Deep latent models of word representation

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April 3, 2018

Word representation

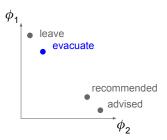
An abstraction that stands for the use of a word

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Word representation

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How do we choose the components?

Distributional Hypothesis

Context can represent the intended use of a word

In the event of a chemical spill, most children know they should **evacuate** as advised by people in charge.

success hinges on discriminative power of available context

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Place words in \mathbb{R}^d as to answer questions like

"Have I seen this word in this context?"

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(Goldberg and Levy, 2014)

- positive examples
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But data only show positive examples

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ambiguity

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Limitations

Meaning representation is an unsupervised problem

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Distributional hypothesis seems pretty strong

• but it fails when context is not sufficiently discriminative

EMBEDALIGN

Generative treatment

- model what we want to induce (i.e. representations)
- learn from positive examples
- learn from richer (less ambiguous) context

Outline

- Embed-Align
- 2 Bayesian Skip-Gram
- 3 Practica

Equivalence through translation

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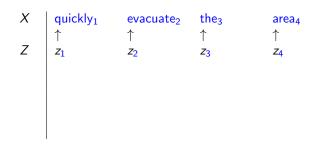
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Observation from WSD community

 foreign text as proxy to sense supervision (Diab and Resnik, 2002)



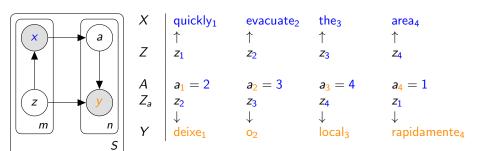




$$egin{array}{c|cccc} X & \mathsf{quickly_1} & \mathsf{evacuate_2} & \mathsf{the_3} & \mathsf{area_4} \\ \uparrow & \uparrow & \uparrow & \uparrow \\ z_1 & z_2 & z_3 & z_4 \\ A & a_1 = 2 & a_2 = 3 & a_3 = 4 & a_4 = 1 \\ \hline \end{array}$$

Χ	quickly ₁	evacuate ₂	the ₃	area ₄
Ζ	↑ z ₁	\uparrow z_2	↑ <i>z</i> ₃	↑ <i>z</i> ₄
Д		$a_2 = 3$	$a_3 = 4$	$a_4 = 1$
Z_a	Z_2	z_3	Z ₄	z_1
Y	↓ deixe ₁	↓ <mark>0</mark> 2	↓ local ₃	↓ rapidamente₄

quickly evacuate the area / deixe o local rapidamente



Marginalising alignments collects additional training data for z

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Z_i(1)

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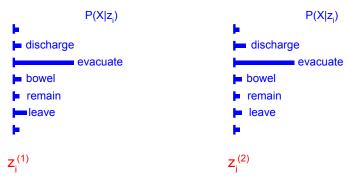
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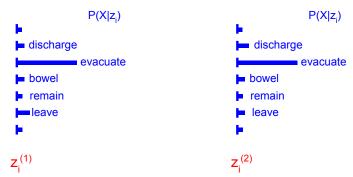
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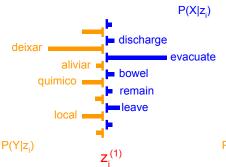


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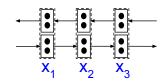
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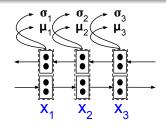
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Read sentence

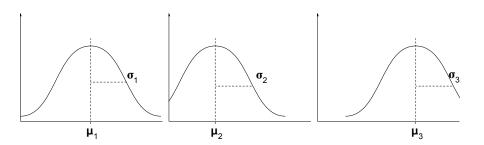


Evacuate₁ the₂ area₃

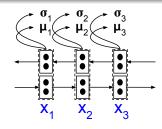
- Read sentence
- 2 Predict posterior mean μ_i and std σ_i



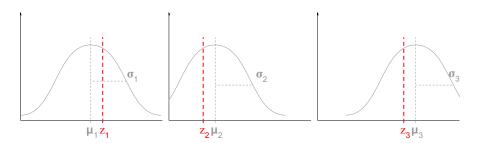
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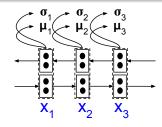
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- **3** Sample $Z_i \sim \mathcal{N}(\mu_i, \sigma_i^2)$



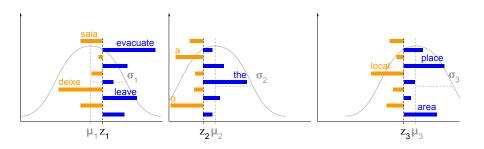
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- Predict categorical distributions

