# ASTR 5550: HW2 - Clean

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### 1. Binomial and Gaussian Distributions

Basically, this problem is a repeat of a previous problem (in HW1). The goal of this problem is to show how a Gaussian distribution is a good approximation of a Poisson or binomial distribution at high expected counts ( $\lambda$  or  $\mu$ ). Furthermore, a Gaussian distribution is continuous whereas the Poisson and binomial distributions can be difficult to evaluate via computer at high n.

In this problem, a well-established count rate of photons from a star (from a given telescope) is 0.1/s. The photons from a star are counted for 100 seconds by a "perfect" CCD that is read out every 1 second (1 second accumulation).

# Part (a)

Start with the binomial distribution. Write a program that calculates P(x; n, p) as a function of x with n = 100 and p = 0.1. Plot your results for  $0 \le x \le 20$ .

#### Jasmine's reminder to self

#### Binomial probability:

$$P_B(x;n,p) = \frac{n!}{(n-x)!x!} p^x q^{n-x}$$

where, q = 1 - p

And other useful:

```
E(x) = \muE(x - \mu)^2 = \sigma^2
```

```
where, E(f(x)) = \sum f(x)P(x)
```

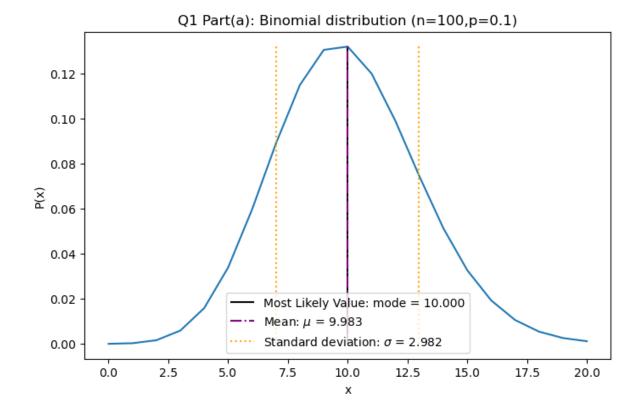
```
# libraries
import math
import numpy as np
import matplotlib.pyplot as plt
import random as rnd
from scipy.special import factorial
from scipy.signal import peak_widths
```

```
class helper:
    def __init__(self,x_min=0,x_max=15,step=1):
        self.step = step
        self.x = np.arange(x_min,x_max+self.step,self.step)
    def Expectation(self,fx,Px):
       E = np.sum(fx*Px)
                                      # E(x) = sum(f(x)*P(x))
        return E
    def calculate_mean_mode_sigma(self,x,Px):
        # calculate mean: mu = E(x)
        self.mu = self.Expectation(fx=x,Px=Px)
        # calculate stand. dev: sigma = sqrt(sigma^2); sigma^2 = E(x-mu)^2
        self.sigma = np.sqrt(self.Expectation(fx=(x-self.mu)**2,Px=Px))
        # calculate mode: x value where P(x) is highest
        self.mode = x[np.argmax(Px)]
    def plot_mark_mean(self,
                       Px : np.array,
                       ymin=0,
                       ymax=None,
                       colors= 'purple',
                       linestyles='dashdot',
                       label = "Mean: "):
        if ymax is None:
```

```
ymax = max(Px)
    # mark mean in plot
    plt.vlines(self.mu, ymin=ymin, ymax=ymax,
               linestyles=linestyles,
               colors=colors,
               label=label + "$\mu$ = {0:.3f}".format(self.mu))
    plt.legend()
def plot_mark_mode(self,
                   Px : np.array,
                   ymin=0,
                   ymax=None,
                   colors='black',
                   linestyles='solid',
                   label='Most Likely Value: '):
    if ymax is None:
        ymax = max(Px)
    # mark mode in plot
    plt.vlines(self.mode,ymin=ymin,ymax=ymax,
               linestyles=linestyles,
               colors=colors,
               label=label+ "mode = {0:.3f}".format(self.mode))
    plt.legend()
def plot_mark_std(self,
                  Px : np.array,
                   ymin=0,
                   ymax=None,
                  colors='orange',
                  linestyles='dotted',
                  label="Standard deviation: "):
    if ymax is None:
        ymax = max(Px)
    # mark standard deviation (wrt mean) in plot
    plt.vlines(self.mu+self.sigma,
               ymin=ymin,ymax=ymax,
               colors=colors,
               linestyles=linestyles,
               label=label + "$\sigma$ = {0:.3f}".format(self.sigma))
```

```
class binomial(helper):
    # TODO: add documentation to the rest of these functions
    def __init__(self, x_min=0, x_max=15, step=1,n=100,p=0.5):
        """Functions related to using Binomial probability
        Parameters
        -----
        x min : numeric
           minimum x value of x array
        x_max : numeric
           maximum x value of x array
        step : numeric
           step size between each element in x array
        n : numeric
           number of 'trials'
        p : float
            probability of 'success' in single trial
        super().__init__(x_min, x_max, step)
        self.n = n
        self.p = p
        self.q = 1-self.p
    def P_binomial(self, x, n, p, q):
        """Function to calculate Binomial Probability P(x;n,p)
        Pb(x;n,p) = \{(n!)/[(n-x)!(x!)]\} * p^x * q^(n-x)
        11 11 11
        # formula components
        numerator = factorial(n)
                                                                 # fraction numerator: n!
                                                                 # fraction denominator: (n-x
        denominator = factorial(n-x)*factorial(x)
        fraction = numerator/denominator
        px = p**x
                                                                 # p^x
                                                                 \# q^{(n-x)} = (1-p)^{(n-x)}
        qnx = q**(n-x)
        # Binomial probability
```

```
part_a = binomial(x_min=0,x_max=20,n=100,p=0.1)
plt.figure(figsize=(8,5))
part_a.complete_ind_binomial_plot(title="Q1 Part(a): Binomial distribution (n={},p={})".form.
```



