



Graphical models II

Elements of Machine Learning

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with content and images from

Pattern Recognition and Machine Learning, Christopher M. Bishop



About this lecture

Methods for inference in graphical models

- Factor graphs
- Sum-Product algorithm
- Max-Sum algorithm
- Sampling methods



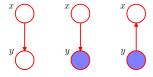
Inference in graphical models

- Graphical model structure allow efficient algorithms for inference
- By for instance passing local messages around the graph



A basic example

Graphical representation of Bayes' theorem



Suppose we are interested in the posterior distribution of x given y (right figure).

We can write the joint distribution as p(x, y) = p(x)p(y|x) and using the sum and product rules of probability

$$p(y) = \sum_{x'} p(y|x')p(x')$$

which we can combine with Bayes' to give us

$$p(x|y) = \frac{p(y|x)p(x)}{p(y)}$$



Joint distribution

$$p(\mathbf{x}) = \frac{1}{Z} \psi_{1,2}(x_1, x_2) \psi_{2,3}(x_2, x_3) \dots \psi_{N-1,N}(x_{N-1}, x_N)$$

Marginal probability of x_n by definition, sum over all variables except x_n

$$p(x_n) = \sum_{x_1} \cdots \sum_{x_{n-1}} \sum_{x_{n+1}} \cdots \sum_{x_N} p(\mathbf{x})$$

which is exponential in N.

With chain we can do better, because each potential $\psi_{N-1,N}(x_{N-1},x_N)$ is the only one that depends on x_N . So we only need this term in the summation.

$$\sum_{\mathsf{X}_{\mathsf{N}}} = \psi_{\mathsf{N}-1,\mathsf{N}}(\mathsf{X}_{\mathsf{N}-1},\mathsf{X}_{\mathsf{N}})$$



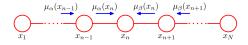
We can proceed recursively from both ends of the chain towards x_n to give

$$\rho(x_n) = \frac{1}{Z} \underbrace{\left(\sum_{x_{n-1}} \psi_{n-1,n}(x_{n-1}, x_n) \dots \left(\sum_{x_2} \psi_{2,3}(x_2, x_3) \left(\sum_{x_1} \psi_{1,2}(x_1, x_2)\right)\right) \dots\right)}_{\mu_{\alpha}(x_n)} \underbrace{\left(\sum_{x_{n+1}} \psi_{n,n+1} \dots \left(\sum_{x_n} \psi_{N-1,N}(x_{N_1}, x_N)\right) \dots\right)}_{\mu_{\beta}(x_n)}$$

Message passing

- $\mu_{alpha}(x_n)$ forward through the chain
- $\mu_{beta}(x_n)$ backward through the chain





Forward messages

$$\mu_{\alpha}(x_n) = \begin{cases} \sum_{x_1} \psi_{1,2}(x_1, x_2) & \text{for } n = 2\\ \sum_{x_{n-1}} \psi_{n-1,n}(x_{n-1}, x_n) \mu_{\alpha}(x_{n-1}) & \text{else.} \end{cases}$$

Backward messages

$$\mu_{\beta}(x_n) = \begin{cases} \sum_{x_N} \psi_{N-1,N}(x_{N-1}, x_N) & \text{for } n = N \\ \sum_{x_{n+1}} \psi_{n+1,n}(x_{n+1}, x_n) \mu_{\beta}(x_{n+1}) & \text{else.} \end{cases}$$

Normalization constant (choose any node n and sum messages)

$$Z = \sum_{x_n} \mu_{\alpha}(x_n) \mu_{\beta}(x_n)$$



What if want to compute marginals $p(x_n)$ for all n?

wasteful as many messages are repeated

Solution:

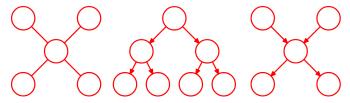
- **1** compute and store all forward messages, $\mu_{\alpha}(x_n)$
- **2** compute and store all backward messages, $\mu_{\beta}(x_n)$
- **3** compute Z at any node x_m



Trees

Efficient inference possible in tree-structured graphs as well...

- Undirected tree one and only one path between any pair of nodes (no loops)
- Directed tree a single node (root) with no parents, and all other nodes have one parent
 - Moralization leads to an undirected tree
- Directed polytree nodes with more than one parent exist but only one path (ignoring edge direction) between any pair of nodes

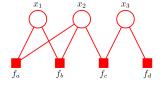


Undirected tree, directed tree and a directed polytree



Factor graphs¹²

A factor graph is a bipartite graph representing the factorization of a function



- one variable node for each variable
- one factor node for each function
- one edge between x_i and f_j if x_i is a variable in f_j

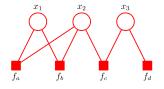


¹wikipedia.org

 $^{{}^2 \}verb|https://www.doc.ic.ac.uk/-mpd37/teaching/ml_tutorials/2016-11-09-Svensson-BP.pdf| Jens Petersen — Graphical models II$

Factor graphs¹²

A factor graph is a bipartite graph representing the factorization of a function



Write distribution as product of factors

$$p(\mathbf{x}) = \prod_{s} f_s(\mathbf{x}_s)$$

where x_s denotes subset of the variables. Example:

$$p(\mathbf{x}) = f_a(x_1, x_2) f_b(x_1, x_2) f_c(x_2, x_3) f_d(x_3)$$



 $^{^{1}}$ wikipedia.org

https://www.doc.ic.ac.uk/-mpd37/teaching/ml_tutorials/2016-11-09-Svensson-BP.pdf Jens Petersen — Graphical models II
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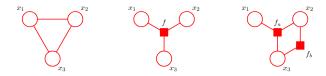
Why factor graphs?