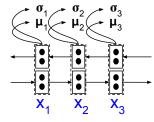


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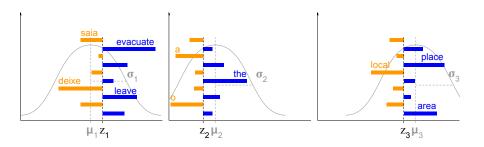


Tractable inference

- Read sentence
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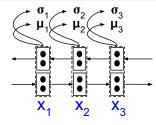


Evacuate₁ the₂ area₃

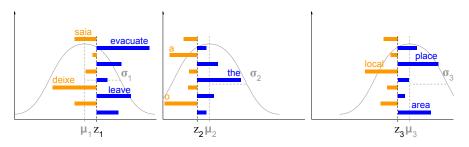


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 Evacuate₁ the₂ area₃ / Deixe₁ o₂ local₃
- Maximise a lowerbound on likelihood (Kingma and Welling, 2014)

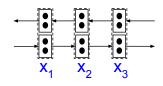


Evacuate₁ the₂ area₃



What's special about it?

The model reads English text and

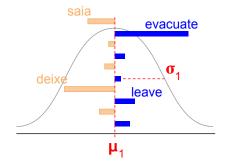


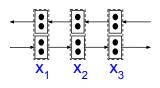
 $\mathsf{Evacuate}_1 \ \mathsf{the}_2 \ \mathsf{area}_3$

What's special about it?

The model reads English text and

- predicts uncertainty
- describes "sense" using Portuguese words





Evacuate₁ the₂ area₃

$$Z_i \sim \mathcal{N}(0, I)$$

$$egin{aligned} Z_i &\sim \mathcal{N}(0,I) \ X_i | z_i &\sim \mathsf{Cat}(\mathbf{f}_i) \ \mathbf{f}_i &= \mathsf{softmax}(\mathsf{affine}_{ heta}(z_i)) \end{aligned}$$

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Generative model: for i = 1, ..., m and j = 1, ..., n

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Inference model: for $i = 1, \ldots, m$

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$$Z_i|x_1^m \sim \mathcal{N}(\mathbf{u}_i, \operatorname{diag}(\mathbf{s}_i \odot \mathbf{s}_i))$$

Generative model: for
$$i=1,\ldots,m$$
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Rios et al. (2018)

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ELBO

ELBO

$$\mathbb{E}_{q_{\lambda}(z_1^m|x_1^m)}\left[\log P_{\theta}(x_1^m,y_1^n|z_1^m)\right] - \mathsf{KL}\left(q_{\lambda}(z_1^m|x_1^m) \mid\mid \rho(z)\right)$$

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KL term

$$\mathsf{KL}\left(q_{\lambda}(\mathsf{z}_1^m|\mathsf{x}_1^m)\mid\mid p(\mathsf{z})\right) = \underbrace{\sum_{i=1}^m \mathsf{KL}\left(q_{\lambda}(\mathsf{z}_i|\mathsf{x}_1^m)\mid\mid p(\mathsf{z})\right)}_{\mathsf{mean field}}$$

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$$= \underbrace{\sum_{i=1}^{m} \mathsf{KL}\left(\underbrace{\mathcal{N}(\mathbf{u}_{i}, \mathsf{diag}(\mathbf{s}_{i} \odot \mathbf{s}_{i}))}_{\mathsf{inference model}} \mid\mid \underbrace{\mathcal{N}(0, I)}_{\mathsf{prior}}\right)}_{\mathsf{prior}}$$

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$$\mathbb{E}_{q_{\lambda}(z_{1}^{m}|x_{1}^{m})}\left[\log P_{\theta}(x_{1}^{m},y_{1}^{n}|z_{1}^{m})\right] - \mathsf{KL}\left(q_{\lambda}(z_{1}^{m}|x_{1}^{m}) \mid\mid p(z)\right)$$

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Likelihood term

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conditional independence

ELBO

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Likelihood term

$$\begin{split} & \mathbb{E}_{q_{\lambda}(\mathbf{z}_{1}^{m}|\mathbf{x}_{1}^{m})}\left[\log P_{\theta}(\mathbf{x}_{1}^{m}, \mathbf{y}_{1}^{n}|\mathbf{z}_{1}^{m})\right] \\ & = \underbrace{\mathbb{E}_{q_{\lambda}(\mathbf{z}_{1}^{m}|\mathbf{x}_{1}^{m})}\left[\log P_{\theta}(\mathbf{x}_{1}^{m}|\mathbf{z}_{1}^{m})\right] + \mathbb{E}_{q_{\lambda}(\mathbf{z}_{1}^{m}|\mathbf{x}_{1}^{m})}\left[\log P_{\theta}(\mathbf{y}_{1}^{n}|m, \mathbf{z}_{1}^{m})\right]}_{\text{conditional independence}} \end{split}$$

