# CSCI3230 (ESTR3108) Fundamentals of Artificial Intelligence

### Tutorial 1

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### Outline

Part 1. Review linear algebra

Part 2. Least squares

Part 3. Bias-variance decomposition



Part 1. Review linear algebra

#### Vector

• Notation for vectors in  $\mathbb{R}^n$ :

$$X = \begin{pmatrix} x_1 \\ \vdots \\ x_n \end{pmatrix} \quad \underbrace{\text{column vector}}$$

$$X' = (x_1, \cdots, x_n)$$
 row vector

Transpose: column vector ↔ row vector

$$X' = X^T, \quad X'^T = X$$

 In the context of vectors, column vectors and row vectors are the same. But when putting them with matrices, they are totally different.

#### Vectors

- Scaling:  $a \in \mathbb{R}$ ,  $X = (x_1, \dots, x_n) \in \mathbb{R}$ ,  $aX = (ax_1, \dots, ax_n)$
- Addition:  $X = (x_1, \dots, x_n)$ ,  $Y = (y_1, \dots, y_n)$ ,  $X + Y = (x_1 + y_1, \dots, x_n + y_n)$
- Suppose we have m vectors:  $X_1 ... X_m$ , linear combination of these vectors: for any m scalars  $a_1 ... a_m$ ,

$$a_1X_1 + \cdots + a_mX_m$$

• Inner product (dot product):  $\forall X \in \mathbb{R}^n, Y \in \mathbb{R}^n$ 

$$X \cdot Y = X^T Y = \sum_{i=1}^n x_i y_i$$

• Euclidean norm:  $\|X\|_2 = \sqrt{X^T X} = \sqrt{\sum_{i=1}^n x_i^2}$ 

### Matrix

• Notation for  $m \times n$  matrices:

$$\mathbf{A} = \begin{pmatrix} a_{1,1} & \cdots & a_{1,n} \\ \vdots & \ddots & \vdots \\ a_{m,1} & \cdots & a_{m,n} \end{pmatrix}$$

 $a_{i,j}$  is the element at the i-th row, j-th column of A.

- Denote the i-th column of A as  $A_i$  (column vectors), and j-th row of A as  $A^{(j)}$  (row vectors).
- Transpose: columns ↔ rows

$$\mathbf{A}^T = \begin{pmatrix} a_{1,1} & \cdots & a_{m,1} \\ \vdots & \ddots & \vdots \\ a_{1,n} & \cdots & a_{m,n} \end{pmatrix} = \begin{bmatrix} A^{(1)^T}, \dots, A^{(m)^T} \end{bmatrix}$$

### Matrix

- ullet n-dimensional row vector:  $1 \times n$  matrix
- n-dimensional column vector:  $n \times 1$  matrix  $n \in \mathbb{N}$

• Scaling: 
$$c \in \mathbb{R}$$
,  $\mathbf{A} = \begin{pmatrix} a_{1,1} & \cdots & a_{1,n} \\ \vdots & \ddots & \vdots \\ a_{m,1} & \cdots & a_{m,n} \end{pmatrix}$ ,  $c\mathbf{A} = \begin{pmatrix} ca_{1,1} & \cdots & ca_{1,n} \\ \vdots & \ddots & \vdots \\ ca_{m,1} & \cdots & ca_{m,n} \end{pmatrix}$ 

 ${\bf A}$  and  ${\bf B}$  have the same shape:  $\mathbb{R}^{m \times n}$ 

# Special matrices

Diagonal matrix in  $\mathbb{R}^{n\times n}$ :

$$\Lambda = \begin{pmatrix} \Lambda_1 & 0 & \cdots & 0 \\ 0 & \Lambda_2 & \cdots & 0 \\ \vdots & \vdots & & \vdots \\ 0 & 0 & \cdots & \Lambda_n \end{pmatrix} \qquad \begin{array}{c} \mathbb{Z} \cdot 3 = 3 \\ \mathbb{Z} \cdot A = A \end{array}$$

- Identity matrix in  $\mathbb{R}^{n \times n}$ :  $I = \begin{pmatrix} 1 & 0 & \cdots & 0 \\ 0 & 1 & \cdots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \cdots & 1 \end{pmatrix}$
- Symmetric matrix:  $A = A^T$