Allows one to

- visualize the problem structure
 - clarify dependencies as well as how a complicated problem can be split into smaller functions of fewer variables
- efficiently compute marginals through algorithms such as the Sum-Product algorithm



Factor graphs from undirected graphs



- left: undirected graph with single clique potential $\psi(x_1, x_2, x_3)$
- middle: factor graph representing the same distribution with $f = \psi(x_1, x_2, x_3)$
- right: another factor graph representing the same distribution with $f_a(x_1, x_2, x_3) f_b(x_1, x_2) = \psi(x_1, x_2, x_3)$.



Factor graphs from undirected graphs

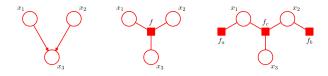


- left: undirected graph with single clique potential $\psi(x_1, x_2, x_3)$
- middle: factor graph representing the same distribution with $f = \psi(x_1, x_2, x_3)$
- right: another factor graph representing the same distribution with $f_a(x_1, x_2, x_3) f_b(x_1, x_2) = \psi(x_1, x_2, x_3)$.

Convertion: create variable nodes corresponding to nodes in the original graph, and then create additional factor nodes corresponding to the maximal cliques \mathbf{x}_s , with the factors $f_s(\mathbf{x}_s)$ equal to the clique potentials.



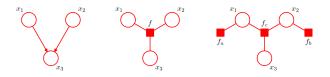
Factor graphs from directed graphs



- left: directed graph with factorization $p(x_1)p(x_2)p(x_3|x_1,x_2)$
- middle: factor graph of the same distribution with $f(x_1, x_2, x_3) = p(x_1)p(x_2)p(x_3|x_1, x_2)$
- right: another factor graph of the same distribution with $f_a(x_1) = p(x_1)$, $f_b(x_2) = p(x_2)$ and $f_c(x_1, x_2, x_3) = p(x_3 | x_1, x_2)$.



Factor graphs from directed graphs



- left: directed graph with factorization $p(x_1)p(x_2)p(x_3|x_1,x_2)$
- middle: factor graph of the same distribution with $f(x_1, x_2, x_3) = p(x_1)p(x_2)p(x_3|x_1, x_2)$
- right: another factor graph of the same distribution with $f_a(x_1) = p(x_1)$, $f_b(x_2) = p(x_2)$ and $f_c(x_1, x_2, x_3) = p(x_3 | x_1, x_2)$.

Convertion: create variable nodes corresponding to nodes in the original graph, and then create factor nodes corresponding to the conditional distributions, and then finally add the appropriate links.



The algorithm known as Belief Propagation is a special case of the Sum-Product algorithm

- Goal: to obtain an efficient, exact inference algorithm for finding (multiple) marginals
 - Messages from variable to factor nodes
 - Messages from factor to variable nodes

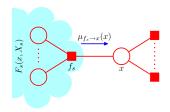


The algorithm known as Belief Propagation is a special case of the Sum-Product algorithm

- Goal: to obtain an efficient, exact inference algorithm for finding (multiple) marginals
 - Messages from variable to factor nodes
 - Messages from factor to variable nodes

Why is it called belief propagation? A message to a variable node can be thought of as the message senders 'beliefs' about the state of that variable node

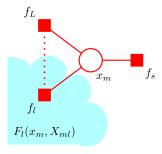




Each message from a factor node (f_s) to variable node (x) is the product of the factor with messages from all neighboring nodes marginalized over all variables except the one associated with x.

$$\mu_{f_s \to x}(x) = \sum_{x_1} \cdots \sum_{x_M} f_s(x, x_1, \dots, x_M) \prod_{m \in ne(f_s) \setminus x} \mu_{x_m \to f_s}(x_m)$$



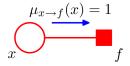


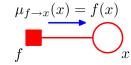
Each message from a variable node (x) to a factor node (f_s) is the product of all messages arriving from neighboring factor nodes except the recipient.

$$\mu_{\mathsf{x}_m \to \mathsf{f}_{\mathsf{s}}}(\mathsf{x}_m) = \prod_{\mathsf{f} \in \mathsf{ne}(\mathsf{x}_m) \setminus \mathsf{f}_{\mathsf{s}}} \mu_{\mathsf{f}_{\mathsf{f}} \to \mathsf{x}_m}(\mathsf{x}_m)$$



Initialization





Initialize messages at the leaves of the graph

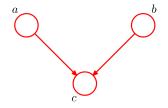


To compute local marginals

- pick an arbitrary node as the root
- 2 compute and propagate messages from the leaf nodes towards the root, storing received messages at every node
 - a node can send a messages towards the root once it has received messages from all of its other neighbours
- compute and propagate messages from the root towards the leaf nodes, storing received messages at every node
- Ocompute the product of received messages at each node for which the marginal is required, and normalize if needed (how do we normalize? Remember the procedure for the chain graph.)