L1 term

$$\mathbb{E}_{q_{\lambda}(z_{1}^{m}|x_{1}^{m})}[\log P_{\theta}(x_{1}^{m}|z_{1}^{m})] = \sum_{i=1}^{m} \mathbb{E}_{q_{\lambda}(z_{i}|x_{1}^{m})}[\log P_{\theta}(x_{i}|z_{i})]$$

ELBO

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$$= \sum_{i=1}^m \mathbb{E}_{q_{\lambda}(z_i|x_1^m)}\left[\log \operatorname{Cat}(x_i|\mathbf{f}_i)\right]$$

$$\mathbb{E}_{q_{\lambda}\left(z_1^m|x_1^m\right)}\left[\log P_{\theta}(y_1^n|m,z_1^m)\right]$$

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$$\begin{split} \mathbb{E}_{q_{\lambda}(z_{1}^{m}|x_{1}^{m})}\left[\log P_{\theta}(y_{1}^{n}|m,z_{1}^{m})\right] &= \mathbb{E}_{q_{\lambda}(z_{1}^{m}|x_{1}^{m})}\left[\log \prod_{j=1}^{n} P_{\theta}(y_{j}|m,z_{1}^{m})\right] \\ &= \mathbb{E}_{q_{\lambda}(z_{1}^{m}|x_{1}^{m})}\left[\sum_{j=1}^{n} \log P_{\theta}(y_{j}|m,z_{1}^{m})\right] \\ &= \mathbb{E}_{q_{\lambda}(z_{1}^{m}|x_{1}^{m})}\left[\sum_{j=1}^{n} \log \sum_{a_{j}}^{m} P_{\theta}(y_{j},a_{j}|m,z_{1}^{m})\right] \\ &= \mathbb{E}_{q_{\lambda}(z_{1}^{m}|x_{1}^{m})}\left[\sum_{j=1}^{n} \log \sum_{a_{j}}^{m} P(a_{j}|m)P_{\theta}(y_{j}|z_{a_{j}})\right] \\ &= \mathbb{E}_{q_{\lambda}(z_{1}^{m}|x_{1}^{m})}\left[\sum_{j=1}^{n} \log \sum_{a_{j}}^{m} \mathcal{U}(a_{j}|1/m)\operatorname{Cat}(y_{j}|\mathbf{g}_{a_{j}})\right] \end{split}$$

$$\mathbb{E}_{q_{\lambda}(z_{1}^{m}|x_{1}^{m})}[\log P_{\theta}(y_{1}^{n}|m,z_{1}^{m})] = \mathbb{E}_{q_{\lambda}(z_{1}^{m}|x_{1}^{m})}\left|\sum_{j=1}^{n}\log \sum_{a_{j}}^{m}P(a_{j}|m)P_{\theta}(y_{j}|z_{a_{j}})\right|$$

$$\begin{split} \mathbb{E}_{q_{\lambda}(z_1^m|x_1^m)}\left[\log P_{\theta}(y_1^n|m,z_1^m)\right] &= \mathbb{E}_{q_{\lambda}(z_1^m|x_1^m)}\left[\sum_{j=1}^n \log \sum_{a_j}^m P(a_j|m)P_{\theta}(y_j|z_{a_j})\right] \\ &= \mathbb{E}_{q_{\lambda}(z_1^m|x_1^m)}\left[\sum_{j=1}^n \log \mathbb{E}_{P(a_j|m)}\left[P_{\theta}(y_j|z_{a_j})\right]\right] \end{split}$$

$$\begin{split} \mathbb{E}_{q_{\lambda}(z_1^m|x_1^m)}\left[\log P_{\theta}(y_1^n|m,z_1^m)\right] &= \mathbb{E}_{q_{\lambda}(z_1^m|x_1^m)}\left[\sum_{j=1}^n \log \sum_{a_j}^m P(a_j|m)P_{\theta}(y_j|z_{a_j})\right] \\ &= \mathbb{E}_{q_{\lambda}(z_1^m|x_1^m)}\left[\sum_{j=1}^n \log \mathbb{E}_{P(a_j|m)}\left[P_{\theta}(y_j|z_{a_j})\right]\right] \\ &\stackrel{\mathsf{JI}}{\geq} \mathbb{E}_{q_{\lambda}(z_1^m|x_1^m)}\left[\sum_{j=1}^n \mathbb{E}_{P(a_j|m)}[\log P_{\theta}(y_j|z_{a_j})]\right] \end{split}$$