# Part (b)

Overplot the Poisson distribution; use a different color or line style to distinguish the two. What is the mean and sigma?

### Jasmine's reminder to self

### Poisson probability:

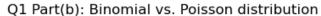
$$P_P(x;\lambda) = \frac{\lambda^x}{x!} e^{-\lambda}$$

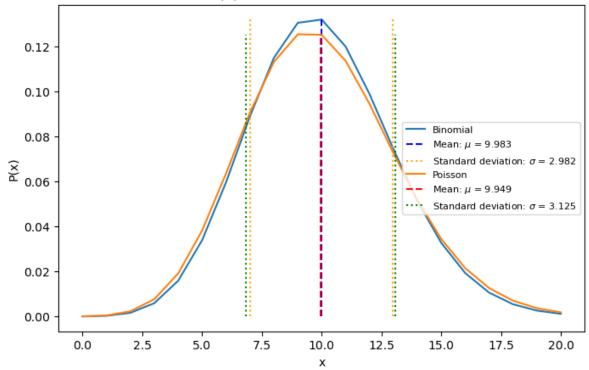
```
# function to calculate poisson probability P(x;lambda)
def P_poisson(self,x : int, mu : float):
    """Function to calculate Poisson probability P(x;mu)
   Parameters
    _____
       x value we want to get probability of
   mu : float
       mean occurence of value x
       This one especially HAS to be a float for this probability function,
       otherwise the `np.power(mu,x)` (to calculate mu^x portion) doesn't work properly
       will end up with weird (like negative) probabilities
   Returns
    _____
   Pp : float
       Poisson probability of x
   # formula components
   numerator = np.power(mu,x)
                                                # fraction numerator: mu^x (this one can
   denominator = factorial(x)
                                                # fraction denominator: x!
   fraction = numerator/denominator
                                                # e^(-mu)
   e_mu = np.exp(-mu)
   # Poisson probability
   Pp = fraction*e_mu
                                                # P(x;mu) = [(mu^x)/x!] * e^(-mu)
   return Pp
def plot_poisson(self,title="Poisson Distribution",label="Poisson",color='tab:orange'):
   self.Px = self.P_poisson(x=self.x,mu=self.mu)
   plt.plot(self.x,self.Px, label=label,c=color)
   plt.title(title)
   plt.xlabel("x")
   plt.ylabel("P(x)")
def complete_ind_poisson_plot(self,title="Poisson Distribution"):
   self.plot_poisson(title=title,label=None)
   self.calculate_mean_mode_sigma(x=self.x,Px=self.Px)
   self.plot_mark_mean(Px=self.Px)
```

```
self.plot_mark_mode(Px=self.Px)
self.plot_mark_std(Px=self.Px)
```

```
part_b = poisson(mu=part_a.mu,x_max=20) #instantiation

# comparison plot between Binomial and Poisson
plt.figure(1,figsize=(8,5))
#Binomial
part_a.plot_binomial(label="Binomial")
part_a.plot_mark_mean(Px=part_a.Px,colors='Blue',linestyles='dashed')
part_a.plot_mark_std(Px=part_a.Px,colors="orange")
#Poisson
part_b.plot_poisson(title="Q1 Part(b): Binomial vs. Poisson distribution")
part_b.calculate_mean_mode_sigma(x=part_b.x,Px=part_b.Px)
part_b.plot_mark_mean(Px=part_b.Px,colors='red',linestyles='dashed')
part_b.plot_mark_std(Px=part_b.Px,colors="green")
plt.legend(loc='center right',fontsize=8)
```





### Part (c)

Overplot the Gaussian distribution using a fine x-axis. How close are the two functions? Do they peak at the same value? Is it reasonable to assume a Gaussian parent distribution in this case?