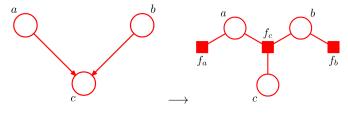
Suppose we are given the following graph and corresponding joint distribution



$$p(a,b,c) = p(c|a,b)p(a)p(b),$$



We can convert this into the following factor graph



$$p(a,b,c) = p(c|a,b)p(a)p(b),$$



Additionally, the distributions for a and b and the conditional distribution of c given a and b are

	а	Ь	p(c=0 a,b)	p(c=1 a,b)	$a \longrightarrow b$
h n(h)	0	0	0.1	0.9	$\left(\begin{array}{c} \downarrow \\ \downarrow \\ f_{a} \end{array}\right)$
$a \mid p(a) = 0 \mid p(b) = 0$	0	1	0.2	0.8	
0 0.7 1 0.5	0	2	0.3	0.7 f	f_b
1 0.3 0.2	1	0	0.3	0.7	
2 0.3	1	1	0.2	0.8	()
	1	2	0.1	0.9	c



	а	Ь	p(c=0 a,b)	p(c=1 a,b)	$a \longrightarrow b$
$b \mid p(b)$	0	0	0.1	0.9	$\begin{cases} \\ \\ \\ \end{cases} \end{cases} f_c \end{cases} $
$a \mid p(a) \stackrel{b}{\longrightarrow} p(b)$	0	1	0.2	0.8	
0 0.7 1 0.3	0	2	0.3	$0.7 extit{f}_{c}$	f_b
1 0.2	1	0	0.3	0.7	
2 0.3	1	1	0.2	0.8	()
	1	2	0.1	0.9	c

We start by selecting a node as root, lets choose c and begin passing messages from the leafs towards the root.

$$\mu_{f_a \to a}(a) = \begin{array}{c|c} a & f_a \\ \hline 0 & 0.7 \\ 1 & 0.3 \end{array}$$

The message from factor f_a to variable a



$$\mu_{f_a \to a}(a) = \begin{array}{c|c} a & f_a \\ \hline 0 & 0.7 \\ 1 & 0.3 \\ \\ \mu_{f_b \to b}(b) = \begin{array}{c|c} b & f_b \\ \hline 0 & 0.5 \\ 1 & 0.2 \\ 3 & 0.3 \end{array}$$

The message from factor f_b to variable b



$$\mu_{f_a \to a}(a) = \frac{\begin{array}{c|c} a & f_a \\ \hline 0 & 0.7 \\ \hline 1 & 0.3 \end{array}$$

$$\mu_{f_b \to b}(b) = \begin{array}{c|c} \frac{b & f_b \\ \hline 0 & 0.5 \\ \hline 1 & 0.2 \\ \hline 3 & 0.3 \end{array}$$

$$\mu_{a \to f_c}(a) = \mu_{f_a \to a}(a) = \begin{array}{c|c} \frac{a & f_a \\ \hline 0 & 0.7 \\ \hline 1 & 0.3 \end{array}$$

The message from variable a to factor f_c



$$\mu_{f_b \to b}(b) = \begin{array}{c|c} b & f_b \\ \hline 0 & 0.5 \\ 1 & 0.2 \\ 3 & 0.3 \end{array}$$

$$\mu_{a \to f_c}(a) = \mu_{f_a \to a}(a) = \begin{array}{c|c} a & f_a \\ \hline 0 & 0.7 \\ 1 & 0.3 \end{array}$$

$$\mu_{b \to f_c}(b) = \mu_{f_b \to b}(b) = \begin{array}{c|c} b & f_b \\ \hline 0 & 0.5 \\ 1 & 0.2 \\ 3 & 0.3 \end{array}$$

The message from variable b to factor f_c



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$$\begin{split} \mu_{a \to f_{c}}(a) = & \mu_{f_{a} \to a}(a) = \frac{\begin{array}{ccc} a & f_{a} \\ \hline 0 & 0.7 \\ 1 & 0.3 \\ \end{array} \\ \mu_{b \to f_{c}}(b) = & \mu_{f_{b} \to b}(b) = \begin{array}{ccc} \frac{b & f_{b} \\ \hline 0 & 0.5 \\ 1 & 0.2 \\ \hline 3 & 0.3 \\ \end{array} \\ \mu_{f_{c} \to c}(c) = & \sum_{a} \sum_{b} f_{c}(a, b, c) \mu_{a \to f_{c}}(a) \mu_{b \to f_{c}}(b) \\ & \text{The message from factor } f_{c} \text{ to variable } c \end{split}$$



When multiplying factors, think combining probabilities of events!

