

Assignment II

Date :

*Problem 1:

⇒ According to the question, We will first execute its mean,

$$\begin{aligned} E[X] &= \int_0^{\infty} x \cdot \text{Gamma}(x|a, b) dx \\ &= \int_0^{\infty} x \cdot \frac{b^a}{\Gamma(a)} x^{a-1} e^{-bx} dx \\ &= \int_0^{\infty} x \frac{b^a}{\Gamma(a)} x^{a-1} e^{-bx} dx \\ &= \frac{b^a}{\Gamma(a)} \int_0^{\infty} x \cdot x^{a-1} \cdot e^{-bx} dx \\ &= \frac{b^a}{\Gamma(a)} \int_0^{\infty} x^a e^{-bx} dx \end{aligned}$$

By the property we know that,

$$\int_0^{\infty} x^{a-1} e^{-bx} dx = \frac{\Gamma(a)}{b^a}$$

$$\text{So, } \int_0^{\infty} x^a e^{-bx} dx = \frac{\Gamma(a+1)}{b^{a+1}}$$

$$E[X] = \frac{b^a}{\Gamma(a)} \frac{\Gamma(a+1)}{b^{a+1}},$$

By the property of $\Gamma(a+1) = a\Gamma(a)$

$$E[x] = \frac{b^a}{\Gamma(a)} \frac{\Gamma(a)}{b^a b}$$

$$= a/b$$

Now we need to calculate the Variance,

$$\text{Var}[x] = E[x^2] - (E[x])^2$$

$$\text{So, } E[x^2] = \int_0^{\infty} x^2 \frac{b^a}{\Gamma(a)} x^{a-1} e^{-bx} dx$$

$$= \frac{b^a}{\Gamma(a)} \int_0^{\infty} x^2 x^{a-1} e^{-bx} dx$$

$$= \frac{b^a}{\Gamma(a)} \int_0^{\infty} x^2 x^{a-1} e^{-bx} dx$$

$$= \frac{b^a}{\Gamma(a)} \int_0^{\infty} x^{a+1} e^{-bx} dx$$

By Using the same properties we used for mean,

$$E[x^2] = \frac{b^a}{\Gamma(a)} \frac{\Gamma(a+2)}{b^{a+2}}$$

$$= \frac{b^a}{\Gamma(a)} \frac{\Gamma(a+2)}{b^{a+2}}$$

$$= \frac{b^a}{\Gamma(a)} \frac{(a+1)a\Gamma(a)}{b^{a+1}}$$

$$= \frac{a(a+1)}{b^a}$$

$$\therefore \text{Var}[x] = \frac{a(a+1)}{b^a} - \frac{a^2}{b^a}$$

$$= \frac{a}{b^a}$$

*Problem 2: Given that,

$$\begin{pmatrix} A & B \\ C & D \end{pmatrix}^{-1} = \begin{pmatrix} M & -MBD^{-1} \\ -D^{-1}CM & D^{-1} + D^{-1}CMBD^{-1} \end{pmatrix}$$

$$M = (A - BD^{-1}C)^{-1}$$

We have to prove, $Y^{-1} = X$

Also, $XY = I$ [I = Identity Matrix]

$$\begin{pmatrix} A & B \\ C & D \end{pmatrix} \begin{pmatrix} M & -MBD^{-1} \\ -D^{-1}CM & D^{-1} + D^{-1}CMBD^{-1} \end{pmatrix} = \begin{pmatrix} M_{11} & M_{12} \\ M_{21} & M_{22} \end{pmatrix}$$

$$M_{11} = AM - BD^{-1}CM = M(A - BD^{-1}C) = (A - BD^{-1}C)^{-1} \cdot (A - BD^{-1}C) = I$$

$$\begin{aligned} M_{12} &= -AMB D^{-1} + BD^{-1} + BD^{-1}CMBD^{-1} \\ &= BD^{-1}(-AM + I + CMBD^{-1}) \end{aligned}$$

