

Probability and Statistics: Lecture-16

Monsoon-2020

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» Example of Poisson Distribution...

$n \gg 1$ Binomial (n, p)

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$p \ll 1$

$\begin{cases} n \rightarrow \infty \\ p \rightarrow \text{small} \end{cases}$

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$$P(X=0) = \frac{\lambda^0}{0!} e^{-\lambda} = \frac{0.01^0}{0!} e^{-0.01} \approx 0.9900498$$

- * We could have solved this using Binomial $X \sim \text{Binomial}(10^4, 10^{-6})$. This may have been harder to compute!

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$$\begin{aligned} P(X = 0) &= \frac{\lambda^i}{i!} e^{-\lambda} \\ &= \frac{0.01^0}{0!} e^{-0.01} \\ &\approx 0.9900498 \end{aligned}$$

- * We could have solved this using Binomial $X \sim \text{Binomial}(10^4, 10^{-6})$. This may have been harder to compute!
- * When n large, and p small: can use Poisson!

» Another Example of Poisson Distribution...

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The Poisson distribution is often used to model the number of events that occur independently at any time in an interval of time or space, with a **constant** average rate. Earthquakes are a good example of this. Suppose there are an average of 2.8 major earthquakes in the world each year. What is the probability of getting more than one major earthquake next year?

$$\begin{aligned} X &\sim \text{Poisson}(2.8), \quad X = \# \text{ earthquakes} \\ P(X > 1) \\ &= 1 - P(X = 0) - P(X = 1) \end{aligned}$$

Handwritten notes and diagrams:

- A box around $X > 1$ with an arrow pointing to it from the text $X \sim \text{Poisson}(2.8)$.
- A diagram showing a horizontal line with tick marks representing values 0, 1, 2, The tick mark for 0 is circled, and the tick mark for 1 is marked with an 'X'. The text $= 1 - P(X=0) - P(X=1)$ is written below the line, with arrows pointing to the 0 and 1 tick marks.

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- * Let $X \sim \text{Poisson}(2.8)$ be the number of major earthquakes next year. We want $P(X > 1)$.

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- * We can use

$$\underline{P(X > 1)} = 1 - \underline{P(X = 0)} - \underline{P(X = 1)}.$$

Using PMF for Poisson

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Using PMF for Poisson

$$\begin{aligned}P(X > 1) &= 1 - P(X = 0) - P(X = 1). \\P(X > 1) &= 1 - P(X = 0) - P(X = 1) \\&= 1 - e^{-2.8} \frac{2.8^0}{0!} - e^{-2.8} \frac{2.8^1}{1!} \\&= 1 - e^{-2.8} - 2.8e^{-2.8} \\&\approx 1 - 0.06 - 0.17 = 0.77\end{aligned}$$

77%

» Expectation of Poisson Distribution...

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Expectation of Poisson

Let $X \sim \text{Poisson}(\lambda)$, with PMF given by

$$p_X(k) = \begin{cases} \frac{e^{-\lambda} \lambda^k}{k!}, \\ 0 \end{cases}$$

$k \in \mathbb{R}_X$
otherwise

$$E[X] = \sum x p(x)$$

Show that the expectation $E[X] = \lambda$.

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Let $X \sim \text{Poisson}(\lambda)$, with PMF given by

$$P_X(k) = \begin{cases} \frac{e^{-\lambda} \lambda^k}{k!}, & k \in R_X \\ 0 & \text{otherwise} \end{cases}$$

Show that the expectation $E[X] = \lambda$.