$$\mu_{a \to f_c}(a) = \mu_{f_a \to a}(a) = \frac{\begin{array}{c} a & f_a \\ \hline 0 & 0.7 \\ 1 & 0.3 \end{array}}{\mu_{b \to f_c}(b) = \mu_{f_b \to b}(b) = \begin{array}{c} b & f_b \\ \hline 0 & 0.5 \\ 1 & 0.2 \\ 3 & 0.3 \end{array}}$$

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$$\mu_{f_c \rightarrow c}(c) = \sum_{a} \sum_{b} f_c(a,b,c) \mu_{a \rightarrow f_c}(a) \mu_{b \rightarrow f_c}(b) = \sum_{a} \sum_{b} f_c(a,b,c) \begin{array}{ccc} a & b & f_a \cdot f_b \\ \hline 0 & 0 & 0.7 \cdot 0.5 \\ 0 & 1 & 0.7 \cdot 0.2 \\ \hline 0.7 \cdot 0.2 \\ 0.7 \cdot 0.3 \\ 1 & 0 & 0.3 \cdot 0.5 \\ \hline 1 & 1 & 0.3 \cdot 0.2 \\ \end{array}$$
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 $0.3 \cdot 0.3$

Leading to...

$$\mu_{a \to f_c}(a) = \mu_{f_a \to a}(a) = \begin{array}{c|c} a & f_a \\ \hline 0 & 0.7 \\ 1 & 0.3 \end{array}$$

$$\mu_{b \to f_c}(b) = \mu_{f_b \to b}(b) = \begin{array}{c|c} b & f_b \\ \hline 0 & 0.5 \\ 1 & 0.2 \\ 3 & 0.3 \end{array}$$

$$\mu_{f_C \to c}(c) = \sum_{a} \sum_{b} f_c(a,b,c) \mu_{a \to f_C}(a) \mu_{b \to f_C}(b) = \sum_{a} \sum_{b} f_c(a,b,c) \begin{array}{ccc} a & b & f_a \cdot f_b \\ \hline 0 & 0 & 0.35 \\ 0 & 1 & 0.14 \\ 1 & 0.14 \\ 1 & 0 & 0.15 \\ 1 & 1 & 0.06 \\ 1 & 2 & 0.09 \\ \hline \text{Slide } 21/47 \\ \end{array}$$



	а	Ь	p(c=0 a,b)	p(c=1 a,b)	$a \longrightarrow b$
b p(b)	0	0	0.1	0.9	$\left(\begin{array}{c} \left(\begin{array}{c} \left(\right) \\ \left(\begin{array}{c} \left(\end{array}\right) \\ \end{array} \right) \end{array} \right) \end{array} \right) \end{array}\right) \end{array}\right)$
$a \mid p(a) \stackrel{b}{\longrightarrow} p(b)$	0	1	0.2	0.8	
0 0.7 1 0.3	()	2	0.3	0.7	f_a f_b
$1 \mid 0.3 \frac{1}{3} \mid \frac{0.2}{0.3}$	1	0	0.3	0.7	
3 0.3	1	1	0.2	0.8	()
	1	2	0.1	0.9	c

Writing out the factor f_c

	a	Ь	p(c=0 a,b)	p(c=1 a,b)		a	Ь	$f_a \cdot f_b$
	0	0	0.1	0.9	_	0	0	0.35
$\mu_{f_c \to c}(c) = \sum \sum$	0	1	0.2	8.0	(0	1	0.14
	0	2	0.3	0.7	٠ (0	2	0.21
a b	1	0	0.3	0.7		1	0	0.15
	1	1	0.2	8.0		1	1	0.06
	1	2	0.1	0.9		1	2	0.09



				а	Ь	p(c=0 a,b)	p(c=1 a,b)	$a \longrightarrow b$
		b n(b)	n(h)	0	0	0.1	0.9	$\left(\begin{array}{c} \downarrow \\ \downarrow \\ f_c \end{array}\right)$
a µ	o(a)		0.5	0	1	0.2	0.8	
$ \begin{array}{c cccc} \hline 0 & 0.7 & 1 \\ 1 & 0.3 & 3 \end{array} $	1	1 0.2	0	2	0.3	0.7	$\overline{f_a}$ $\overline{f_b}$	
	J.		1	0	0.3	0.7		
			0.3	1	1	0.2	0.8	()
				1	2	0.1	0.9	c

Product of f_c , f_a and f_b $\begin{array}{c|cccc}
a & b & p(c = 0, a, b) & p(c = 1, a, b) \\
\hline
0 & 0 & 0.1 \cdot 0.35 & 0.9 \cdot 0.35 \\
0 & 1 & 0.2 \cdot 0.14 & 0.8 \cdot 0.14 \\
\mu_{f_c \to c}(c) = \sum_{a} \sum_{b} \sum_{1} 0 & 2 & 0.3 \cdot 0.21 & 0.7 \cdot 0.21 \\
1 & 1 & 0.3 \cdot 0.15 & 0.7 \cdot 0.15 \\
1 & 1 & 0.2 \cdot 0.06 & 0.8 \cdot 0.06 \\
1 & 2 & 0.1 \cdot 0.09 & 0.9 \cdot 0.09
\end{array}$



				a	Ь	p(c=0 a,b)	p(c=1 a,b)	$a \longrightarrow b$
	Ь	b p(b)	0	0	0.1	0.9	$\begin{cases} \\ \\ \end{cases} f_c \end{cases} $	
	$\begin{array}{c ccccccccccccccccccccccccccccccccccc$	$\frac{D}{D}$	0	1	0.2	0.8		
		1 0.2	0	2	0.3	0.7	f_a f_b	
1 0.3 3	J		1	0	0.3	0.7		
	3 0	0.5	1	1	0.2	0.8	()	
				1	2	0.1	0.9	c