#### Jasmine's reminder to self

#### Gaussian probability:

$$P_G(x; \lambda, \sigma) = \frac{1}{\sigma \sqrt{2\pi}} \exp\left(\frac{-(x - \lambda)^2}{2\sigma^2}\right)$$

```
class gaussian(helper):
   # TODO: add documentation to the rest of these functions
   def __init__(self, mu, sigma, x_min=0, x_max=15, step=1):
       super().__init__(x_min, x_max, step)
       self.mu = mu
       self.sigma = sigma
   def P_gaussian(self,x,mu,sigma:float):
        """Function to calculate Poisson probability P(x;mu)
       Parameters
        -----
       x : int
           x value we want to get probability of
       mu : float
           mean occurence of value x
       sigma : float
           standard deviation of x
       Returns
        _____
       Pg : float
           Gaussian probability of x
       # formula components
       pow_num = -((x-mu)**2)
                                                # numerator of exponential power
                                                # denominator of exponential power
       pow_den = 2*(sigma**2)
```

```
power = pow_num/pow_den  # power of exponential: -(x-<x>)^2/2(sigma^2)

denominator = sigma*np.sqrt(2*np.pi)  # fraction denominator: sigma*squareroot(2*p

# Gaussian probability
Pg = (1/denominator)*np.exp(power)  # P(x;mu,sigma)
return Pg

def plot_gaussian(self,title = 'Gaussian Distribution',label='Gaussian',color = "tab:greself.Px = self.P_gaussian(x=self.x,mu=self.mu,sigma=self.sigma)

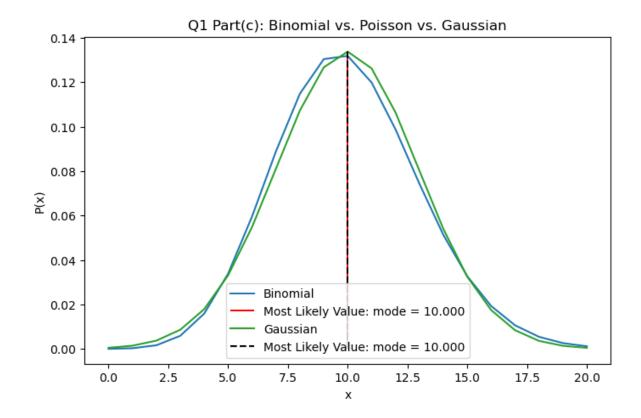
plt.plot(self.x,self.Px, label=label,c=color)
plt.title(title)
plt.xlabel("x")
plt.ylabel("P(x)")
```

```
part_c = gaussian(mu=part_a.mu,sigma=part_a.sigma,x_max=20) #instantiation

# comparison plot between Binomial and Poisson
plt.figure(1,figsize=(8,5))

#Binomial
part_a.plot_binomial(label="Binomial")
part_a.plot_mark_mode(Px=part_a.Px,colors='red')

#Gaussian
part_c.plot_gaussian(title="Q1 Part(c): Binomial vs. Poisson vs. Gaussian")
part_c.calculate_mean_mode_sigma(x=part_c.x,Px=part_c.Px)
part_c.plot_mark_mode(Px=part_c.Px,linestyles='dashed')
plt.legend()
```



#### JK written answer

They both seem to be pretty similar, and both peak at the same value.

```
print("Not asked for, I'm just curious to see them all together")
part_c = gaussian(mu=part_a.mu,sigma=part_a.sigma,x_max=20)  #instantiation

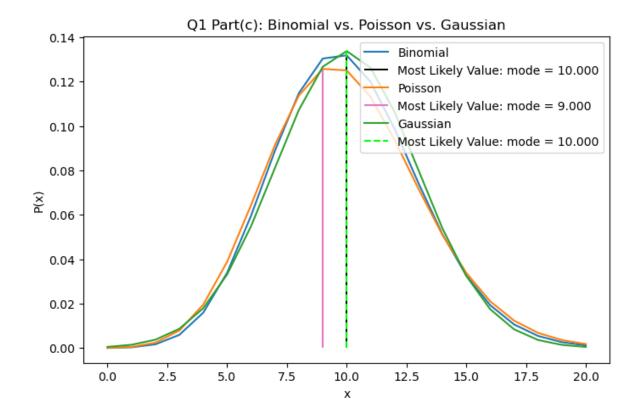
# comparison plot between Binomial and Poisson
plt.figure(1,figsize=(8,5))

#Binomial
part_a.plot_binomial(label="Binomial")
# part_a.plot_mark_mean(colors='Blue',linestyles='dashed')
# part_a.plot_mark_std(colors="orange")
part_a.plot_mark_mode(Px=part_a.Px)

# #Poisson
part_b.plot_poisson(label='Poisson')
part_b.calculate_mean_mode_sigma(x=part_b.x,Px=part_b.Px)
```

```
# # part_b.plot_mark_mean(colors='red',linestyles='dashed')
# # part_b.plot_mark_std(colors="green")
part_b.plot_mark_mode(Px=part_b.Px,colors='tab:pink')
#Gaussian
part_c.plot_gaussian(title="Q1 Part(c): Binomial vs. Poisson vs. Gaussian")
part_c.calculate_mean_mode_sigma(x=part_c.x,Px=part_c.Px)
# part_c.plot_mark_mean(colors='lime',linestyles='dashed')
# part_c.plot_mark_std(colors = 'cyan')
part_c.plot_mark_mode(Px=part_c.Px,linestyles='dashed',colors='lime')
plt.legend()
```