### Matrix-vector multiplication

Matrix-vector multiplication

$$\mathbf{A}X = \begin{pmatrix} a_{1,1} & \cdots & a_{1,n} \\ \vdots & \ddots & \vdots \\ a_{m,1} & \cdots & a_{m,n} \end{pmatrix} \begin{pmatrix} x_1 \\ \vdots \\ x_n \end{pmatrix}$$

$$= \begin{pmatrix} a_{1,1} \\ \vdots \\ a_{m,1} \end{pmatrix} x_1 + \begin{pmatrix} a_{1,2} \\ \vdots \\ a_{m,2} \end{pmatrix} x_2 + \cdots \begin{pmatrix} a_{1,n} \\ \vdots \\ a_{m,n} \end{pmatrix} x_n$$

$$= \mathbf{A}_1 x_1 + \cdots + \mathbf{A}_n x_n$$

- Notice the shape:  $\mathbf{A} \in \mathbb{R}^{m \times n}$ ,  $X \in \mathbb{R}^n$ ,  $\mathbf{A}X \in \mathbb{R}^m$ .
- Matrix-vector multiplication = linear combination of A's columns.

### Matrix-vector multiplication

Another perspective of matrix-vector multiplication:

Matrix-vector multiplication

$$\begin{aligned} \mathbf{A}X &= \begin{pmatrix} \mathbf{A}^{(1)} \\ \vdots \\ \mathbf{A}^{(m)} \end{pmatrix} X \\ &= \begin{pmatrix} \mathbf{A}^{(1)} \cdot X \\ \vdots \\ \mathbf{A}^{(m)} \cdot X \end{pmatrix}^{2} \text{ Sinfer} \end{aligned}$$

Recall that  $A^{(i)}$  is the i-th row of A.

• Matrix-vector multiplication = dot product of  $\mathbf{A}$ 's rows and X.

# Matrix-matrix multiplication

• Matrix-matrix multiplication

$$\mathbf{AX} = \begin{pmatrix} a_{1,1} & \cdots & a_{1,n} \\ \vdots & \ddots & \vdots \\ a_{m,1} & \cdots & a_{m,n} \end{pmatrix} \begin{pmatrix} x_{1,1} & \cdots & x_{1,l} \\ \vdots & \ddots & \vdots \\ x_{n,1} & \cdots & x_{n,l} \end{pmatrix}$$

• Let  $X_i$  be the i-th column of X, the i-th column of AX is  $AX_i$ :

$$\mathbf{AX} = [\mathbf{A}X_1, \dots, \mathbf{A}X_l]$$

- Notice the shape:  $\mathbf{A} \in \mathbb{R}^{m \times n}$ ,  $\mathbf{X} \in \mathbb{R}^{n \times l}$ ,  $\mathbf{A}\mathbf{X} \in \mathbb{R}^{m \times l}$ .
- $\bullet$  Multiplication for block matrices:  $\mathbf{A} = [\mathbf{A}_1, \mathbf{A}_2]$

$$\mathbf{B}\mathbf{A} = \mathbf{B}[\mathbf{A}_1, \mathbf{A}_2] = [\mathbf{B}\mathbf{A}_1, \mathbf{B}\mathbf{A}_2]$$

#### Inverse matrix

ullet For squared matrix  $\mathbf{A} \in \mathbb{R}^{n \times n}$ , if there is a  $\mathbf{B} \in \mathbb{R}^{n \times n}$  such that

$$AB = I = BA$$

A is invertible and  $A^{-1} = B$ 

- How to find inverse matrix of A (if invertible):
  - $\textbf{ Gauss-Jordan elimination: find a row-operating matrix } \mathbf{B} \text{ which transforms } \mathbf{A} \text{ to } \mathbf{I}. \text{ (More feasible for human begings)}$
  - 2 Use eign-decomposition:  $\mathbf{A} = U\Lambda U^{-1}$ ,  $\mathbf{A}^{-1} = U\Lambda^{-1}U^{-1}$
  - Use the analytic solution:

$$\mathbf{A}^{-1} = \frac{1}{|\mathbf{A}|} \mathbf{C}^T$$

where C is the adjugate matrix of A.

### Matrix calculus

- Univariate real-valued function  $f(x): \mathbb{R} \to \mathbb{R}$
- Multi-variate real-valued function  $f(x_1, \dots, x_n) : \mathbb{R}^n \to \mathbb{R}$ . For convenience, we use  $X = (x_1, \dots, x_n)^T$  to represent the independent variable. So we can write it as  $f(X): \mathbb{R}^n \to \mathbb{R}$ .
- We use denominator layout to find the derivative w.r.t. X.

