

Лекция. Матстат(11.10.24)

$F(x; \theta)$ F — известна, θ неизвестна

$$H_0 : \theta = \theta_0$$

$$H_1 : \theta \neq \theta_0$$

$$\lambda = \frac{L_x(\theta_0)}{\max_{\theta} L_x(\theta)} = \frac{L_x(\theta_0)}{L_x(\hat{\theta})}$$

Отвергается H_0 , если $\lambda \leq \lambda_{\alpha}$ α - уровень значимости

$$P(\lambda \leq \lambda_{\alpha} \mid H_0 \text{ верна}) \leq \alpha$$

$X \in N(a, \sigma^2)$, σ^2 - известна

$$H_0 : a = a_0$$

$$H_1 : a \neq a_0$$

$$\frac{\sqrt{n} |\bar{x} - a_0|}{\sigma} > z_{1-\alpha/2}$$

$F(x)$ - ф.р.

$$H_0 : F(x) = F_0(x)$$

$$H_1 : F(x) \neq F_0(x)$$

Гистограмма $p_{0i} = F_0(a_i) - F_0(a_{i-1})$

Выборка объёма n

π_i - истинные вер. (π_1, \dots, π_k)

(m_1, \dots, m_k - имеет полином p)

$$p(x_1 = m_1, \dots, x_k = m_k) = \frac{n!}{m_1! \dots m_k!} \prod_{i=1}^k \pi_i^{m_i} = L_x(\pi_1, \dots, \pi_k)$$

При $H_0 : \pi_i = p_{i0}, i = \overline{1, k}$

$\max_{\pi_1, \dots, \pi_k} L_x(\pi, \dots, \pi_k) =$ При $\pi_i = \frac{m_i}{n}$

$$\lambda = \frac{L_x(p_{10}, \dots, p_{k0})}{L_x(\frac{m_1}{n}, \dots, \frac{m_k}{n})} = \frac{\prod_{i=1}^k (p_{i0}^{m_i})}{\prod_{i=1}^k (\frac{m_i}{n})^{m_i}} = \prod_{i=1}^k \left(\frac{p_{i0}}{m_i} \right)^{m_i}$$

$$-2 \ln \lambda = 2 \sum_{i=1}^k m_i \ln \frac{m_i}{np_{i0}} = 2 \sum_{i=1}^k ((m_i - np_{i0}) + np_{i0}) \ln(1 + \Delta_i) = ((m_i - np_{i0} + p_{i0}) (\Delta_i - \frac{\Delta_i^2}{2}))$$

$$\Delta_i = \frac{m_i - np_{i0}}{np_{i0}} = \frac{m_i}{np_{i0}} - 1$$

$$\ln(1+x) = x - \frac{x^2}{2} + o(x^2)$$

$$\sum_{i=1}^k p_{i0} \Delta_i = \sum_{i=1}^k \frac{m_i - np_{i0}}{n} = \frac{1}{n} \left(\sum_{i=1}^k m_i - n \sum_{i=1}^k p_{i0} \right) = 0$$

$$= 2(\chi^2 - \frac{1}{2}\chi^2) = 2\frac{\chi}{2} = \chi^2$$

$$n \rightarrow \infty \quad -2 \ln \lambda \sim \chi^2(k-1)$$

$$f(x; \theta_1, \theta_2)$$

$$H_0 : \theta_1 = \theta_{10}$$

$$H_1 : \theta_1 \neq \theta_{10}$$

θ_2 - неизвестен (мешающий)

$$\lambda = \frac{\max_{\theta_2} L_x(\theta_{10}, \theta_2)}{\max_{\theta_1, \theta_2} L_x(\theta_1, \theta_2)}$$

$$\lambda \leq \lambda_2 \text{ - отв } H_0$$

ПОМОГИТЕ

$$X \in N(a, \sigma^2), \sigma^2 - \text{ неизв}$$

$$H_0 : a = a_0$$

$$H_1 : a \neq a_0$$

$$\theta = (a, \sigma^2)$$

$$L_x(\theta) = (2\pi\sigma^2)^{-\frac{n}{2}} \exp\left(-\frac{1}{2\sigma^2} \sum_{i=1}^n (x_i - a)^2\right)$$

При гипотезе

$$\max_{\sigma^2} L_x(a_0, \sigma^2) = L_x(a_0, \tilde{\sigma}^2)$$

$$\tilde{\sigma}^2 = \frac{1}{n} \sum_{i=1}^n (x_i - a_0)^2 = s^2 + (\bar{x} - a_0)^2$$

$$s^2 = \frac{1}{n} \sum_{i=1}^n (x_i - \bar{x})^2$$

$$\max_{a, \sigma^2} L_x(a, \sigma^2) = L_x(\bar{x}, s^2) = (2\pi s^2)^{-\frac{n}{2}} \exp\left(-\frac{n}{2}\right)$$

$$\lambda = \left(\frac{s^2}{s^2 + (\bar{x} - a_0)^2}\right)^2 = \left(\frac{1}{1 + \frac{(\bar{x} - a_0)^2}{s^2}}\right)^{\frac{n}{2}} = \left(\frac{1}{1 + \frac{t^2}{n}}\right)^{\frac{n}{2}}$$

$$t = \frac{\sqrt{n} (\bar{x} - a_0)}{s}$$

$$\frac{\sqrt{n} |\bar{x} - a_0|}{s} \geq c_\alpha = c_{1 - \frac{\alpha}{2}}$$

$$\frac{\sqrt{n} |\bar{x} - a_0|}{s} \in t(n-1)$$

$$\text{по } \alpha \text{ и } \nu = n-1$$

$$t_{1 - \frac{\lambda}{2}}(\nu) - \text{квантиль пор } 1 - \frac{\alpha}{2}$$

Дисперсионный анализ $H_0 : \sigma_1 = \sigma_2 = \dots = \sigma_k$ (самим)

Доверительный интервал

$$\hat{\theta}_n - \text{оценка}$$

Использовать неравенство Чебышева

$$E_\theta(\hat{\theta}_n) = \theta$$

$$D_\theta(\hat{\theta}_n) = \sigma^2$$

$$P(|\hat{\theta}_n - \theta_0| > 36) \leq \frac{1}{9}$$

$$|\hat{\theta}_n - \theta_0| \leq 36$$

$$X \in N(a, \sigma^2), \sigma^2 - \text{изв}$$

$$P_{a_0}\left(\frac{\sqrt{n} |\bar{x} - a_0|}{\sigma} \geq z_{1-\alpha/2}\right) \leq \alpha$$

$$P_{a_0}\left(\frac{\sqrt{n} |\bar{x} - a_0|}{\sigma} < z_{1-\alpha/2}\right) \geq 1 - \alpha$$

$$P_{a_0}\left(\bar{x} - z_{1-\alpha/2} \frac{\sigma}{\sqrt{n}} < a_0 < \bar{x} + z_{1-\alpha/2} \frac{\sigma}{\sqrt{n}}\right) \geq 1 - \alpha$$

Def Доверительным интервалом с границами $(\hat{\theta}_1(x), \hat{\theta}_2(x))$

$$P((\hat{\theta}_1(x), \hat{\theta}_2(x)) \ni \theta) \geq \gamma$$

Качество статистических критериев