Proof

$$E[X] = \sum_{x=0}^{\infty} x \frac{e^{-\lambda} \lambda^x}{x!} = \sum_{x=1}^{\infty} x \frac{e^{-\lambda} \lambda^x}{x!} = e^{-\lambda} \sum_{x=1}^{\infty} \frac{\lambda^x}{(x-1)!}$$

$$= e^{-\lambda} \sum_{z=0}^{\infty} \frac{\lambda^{z+1}}{z!} = \lambda e^{-\lambda} \sum_{z=0}^{\infty} \frac{\lambda^z}{z!} = \lambda$$

Recall Calculus / Real Analysis

If $f: \mathbb{R} \rightarrow \mathbb{R}$ is infinitely diff., then f can be expanded around $x=a$

$$f(x) = f(a) + \frac{f'(a)}{1!}(x-a) + \frac{f''(a)}{2!}(x-a)^2 + \dots$$

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$x=0$

$$e^x = e^0 + \frac{e^0}{1!}x + \frac{e^0}{2!}x^2 + \dots$$

Taylor's expansion

» Expectation of Binomial Distribution...

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Expectation of a Binomial Distribution

Show that $E[X] = np$, where $X \sim \text{Binomial}(n, p)$, $0 < p < 1$ with PMF given as follows

$$P_X(k) = \begin{cases} \binom{n}{k} p^k (1-p)^{n-k} & \text{for } k = 0, 1, 2, \dots, n \\ 0 & \text{otherwise,} \end{cases}$$

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Proof

$$E[X] = \sum_{x=0}^n x \binom{n}{x} p^x q^{n-x} = \sum_{x=1}^n x \binom{n}{x} p^x q^{n-x}.$$

Handwritten notes: A '0' is written above the sum. 'x PMF' is written below the binomial coefficient.

$$\left[\text{For } 0 < x \leq n, \binom{n}{x} = \frac{n!}{(n-x)!x!} = \frac{n!}{(n-x)!(x-1)!} = n \binom{n-1}{x-1} \right]$$

Handwritten notes: The binomial coefficient and the final term are circled in orange. Arrows indicate the cancellation of $n!$ and the shift of indices.

$$\Rightarrow E[X] = \sum_{x=1}^n n \binom{n-1}{x-1} p^x q^{n-x} = \sum_{z=0}^{n-1} n \binom{n-1}{z} p^{z+1} q^{n-1-z} = np.$$

Handwritten notes: The binomial coefficient $\binom{n-1}{z}$ is circled in orange. The final result np is circled in orange. A note 'take n & p common' with an arrow points to the n and p in the sum. Another note 'z-1-->0' is written below the sum.

Motivation for Variance

- * The expectation value $E(X)$ (also called the mean) of a discrete random variable is a rather **coarse** measure of how X is distributed. For example, consider the following three situations:

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- * Let X be the discrete random variable corresponding to your winnings. In all three cases, $E(X) = \text{Rs } 1000$, but you will agree that **your chances of actually getting any money are quite different in all three cases.**

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- * The three cases differ in the spread of PMF. This can be measured by **Variance**.

» Variance and Standard Deviation...

Variance

Let $\mu = E[X]$. The **variance** $\text{Var}(X)$ of X is defined by

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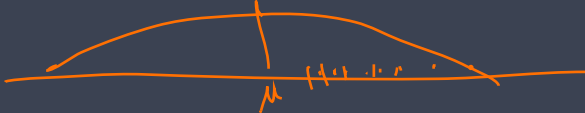
$$E[\underbrace{g(x)}] = \sum_x g(x) p_x(x)$$

Variance

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$$\text{Var}(X) = E(\underbrace{(X - \mu)^2}_{\rightarrow g(x)}) = \sum_x \underbrace{(x - \mu)^2}_{\leftarrow \text{sometimes use } p_x \text{ or } f_x} \underbrace{f_X(x)}$$

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Variance

Let $\mu = E[X]$. The variance $\text{Var}(X)$ of X is defined by

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The positive square root is called the standard deviation and it is usually denoted by σ , hence


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$$\sigma(X) = \sqrt{\sum_x (x - \mu)^2 f_X(x)} = \sqrt{\text{Var}(X)}.$$

» Example of Computing Variance...

