



Count outcomes, Poisson GLMs

Regression Models

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Key ideas

- Many data take the form of counts
 - Calls to a call center
 - Number of flu cases in an area
 - Number of cars that cross a bridge
- Data may also be in the form of rates
 - Percent of children passing a test
 - Percent of hits to a website from a country
- Linear regression with transformation is an option

Poisson distribution

- The Poisson distribution is a useful model for counts and rates
- Here a rate is count per some monitoring time
- Some examples uses of the Poisson distribution
 - Modeling web traffic hits
 - Incidence rates
 - Approximating binomial probabilities with small p and large n
 - Analyzing contingency table data

The Poisson mass function

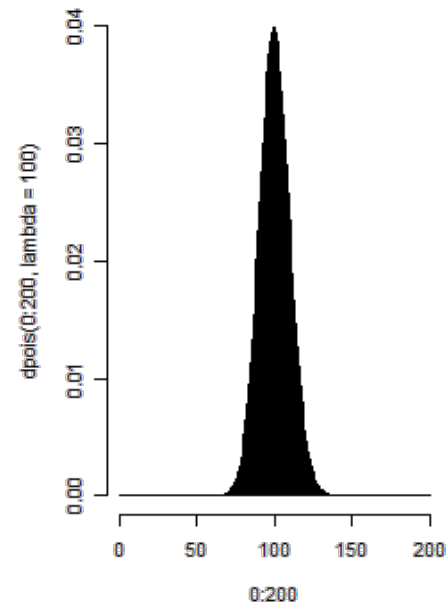
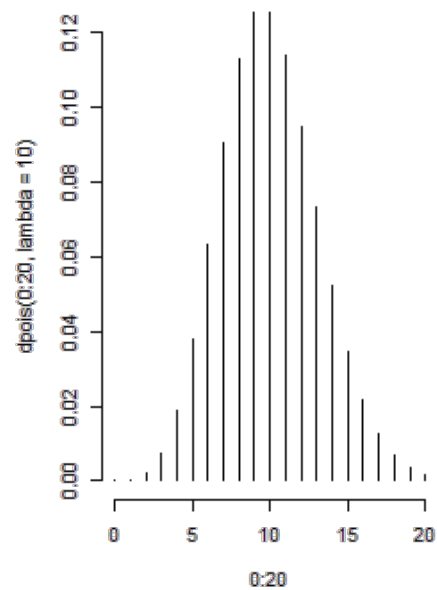
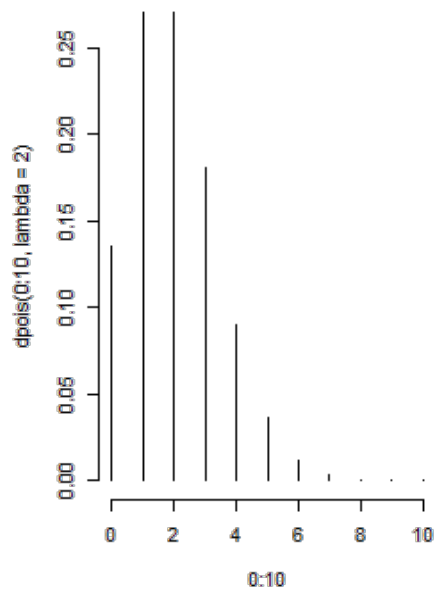
- $X \sim \text{Poisson}(t\lambda)$ if

$$P(X = x) = \frac{(t\lambda)^x e^{-t\lambda}}{x!}$$

For $x = 0, 1, \dots$

- The mean of the Poisson is $E[X] = t\lambda$, thus $E[X/t] = \lambda$
- The variance of the Poisson is $\text{Var}(X) = t\lambda$.
- The Poisson tends to a normal as $t\lambda$ gets large.

```
par(mfrow = c(1, 3))  
plot(0 : 10, dpois(0 : 10, lambda = 2), type = "h", frame = FALSE)  
plot(0 : 20, dpois(0 : 20, lambda = 10), type = "h", frame = FALSE)  
plot(0 : 200, dpois(0 : 200, lambda = 100), type = "h", frame = FALSE)
```



