Practical: Metropolis-Hastings Algorithm

Instructors: Kari Auranen, Elizabeth Halloran and Vladimir Minin July 11 – July 13, 2018

Sampling from the standard normal distribution

Suppose our target is a univariate standard normal distribution with density $f(x) = 1/(\sqrt{2\pi})e^{-x^2/2}$. Given current state $x^{(t)}$, we generate two uniform r.v.s $U_1 \sim U[-\delta, \delta]$ and $U_2 \sim U[0, 1]$. Then set

$$x^{(t+1)} = \begin{cases} x^{(t)} + U_1 & \text{if } U_2 \le \min \left\{ e^{\left[\left(x^{(t)} \right)^2 - \left(x^{(t)} + U_1 \right)^2 \right] / 2}, 1 \right\} \\ x^{(t)} & \text{otherwise.} \end{cases}$$

 δ is a tuning parameter. Large δ leads to small acceptance rate, small δ leads to slow exploration of the state space. The rule of thumb for random walk proposals is to keep acceptance probabilities around 30-40%. If your proposal is close to the target, then higher acceptance rates are favorable.

Your task

Implement the above algorithm. Experiment with the tuning parameter δ and report empirically estimate acceptance probabilities for different values of this parameter.

Distribution of the time of infection

Consider a two state continuous-time Markov SIS model, where the disease status X_t cycles between the two states: 1=susceptible, 2=infected. Let the infection rate be λ_1 and clearance rate be λ_2 . Suppose that an individual is susceptible at time 0 ($X_0 = 1$) and infected at time T ($X_T = 2$). We don't know anything else about the disease status of this individual during the interval [0, T]. If T is small enough, it is reasonable to assume that the individual was infected only once during this time interval. We would like to obtain the distribution of the time of infection I, conditional on the information we have.

Your task

Implement a Metropolis-Hastings sampler to draw realizations from the distribution

$$\Pr(I \mid X_0 = 1, X_t = 2, N_t = 1) \propto \Pr(0 < t < I : X_t = 1, I < t < T : X_t = 2),$$

where N_t is the number of infections. Since X_t is a continuous-time Markov chain, the last probability (it is actually a density) can be written as

$$\Pr(0 < t < I : X_t = 1, I < t < T : X_t = 2) = \underbrace{\lambda_1 e^{-\lambda_1 I}}_{\text{density of waiting time until infection}} \times \underbrace{e^{-\lambda_2 (T-I)}}_{\text{prob of staying infected}}.$$

To make things concrete, set $\lambda_1 = 0.1$, $\lambda_2 = 0.2$ and T = 1.0. For your proposal distribution, use a uniform random walk with reflective boundaries 0 and T. In other words, given a current value of the infection time t_c , generate $u = \text{Unif}_{[t_c - \delta, t_c + \delta]}$ ($2\delta < T$) and then make a proposal value

$$t_p = \begin{cases} u & \text{if } 0 < u < T, \\ 2T - u & \text{if } u > T, \\ -u & \text{if } u < 0. \end{cases}$$

This is a symmetric proposal, so your M-H ratio will contain only the ratio of target densities:

$$\frac{\lambda_1 e^{-\lambda_1 t_p} e^{-\lambda_2 (T - t_p)}}{\lambda_1 e^{-\lambda_1 t_c} e^{-\lambda_2 (T - t_c)}} = e^{-\lambda_1 (t_p - t_c) - \lambda_2 (t_c - t_p)} = e^{(t_p - t_c)(\lambda_2 - \lambda_1)}.$$

Plot the histogram of the posterior distribution of the infection time. Try a couple of sets of values for λ_1 and λ_2 and examine the effect of these changes on the posterior distribution of the infection time.