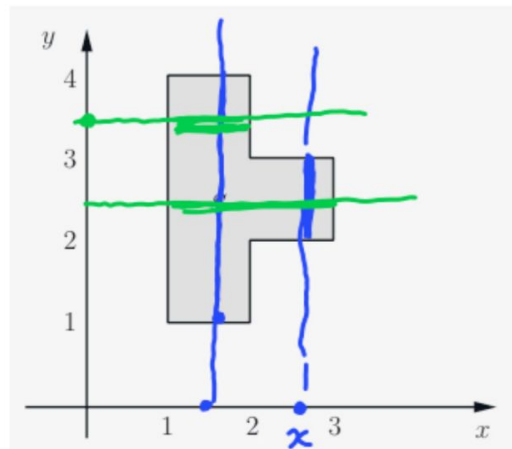
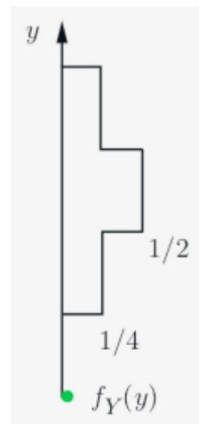
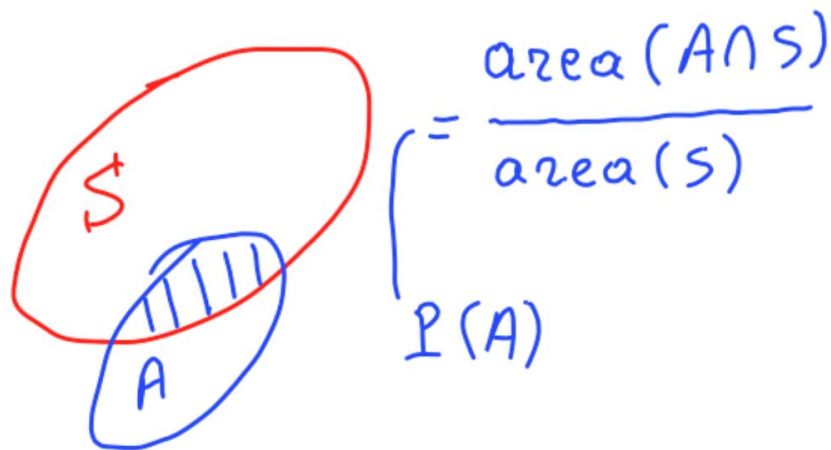
$$f_X(x) = \int f_{X,Y}(x,y) dy$$

$$f_Y(y) = \int f_{X,Y}(x,y) dx$$



Uniform joint PDF on a set S

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$$f_{X,Y} = \frac{1}{4}$$

Conditional PDFs, given another r.v.

$$p_{X|Y}(x | y) = \mathbf{P}(X = x | Y = y) = \frac{p_{X,Y}(x, y)}{p_Y(y)}, \quad \text{if } p_Y(y) > 0$$

$p_{X,Y}(x, y)$	$f_{X,Y}(x, y)$
$p_{X A}(x)$	$f_{X A}(x)$
$p_{X Y}(x y)$	$f_{X Y}(x y)$

Definition: $f_{X|Y}(x | y) = \frac{f_{X,Y}(x, y)}{f_Y(y)}$ if $f_Y(y) > 0$

$$\mathbf{P}(x \leq X \leq x + \delta | A) \approx f_{X|A}(x) \cdot \delta, \quad \text{where } \mathbf{P}(A) > 0$$