Giving... Now all that is needed is to marginalize over a and b

	а	Ь	p(c=0,a,b)	p(c=1,a,b)	
	0	0	0.035	0.315	
	0	1	0.028	0.112	
$\mu_{f_c \to c}(c) = \sum_{c} \sum_{c}$	0	2	0.063	0.147	
a b	1	0	0.045	0.105	
	1	1	0.012	0.048	
	1	2	0.009	0.081	



				а	Ь	p(c=0 a,b)	p(c=1 a,b)	$a \longrightarrow b$
		h n	$b \mid p(b)$	0	0	0.1	0.9	$\left(\begin{array}{c} 1 \\ 1 \end{array}\right) f_{\alpha} \left(\begin{array}{c} 1 \\ 1 \end{array}\right)$
а	p(a)	-0	ρ(Β)	0	1	0.2	0.8	
0	0.7	1	0.5	0	2	0.3	0.7	f_a f_b
1 0.3	1	3 0.3	1	0	0.3	0.7	, , , , , , , , , , , , , , , , , , ,	
	3		1	1	0.2	0.8	()	
				1	2	0.1	0.9	c

Leading to...

$$\mu_{f_c \to c}(c) = \sum_{a} \sum_{b=1}^{a} \frac{b \mid p(c=0, a, b) \mid p(c=1, a, b)}{0 \mid 0 \mid 0.035} \frac{0.315}{0.315}$$

$$0 \mid 1 \mid 0.028 \mid 0.112$$

$$0 \mid 2 \mid 0.063 \mid 0.147$$

$$1 \mid 1 \mid 0.012 \mid 0.048$$

$$1 \mid 2 \mid 0.009 \mid 0.081$$

$$= \frac{c \mid p(c)}{0 \mid 0.192}$$

$$1 \mid 0.808$$



The Sum-Product algorithm - example I

At this point we could start sending the messages backwards (root to leafs) to compute the marginal probabilities of the remaining variable nodes, but these are already given in this case.

As an exercise you can convince yourself that doing this indeed does not change the marginal probabilities of a and b.



The Sum-Product algorithm - example I

At this point we could start sending the messages backwards (root to leafs) to compute the marginal probabilities of the remaining variable nodes, but these are already given in this case.

As an exercise you can convince yourself that doing this indeed does not change the marginal probabilities of a and b.

Note what a message is somewhat unclear from this writeup. However, they may for instance be implemented as vectors. One just needs to carefully handle the products and the input and output dimensions.



The Sum-Product algorithm - example II

 f_a f_b f_b

Unnormalized joint distribution

$$\tilde{p}(\mathbf{x}) = f_a(x_1, x_2) f_b(x_2, x_3) f_c(x_2, x_4)$$



The Sum-Product algorithm - example II

Unnormalized joint distribution

$$\tilde{p}(\mathbf{x}) = f_a(x_1, x_2) f_b(x_2, x_3) f_c(x_2, x_4)$$

Pick x_3 as root and compute messages towards it

$$\mu_{x_1 \to f_a}(x_1) = 1$$

$$\mu_{f_a \to x_2}(x_2) = \sum_{x_1} f_a(x_1, x_2)$$

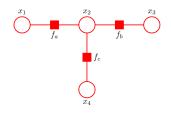
$$\mu_{x_4 \to f_c}(x_4) = 1$$

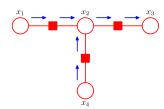
$$\mu_{f_c \to x_2}(x_2) = \sum_{x_4} f_c(x_2, x_4)$$

$$\mu_{x_2 \to f_b}(x_2) = \mu_{f_a \to x_2}(x_2) \mu_{f_c \to x_2}(x_2)$$

$$\mu_{X_2 \to f_b(X_2)} - \mu_{f_a \to X_2}(X_2)\mu_{f_c \to X_2}(X_2)$$

$$\mu_{f_b \to X_3}(X_3) = \sum_{X_2} f_b(X_2, X_3)\mu_{X_2 \to f_b}(X_2)$$
Petersen – Graphical models II







The Sum-Product algorithm - example II

Unnormalized joint distribution

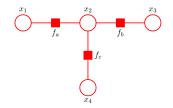
$$\tilde{p}(\mathbf{x}) = f_a(x_1, x_2) f_b(x_2, x_3) f_c(x_2, x_4)$$