So, therefore,

$$-AM + CMBD^{-1} = -I, \quad BD^{-1}(I - I) = 0$$

$$M_{21} = CM - DD^{-1}CM = CM - CM = 0$$

$$\begin{aligned} M_{22} &= -CMBD^{-1} + DD^{-1} + DD^{-1}CMBD^{-1} \\ &= -CMBD^{-1} + I + CMBD^{-1} \\ &= I \end{aligned}$$

Hence,

$$\begin{pmatrix} A & B \\ C & D \end{pmatrix} \begin{pmatrix} M & -MBD^{-1} \\ -D^{-1}CM & D^{-1} + D^{-1}CMBD^{-1} \end{pmatrix} = \begin{pmatrix} I & 0 \\ 0 & I \end{pmatrix}$$

$$\text{Also, } \begin{pmatrix} A & B \\ C & D \end{pmatrix}^{-1} = \begin{pmatrix} M & -MBD^{-1} \\ -D^{-1}CM & D^{-1} + D^{-1}CMBD^{-1} \end{pmatrix}$$

*Problem 3:

$$\Rightarrow R = \begin{pmatrix} I + A^T L A & -A^T L \\ -L A & I \end{pmatrix}$$

By using the property,

$$\begin{pmatrix} A & B \\ C & D \end{pmatrix}^{-1} = \begin{pmatrix} M & -M B D^{-1} \\ -D^{-1} C M & D^{-1} + D^{-1} C M B D^{-1} \end{pmatrix}$$

$$M = (I + A^T L A - (A^T L L^{-1} - L A))^{-1}$$

$$= (I + A^T L A - (-A^T - L A))^{-1}$$

$$= I^{-1}$$

$$\therefore R^{-1} = \begin{pmatrix} M_{11} & M_{12} \\ M_{21} & M_{22} \end{pmatrix}$$

$$M_{11} = I^{-1}$$

$$M_{12} = -I^{-1} - A^T L L^{-1} = -I^{-1} A^T$$

$$M_{21} = (-L^{-1}) (-L A) (I^{-1}) = A I^{-1}$$

$$M_{22} = L^{-1} + (L^{-1}) (-L A I^{-1}) (-A^T) L L^{-1}$$

$$= L^{-1} A I^{-1} A^T$$

$$\text{Hence, } \text{cov}(z) = R^{-1} = \begin{pmatrix} I^{-1} & -I^{-1} A^T \\ A I^{-1} & L^{-1} A I^{-1} A^T \end{pmatrix}$$

* Problem 4: Given that,

$$\Rightarrow P(x) = N(x | \mu_x, \Sigma_x)$$

$$P(z) = N(z | \mu_z, \Sigma_z)$$

$$y = x + z$$

We know that, $E[x] = \mu_x$

$$\text{Cov}[x] = E[x^2] - (E[x])^2 = 0$$

therefore,

$$\begin{aligned} \mu_{y|x} &= E[x] + E[z] \\ &= \mu_x + \mu_z \end{aligned}$$

$$\Sigma_{y|x} = \text{Cov}[x] + \text{Cov}[z] = \Sigma_z$$

$$\text{Also, } P(y|x) = N(y | \mu_x + \mu_z, \Sigma_z)$$

By comparing the equation,

$$P(x) = N(x | \mu, A^{-1})$$

$$P(y|x) = N(y | Ax + b, L^{-1})$$

$$P(y) = N(y | A\mu + b, L^{-1} + A\mu A^T)$$

$$\mu = \mu_x$$

$$A = I$$

$$b = \mu_z$$

$$A^{-1} = \Sigma_x$$

$$L^{-1} = \Sigma_z$$

$$\begin{aligned} P(y) &= N(y | \mu_x + \mu_z, \Sigma_z + \Sigma_x A^T A) \\ &= N(y | \mu_x + \mu_z, \Sigma_z + \Sigma_x) \end{aligned}$$

Assignment 2

Date

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→ In this problem, we are requested to approach the problem from a different angle in this question. We need to write the joint distribution $p(x, y)$ and then embed over to obtain the marginal distribution $p(y)$.