$$\mathbf{P}(x \leq X \leq x + \delta | y \leq Y \leq y + \epsilon)$$

Definition: $\mathbf{P}(X \in A | Y = y) = \int_A f_{X|Y}(x | y) dx$

Comments on conditional PDFs

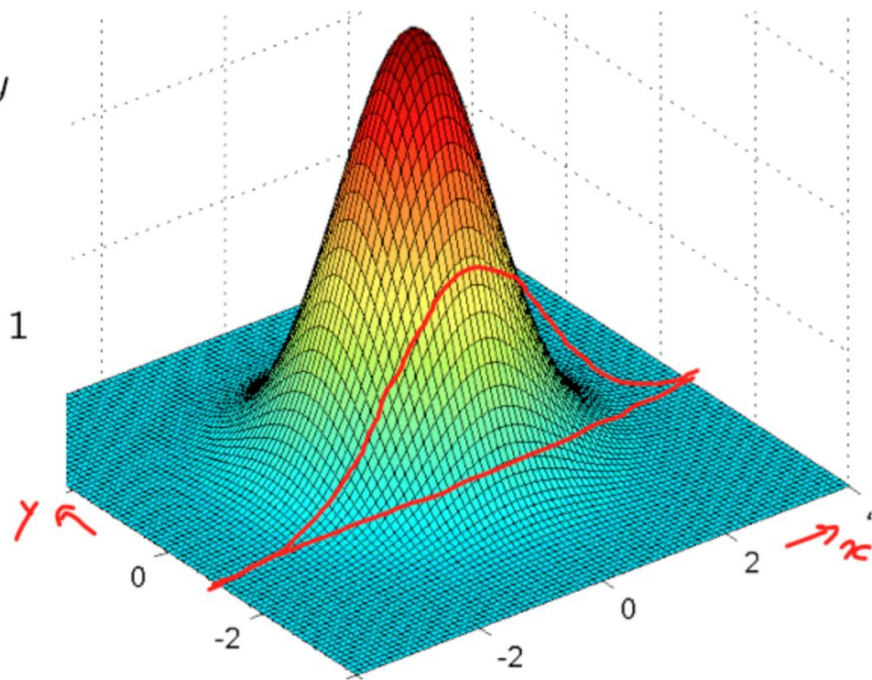
$$f_{X|Y}(x | y) = \frac{f_{X,Y}(x, y)}{f_Y(y)} \quad \bullet \quad f_{X|Y}(x | y) \geq 0$$

- Think of value of Y as fixed at some y
shape of $f_{X|Y}(\cdot | y)$: slice of the joint

$$\bullet \quad \int_{-\infty}^{\infty} f_{X|Y}(x | y) dx = \frac{\int_{-\infty}^{\infty} f_{X,Y}(x, y) dx}{f_Y(y)} = 1$$

- Multiplication rule:

$$\begin{aligned} f_{X,Y}(x, y) &= f_Y(y) \cdot f_{X|Y}(x | y) \\ &= f_X(x) \cdot f_{Y|X}(y | x) \end{aligned}$$



Independence

$$p_{X,Y}(x,y) = p_X(x) p_Y(y), \quad \text{for all } x, y$$

$$f_{X,Y}(x,y) = f_X(x) f_Y(y), \quad \text{for all } x \text{ and } y$$

$$f_{X,Y}(x,y) = f_{X|Y}(x|y) f_Y(y)$$

- equivalent to: $f_{X|Y}(x|y) = f_X(x)$, for all y with $f_Y(y) > 0$ and all x

If X, Y are **independent**: $\mathbf{E}[XY] = \mathbf{E}[X]\mathbf{E}[Y]$

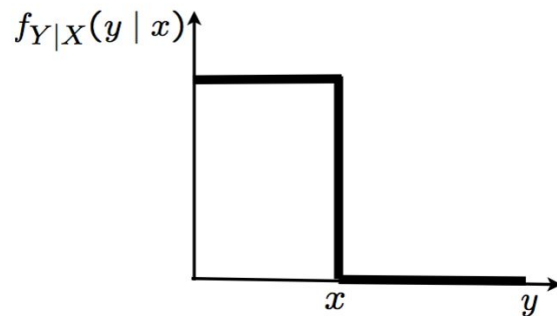
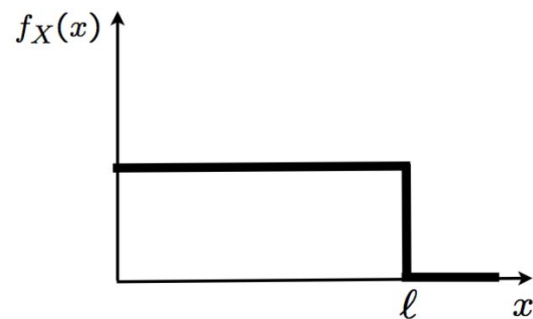
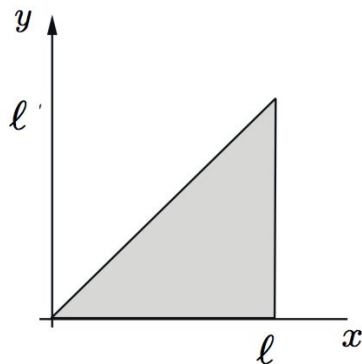
$$\mathbf{var}(X + Y) = \mathbf{var}(X) + \mathbf{var}(Y)$$