Poisson distribution

Sort of, showing that the mean and variance are equal

```
x <- 0 : 10000; lambda = 3
mu <- sum(x * dpois(x, lambda = lambda))
sigmasq <- sum((x - mu)^2 * dpois(x, lambda = lambda))
c(mu, sigmasq)
```

```
[1] 3 3
```

Example: Leek Group Website Traffic

- Consider the daily counts to Jeff Leek's web site

<http://biostat.jhsph.edu/~jleek/>

- Since the unit of time is always one day, set $t = 1$ and then the Poisson mean is interpreted as web hits per day. (If we set $t = 24$, it would be web hits per hour).

Website data

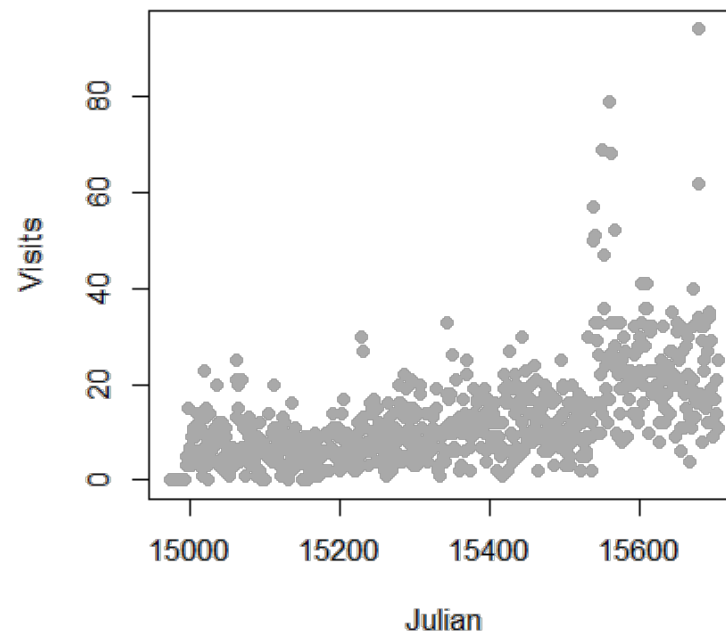
```
download.file("https://dl.dropboxusercontent.com/u/7710864/data/gaData.rda",destfile="./data/gaData.rda")
load("./data/gaData.rda")
gaData$julian <- julian(gaData$date)
head(gaData)
```

| | date | visits | simplystats | julian |
|---|------------|--------|-------------|--------|
| 1 | 2011-01-01 | 0 | 0 | 14975 |
| 2 | 2011-01-02 | 0 | 0 | 14976 |
| 3 | 2011-01-03 | 0 | 0 | 14977 |
| 4 | 2011-01-04 | 0 | 0 | 14978 |
| 5 | 2011-01-05 | 0 | 0 | 14979 |
| 6 | 2011-01-06 | 0 | 0 | 14980 |

<http://skardhamar.github.com/rga/>

Plot data

```
plot(gaData$julian,gaData$visits,pch=19,col="darkgrey",xlab="Julian",ylab="Visits")
```



Linear regression

$$NH_i = b_0 + b_1 JD_i + e_i$$

NH_i - number of hits to the website

JD_i - day of the year (Julian day)