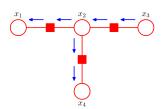
Compute messages from x_3 towards the leafs

$$\begin{split} &\mu_{x_3 \to f_b}(x_3) = 1 \\ &\mu_{f_b \to x_2}(x_2) = \sum_{x_3} f_b(x_2, x_3) \\ &\mu_{x_2 \to f_a}(x_2) = \mu_{f_b \to x_2}(x_2) \mu_{f_c \to x_2}(x_2) \\ &\mu_{f_a \to x_1}(x_1) = \sum_{x_2} f_a(x_1, x_2) \mu_{x_2 \to f_a}(x_2) \end{split}$$

$$\mu_{\mathsf{X}_2 o f_c}(\mathsf{X}_2) = \mu_{f_{\mathsf{a}} o \mathsf{X}_2}(\mathsf{X}_2) \mu_{f_{\mathsf{b}} o \mathsf{X}_2}(\mathsf{X}_2)$$

$$\mu_{f_c o \mathsf{X}_4}(\mathsf{X}_4) = \sum_{\mathsf{X}_2} f_c(\mathsf{X}_2, \mathsf{X}_4) \mu_{\mathsf{X}_2 o f_c}(\mathsf{X}_2)$$
Jens Petersen – Graphical models II







The Sum-Product algorithm - example II - check

$$\widetilde{p}(x_2) = \mu_{f_a \to x_2}(x_2) \mu_{f_b \to x_2}(x_2) \mu_{f_c \to x_2}(x_2)
= \left[\sum_{x_1} f_a(x_1, x_2) \right] \left[\sum_{x_3} f_b(x_2, x_3) \right] \left[\sum_{x_4} f_c(x_2, x_4) \right]
= \sum_{x_1} \sum_{x_2} \sum_{x_4} f_a(x_1, x_2) f_b(x_2, x_3) f_c(x_2, x_4)
= \sum_{x_1} \sum_{x_3} \sum_{x_4} \widetilde{p}(\mathbf{x})$$



The maximum of the joint distribution is not necessarily at same values as the maxima of each marginal.

Example of a joint distribution over two binary variables for which the maximum of the joint distribution occurs for different variable values compared to the maxima of the two marginals.

	x = 0	x = 1
y = 0	0.3	0.4
y = 1	0.3	0.0



Goal: we want to find the most likely variable configuration and how likely it is.

$$\mathbf{x}^{\text{max}} = \arg \max_{\mathbf{x}} p(\mathbf{x})$$
 $p(\mathbf{x}^{\text{max}}) = \max_{\mathbf{x}} p(\mathbf{x})$

Approach: write out the max in terms of components of x

$$\max_{\mathbf{x}} p(\mathbf{x}) = \max_{x_1} \dots \max_{x_M} p(\mathbf{x})$$

and exploit

$$max(ab, ac) = a max(b, c),$$

where M is the number of variables and $a \ge 0$, to allow the exchange of products with maximizations across each factor.



The Max-Sum algorithm - chain

$$\max_{\mathbf{x}} p(\mathbf{x}) = \frac{1}{Z} \max_{x_1} \dots \max_{x_N} (\psi_{x_1, x_2}(x_1, x_2) \dots \psi_{N-1, N}(x_{N-1}, x_N))$$

$$= \frac{1}{Z} \max_{x_1} \left(\max_{x_2} (\psi_{1, 2}(x_1, x_2) (\dots \max_{x_N} \psi_{N-1, N}(x_{N-1}, x_N))) \right)$$



Generalizes to tree-structured factor graph

$$\max_{\mathbf{x}} p(\mathbf{x}) = \max_{X_n} \prod_{f_s \in ne(X_n)} \max_{X_s} f_s(X_n, X_s)$$

maximizing as close to the leaf nodes as possible

This could be called the Max-Product algorithm, but computing the product of a lot of small numbers may lead to numerical issues, so use

$$\ln\left(\max_{\mathbf{x}}p(\mathbf{x})\right)=\max_{\mathbf{x}}\ln p(\mathbf{x})$$

and again, use distributive law

$$\max(a+b,a+c) = a + \max(b,c)$$

to replace all the products with sums.



As with Sum-Product, we can now choose a root node and... initialization (leaf nodes)

• in Sum-Product we have

$$\mu_{x \to f}(x) = 1$$

$$\mu_{f \to x}(x) = f(x)$$

• in Max-Sum we have

$$\mu_{x \to f}(x) = 0$$

$$\mu_{f \to x}(x) = \ln f(x)$$



and recursion

in Sum-Product we have

$$\mu_{f \to x}(x) = \sum_{x_1} \cdots \sum_{x_M} f(x, x_1, \dots, x_M) \prod_{m \in ne(f) \setminus x} \mu_{x_m \to f}(x_m)$$

$$\mu_{x \to f}(x) = \prod_{l \in ne(x_m) \setminus f} \mu_{f_l \to x}(x)$$

• in Max-Sum we have

$$\mu_{f \to x}(x) = \max_{x_1, \dots, x_M} \left(\ln f(x, x_1, \dots, x_M) + \sum_{m \in ne(f_s) \setminus x} \mu_{x_m \to f}(x_m) \right)$$

$$\mu_{x \to f}(x) = \sum_{l \in ne(x) \setminus f} \mu_{f_l \to x}(x)$$



Termination at root node

• in Sum-Product we have

$$p(x) = \prod_{s \in ne(x)} \mu_{f_s \to x}(x)$$

in Max-Sum we have

$$\begin{split} p^{\text{max}} &= \max_{x} \big(\sum_{s \in \textit{ne}(x)} \mu_{\textit{f}_{s \to x}}(x) \big) \\ x^{\text{max}} &= \arg \max_{x} \big(\sum_{s \in \textit{ne}(x)} \mu_{\textit{f}_{s \to x}}(x) \big) \end{split}$$



How to find the remaining variables?