Not asked for, I'm just curious to see them all together



### 2. How Old is This Volcano?

The age of a volcanic eruption on Mars can be inferred from counting the surface density of craters of a given diameter or greater. It is established that the impact rate r = 0.01 craters

 $\mathrm{km}^{-2}\mathrm{Myr}^{-1}$ . You survey an area (A) of 10  $\mathrm{km}^2$  and find 3 craters (x).

# Part (a)

Start by making a simple estimate of the volcano's age and a standard deviation.

#### Jasmine's written answer

### Simple estimate of age

We have - established impact rate:  $r=0.01\frac{\rm craters}{\rm km^2Myr}$  - number of craters observed: x=3 craters - survey area:  $A=10~{\rm km}^2$ 

And let's call the simple estimate of age, y.

Therefore,

$$r = \frac{x}{A*t}$$
 
$$\Rightarrow t = \frac{x}{A*r}$$

Plug in numbers:

$$t = \frac{x}{A * r} = \frac{3}{(10)(0.01)}$$

```
#approx age
x = 3
r = 0.01
A = 10

age = x/(A*r)
print("Simple estimate of age:",age,"Myr")
```

Simple estimate of age: 30.0 Myr

$$t = \frac{x}{A * r} = \frac{3}{(10)(0.01)}$$

$$t = 30 \text{ Myr}$$

#### Standard deviation

$$\sigma = \sqrt{3/10} \approx 0.5477$$

(Honestly, I'm not sure if I chose an overly simplistic way to calculate the standard deviation.)

## Part (b)

The simple estimation of the age and uncertainty is often adequate. One can explore this problem further by calculating the *posterior* probability, P(x;t) over a range of times (say, t=0 to 150 in Myr intervals) assuming a Poisson parent distribution and knowing x=3. Do this problem on your computer.

Important point: the initial calculation yields a set of *relative* probabilities; under Bayes' law:

$$P(t;x) = \frac{P(x;t)P(t)}{P(x)}$$

Since we have no prior knowledge in this case, we treat P(t)/P(x) as a constant; You must normalize P(t;x) so that the total probability is 1. P(t;x) will have units of "probability/Myr". Plot your results. Mark the mean and mode (most likely age) and directly compute the standard deviation. Compare these values with your simple estimate.

#### JK written answer

Poisson probability is often used for counting statistics

$$P_P(x;\lambda) = \frac{\lambda^x}{x!} e^{-\lambda}$$

Thus,  $P(x;t) = P_P(x,\lambda)$ 

And, we're treating P(t)/P(x) as a constant, such that  $\sum_{i=0}^{t_f} P(t_i;x) = 1$ 

Therefore, we want

$$\begin{split} 1 &= \sum_{i=0}^{t_f} P(t_i; x) = \sum_{i=0}^{t_f} P_p(x; t) \frac{P(t)}{P(x)} \\ \Rightarrow 1 &= \sum_{i=0}^{t_f} P_p(x; t) \frac{P(t)}{P(x)} \end{split}$$

Since we're treating P(t)/P(x) as a constant, perhaps we just call it k, and perhaps move it in front of the summation. (We don't *have* to rename it a variable, but I just find it easier to think of it as a constant that way.)

$$\begin{split} \frac{P(t)}{P(x)} &= k \\ 1 &= \sum_{i=0}^{t_f} P_p(x;t) \frac{P(t)}{P(x)} \Rightarrow 1 = k \sum_{i=0}^{t_f} P_p(x;t) \\ &\Rightarrow k = \frac{1}{\sum P_p(x;t)} \end{split}$$