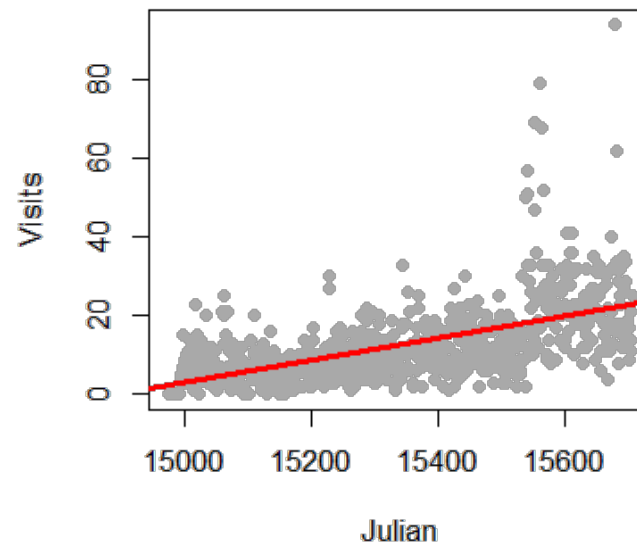
b_0 - number of hits on Julian day 0 (1970-01-01)

b_1 - increase in number of hits per unit day

e_i - variation due to everything we didn't measure

Linear regression line

```
plot(gaData$julian,gaData$visits,pch=19,col="darkgrey",xlab="Julian",ylab="Visits")  
lm1 <- lm(gaData$visits ~ gaData$julian)  
abline(lm1,col="red",lwd=3)
```



Aside, taking the log of the outcome

- Taking the natural log of the outcome has a specific interpretation.
- Consider the model

$$\log(\text{NH}_i) = b_0 + b_1 \text{JD}_i + e_i$$

NH_i - number of hits to the website

JD_i - day of the year (Julian day)

b_0 - log number of hits on Julian day 0 (1970-01-01)

b_1 - increase in log number of hits per unit day

e_i - variation due to everything we didn't measure

Exponentiating coefficients

- $e^{E[\log(Y)]}$ geometric mean of Y .
 - With no covariates, this is estimated by $e^{\frac{1}{n} \sum_{i=1}^n \log(y_i)} = (\prod_{i=1}^n y_i)^{1/n}$
- When you take the natural log of outcomes and fit a regression model, your exponentiated coefficients estimate things about geometric means.
- e^{β_0} estimated geometric mean hits on day 0
- e^{β_1} estimated relative increase or decrease in geometric mean hits per day
- There's a problem with logs with you have zero counts, adding a constant works

```
round(exp(coef(lm(I(log(gaData$visits + 1)) ~ gaData$julian))), 5)
```

```
(Intercept) gaData$julian  
0.000      1.002
```

Linear vs. Poisson regression

Linear

$$NH_i = b_0 + b_1 JD_i + e_i$$

or

$$E[NH_i | JD_i, b_0, b_1] = b_0 + b_1 JD_i$$

Poisson/log-linear

$$\log(E[NH_i | JD_i, b_0, b_1]) = b_0 + b_1 JD_i$$

or

$$E[NH_i | JD_i, b_0, b_1] = \exp(b_0 + b_1 JD_i)$$

Multiplicative differences

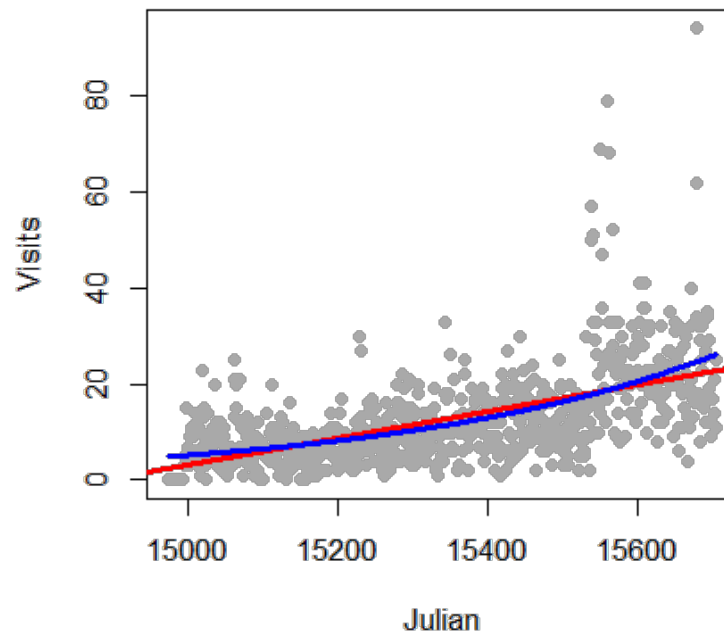
$$E[NH_i | JD_i, b_0, b_1] = \exp(b_0 + b_1 JD_i)$$