Reference: Wikipedia - Matrix calculus

### Matrix calculus

$$\forall X \in \mathbb{R}^n, Y \in \mathbb{R}^n$$
:

$$\frac{\partial Y^T X}{\partial X} = Y$$

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$$\frac{\partial X^T Y}{\partial X} = X$$

$$\frac{9x}{9\lambda_{1}x} = \left(\frac{9x!}{9\lambda_{1}x} \cdots \frac{9x^{n}}{9\lambda_{1}x}\right) = \lambda$$

$$\lambda_{1}x = \sum_{i=1}^{3} \lambda_{i}x_{i} = \lambda_{i}x_{i} + \lambda_{i}y_{i} + \cdots + \lambda_{i}x_{i}$$

### Matrix calculus

$$\forall X \in \mathbb{R}^n, \mathbf{A} \in \mathbb{R}^{n \times n}$$
:

$$f(x): X^{T}AX$$

$$\frac{\partial X^T \mathbf{A} X}{\partial X} = (\mathbf{A} + \mathbf{A}^T) X \qquad \text{f(x): } \underline{\mathbf{a}} \mathbf{x}^\mathbf{1} + \underline{\mathbf{b}} \mathbf{x} + \underline{\mathbf{c}}$$

f'(x): 2ax + b



$$\frac{\partial X^{T} A X}{\partial X_{I}} = 2 \sum_{A \in I} A_{I} X_{I}$$

$$(A + A^{T}) X$$

## Linear algebra materials

- Introduction to Linear Algebra. Gilbert Strang.
- Matrix Analysis and Applied Linear Algebra. Carl D. Meyer.
- Advanced Linear Algebra. Steven Roman.
- Linear Algebra and Its Applications. Manolis C. Tsakiris.



# Part 2. Least squares

## Problem settings

• Recall that our data matrix X and observed labels Y:

$$\mathbf{X} = \begin{pmatrix} 1 & x_1^{(1)} & \dots & x_n^{(1)} \\ \vdots & \vdots & \ddots & \vdots \\ 1 & x_1^{(m)} & \dots & x_n^{(m)} \end{pmatrix} \quad Y = \begin{pmatrix} y^{(1)} \\ \vdots \\ y^{(m)} \end{pmatrix}$$

ullet We use linear function to fit a linear mapping from  ${f X}$  to Y:

$$\hat{Y} = \mathbf{X}\Theta$$

where 
$$\Theta = \begin{pmatrix} \theta_0 \\ \theta_1 \\ \vdots \\ \theta_n \end{pmatrix}$$
, such that  $\hat{Y}$  is very close to  $Y$ .

## Ordinary least squares

- $\bullet$  Recall that we use  $\|\hat{Y}-Y\|_2^2$  to measure the distance between real labels and estimated labels.
- Note that:

$$\|\hat{Y} - Y\|_2^2 = \sum_{i=0}^m (\hat{y}^{(i)} - y^{(i)})^2$$

which is exactly the residual sum of squares (RSS).

• Ordinary least squares (OLS) estimator:

$$\hat{\Theta} = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T Y$$

is the optimal solution that minimizes  $\|\hat{Y} - Y\|_2$ .

#### Derivation

#### Analytic solution:

$$\begin{split} J(\Theta) &= \| \hat{f}_{\Theta}(\mathbf{X}) - Y \|_2^2 = (\mathbf{X}\Theta - Y)^T (\mathbf{X}\Theta - Y) \\ &= \Theta^T \mathbf{X}^T \mathbf{X}\Theta - Y^T \mathbf{X}\Theta - \Theta^T \mathbf{X}^T Y - Y^T Y \\ &\frac{\partial J(\Theta)}{\partial \Theta} = 2 \mathbf{X}^T \mathbf{X}\Theta - \mathbf{X}^T Y - \mathbf{X}^T Y \\ &= 2 \mathbf{X}^T (\mathbf{X}\Theta - Y) = 0 \quad \cdots \cdots \quad (1) \\ &\frac{\partial^2 J(\Theta)}{\partial \Theta^2} = 2 \mathbf{X}^T \mathbf{X} \succeq 0 \quad \text{is for true.} \end{split}$$
 By (1),  $\mathbf{X}^T \mathbf{X}\Theta = \mathbf{X}^T Y \implies \hat{\Theta} = \Theta^* = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T Y$ 

### Example

We have 3 samples with bi-variate features:

$$X^{(1)} = \begin{pmatrix} 0 \\ 1 \end{pmatrix} \quad X^{(2)} = \begin{pmatrix} 1 \\ 0 \end{pmatrix} \quad X^{(3)} = \begin{pmatrix} 0 \\ 0 \end{pmatrix}$$

with corresponding observed labels:  $y^{(1)}=1, y^{(2)}=2, y^{(3)}=3$ . Suppose we use linear regression model to predict the relationship between the featuers and labels. Please use OLS to find the  $\hat{\Theta}$ .