How to find the remaining variables?

- Because multiple configurations may give rise to the same p^{max} we cannnot simply send messages back towards the leafs
- We can store the variables giving rise to maximum in the forward pass for each variable

$$\phi(x) = \arg\max_{x_1, \dots, x_M} \left(\ln f(x, x_1, \dots, x_M) + \sum_{m \in ne(f_s) \setminus x} \mu_{x_m \to f}(x_m) \right)$$

and follow them backwards



Inference in general graphs

Exact

- Junction tree algorithm
 - Often not feasible

Approximate

- Variational methods (covered later in course)
 - Deterministic
- Sampling methods
 - Stochastic
- Loopy belief propagation
 - Handling loops by repeatedly passing messages around until convergence



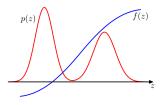
Sampling has been mentioned multiple times

- Ancestral sampling
- Generative models

Sampling distributions can be used to approximate

- the posterior distribution $p(\mathbf{z}|\mathbf{x})$.
- the expectation step in EM.
- exemplify joint distributions





Goal: find the expectation of some function $f(\mathbf{x})$ with respect to a probability distribution

$$E[f] = \int f(\mathbf{z})p(\mathbf{z})d\mathbf{z}$$

Given samples $\mathbf{z}^{(l)}, l \in \{1, \dots L\}$ drawn independently from $p(\mathbf{z})$, we can approximate by

$$\hat{f} = \frac{1}{L} \sum_{l=1}^{L} f(\mathbf{z}^{(l)})$$



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$$E[\hat{\mathbf{z}}] = E[\mathbf{z}]$$



Given samples $\mathbf{z}^{(l)}, l \in \{1, \dots L\}$ drawn independently from $p(\mathbf{z})$, we can approximate by

$$\hat{f} = \frac{1}{L} \sum_{l=1}^{L} f(\mathbf{z}^{(l)})$$

We have

$$E[\hat{\mathbf{z}}] = E[\mathbf{z}]$$

and

$$\operatorname{var}[\hat{f}] = \frac{1}{I}[(f - E[f])^2]$$

Note the variance does not depend on the dimension of \mathbf{x} , but we need independent samples from $p(\mathbf{z})!$



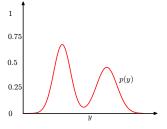
Sampling methods - inverse transform method

Given: a uniformly sampled distribution of z on the interval (0,1), and a distribution of y which is related to z through y = f(z)

Goal: sample y



Sampling methods - inverse transform method Using the cumulative distribution

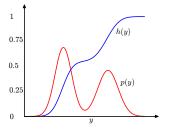


Here p(y) is the distribution we want to sample, where y = f(z).



Sampling methods - inverse transform method

Using the cumulative distribution

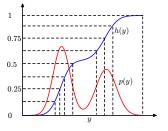


Here p(y) is the distribution we want to sample, where y = f(z). The function h(y) is the cumulative distribution of p(y), that is $h(y) = \int_{-\infty}^{y} p(y)$.



Sampling methods - inverse transform method

Using the cumulative distribution



Here p(y) is the distribution we want to sample, where y=f(z). The function h(y) is the cumulative distribution of p(y), that is $h(y)=\int_{-\infty}^{y}p(y)$. So if we can compute $h^{-1}(z)$, then we can sample p(y) as $y=h^{-1}(z)$.



Sampling methods - inverse transform example

Exponential distribution

$$p(y) = \lambda \exp(-\lambda y)$$

where $0 \le y < \infty$.



Sampling methods - inverse transform example

Exponential distribution

$$p(y) = \lambda \exp(-\lambda y)$$

where $0 \le y < \infty$. Integrating gives

$$h(y) = \int_0^\infty p(y) = 1 - \exp(-\lambda y)$$



Sampling methods - inverse transform example

Exponential distribution

$$p(y) = \lambda \exp(-\lambda y)$$

where $0 \le y < \infty$. Which can be inverted to give

$$h^{-1}(z) = -\lambda^{-1} \ln(1-z)$$

So if we sample \hat{z} uniformly in (0,1) and compute $\hat{y}=h^{-1}(\hat{z})$ then \hat{y} will be sampled according to an exponential distribution.