$$E[NH_i | JD_i, b_0, b_1] = \exp(b_0) \exp(b_1 JD_i)$$

If JD_i is increased by one unit, $E[NH_i | JD_i, b_0, b_1]$ is multiplied by $\exp(b_1)$

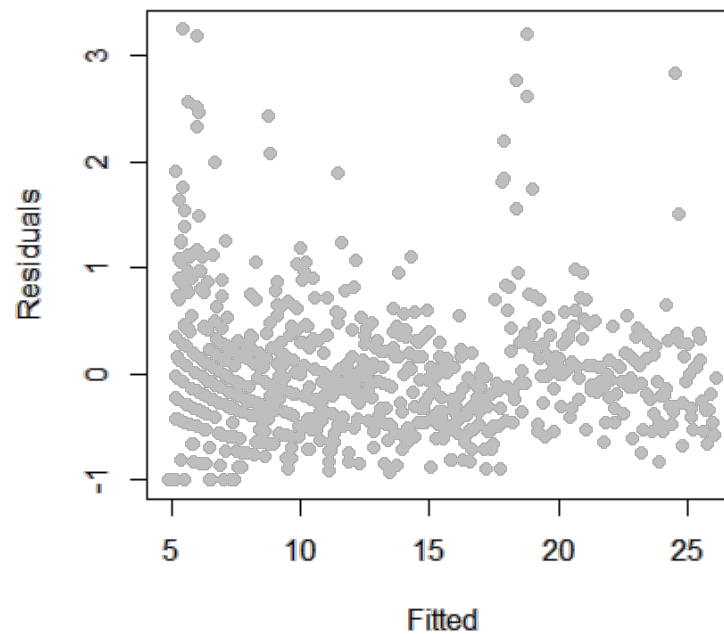
Poisson regression in R

```
plot(gaData$julian,gaData$visits,pch=19,col="darkgrey",xlab="Julian",ylab="Visits")  
glm1 <- glm(gaData$visits ~ gaData$julian,family="poisson")  
abline(lm1,col="red",lwd=3); lines(gaData$julian,glm1$fitted,col="blue",lwd=3)
```



Mean-variance relationship?

```
plot(glm1$fitted,glm1$residuals,pch=19,col="grey",ylab="Residuals",xlab="Fitted")
```



Model agnostic standard errors

```
library(sandwich)
confint.agnostic <- function (object, parm, level = 0.95, ...)
{
  cf <- coef(object); pnames <- names(cf)
  if (missing(parm))
    parm <- pnames
  else if (is.numeric(parm))
    parm <- pnames[parm]
  a <- (1 - level)/2; a <- c(a, 1 - a)
  pct <- stats::format.perc(a, 3)
  fac <- qnorm(a)
  ci <- array(NA, dim = c(length(parm), 2L), dimnames = list(parm,
                                                                pct))

  ses <- sqrt(diag(sandwich::vcovHC(object)))[parm]
  ci[] <- cf[parm] + ses %0% fac
  ci
}
```

<http://stackoverflow.com/questions/3817182/vcovhc-and-confidence-interval>

Estimating confidence intervals

```
confint(glm1)
```

| | 2.5 % | 97.5 % |
|----------------|-----------|------------|
| (Intercept) | -34.34658 | -31.159716 |
| gaData\$julian | 0.00219 | 0.002396 |

```
confint.agnostic(glm1)
```

| | 2.5 % | 97.5 % |
|----------------|------------|------------|
| (Intercept) | -36.362675 | -29.136997 |
| gaData\$julian | 0.002058 | 0.002528 |