In other words,

$$k = \frac{P(t)}{P(x)} = \frac{1}{\sum P_p(x;t)}$$

(I believe this is how we are 'normalizing' the probability)

```
class Posterior_Probability(poisson):
    # TODO: add documentation to this class and its functions
    def __init__(self,t0=0,tf=150,step=1,obs_area=10):
        super().__init__(step)
        self.t = np.arange(t0,tf+step,step)
        self.obs_area = obs_area

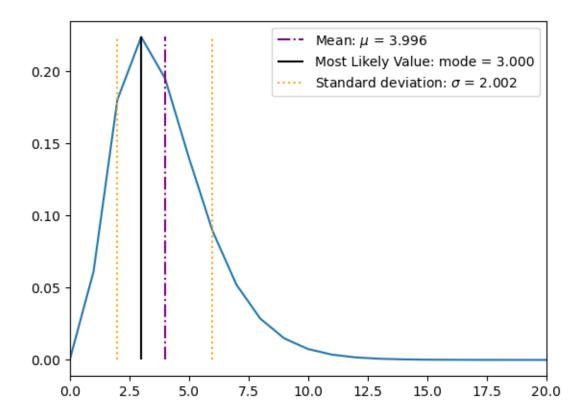
def post_prob(self,x):
    # P(x;t)/A => probability/km^2
    self.Pp = self.P_poisson(x=x,mu=self.t.astype('float'))/self.obs_area

#constant P(t)/P(x) = 1/sum(P(x;t))
    self.Pt_Px = 1/np.sum(self.Pp)

#posterior probability
    self.posterior = self.Pp*self.Pt_Px
    return self.posterior
```

```
ppb = Posterior_Probability(tf=150)

plt.plot(ppb.t,ppb.post_prob(x=3))
ppb.calculate_mean_mode_sigma(ppb.t,ppb.posterior)
ppb.Px = ppb.posterior
ppb.plot_mark_mean(Px=ppb.Px)
ppb.plot_mark_mode(Px=ppb.Px)
ppb.plot_mark_std(Px=ppb.Px)
ppt.xlim([0,20])
```



# Part (c)

A low count rate creates two issues. One issue is, of course, that the standard deviation is large regardless of how it is calculated. The other issue is that the probability distributions are often not symmetric, which creates a bit of a conundrum. It is natural to choose the most likely value of P(t;x) as the age of the volcano but one can see that there is a greater probability that the volcano is older than the likely age rather than younger (if your plot is correct at this point). So how does one determine the uncertainty?

There are several methods that can assign the uncertainty, one of which (relatively easy) is to calculate the full width at half the maximum (FWHM) of P(t;x) and apply the Gaussian formula that  $FWHM = \sigma \sqrt{8ln(2)}$  to calculate  $\sigma$ . What does the does this method yield  $[t_{min}, t_{max}]$ ? You now have calculated sigma three times. Is there a significant difference between them? (More to come on this!)

```
from scipy.signal import peak_widths

def get_fwhm(pdf):
    pdf_max = max(pdf)
```

```
where_max = np.where(pdf == pdf_max)[0]
fwhm = peak_widths(pdf,where_max,rel_height=0.5)[0]
return fwhm
```

```
FWHM = get_fwhm(ppb.posterior)[0]
print('FWHM =',FWHM)
```

FWHM = 4.128795671917094

We're using FWHM =  $\sigma \sqrt{8ln(2)}$  to get  $\sigma$ , therefore

$$\begin{aligned} \text{FWHM} &= \sigma \sqrt{8ln(2)} \\ \Rightarrow \sigma &= \frac{\text{FWTH}}{\sqrt{8\ln(2)}} \end{aligned}$$

```
print('FWHM =',FWHM)

Q2pc_sigma = FWHM/np.sqrt(8*np.log(2))
print('sigma =',Q2pc_sigma)
```

FWHM = 4.128795671917094 sigma = 1.7533380865470036

# Part (d)

Now suppose that 34 impact craters are identified over a 100 km<sup>2</sup> area. Redo parts (a), (b), and (c). Plot your results and compare.

```
# Q2 - Part (d): redo Part (a)
#approx age
x = 34
r = 0.01
A = 100
age = x/(A*r)
print("Simple estimate of age:",age,"Myr")
```