$$\begin{split} \mathbb{E}_{q_{\lambda}(z_{1}^{m}|x_{1}^{m})}\left[\log P_{\theta}(y_{1}^{n}|m,z_{1}^{m})\right] &= \mathbb{E}_{q_{\lambda}(z_{1}^{m}|x_{1}^{m})}\left[\sum_{j=1}^{n}\log \sum_{a_{j}}^{m}P(a_{j}|m)P_{\theta}(y_{j}|z_{a_{j}})\right] \\ &= \mathbb{E}_{q_{\lambda}(z_{1}^{m}|x_{1}^{m})}\left[\sum_{j=1}^{n}\log \mathbb{E}_{P(a_{j}|m)}\left[P_{\theta}(y_{j}|z_{a_{j}})\right]\right] \\ &\stackrel{\text{JI}}{\geq} \mathbb{E}_{q_{\lambda}(z_{1}^{m}|x_{1}^{m})}\left[\sum_{j=1}^{n}\mathbb{E}_{P(a_{j}|m)}[\log P_{\theta}(y_{j}|z_{a_{j}})]\right] \\ &= \mathbb{E}_{q_{\lambda}(z_{1}^{m}|x_{1}^{m})}\left[\sum_{j=1}^{n}\mathbb{E}_{\mathcal{U}(a_{j}|m)}[\log \mathsf{Cat}(y_{j}|\mathbf{g}_{a_{j}})]\right] \end{split}$$

L2 term

$$\begin{split} \mathbb{E}_{q_{\lambda}(z_{1}^{m}|x_{1}^{m})}\left[\log P_{\theta}(y_{1}^{n}|m,z_{1}^{m})\right] &= \mathbb{E}_{q_{\lambda}(z_{1}^{m}|x_{1}^{m})}\left[\sum_{j=1}^{n}\log \sum_{a_{j}}^{m}P(a_{j}|m)P_{\theta}(y_{j}|z_{a_{j}})\right] \\ &= \mathbb{E}_{q_{\lambda}(z_{1}^{m}|x_{1}^{m})}\left[\sum_{j=1}^{n}\log \mathbb{E}_{P(a_{j}|m)}\left[P_{\theta}(y_{j}|z_{a_{j}})\right]\right] \\ &\stackrel{\text{JI}}{\geq} \mathbb{E}_{q_{\lambda}(z_{1}^{m}|x_{1}^{m})}\left[\sum_{j=1}^{n}\mathbb{E}_{P(a_{j}|m)}[\log P_{\theta}(y_{j}|z_{a_{j}})\right]\right] \\ &= \mathbb{E}_{q_{\lambda}(z_{1}^{m}|x_{1}^{m})}\left[\sum_{j=1}^{n}\mathbb{E}_{\mathcal{U}(a_{j}|m)}[\log \mathsf{Cat}(y_{j}|\mathbf{g}_{a_{j}})]\right] \end{split}$$

Alignment model $P(a_j|m) = \frac{1}{m}$ is not a function of θ

 we can make an MC estimate by sampling candidate alignments uniformly

The softmax problem

Categorical parameters are expensive to compute

$$X \sim \mathsf{Cat}(\mathbf{f})$$
 $\mathbf{f} = \mathsf{softmax}(\hat{\mathbf{f}})$
 $f_X = \frac{\mathsf{exp}(\hat{f}_X)}{\sum_{X' \in \mathcal{X}} \mathsf{exp}(\hat{f}_{X'})}$

 $v_1=|\mathcal{X}|$ is the size of the vocabulary of L_1 ; $v_2=|\mathcal{Y}|$ is the size of the vocabulary of L_2

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 $\mathbf{f} = \mathsf{softmax}(\hat{\mathbf{f}})$
 $f_X = \frac{\mathsf{exp}(\hat{f}_X)}{\sum_{x' \in \mathcal{X}} \mathsf{exp}(\hat{f}_{x'})}$

- \mathbf{f}_1^m requires normalising m distributions over the vocabulary of L_1 , thus it takes time $O(m \times v_1)$
- \mathbf{g}_1^m requires normalising m distributions over the vocabulary of L_2 , thus it takes time $O(m \times v_2)$

 $v_1=|\mathcal{X}|$ is the size of the vocabulary of $L_1;~v_2=|\mathcal{Y}|$ is the size of the vocabulary of L_2

Efficient softmax

Logistic regression

$$P(X = x | z) = \frac{\exp(u(z, x))}{\sum_{x' \in \mathcal{X}} \exp(u(z, x'))}$$