Let's start by writing the quadratic form of $p(x, y)$ in exponential form:

$$-\frac{1}{2} (x-u)^T \Lambda (x-u) - \frac{1}{2} (y-Ax-b)^T L (y-Ax-b)$$

Those terms involving x are retrieved:

$$= -\frac{1}{2} x^T (\Lambda + A^T L A) x + x^T [\Lambda u + A^T L (y-b)] + \text{constant}$$

$$= -\frac{1}{2} (x-m)^T (\Lambda + A^T L A) (x-m) + \frac{1}{2} m^T (\Lambda + A^T L A) m + \text{constant}$$

Here's what we have come up with:

$$m = (\Lambda + A^T L A)^{-1} [\Lambda u + A^T L (y-b)]$$

When we integrate over x , we can observe that the first term disappeared to a constant, then we extract the remaining terms that include y .

We can get the following parts:

$$= \frac{1}{2} y^T [L - LA(\lambda + A^T L A)^{-1} A^T L] y + y^T \{ [L - LA(\lambda + A^T L A)^{-1} A^T L] b + LA(\lambda + A^T L A)^{-1} \lambda y \}$$

By taking the inverse of y^T coefficient, we can get the covariance matrix $\text{cov}[y]$

$$L - LA(\lambda + A^T L A)^{-1} A^T L$$

By utilizing Woodbury inversion formula:

$$(X + YXU)^{-1} = X^{-1} - X^{-1}Y(X^{-1} + UX^{-1}Y)^{-1}UX^{-1}$$

$$\text{where, } X^{-1} = L, Y = A, X^{-1} = \lambda, U = A^T$$

so that, we can perform

$$\text{cov}[y] = (L^{-1} + A\lambda^{-1}A^T)^{-1}$$

Coefficient y^T must be equal to $E[y] (\text{cov}[y])^{-1}$

$$E[y] (L^{-1} + A\lambda^{-1}A^T)^{-1} = \{ [L - LA(\lambda + A^T L A)^{-1} A^T L] b + LA(\lambda + A^T L A)^{-1} \lambda y \}$$

$$E[y] = (L^{-1} + A\lambda^{-1}A^T) \{ [L - LA(\lambda + A^T L A)^{-1} A^T L] b + LA(\lambda + A^T L A)^{-1} \lambda y \}$$

$$E[y] = (L^{-1} + A\lambda^{-1}A^T) \{ (L^{-1} + A\lambda^{-1}A^T)^{-1} b + LA(\lambda + A^T L A)^{-1} \lambda y \}$$

$$E[y] = \{ b + (L^{-1} + A\lambda^{-1}A^T) LA(\lambda + A^T L A)^{-1} \lambda y \}$$

By utilizing Woodbury inversion method, we can perform

$$\begin{aligned}
 (\Lambda + A^T \Lambda^{-1} A)^{-1} &= \Lambda^{-1} - \Lambda^{-1} A^T (L^{-1} + A \Lambda^{-1} A^T)^{-1} A \Lambda^{-1} \\
 E[y] &= b + (L^{-1} + A \Lambda^{-1} A^T) \Lambda A (\Lambda^{-1} - \Lambda^{-1} A^T) (L^{-1} + A \Lambda^{-1} A^T)^{-1} A \Lambda^{-1} \\
 &= b + (L^{-1} + A \Lambda^{-1} A^T) (A \Lambda^{-1} - (L^{-1} + A \Lambda^{-1} A^T) \Lambda A \Lambda^{-1} A^T (L^{-1} + A \Lambda^{-1} A^T)^{-1} A \Lambda^{-1}) \\
 &= b + (L^{-1} + A \Lambda^{-1} A^T) (A \Lambda^{-1} - L A \Lambda^{-1} A^T \Lambda^{-1}) \\
 &= b + L^{-1} L A \Lambda^{-1} + A \Lambda^{-1} A^T L A \Lambda^{-1} - A \Lambda^{-1} A^T L A \Lambda^{-1} \\
 &= b + A \Lambda^{-1}
 \end{aligned}$$