Simple estimate of age: 34.0 Myr

$$t = \frac{x}{A * r}$$
$$= \frac{34}{(100)(0.01)}$$

$$t = 34 \text{ Myr}$$

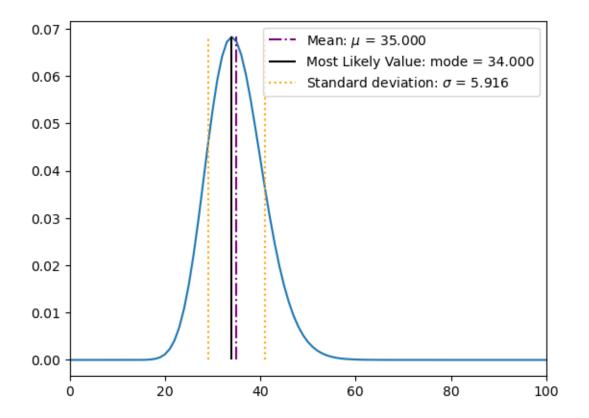
#### Standard deviation

$$\sigma = \sqrt{34/100} \approx 0.5830$$

(Well, I'm starting to believe that the standard deviation calculation is overly simple because that seems way too small.)

```
# Q2 - Part (d): redo Part (b)
ppd = Posterior_Probability(tf=150,obs_area=100)

plt.plot(ppd.t,ppd.post_prob(x=34))
ppd.calculate_mean_mode_sigma(ppd.t,ppd.posterior)
ppd.Px = ppd.posterior
ppd.plot_mark_mean(Px=ppd.Px)
ppd.plot_mark_mode(Px=ppd.Px)
ppd.plot_mark_std(Px=ppd.Px)
ppd.plot_mark_std(Px=ppd.Px)
ppt.xlim([0,100])
```



```
# Q2 - Part (d): redo Part (c)

FWHM_d = get_fwhm(ppd.posterior)[0]
print('FWHM =',FWHM_d)

Q2pd_sigma = FWHM_d/np.sqrt(8*np.log(2))
print('sigma =',Q2pd_sigma)
```

FWHM = 13.759478801950983 sigma = 5.843112653548922

### JK written answer/comment:

Well, I might've missed something because I was expecting Parts (a), (b), and (c) to be similar (and similarly in all redone parts in part (d)). I'm guessing I overly simplified part (a) or messed something up in my coding functions in parts (b) and (c). (Or worse, both.)

### 3. Random Walk and the Central Limit Theorem

### Part (a)

The random walk is often used to eludicate the central limit theorem. Using a classical but perhaps no longer politically correct analogy, a drunk at a lamp post (x = 0) takes a step either to the left or to the right with equal probability. His/her position (x) after n steps, can be represented with the binomial distribution:

$$P_B(x; n, p) = \frac{n!}{(n-x)! x!} p^{x+1} q^{n-x}$$

Set p = 1/2 for a step to the right (positive) of one unit, otherwise the step is to the left. The distance traveled is the d = 2x - n. We know from the central limit theorem that  $P_B(d; n, p)$  approaches a Gaussian at large n. Numerically calculate and plot  $P_B(d; n, p)$  for n = 100.

Overplot:

$$P_G(d) = \frac{C}{\sigma\sqrt{2\pi}}e^{-\frac{d^2}{2\sigma^2}}$$

Careful with normalization!  $P_B(d;n,p)$  is normalized so that its total is 1. When using d=2x-n with n as an even number, d must be an even number. There are only 101 possible positions. On the other hand,  $P_G(d)$  does not require d to be even (or an integer for that matter) and has twice as many possible positions.

I recommend that the plot is  $\log/\text{linear}$  with the vertical axis (probability) range of  $10^{-5}$  to 1. Don't print it out yet. See below.

```
class rand_walk(binomial,gaussian):
    # TODO: add documentation to these functions

def __init__(self, step=1, n=100, p=0.5):
    super().__init__(step, n, p)
    self.x = np.arange(0,self.n+step,step)
    self.d = self.distance(x=self.x,n=self.n)

    self.P_right = p

def distance(self,x,n):
    d = 2*x -n
    return d

def get_xn(self,ntrial,normalize = True):
    self.xn = np.zeros(len(self.d))
```