Efficient softmax

Logistic regression

$$P(X = x|z) = \frac{\exp(u(z,x))}{\sum_{x' \in \mathcal{X}} \exp(u(z,x'))}$$

Define

- a set C(x) such that $x \in C$
- a set $\mathcal{N}(x)$ such that $\mathcal{C}(x) \cap \mathcal{N}(x) = \emptyset$

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Define

- a set C(x) such that $x \in C$
- a set $\mathcal{N}(x)$ such that $\mathcal{C}(x) \cap \mathcal{N}(x) = \emptyset$

Re-express normaliser for P(X = x|z)

$$\sum_{x' \in \mathcal{X}} \exp(u(z, x')) = \sum_{x' \in \mathcal{C}(x)} \exp(u(z, x')) + \sum_{x' \in \mathcal{N}(x)} \kappa(x') \exp(u(z, x'))$$

• $\kappa(x') = \frac{1}{q(x')}$ and q(x') is an importance distribution

Approximate $P_{X|Z}$

Logistic regression

$$P_{\theta}(x|z) = \frac{\exp(u_{\theta}(z,x))}{\sum_{x' \in \mathcal{X}} \exp(u_{\theta}(z,x'))}$$

Approximate $P_{X|Z}$

Logistic regression

$$P_{\theta}(x|z) = \frac{\exp(u_{\theta}(z,x))}{\sum_{x' \in \mathcal{X}} \exp(u_{\theta}(z,x'))}$$

Build

- a set C containing all L_1 words in batch
- ullet a set ${\mathcal N}$ sampling uniformly without replacement from ${\mathcal X}\setminus {\mathcal C}$

Approximate $P_{X|Z}$

Logistic regression

$$P_{\theta}(x|z) = \frac{\exp(u_{\theta}(z,x))}{\sum_{x' \in \mathcal{X}} \exp(u_{\theta}(z,x'))}$$

Build

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- ullet a set ${\mathcal N}$ sampling uniformly without replacement from ${\mathcal X}\setminus {\mathcal C}$

Approximate normaliser for $P_{\theta}(x|z)$

$$\sum_{x' \in \mathcal{X}} \exp(u(z, x')) \approx \sum_{x' \in \mathcal{C}} \exp(u(z, x')) + \sum_{x' \in \mathcal{N}} \frac{|\mathcal{X} \setminus \mathcal{C}|}{|\mathcal{N}|} \exp(u(z, x'))$$

• $u_{\theta}(z, x) = z^{\top} \mathbf{c}_{x} + b_{x}$ \mathbf{c}_{x} is a deterministic embedding, b_{x} a bias term

Approximate $P_{Y|Z}$

Logistic regression

$$P_{\theta}(y|z) = rac{\exp(u_{\theta}(z,y))}{\sum_{y' \in \mathcal{Y}} \exp(u_{\theta}(z,y'))}$$

Approximate $P_{Y|Z}$

Logistic regression

$$P_{\theta}(y|z) = \frac{\exp(u_{\theta}(z,y))}{\sum_{y' \in \mathcal{Y}} \exp(u_{\theta}(z,y'))}$$

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Build

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$$\sum_{y' \in \mathcal{X}} \exp(u(z, y')) \approx \sum_{y' \in \mathcal{C}} \exp(u(z, y')) + \sum_{y' \in \mathcal{N}} \frac{|\mathcal{Y} \setminus \mathcal{C}|}{|\mathcal{N}|} \exp(u(z, y'))$$

• $u_{\theta}(z, y) = z^{\top} \mathbf{c}_{y} + b_{y}$ \mathbf{c}_{y} is a deterministic embedding, b_{y} a bias term

Complementary Sum Sampling (CSS) - Summary

The approximation effectively reduces the size of the support of the categorical variable

- in each batch, the support is made of the word types in the batch
- along with a random subset of "negative words"
- this is similar to "negative sampling" but improves on asymptotic behaviour
- it only affects the softmax: the model remains generative

Outline

- 1 Embed-Align
- 2 Bayesian Skip-Gram
- 3 Practica

Bayesian Skip-Gram

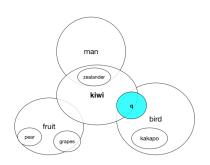


Figure 1: An idealized illustration of density embeddings. Unshaded ellipsoids encode prior densities of Gaussians. The shaded ellipsoid corresponds to the posterior for the word 'kiwi' when it appears in a context indicating that 'kiwi' refers to a bird.

"Representing a word as a distribution provides many potential benefits. For example, such embeddings let us encode generality of terms (e.g., 'kakapo' is a type of 'bird'), characterize uncertainty about semantic properties of the corresponding referent (e.g., a proper noun, such as 'John', encodes little about the person it refers to) or represent polysemy (e.g., 'kiwi' may refer to a fruit, a bird or a New Zealander)."