$g(X)$ and $h(Y)$ are also independent: $\mathbf{E}[g(X)h(Y)] = \mathbf{E}[g(X)] \cdot \mathbf{E}[h(Y)]$

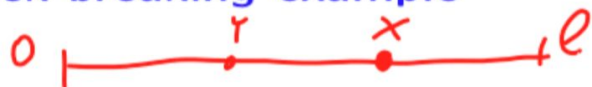
Stick-breaking example

- Break a stick of length ℓ twice
 - first break at X : uniform in $[0, \ell]$
 - second break at Y : uniform in $[0, X]$

$$f_{X,Y}(x,y) = f_X(x)f_{Y|X}(y|x) =$$



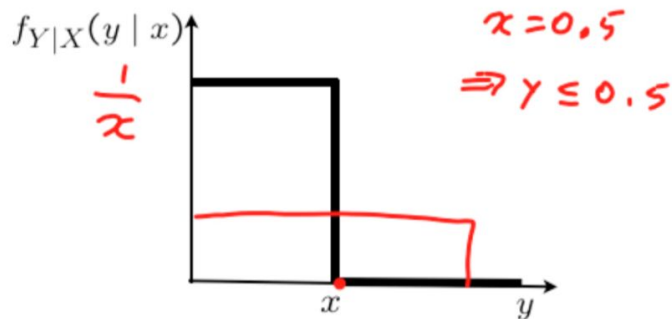
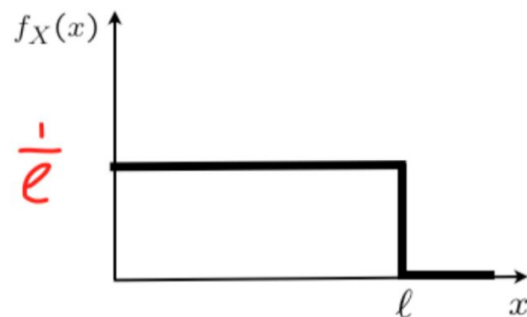
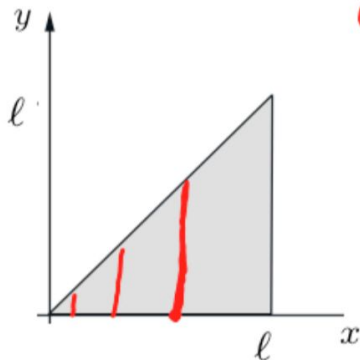
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$$f_{X,Y}(x,y) = f_X(x)f_{Y|X}(y|x) = \frac{1}{\ell x}$$

$$0 \leq y \leq x \leq \ell$$



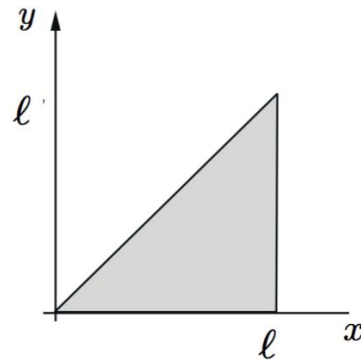
Stick-breaking example

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$$f_Y(y) =$$

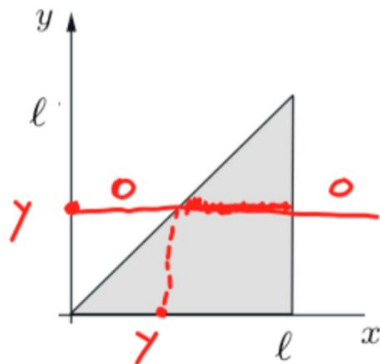
$$\mathbf{E}[Y] =$$

- Using total expectation theorem:



Stick-breaking example

$$f_{X,Y}(x,y) = \frac{1}{\ell x}, \quad 0 \leq y \leq x \leq \ell$$



$$f_Y(y) = \int_y^{\ell} f_{X,Y}(x,y) dx = \int_y^{\ell} \frac{1}{\ell x} dx = \frac{1}{\ell} \log\left(\frac{\ell}{y}\right)$$

$$E[Y] = \int_0^{\ell} y \frac{1}{\ell} \log\left(\frac{\ell}{y}\right) dy$$

- Using total expectation theorem:

$$E[Y] = \int_0^{\ell} \frac{1}{\ell} E[Y|X=x] dx = \int_0^{\ell} \left(\frac{1}{\ell}\right) \frac{x}{2} dx = \frac{1}{2} E[X] = \frac{1}{2} \cdot \frac{\ell}{2} = \frac{\ell}{4}$$

