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Sec 2: Bayes Rule and Probability Review

Manipulate Bayes Rule

We have the bayes rule as follows:

$$p(A|B) = \frac{p(B|A)P(A)}{P(B)}$$

If we don't have $p(B)$ directly, we can calculate it from the numerator, i.e.

$$P(B) = \int_A P(A, B) = \int_A p(B|A)P(A)$$

We can think of the bottom term as a "normalization constant" so that the distribution sums up to 1

$$p(A|B) = \frac{p(B|A)P(A)}{\int_A p(B|A)P(A)} = \frac{p(B|A)P(A)}{Z} \propto p(B|A)P(A)$$

as Z is independent of A . This could be useful if we want to optimize over A , i.e.,

$$\operatorname{argmax}_A P(A|B) = \operatorname{argmax}_A P(B|A)P(A)$$

CTR and Bernoulli distribution

Think of click through rate (CTR) as coin tosses, which has a [Bernoulli distribution](#) for each click/don't click choice $Bern(p)$. Let H = click, T = No click, their corresponding # be N_H and N_T . The likelihood is

$$L(N_H, N_T) = p^{N_H} (1 - p)^{N_T}$$

To estimate p , we maximize the log likelihood

$$L = \log[L(N_H, N_T)] = N_H \log p + N_T \log(1 - p)$$

so

$$\frac{\partial L}{\partial p} = \frac{N_H}{p} - \frac{N_T}{1 - p} = 0 \iff p = \frac{N_H}{N_H + N_T}$$

Central Limit Theorem

- If X_i 's are *iid* random variables with $E[X_i] = \mu$, $Var[X_i] = \sigma^2 < \infty$. Let $\bar{X} = \frac{1}{N} \sum_{i=1}^N X_i$, CLT states that

$$\bar{X} \xrightarrow{d} N\left(\mu, \frac{\sigma^2}{N}\right)$$

Note that in practice, we don't actually know σ , instead we could approximate it by

$$\hat{\sigma} = \sqrt{\frac{1}{N} \sum_{i=1}^N (X_i - \bar{X})^2}$$

Confidence Intervals

Gaussian approximated CI

- Note that $\frac{\sqrt{N}}{\hat{\sigma}}(\bar{X} - \mu) \sim N(0, 1)$ approximately (As we are using the empirical variance)

- Let $\hat{\mu} = \bar{X}$, $\hat{\sigma} = \sqrt{\frac{1}{N} \sum_{i=1}^N (X_i - \bar{X})^2}$. Given X_1, \dots, X_N , we calculate the level $-\alpha$ (e.g., .05) C.I. of μ as

$$\left[\hat{\mu} + z_{\alpha/2} \frac{\hat{\sigma}}{\sqrt{N}}, \hat{\mu} + z_{1-\alpha/2} \frac{\hat{\sigma}}{\sqrt{N}} \right]$$

Where $z_{\alpha/2} = \Phi^{-1}(\alpha/2)$ and $z_{1-\alpha/2} = \Phi^{-1}(1 - \alpha/2) = -z_{\alpha/2}$

Non-approximated CI with Student's t -distribution

- To find the non-approximated version, we need to use [t-distribution](#). Specifically, let

$S^2 = \frac{1}{N-1} \sum_{i=1}^N (X - \bar{X})^2$, the exact distribution of $\frac{\bar{X} - \mu}{S/\sqrt{N}}$ is a Student's t -distribution with $N - 1$ degrees of freedom.

- Note the difference between S^2 and $\hat{\sigma}^2$

Sec 3: Traditional A/B Testing

Test the Mean Difference Between Two Groups (t Procedure)

Assuming all data are *iid* Gaussian, we have $X_1 = \{x_1^{(1)}, \dots, x_1^{(N_1)}\}$, $X_2 = \{x_2^{(1)}, \dots, x_2^{(N_2)}\}$, and we want to test whether or not group 1 is different than group 1 on average

- Null hypothesis $H_0 : \mu_1 - \mu_2 \neq 0$
- The test statistic $t = \frac{(\bar{X}_1 - \bar{X}_2) - (\mu_1 - \mu_2)}{\sqrt{\frac{S_1^2}{N_1} + \frac{S_2^2}{N_2}}}$
- If the null hypothesis is true, $t \sim T_{N_1+N_2-2}$
- The C.I. for the difference in means $\mu_1 - \mu_2$ is given by $(\bar{X}_1 - \bar{X}_2) \pm t^* \sqrt{\frac{S_1^2}{N_1} + \frac{S_2^2}{N_2}}$ where t^* is the $1 - \alpha/2$ (e.g., 0.975) critical value.

Pooled t Procedure

Note that this is to assume that the two group has different variance, if we know that the two groups have the same variance (standard deviation), we use the [pooled t procedure](#) where we calculate the *pooled sample variance*

$$S_p^2 = \frac{(N_1 - 1)S_1^2 + (N_2 - 1)S_2^2}{N_1 + N_2 - 2}$$

, and the test statistic becomes

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$$t = \frac{(\bar{X}_1 - \bar{X}_2) - (\mu_1 - \mu_2)}{S_p \sqrt{\frac{1}{N_1} + \frac{1}{N_2}}}$$

P-Value

- **Definition:** The probability of obtaining a result equal to or 'more extreme' than what was actually observed, when the null hypothesis is true.
- Since we use a significance level α , we reject the null hypothesis if p-value $\leq \alpha$