Bayesian Skip-Gram

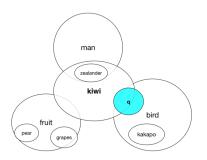


Figure 1: An idealized illustration of density embeddings. Unshaded ellipsoids encode prior densities of Gaussians. The shaded ellipsoid corresponds to the posterior for the word 'kiwi' when it appears in a context indicating that 'kiwi' refers to a bird.

"In principle, using densities to represent words provides a natural way of encoding entailment: the decision regarding entailment relation can be made by testing the level sets of the distributions for 'soft inclusion'. For example, in Figure 1, the ellipse for 'kakapo' lies within the ellipse for 'bird'."

Generative model: for i = 1, ..., m

$$Z_i | x_i \sim \mathcal{N}(\mu_i, \operatorname{diag}(\sigma_i^2))$$

 $\mu_i = \operatorname{location}_{\theta}(x_i)$
 $\sigma_i = \operatorname{softplus}(\operatorname{scale}_{\theta}(x_i))$

Generative model: for i = 1, ..., m

$$Z_i|x_i \sim \mathcal{N}(oldsymbol{\mu}_i, \operatorname{diag}(oldsymbol{\sigma}_i^2))$$
 $oldsymbol{\mu}_i = \operatorname{location}_{ heta}(x_i)$ $oldsymbol{\sigma}_i = \operatorname{softplus}(\operatorname{scale}_{ heta}(x_i))$ for $k \in \mathcal{K}_i = \{\underbrace{i-n, \ldots, i-1, i+1, \ldots, i+n}_{n \text{ words on each side of } x_i}\}$

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$$Z_i|x_i \sim \mathcal{N}(oldsymbol{\mu}_i, \operatorname{diag}(oldsymbol{\sigma}_i^2))$$
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Generative model: for i = 1, ..., m

$$Z_i|x_i \sim \mathcal{N}(oldsymbol{\mu}_i, \operatorname{diag}(oldsymbol{\sigma}_i^2))$$
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Inference model

$$Z_i|x_{i-n}^{i+n} \sim \mathcal{N}(\mathbf{u}_i, \operatorname{diag}(\mathbf{s}_i \odot \mathbf{s}_i))$$
 $\mathbf{h}_i = \sum_{k \in \mathcal{K}_i} \operatorname{relu}(\operatorname{affine}_{\lambda}([x_k, x_i]))$
 $\mathbf{u}_i = \operatorname{affine}_{\lambda}(h_i)$
 $\mathbf{s}_i = \operatorname{softplus}(\operatorname{affine}_{\lambda}(h_i))$

ELBO

$$\mathbb{E}_{q_{\lambda}(z_1^m|x_1^m)}\left[\log\prod_{i=1}^m\prod_{k\in\mathcal{K}_i}P(x_k|z_i)\right]-\mathsf{KL}\left(q_{\lambda}(z_1^m|x_1^m)\mid\mid p_{\theta}(z_1^m|x_1^m)\right)$$

 observation model is trained discriminatively latent variables generate overlapping subsets of observations

ELBO - KL term

$$\mathbb{E}_{q_{\lambda}(z_1^m|x_1^m)}\left[\log\prod_{i=1}^m\prod_{k\in\mathcal{K}_i}P_{\theta}(x_k|z_i)\right]-\mathsf{KL}\left(q_{\lambda}(z_1^m|x_1^m)\mid\mid p_{\theta}(z_1^m|x_1^m)\right)$$

KL term

$$\mathsf{KL}\left(q_{\lambda}(z_1^m|x_1^m)\mid\mid p_{\theta}(z_1^m|x_1^m)\right)$$

ELBO - KL term

$$\mathbb{E}_{q_{\lambda}(z_1^m|x_1^m)}\left[\log\prod_{i=1}^m\prod_{k\in\mathcal{K}_i}P_{\theta}(x_k|z_i)\right]-\mathsf{KL}\left(q_{\lambda}(z_1^m|x_1^m)\mid\mid p_{\theta}(z_1^m|x_1^m)\right)$$

KL term

$$\mathsf{KL}\left(q_{\lambda}(z_1^m|x_1^m) \mid\mid p_{\theta}(z_1^m|x_1^m)\right)$$

$$= \sum_{i=1}^m \mathsf{KL}\left(q_{\lambda}(z_i|x_{i-n}^{i+n}) \mid\mid p_{\theta}(z_i|x_i)\right)$$