$$\begin{aligned}
 \therefore E[y] &= b + A \Lambda^{-1} \mu \\
 &= A^{-1} b
 \end{aligned}$$


```
In [19]: N = 1000
Pw0 = Pw1 = Pw2 = 1/3          # A priori probabilities for each class are equal

m1 = np.array([0, 0, 0])
m2 = np.array([1, 2, 2])
m3 = np.array([3, 3, 4])

#S = S0 = S1 = S2 = 0.8*np.eye(3)

S1 = np.array([[0.8,0.2,0.1],
               [0.2,0.8,0.2],
               [0.1,0.2,0.8]])
S2 = np.array([[0.6,0.01,0.01],
               [0.01,0.8,0.01],
               [0.01,0.01,0.6]])
S3 = np.array([[0.6,0.1,0.1],
               [0.1,0.6,0.1],
               [0.1,0.1,0.6]])

# Creating a training set
Xtr_w0 = np.random.multivariate_normal(m1, S1, 333) # a set of vectors for c
ytr_w0 = 0*np.ones((333, 1))                       # class 0's labels

Xtr_w1 = np.random.multivariate_normal(m2, S2, 333) # a set of vectors for c
ytr_w1 = 1*np.ones((333, 1))                       # class 1's labels

Xtr_w2 = np.random.multivariate_normal(m3, S3, 333) # a set of vectors for c
ytr_w2 = 2*np.ones((333, 1))                       # class 2's labels

# Data and labels are collected in a single set.
Xtr = np.concatenate((Xtr_w0, Xtr_w1, Xtr_w2), axis = 0)
ytr = np.concatenate((ytr_w0, ytr_w1, ytr_w2), axis = 0)

## creating a test set

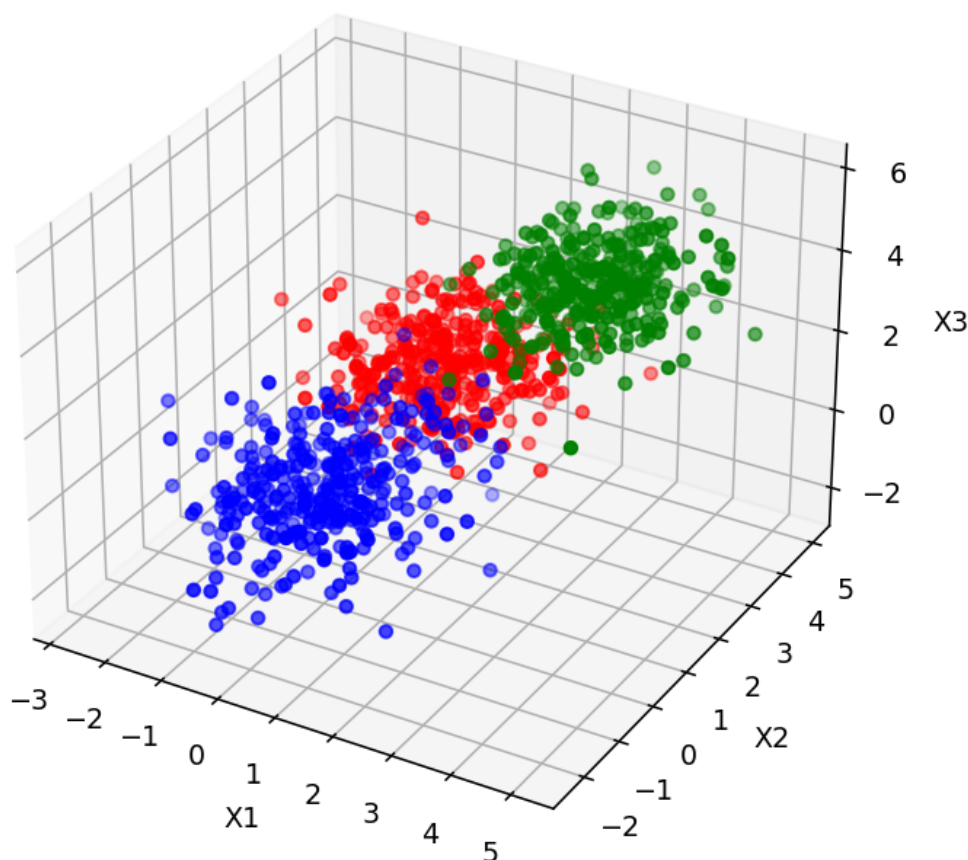
Xte_w0 = np.random.multivariate_normal(m1, S1, 333) # a set of vectors for c
yte_w0 = 0*np.ones((333, 1))                       # class 0's labels

Xte_w1 = np.random.multivariate_normal(m2, S2, 333) # a set of vectors for c
yte_w1 = 1*np.ones((333, 1))                       # class 1's labels

Xte_w2 = np.random.multivariate_normal(m3, S3, 333) # a set of vectors for c
yte_w2 = 2*np.ones((333, 1))                       # class 2's labels

# Data and labels are collected in a single set.
Xte = np.concatenate((Xte_w0, Xte_w1, Xte_w2), axis = 0)
yte = np.concatenate((yte_w0, yte_w1, yte_w2), axis = 0)
```