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$$R_X = \{1, 2, 3, 4, 5, 6\}$$

Example

Let X be the value on one roll of a 6-sided die. Recall that $E[X] = 7/2$. What is $\text{Var}(X)$?

$$\text{Var}(X) = E[(X - \mu)^2]$$

$$= \sum_x (x - 7/2)^2 \underbrace{P_X(x)}$$

$$= \sum_x (x - 7/2)^2 \cdot \frac{1}{6}$$

$$= \left(1 - 7/2\right)^2 \cdot \frac{1}{6} + \left(2 - 7/2\right)^2 \cdot \frac{1}{6} + \dots + \left(6 - 7/2\right)^2 \cdot \frac{1}{6}$$

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Calculate Variance

- * **First case:** There is only one outcome and it is the mean, variance is 0

» Variance

$$R_T = \begin{cases} 0 \\ 2000 \end{cases}$$

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- * **First case:** There is only one outcome and it is the mean, variance is 0
- * **Second case:** $\text{Var}(X) = \frac{1}{2}(\underbrace{2000 - 1000}_{n - \mu})^2 + \frac{1}{2}(\underbrace{0 - 1000}_{\mu - n})^2 = 10^6$

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Calculate Variance

- * **First case:** There is only one outcome and it is the mean, variance is 0
- * **Second case:** $\text{Var}(X) = \frac{1}{2}(2000 - 1000)^2 + \frac{1}{2}(0 - 1000)^2 = 10^6$
- * **Third case:** $\text{Var}(X) = 10^{-3}(10^6 - 10^3)^2 + 999 \times 10^{-3}(0 - 10^3)^2 \approx 10^9$

» Remarks on Expectation, Variance, Standard Deviation...

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Same Standard Deviation Does Not Imply Same Data set

For example, the data sets 199, 200, 201 and 0, 200, 400 both have the same average (200) yet they have very different standard deviations.

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Same Standard Deviation Does Not Imply Same Dataset

For example, the data sets 199, 200, 201 and 0, 200, 400 both have the same average (200) yet they have very different standard deviations. The first data set has a very small standard deviation ($s=1$) compared to the second data set ($s=200$).

» Another expression for the variance

Theorem (Another Expression for Variance)

If X is a discrete random variable with mean μ , then

$$\text{Var}(X) = \underbrace{E[X^2]} - \underbrace{\mu^2}$$

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Theorem (Another Expression for Variance)

If X is a discrete random variable with mean μ , then

$$\text{Var}(X) = E[X^2] - \mu^2$$

Proof

$$\begin{aligned}\text{Var}(X) &= \sum_x (x - \mu)^2 p_X(x) = \sum_x (x^2 - 2\mu x + \mu^2) p_X(x) \\ &= \sum_x x^2 p_X(x) - 2\mu \sum_x x p_X(x) + \mu^2 \sum_x p_X(x) \\ &= E[X^2] - 2\mu E[X] + \mu^2 = E[X^2] - \mu^2\end{aligned}$$

» Properties of Variance...

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Theorem

Let X be a discrete random variable and α a constant. Then

$$\text{Var}(\alpha X) = \alpha^2 \text{Var}(X) \quad \text{and} \quad \text{Var}(X + \alpha) = \text{Var}(X)$$

$$E[X] = \mu$$

$$E[\alpha X] = \alpha \mu$$

$$\begin{aligned} \rightarrow \text{Var}(\alpha X) &= E[(\alpha X)^2] - \frac{(\alpha \mu)^2}{\alpha^2} \\ &= \alpha^2 E[X^2] - \alpha^2 \mu^2 \\ &= \alpha^2 (E[X^2] - \mu^2) \\ &= \alpha^2 \text{Var}(X) \end{aligned}$$

$$\rightarrow \text{Var}(X + \alpha) = E[(X + \alpha)^2] - \frac{(E[X + \alpha])^2}{1}$$