ELBO - KL term

$$\mathbb{E}_{q_{\lambda}(z_1^m|x_1^m)}\left[\log\prod_{i=1}^m\prod_{k\in\mathcal{K}_i}P_{\theta}(x_k|z_i)\right]-\mathsf{KL}\left(q_{\lambda}(z_1^m|x_1^m)\mid\mid p_{\theta}(z_1^m|x_1^m)\right)$$

KL term

$$\begin{split} & \mathsf{KL}\left(q_{\lambda}(z_{1}^{m}|\mathsf{x}_{1}^{m}) \mid\mid p_{\theta}(z_{1}^{m}|\mathsf{x}_{1}^{m})\right) \\ &= \sum_{i=1}^{m} \mathsf{KL}\left(q_{\lambda}(z_{i}|\mathsf{x}_{i-n}^{i+n}) \mid\mid p_{\theta}(z_{i}|\mathsf{x}_{i})\right) \\ &= \sum_{i=1}^{m} \mathsf{KL}\left(\underbrace{\mathcal{N}(\mathbf{u}_{i}, \mathsf{diag}(\mathbf{s}_{i} \odot \mathbf{s}_{i}))}_{\mathsf{inference model}} \mid\mid \underbrace{\mathcal{N}(\boldsymbol{\mu}_{i}, \mathsf{diag}(\boldsymbol{\sigma}_{i}^{2}))}_{\mathsf{prior}}\right) \end{split}$$

Empirical Bayes: point estimate prior parameters

$$\mathbb{E}_{q_{\lambda}(z_1^m|x_1^m)}\left[\log\prod_{i=1}^m\prod_{k\in\mathcal{K}_i}P_{\theta}(x_k|z_i)\right]-\mathsf{KL}\left(q_{\lambda}(z_1^m|x_1^m)\mid\mid p_{\theta}(z_1^m|x_1^m)\right)$$

$$\mathbb{E}_{q_{\lambda}(z_1^m|x_1^m)}\left|\log\prod_{i=1}^m\prod_{k\in\mathcal{K}_i}P_{\theta}(x_k|z_i)\right|-\mathsf{KL}\left(q_{\lambda}(z_1^m|x_1^m)\mid\mid p_{\theta}(z_1^m|x_1^m)\right)$$

$$\mathbb{E}_{q_{\lambda}(z_1^m|x_1^m)}\left[\log\prod_{i=1}^m\prod_{k\in\mathcal{K}_i}P_{\theta}(x_k|z_i)\right]$$

$$\mathbb{E}_{q_{\lambda}(z_1^m|x_1^m)}\left[\log\prod_{i=1}^m\prod_{k\in\mathcal{K}_i}P_{\theta}(x_k|z_i)\right]-\mathsf{KL}\left(q_{\lambda}(z_1^m|x_1^m)\mid\mid p_{\theta}(z_1^m|x_1^m)\right)$$

$$\mathbb{E}_{q_{\lambda}(z_1^m|x_1^m)}\left[\log\prod_{i=1}^m\prod_{k\in\mathcal{K}_i}P_{\theta}(x_k|z_i)\right] = \mathbb{E}_{q_{\lambda}(z_1^m|x_1^m)}\left[\sum_{i=1}^m\sum_{k\in\mathcal{K}_i}\log P_{\theta}(x_k|z_i)\right]$$

$$\mathbb{E}_{q_{\lambda}(z_1^m|x_1^m)}\left[\log\prod_{i=1}^m\prod_{k\in\mathcal{K}_i}P_{\theta}(x_k|z_i)\right]-\mathsf{KL}\left(q_{\lambda}(z_1^m|x_1^m)\mid\mid p_{\theta}(z_1^m|x_1^m)\right)$$

$$\mathbb{E}_{q_{\lambda}(z_{1}^{m}|x_{1}^{m})}\left[\log\prod_{i=1}^{m}\prod_{k\in\mathcal{K}_{i}}P_{\theta}(x_{k}|z_{i})\right] = \mathbb{E}_{q_{\lambda}(z_{1}^{m}|x_{1}^{m})}\left[\sum_{i=1}^{m}\sum_{k\in\mathcal{K}_{i}}\log P_{\theta}(x_{k}|z_{i})\right]$$
$$= \sum_{i=1}^{m}\sum_{k\in\mathcal{K}_{i}}\mathbb{E}_{q_{\lambda}(z_{i}|x_{i-n}^{i+n})}\left[\log P_{\theta}(x_{k}|z_{i})\right]$$

