

$\theta_1$  наименьшее  
сост  $\theta_3$  среднее  
сост  $\theta'_3$  наибольшее  
сост

$$D[\theta_1] = \frac{\theta^2}{3n} \quad D[\tilde{\theta}'_3] = \frac{\theta^2}{n(n+2)}$$

$$\forall \theta > 0 \quad \frac{\theta^2}{n(n+2)} < \frac{\theta^2}{3n} \Rightarrow \tilde{\theta}'_3 \text{ самая эффективная}$$

ТЗ

$$P(x) = \begin{cases} \frac{1}{\theta} e^{-\frac{x}{\theta}}, & x \geq 0, \theta > 0 \\ 0, & x < 0 \end{cases} \quad \text{по заданию } n=3$$

$$\tilde{\theta}_1 = \bar{x}, \quad \tilde{\theta}_2 = x_{(2)}$$

$$\begin{aligned} M[\tilde{\theta}_1] &= \int_{-\infty}^{+\infty} x P(x) dx = \int_0^{+\infty} \frac{x}{\theta} e^{-\frac{x}{\theta}} dx = \int_0^{+\infty} x e^{-\frac{x}{\theta}} dx = \int_0^{+\infty} x e^{-\frac{x}{\theta}} dx = \int_0^{+\infty} x e^{-\frac{x}{\theta}} dx \\ &= \frac{1}{\theta} \left( x \cdot (-\theta e^{-\frac{x}{\theta}}) + \int_0^{+\infty} \theta e^{-\frac{x}{\theta}} dx \right) = \\ &= \frac{1}{\theta} \left( -\theta^2 e^{-\frac{x}{\theta}} \Big|_0^{+\infty} \right) = \frac{1}{\theta} (+\theta^2) = \theta \end{aligned}$$



$$M[\xi^2] = \int_{-\infty}^{+\infty} x^2 p(x) dx = \int_0^{\infty} \frac{x^2}{\theta} e^{-\frac{x}{\theta}} dx =$$

$$= \theta (x^2 (e^{-\frac{x}{\theta}} (-\theta)) + 2\theta \int_0^{\infty} x e^{-\frac{x}{\theta}} dx =$$

$$= \theta \cdot 2\theta = 2\theta^2$$

$$D[\xi] = 2\theta^2 - \theta^2 = \theta^2$$

$$\tilde{\theta}_1 = \bar{x} = \frac{1}{n} \sum_{i=1}^n x_i$$

$$M[\tilde{\theta}_1] = M[\frac{1}{n} \sum_{i=1}^n x_i] = M[\xi] = \theta$$

$\Rightarrow \tilde{\theta}_1$  несмещ.

$$D[\tilde{\theta}_1] = D[\frac{1}{n} \sum_{i=1}^n x_i] = \frac{1}{n^2} \sum p[\xi] =$$

$$= \frac{\theta^2}{n}$$

$$\tilde{\theta}_2 = x_{(2)}$$

$$x(x) = n \frac{e^{-\frac{x}{\theta}}}{\theta} (n-1) e^{-\frac{(n-2)x}{\theta}} (1 - e^{-\frac{x}{\theta}})$$

$$M[x_{(2)}] = n(n-1) \int_0^{\infty} \frac{x}{\theta} (e^{-\frac{(n-1)x}{\theta}} - e^{-\frac{nx}{\theta}}) dx =$$



$$= n(n-1) \int_0^{\infty} dt (e^{-(n-1)t} - e^{-nt}) dt =$$

$$= n(n-1) \left( \frac{1}{n-1} \int_0^{\infty} e^{-(n-1)t} dt - \frac{1}{n} \int_0^{\infty} e^{-nt} dt \right) =$$

$$= \left( \frac{n(n-1)}{(n-1)^2} - \frac{n(n-1)}{n^2} \right) \Theta = \Theta \left( \frac{n}{n-1} - \frac{n-1}{n} \right)$$

$$= \frac{2n-1}{n(n-1)} \Theta \quad \text{clearly.}$$

$$\tilde{\Theta}_2^1 = \frac{n(n-1)}{2n-1} \tilde{\Theta}_2 - \text{recursion}$$

$$M[\tilde{\Theta}_2^2] = n(n-1) \Theta^2 \int_0^{\infty} t^2 (e^{-(n-1)t} - e^{-nt}) dt =$$

$$= n(n-1) \Theta^2 \left( \frac{2}{n-1} \int_0^{\infty} t e^{-(n-1)t} dt - \frac{2}{n} \int_0^{\infty} t e^{-nt} dt \right) =$$

$$= n(n-1) \Theta^2 \left( \frac{2}{(n-1)^2} \int_0^{\infty} e^{-(n-1)t} dt - \frac{2}{n^2} \int_0^{\infty} e^{-nt} dt \right) =$$