$$\begin{aligned} &= E[X^2 + 2\alpha X + \alpha^2] - (\mu + \alpha)^2 \\ &= E[X^2] + 2\alpha E[X] + \alpha^2 - (\mu^2 + 2\alpha\mu + \alpha^2) \\ &= E[X^2] + 2\alpha\mu + \alpha^2 - \mu^2 - 2\alpha\mu - \alpha^2 \\ &= E[X^2] - \mu^2 \\ &= \text{Var}(X) \end{aligned}$$

$$\text{Var}(X + \alpha) = \text{Var}(X)$$

» Computing Variance: Binomial

Binomial $\binom{n}{k}$

$$E[X] = np$$

Variance of Binomial Distribution

Let $X \sim \text{Binomial}(n, p)$. Then the variance $\text{Var}(X) = np(1 - p)$.

$$\text{Var}(X) = E(X^2) - \mu^2 = E[X(X-1)] + E[X] - E[X]^2$$

$$E[X(X-1)] = \sum_{x=0}^n \underbrace{x(x-1)} \binom{n}{x} p^x (1-p)^{n-x}$$

$$\sum_{x=2}^n \frac{n!}{\underbrace{(x-2)!} (n-x)!} p^x (1-p)^{n-x}$$

» Variance of Binomial Distribution...

$$= \sum_{x=2}^n \frac{n!}{(x-2)! (n-x)!} p^x (1-p)^{n-x}$$

$$= n(n-1)p^2 \sum_{x=2}^n \frac{(n-2)!}{(x-2)! (n-x)!} p^{x-2} (1-p)^{n-x}$$

$$\begin{cases} m = n-2 \\ y = x-2 \end{cases} = n(n-1)p^2 \sum \frac{m!}{y! (m-y)!} p^y (1-p)^{m-y}$$

$$= \underline{n(n-1)p^2} (p + (1-p))^m = n(n-1)p^2 = E[X(X-1)]$$

» Variance of Binomial Distribution...

$$\begin{aligned}\text{Var}(X) &= E[\underline{X(X-1)}] + E[X] - (E[X])^2 \\ &= n(n-1)p^2 + np - (np)^2 \\ &= np(1-p) \\ &= \end{aligned}$$

» Higher Order Moments...

Define n th moment

The n th moment about the mean or n th central moment of a real valued random variable X is defined as follows

$$\mu_n = E[(X - E[X])^n],$$

where E is the expectation operator.

$$E[X]$$

$$\rightarrow E[X - \mu]$$

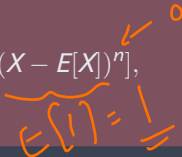
$$E[(X - \mu)^2]$$

$$E[(X - \mu)^3]$$

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where E is the expectation operator.

Handwritten notes in orange ink showing the derivation of the n th central moment formula. It starts with $E[(X - \mu)^n]$ and then shows $E[X] - \mu = 0$, where μ is written below $E[X]$.

- * The **zeroth** central moment μ_0 is 1
- * The **first** central moment μ_1 is 0 (not the same as expected value!)

» Higher Order Moments...

Define n th moment

The n th moment about the mean or n th central moment of a real valued random variable X is defined as follows

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Handwritten orange notation: $E[(X - \mu)^n]$. An arrow points from the μ in the formula to the $E[X]$ in the typed formula above.

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Generating Moments...

Is there a quick way to generate moments?

» Moment Generating Function...

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$$e^{tX} = \underline{g(x)}$$

Moment Generating Function

The moment generating function $M_X(t)$ is the expectation value

$$M_X(t) = E[e^{tX}] = \sum_x e^{tx} p_X(x)$$

$$\sum p_X(x) = 1$$

Lemma

$$* M_X(0) = 1$$

$$* E[X] = M'_X(0), \text{ where ' is the derivative w.r.t. } t$$

$$\frac{d}{dt} M_X(t) = \sum_x \frac{d}{dt} e^{tx} p_X(x)$$

$$= \sum x p_X(x) = \underline{\mu} = E[X]$$