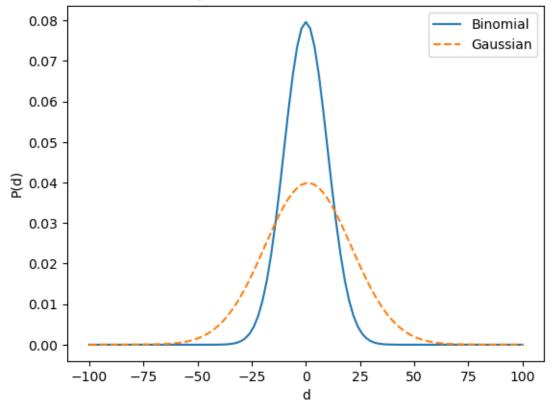
```
self.wh = np.where(self.d == 0.0)
    for _ in range(ntrial):
        i = self.wh[0]
        for _ in range(int(self.n/2)):
            if (rnd.random() <= self.P_right):</pre>
                i = i+self.step
            else:
                i = i-self.step
        self.xn[int(np.round(i))] += 1.0
    if normalize:
        self.xn = self.xn/ntrial
    return self.xn
def get_mu_sig(self):
    self.mu = np.sum(self.x)/self.n
    self.sigma = np.sqrt(self.n)
    return
def plot_rand_walk_BG(self,title='Random Walk (Binomial & Gaussian)'):
    self.get_mu_sig()
    self.Pb = self.P_binomial(x=self.x,n=self.n,p=self.p,q=self.q)
    self.Pg = self.P_gaussian(x=self.x,mu=self.mu,sigma=self.sigma)
    plt.figure()
    plt.plot(self.d,self.Pb,label = 'Binomial')
    plt.plot(self.d,self.Pg,'--', label = 'Gaussian')
    plt.xlabel('d')
    plt.ylabel('P(d)')
    plt.title(title)
    plt.legend()
    plt.show()
def plot_walk_count(self,title='Random Walk Distribution',
                    ntrial=1000,
                    bar_width=15.0,
                    bar_edge_color = 'black',
                    normalize_xn=True):
    self.get_xn(ntrial,normalize=normalize_xn)
    plt.bar(self.d,self.xn,width=bar_width,edgecolor=bar_edge_color)
```

```
plt.xlabel('d')
plt.ylabel('count')
plt.title(title)
plt.show()
```

```
n = 100
p = 0.5

rw = rand_walk(n=n,p=p)
rw.plot_rand_walk_BG(title = "Q3.a: Binomial and Gaussian")
```





# Part (b)

What if, however, the lamppost (x = 0) is at the bottom of a parabolic valley as pictured below. Let the altitude be  $z = 0.005d^2$ , so that the slope = 0.01d. Now suppose the drunk

tends to stagger downhill. Let the probability of a right step be 0.5- slope. With a brute force algorithm, calculate P(d). Overplot your results. What is  $\sigma$ ?

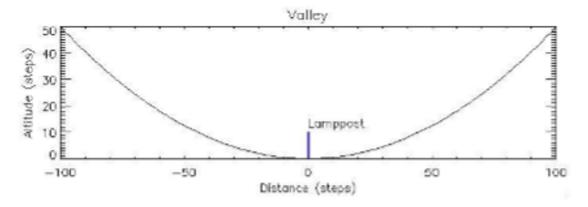


Figure 1: valley

Hint: Implement a random walk program with n=100 steps by calling a uniform random number then comparing to the right-step probability, which is a function of position. Record the end position. Loop ~100,000 times and make a histogram of the end positions.

```
class rand_slope_walk(rand_walk):
    # TODO: add documentation to these functions
    def __init__(self, slope_const=0.01, n=100, step=1, p=0.5):
        super().__init__(step, n, p)
                                                     # defaults to slope = 0.01d
        self.slope = slope_const*self.d
        self.P_right = p - self.slope
    def get_xn(self,ntrial,normalize = True):
        self.xn = np.zeros(len(self.d))
        self.wh = np.where(self.d == 0.0)
        for _ in range(ntrial):
            i = self.wh[0]
            for _ in range(int(self.n/2)):
                if (rnd.random() <= self.P_right[i]):</pre>
                    i = i+self.step
                else:
                    i = i-self.step
            self.xn[int(np.round(i))] += 1.0
        if normalize:
```

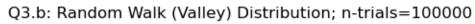
```
self.xn = self.xn/ntrial

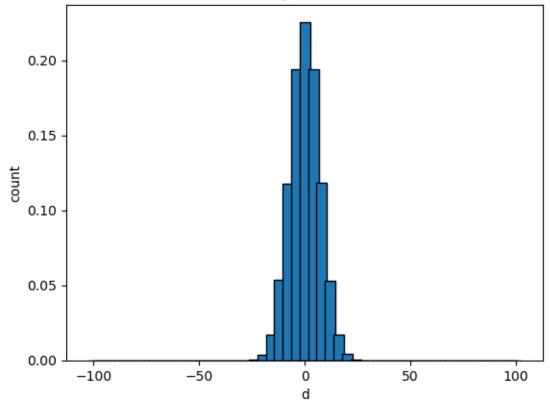
return self.xn

def get_mu_sig(self):
    self.mu = np.sum(self.d*self.xn)
    self.xbar2 = np.sum((self.d.astype('float')**2)*self.xn)
    self.sigma = np.sqrt(self.xbar2 - self.mu)
    return
```

```
%%time
ntrials = 100000
rwb = rand_slope_walk(step=1)

rwb.plot_walk_count(title='Q3.b: Random Walk (Valley) Distribution; n-trials={}'.format(ntrial=ntrials,
bar_width = 5.0)
```





```
CPU times: total: 27 s Wall time: 27.8 s
```

```
rwb.get_mu_sig()
print("sigma =",rwb.sigma)
```