$$= \left( \frac{2n}{(n-1)^2} - \frac{2(n-1)}{n^2} \right) \Theta^2 = 2\Theta^2 \cdot \frac{3n^2 - 3n + 1}{n^2(n+1)^2}$$

$$\begin{aligned}
 \mathcal{D}[\tilde{\Theta}_2] &= \frac{3n^2 - 3n + 1}{n^2(n+1)^2} \mathcal{D}^2 - \frac{4n^2 - 4n + 1}{n^2(n+1)^2} \mathcal{D}^2 \\
 &= \frac{2n^2 - 2n + 1}{n^2(n+1)^2} \mathcal{D}^2
 \end{aligned}$$

$$\mathcal{D}[\tilde{\Theta}_2'] = \frac{2n^2 - 2n + 1}{4n^2 - 4n + 1} \mathcal{D}^2$$



$$a) D[\tilde{\theta}] = \frac{\sigma^2}{n} \quad D[\tilde{\theta}_2'] = \frac{\sigma^2 (2 - \frac{2}{n} + \frac{1}{n^2})}{(n - \frac{4}{n} + \frac{1}{n^2})}$$

$$\Rightarrow \exists n_0: \forall n \geq n_0 \quad \forall \theta \Rightarrow D[\tilde{\theta}] < D[\tilde{\theta}_2']$$

$$\Rightarrow \tilde{\theta}_1 \text{ эффективнее } \tilde{\theta}_2'$$

б) Проверим модель на регрессию

1)  $P(x)$  не пр. гл. на  $\theta > 0$

$$2) \frac{\partial}{\partial \theta} \int_{-\infty}^{+\infty} P(x) dx = \frac{\partial}{\partial \theta} \left( \frac{1}{\theta} \int_{-\infty}^{+\infty} e^{-\frac{x}{\theta}} dx \right) =$$

$$= \frac{\partial}{\partial \theta} \left( \frac{1}{\theta} \cdot \theta \cdot e^{-\frac{x}{\theta}} \right) \Big|_0^{+\infty} = \frac{\partial}{\partial \theta} (1) = 0$$

$$\int_{-\infty}^{+\infty} \frac{\partial}{\partial \theta} P(x) dx = \int_{-\infty}^{+\infty} \frac{\partial}{\partial \theta} \frac{e^{-\frac{x}{\theta}}}{\theta} dx =$$

$$= \int_{-\infty}^{+\infty} \frac{\frac{x}{\theta^2} e^{-\frac{x}{\theta}} \cdot \theta - e^{-\frac{x}{\theta}}}{\theta^2} dx = \int_{-\infty}^{+\infty} \frac{x e^{-\frac{x}{\theta}} - \theta e^{-\frac{x}{\theta}}}{\theta^3} dx$$

$$= \frac{1}{\theta^2} \left( \int_{-\infty}^{+\infty} \frac{x}{\theta} e^{-\frac{x}{\theta}} - \int_{-\infty}^{+\infty} e^{-\frac{x}{\theta}} dx \right) = 0$$



$$3) \ln p(x) = \ln \frac{e^{-\frac{x}{\theta}}}{\theta} = -\frac{x}{\theta} - \ln \theta$$

$$\left( \frac{\partial \ln p(x)}{\partial \theta} \right)^2 = \left( \frac{x}{\theta^2} - \frac{1}{\theta} \right)^2$$

$$I(\theta) = E \left[ \left( \frac{\partial \ln p(x)}{\partial \theta} \right)^2 \right] = \int_0^{\infty} \frac{(x-\theta)^2}{\theta^4} \frac{e^{-\frac{x}{\theta}}}{\theta} dx =$$

$$= \frac{1}{\theta^2} \text{ так как } \theta > 0 \text{ и } I(\theta) > 0$$

регулярность ~~оценки~~  $\Rightarrow$  модель ~~регулярна~~ ~~оценки~~

$\tilde{\theta}_1 = \bar{x}$  модель регулярна,

$\tilde{\theta}_1 = \text{несмещ.}$   $D[\tilde{\theta}_1] = \frac{\theta^2}{n}$  оп.

$\tilde{\theta}_2' = \frac{n(n-1)}{2n-1} \bar{x}(\bar{x})$   $\Rightarrow$  оценка  $\forall n$   $\forall$  параметра  $\theta$  регулярна по ф.р.  $p_{\theta}$   $\Rightarrow$  несмещ.

$$D[\tilde{\theta}_2'] = \theta^2 \left[ \frac{2n^2 - 2n + 1}{4n^2 - 4n + 1} \right] - \text{оп. на } n$$

$\Rightarrow \tilde{\theta}_2' - \text{регулярна}$



Выводимые из условия  
нр-ва Крамера  $\rho \neq 0$

$$\Phi[\theta_1] \geq \frac{1}{3} \frac{1}{\theta^2} = \frac{\theta^2}{3}$$

то  $\tau \in$  единственности  $\theta \Phi$  держат  
 $\theta_1$  эффективна