Rates

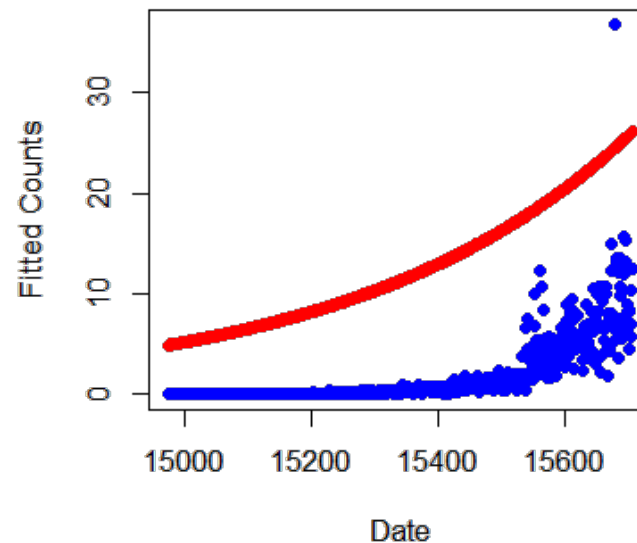
$$E[NHSS_i | JD_i, b_0, b_1] / NH_i = \exp(b_0 + b_1 JD_i)$$

$$\log(E[NHSS_i | JD_i, b_0, b_1]) - \log(NH_i) = b_0 + b_1 JD_i$$

$$\log(E[NHSS_i | JD_i, b_0, b_1]) = \log(NH_i) + b_0 + b_1 JD_i$$

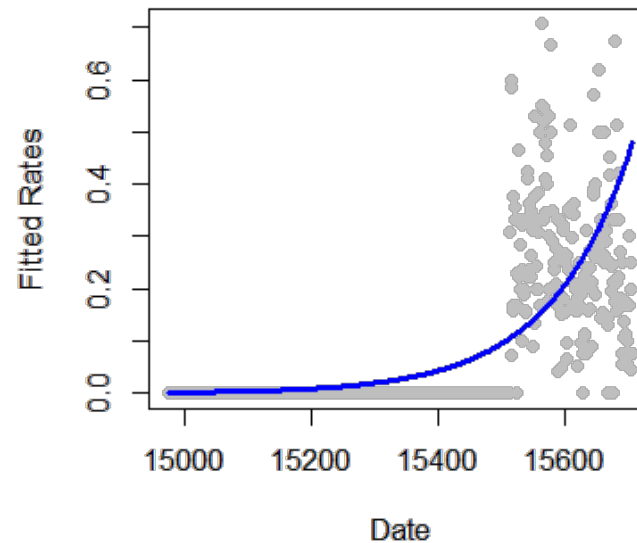
Fitting rates in R

```
glm2 <- glm(gaData$simplystats ~ julian(gaData$date), offset=log(visits+1),  
            family="poisson", data=gaData)  
plot(julian(gaData$date), glm2$fitted, col="blue", pch=19, xlab="Date", ylab="Fitted Counts")  
points(julian(gaData$date), glm1$fitted, col="red", pch=19)
```



Fitting rates in R

```
glm2 <- glm(gaData$simplystats ~ julian(gaData$date), offset=log(visits+1),  
            family="poisson", data=gaData)  
plot(julian(gaData$date), gaData$simplystats/(gaData$visits+1), col="grey", xlab="Date",  
      ylab="Fitted Rates", pch=19)  
lines(julian(gaData$date), glm2$fitted/(gaData$visits+1), col="blue", lwd=3)
```



More information

- [Log-linear models and multiway tables](#)
- [Wikipedia on Poisson regression](#), [Wikipedia on overdispersion](#)
- [Regression models for count data in R](#)
- [pscl package](#) - the function *zeroinfl* fits zero inflated models.