# Example (cont.)

# Example (cont.)

# Example (cont.)



Part 3. Bias-variance decomposition

# Problem settings

- Features: X. True relationship: f(X).
- ullet Observed labels: y=f(X)+arepsilon, where arepsilon is a noise, with  $\mathrm{E}(arepsilon)=0$ .
- ullet  $\hat{f}(X)$  is the estimate of f(X) by some estimators, e.g., OLS.
- Bias-variance decomposition:

$$\mathrm{E}\left[(y-\hat{f})^2\right] = \mathrm{Bias}[\hat{f}]^2 + \sigma^2 + \mathrm{Var}\left[\hat{f}\right]$$

which says the expectation of errors between observed labels and estimated labels is sum of squared bias, squared irreducible errors, and the variance of estimated relationship.

#### Derivation

$$\begin{split} \mathbf{E}\left[(y-\hat{f})^2\right] &= \mathbf{E}\left[(f+\varepsilon-\hat{f})^2\right] \\ &= \mathbf{E}\left[(f+\varepsilon-\hat{f}+\mathbf{E}[\hat{f}]-\mathbf{E}[\hat{f}])^2\right] \\ &= \mathbf{E}\left[(f-\mathbf{E}[\hat{f}])^2\right] + \mathbf{E}[\varepsilon^2] + \mathbf{E}\left[(\mathbf{E}[\hat{f}]-\hat{f})^2\right] + 2\,\mathbf{E}\left[(f-\mathbf{E}[\hat{f}])\varepsilon\right] \\ &+ 2\,\mathbf{E}\left[\varepsilon(\mathbf{E}[\hat{f}]-\hat{f})\right] + 2\,\mathbf{E}\left[(\mathbf{E}[\hat{f}]-\hat{f})(f-\mathbf{E}[\hat{f}])\right] \\ &= (f-\mathbf{E}[\hat{f}])^2 + \mathbf{E}[\varepsilon^2] + \mathbf{E}\left[(\mathbf{E}[\hat{f}]-\hat{f})^2\right] + 2(f-\mathbf{E}[\hat{f}])\,\mathbf{E}[\varepsilon] \\ &+ 2\,\mathbf{E}[\varepsilon]\,\mathbf{E}\left[\mathbf{E}[\hat{f}]-\hat{f}\right] + 2\,\mathbf{E}\left[\mathbf{E}[\hat{f}]-\hat{f}\right](f-\mathbf{E}[\hat{f}]) \end{split}$$

#### Derivation

$$\begin{split} &\mathbf{E}\left[(y-\hat{f})^2\right] = \mathbf{E}\left[(f+\varepsilon-\hat{f})^2\right] \\ &= \mathbf{E}\left[(f+\varepsilon-\hat{f}+\mathbf{E}[\hat{f}]-\mathbf{E}[\hat{f}])^2\right] \\ &= \mathbf{E}\left[(f-\mathbf{E}[\hat{f}])^2\right] + \mathbf{E}[\varepsilon^2] + \mathbf{E}\left[(\mathbf{E}[\hat{f}]-\hat{f})^2\right] + 2\,\mathbf{E}\left[(f-\mathbf{E}[\hat{f}])\varepsilon\right] \\ &+ 2\,\mathbf{E}\left[\varepsilon(\mathbf{E}[\hat{f}]-\hat{f})\right] + 2\,\mathbf{E}\left[(\mathbf{E}[\hat{f}]-\hat{f})(f-\mathbf{E}[\hat{f}])\right] \\ &= (f-\mathbf{E}[\hat{f}])^2 + \mathbf{E}[\varepsilon^2] + \mathbf{E}\left[(\mathbf{E}[\hat{f}]-\hat{f})^2\right] + 2(f-\mathbf{E}[\hat{f}])\,\mathbf{E}[\varepsilon] \\ &+ 2\,\mathbf{E}[\varepsilon]\,\mathbf{E}\left[\mathbf{E}[\hat{f}]-\hat{f}\right] + 2\,\mathbf{E}\left[\mathbf{E}[\hat{f}]-\hat{f}\right](f-\mathbf{E}[\hat{f}]) \\ &= (f-\mathbf{E}[\hat{f}])^2 + \mathbf{E}[\varepsilon^2] + \mathbf{E}\left[(\mathbf{E}[\hat{f}]-\hat{f})^2\right] \\ &= (f-\mathbf{E}[\hat{f}])^2 + \mathbf{Var}[\varepsilon] + \mathbf{Var}\left[\hat{f}\right] \\ &= \mathbf{Bias}[\hat{f}]^2 + \mathbf{Var}[\varepsilon] + \mathbf{Var}\left[\hat{f}\right] \\ &= \mathbf{Bias}[\hat{f}]^2 + \sigma^2 + \mathbf{Var}\left[\hat{f}\right] \end{split}$$