$$\mathbb{E}_{q_{\lambda}(z_1^m|x_1^m)}\left[\log\prod_{i=1}^m\prod_{k\in\mathcal{K}_i}P_{\theta}(x_k|z_i)\right]-\mathsf{KL}\left(q_{\lambda}(z_1^m|x_1^m)\mid\mid p_{\theta}(z_1^m|x_1^m)\right)$$

$$\begin{split} \mathbb{E}_{q_{\lambda}(z_{1}^{m}|x_{1}^{m})} \left[\log \prod_{i=1}^{m} \prod_{k \in \mathcal{K}_{i}} P_{\theta}(x_{k}|z_{i}) \right] &= \mathbb{E}_{q_{\lambda}(z_{1}^{m}|x_{1}^{m})} \left[\sum_{i=1}^{m} \sum_{k \in \mathcal{K}_{i}} \log P_{\theta}(x_{k}|z_{i}) \right] \\ &= \sum_{i=1}^{m} \sum_{k \in \mathcal{K}_{i}} \mathbb{E}_{q_{\lambda}(z_{i}|x_{i-n}^{i+n})} \left[\log P_{\theta}(x_{k}|z_{i}) \right] \\ &= \sum_{i=1}^{m} \sum_{k \in \mathcal{K}_{i}} \mathbb{E}_{q_{\lambda}(z_{i}|x_{i-n}^{i+n})} \left[\log \operatorname{Cat}(x_{k}|\mathbf{f}_{i}) \right] \end{split}$$

The softmax problem

To circumvent an expensive softmax, change the likelihood term

$$P_{ heta}(x|z) = rac{s_{ heta}(z,x)}{\sum_{x'\in\mathcal{X}}s_{ heta}(z,x')} \qquad ext{with } s_{ heta}(\cdot,\cdot) > 0$$

The softmax problem

To circumvent an expensive softmax, change the likelihood term

$$P_{ heta}(x|z) = rac{s_{ heta}(z,x)}{\sum_{x'\in\mathcal{X}}s_{ heta}(z,x')} \qquad ext{with } s_{ heta}(\cdot,\cdot) > 0$$

and re-write the likelihood term

$$\begin{split} \mathbb{E}_{q_{\lambda}(z)}\left[\log P_{\theta}(x|z)\right] &= \mathbb{E}_{q_{\lambda}(z)}\left[\log s_{\theta}(z,x) - \log \sum_{x' \in \mathcal{X}} s_{\theta}(z,x')\right] \\ &= \mathbb{E}_{q_{\lambda}(z)}\left[\log s_{\theta}(z,x)\right] - \mathbb{E}_{q_{\lambda}(z)}\left[\log \sum_{x' \in \mathcal{X}} s_{\theta}(z,x')\right] \end{split}$$

Lowerbound on likelihood term

Then (by design) let

$$s_{ heta}(z,x) = \underbrace{P(x)}_{ ext{fixed}} \mathcal{N}(z|oldsymbol{\mu}_x, ext{diag}(oldsymbol{\sigma}_x^2))$$

Lowerbound on likelihood term

Then (by design) let

$$s_{\theta}(z,x) = \underbrace{P(x)}_{\text{fixed}} \mathcal{N}(z|\mu_x, \text{diag}(\sigma_x^2))$$

And bound $\mathbb{E}_{q_{\lambda}(z)}\left[\log\sum_{x'\in\mathcal{X}}s_{\theta}(z,x')\right]$

$$\begin{split} &= \mathbb{E}_{q_{\lambda}(z)} \left[\log \sum_{x' \in \mathcal{X}} P(x) \mathcal{N}(z | \mu_{x}, \operatorname{diag}(\sigma_{x}^{2})) \right] \\ &= \mathbb{E}_{q_{\lambda}(z)} \left[\log \mathbb{E}_{P(x)} \left[\mathcal{N}(z | \mu_{x}, \operatorname{diag}(\sigma_{x}^{2})) \right] \right] \\ &\geq \mathbb{E}_{q_{\lambda}(z)} \left[\mathbb{E}_{P(x)} \left[\log \mathcal{N}(z | \mu_{x}, \operatorname{diag}(\sigma_{x}^{2})) \right] \right] \end{split}$$

• P(x) does not depend on θ , we can compute an MC estimate e.g. empirical (unigram) distribution

Outline

- 1 Embed-Align
- 2 Bayesian Skip-Gram
- Practical

Practical

- Skip-gram
- Bayesian skip-gram
- Embed-Align

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(Mikolov et al., 2013)
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(Bražinskas et al., 2017)

(Rios et al., 2018)

Comparison

Model	LVM	Gen.	Prior	Softmax
SkipGram				Neg sampling
BSG	\checkmark		type-specific	JI
EmbedAlign	\checkmark	\checkmark	general	CSS

Literature I

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