Classification



```
In [21]: m1_hat = (1.0/(N/3))*np.sum(Xtr_w0, axis = 0)
          S0_hat = (1.0/(N/3))*np.dot((Xtr_w0-m1_hat).T, (Xtr_w0-m1_hat))

          m2_hat = (1.0/(N/3))*np.sum(Xtr_w1, axis = 0)
          S1_hat = (1.0/(N/3))*np.dot((Xtr_w1-m2_hat).T, (Xtr_w1-m2_hat))

          m3_hat = (1.0/(N/3))*np.sum(Xtr_w2, axis = 0)
          S2_hat = (1.0/(N/3))*np.dot((Xtr_w2-m3_hat).T, (Xtr_w2-m3_hat))

          S_hat = (1.0/3.0)*(S0_hat + S1_hat + S2_hat)
```

```
In [22]: inv_S = np.linalg.inv(S_hat)
dm_0 = np.sqrt(np.sum(np.dot((Xte-m1_hat), inv_S)*(Xte-m1_hat), axis = 1))
dm_1 = np.sqrt(np.sum(np.dot((Xte-m2_hat), inv_S)*(Xte-m2_hat), axis = 1))
dm_2 = np.sqrt(np.sum(np.dot((Xte-m3_hat), inv_S)*(Xte-m3_hat), axis = 1))

# Then, using the obtained euclidean distances, I classify the objects.
dm_matrix = np.stack((dm_0, dm_1, dm_2), axis = 1)
Mahal result = np.argmin(dm_matrix, axis = 1)
```

```
In [23]: def multivariate_normal_pdf_v2(x, mean, sigma):
1 = x.shape[1]
det_S = np.linalg.det(sigma)
norm_const = 1.0/((2.0*np.pi)**(1/2.0)*np.sqrt(det_S))
inv_S = np.linalg.inv(sigma)
a1 = np.sum(np.dot(x-mean, inv_S)*(x-mean), axis = 1)

return norm_const*np.exp(-0.5*a1)
```

```
In [24]: baydis_x1 = Pw0*multivariate_normal_pdf_v2(Xte, m1_hat,S_hat)

baydis_x2 = Pw1*multivariate_normal_pdf_v2(Xte, m2_hat,S_hat)

baydis_x3 = Pw2*multivariate_normal_pdf_v2(Xte, m3_hat,S_hat)


de_matrix = np.stack((baydis_x1, baydis_x2, baydis_x3), axis = 1)
Bayes_result = np.argmax(de_matrix, axis = 1)

print(Bayes_result)
```

[illegible]

```
in [25]: #Error Bayesian classifier probability
error_bayesian = 1-np.sum(Bayes_result == yte.flatten())/N # Bayes_result = yte.flatten()

#Error probability of Mahalanobis
error_mahalanobis = 1-np.sum(Mahal_result == yte.flatten())/N

#print(error_bayesian)
print(error_bayesian)
#print(error_mahalanobis)
print(error_mahalanobis)
```

```
#The probability er
```

Comparing

*it is possible to observe that the error using both methods is the same due to the same probability that each class have