Assumptions

• Sampling directly from $p(\mathbf{z})$ is difficult but we are able to evaluate $p(\mathbf{z})$ for any value of \mathbf{z} up to an unknown normalization constant Z_p

$$p(z)=\frac{1}{Z_p}\tilde{p}(z)$$

 We can sample from a simpler proposal distribution q(z) for which there exists a constant k such that

$$kq(z) \geq \tilde{p}(z)$$



Assumptions

• Sampling directly from $p(\mathbf{z})$ is difficult but we are able to evaluate $p(\mathbf{z})$ for any value of \mathbf{z} up to an unknown normalization constant Z_p

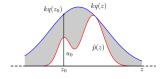
$$p(z) = \frac{1}{Z_p} \tilde{p}(z)$$

 We can sample from a simpler proposal distribution q(z) for which there exists a constant k such that

$$kq(z) \geq \tilde{p}(z)$$

Then we can sample from p(z) using rejection sampling





Samples are rejected if they fall in the grey area between the unnormalized distribution $\tilde{p}(z)$ and the scaled distribution kq(z).

- **1** generate random number z_0 from distribution q(z)
- **2** generate random number u_0 from uniform distribution over $[0, kq(z_0)]$
- 3 if $u_0 > \tilde{p}(z_0)$ then the sample is rejected, otherwise it is retained.

The retained z are distributed according p(z).



$$p(\text{accept}) = \int \frac{\tilde{p}(z)}{kq(z)} q(z) dz$$
$$= \frac{1}{k} \int \tilde{p}(z) dz$$

So k should be as small as possible subject to the limitation that $kq(z) > \tilde{p}(z)$ for all $z \to$ the proposal distribution q should be close to p



$$p(\text{accept}) = \int \frac{\tilde{p}(z)}{kq(z)} q(z) dz$$

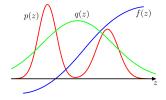
= $\frac{1}{k} \int \tilde{p}(z) dz$

So k should be as small as possible subject to the limitation that $kq(z) > \tilde{p}(z)$ for all $z \to$ the proposal distribution q should be close to p

Rejection sampling does not work well with high dimensional spaces, as the number of rejected samples increases exponentially with dimensions.



Goal: Approximate expectations directly



Reuse the idea of the proposal distribution $q(\mathbf{z})$ and express the expectation in the form of a finite sum over samples $\{\mathbf{z}^{(l)}\}$ drawn from $q(\mathbf{z})$

- No rejection of samples
- Correct bias introduced by sampling from the wrong distribution by weighting samples



$$E[f] = \int f(z)p(z)dz$$
$$= \int f(z)\frac{p(z)}{q(z)}q(z)dz$$



$$E[f] = \int f(\mathbf{z}) \rho(\mathbf{z}) d\mathbf{z}$$

$$= \int f(\mathbf{z}) \frac{\rho(\mathbf{z})}{q(\mathbf{z})} q(\mathbf{z}) d\mathbf{z}$$

$$\simeq \frac{1}{L} \sum_{l=1}^{L} \frac{\rho(\mathbf{z}^{(l)})}{q(\mathbf{z}^{(l)})} f(\mathbf{z}^{(l)})$$

Note we can see this as sampling $f(\mathbf{z}) \frac{p(\mathbf{z})}{q(\mathbf{z})}$ from the $q(\mathbf{z})$ distribution.



$$E[f] = \int f(\mathbf{z}) \rho(\mathbf{z}) d\mathbf{z}$$

$$= \int f(\mathbf{z}) \frac{\rho(\mathbf{z})}{q(\mathbf{z})} q(\mathbf{z}) d\mathbf{z}$$

$$\simeq \frac{1}{L} \sum_{l=1}^{L} \frac{\rho(\mathbf{z}^{(l)})}{q(\mathbf{z}^{(l)})} f(\mathbf{z}^{(l)})$$

 $w_l = p(\mathbf{z}^{(l)})/q(\mathbf{z}^{(l)})$ are known as **importance weights**



- Success depends on how well the proposal distribution $q(\mathbf{x})$ matches the desired distribution $p(\mathbf{x})$.
- Must have $p(z) > 0 \Rightarrow q(z) > 0$.



Sampling and the EM algorithm

If the expectation step of the EM algorithm cannot be performed analytically, we may instead approximate it with sampling.

$$Q(\theta, heta^{ ext{old}}) = \int p(\mathbf{Z}|\mathbf{X}, heta^{ ext{old}}) \ln p(\mathbf{Z}, \mathbf{X}| heta) d\mathbf{Z}$$

Estimate integral with sum over samples drawn from the current estimate of the poster distribution $p(\mathbf{Z}^{(l)}, \mathbf{X}|\theta)$

$$Q(heta, heta^{ ext{old}}) \simeq rac{1}{L} \sum_{l=1}^L \ln p(\mathbf{Z}^{(l)}, \mathbf{X} | heta)$$



Python libraries

- pgmpy Python library for Probabilistic Graphical Models
 - Directed, undirected, factor graphs, etc...
 - Sampling methods, Belief Propagation, Max-Product, etc...
 - And much more.
- Pyro Python/PyTorch based library for Deep Universal Probabilistic Programming
 - Unifying deep learning and graphical models
 - Variational inference, sampling



Questions?