sigma = 7.030166427617486

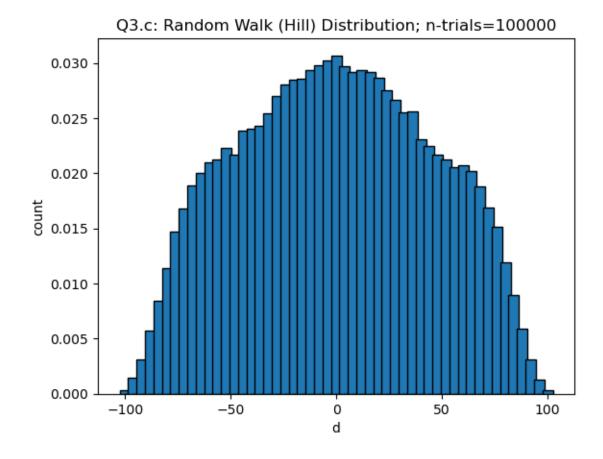
# Part (c)

Now put the lamppost (x = 0) at the top of a parabolic hill so that the slope = -0.01d. Overplot your results. What is  $\sigma$ ?

Note: You will notice that your results look Gaussian. However, the central limit theory does **not** necessarily apply for steps (b) and (c). Save your code for the next problem when we look at power-law tails.

```
%%time
ntrials=100000
rwc = rand_slope_walk(slope_const=-0.01)

rwc.plot_walk_count(title='Q3.c: Random Walk (Hill) Distribution; n-trials={}'.format(ntrials)
ntrial=ntrials,
bar_width=5.0)
```



CPU times: total: 20.6 s Wall time: 20.9 s

rwc.get\_mu\_sig()
print("sigma =",rwc.sigma)

sigma = 45.125772680365266

# 4. Not so Random Walk: Power Law

In this problem, we repeat problem 3 with a new caveat. His/her step size is 0.50 (whatever units) but increases with absolute value of his/her current position:

$$Step = 0.5 + |x| * 0.025$$

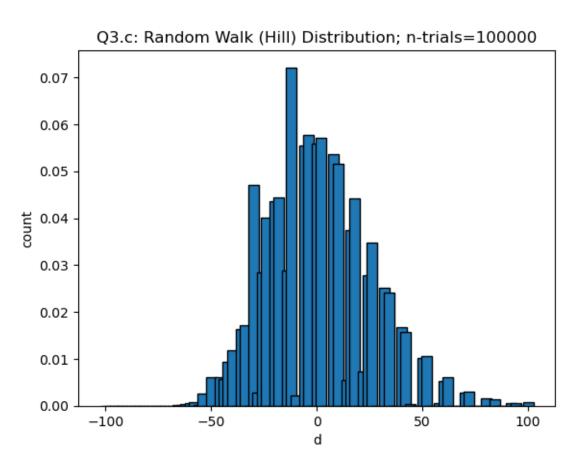
Let the number of steps be n=400 (half the original step size, so four times the number of steps). Implement a random walk program with n=400 steps by calling a uniform random number then comparing to the right-step probability, in this case 1/2. Record the end position. Loop  $\sim 100,000$  times and make a histogram of the end positions.

Calculate  $\sigma$ . The mean should be zero. Overplot a Gaussian with the calculated  $\sigma$ . Is the resulting distribution a Gaussian?

```
class rand_power_law(rand_walk):
    # TODO: add documentation to these functions
    def __init__(self, n=100, p=0.5, step0_size=0.5, step=1.0):
        super().__init__(n, p, step)
        self.x = np.arange(0,self.n+step,step)
        self.d = self.distance(x=self.x,n=self.n)
        self.P_right = p
        self.step0_size = step0_size
    def nearest_value_idx(self,x,value):
        idx = np.abs(np.asarray(x) - value).argmin()
        return idx
    def get_xn(self,ntrial,normalize = True):
        self.xn = np.zeros(len(self.d))
        self.wh = np.where(self.d == 0.0)
        for in range(ntrial):
            i = self.wh[0]
            step = self.step0_size
            for _ in range(int(self.n/2)):
                if (rnd.random() <= self.P_right):</pre>
                    i = i + step
                else:
                    i = i-step
                nearest_x = self.nearest_value_idx(self.x,i.astype('float'))
                step = 0.5 + np.abs(self.x[nearest_x])*0.025
            self.xn[nearest_x] += 1.0
        if normalize:
            self.xn = self.xn/ntrial
        return self.xn
```

```
%%time
ntrials=100000
rwpl = rand_power_law(n=400)

rwpl.plot_walk_count(title='Q3.c: Random Walk (Hill) Distribution; n-trials={}'.format(ntrial)
ntrial=ntrials,
bar_width=5.0)
```



CPU times: total: 37.2 s

Wall time: 37.6 s