

Example of Regime Switching State Space Model

Renhe W.

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1 Kim(1994)-Regime Switching State Space Model

1.1 Regime Switching State Space Model

As an example of a regime switching state space model, Prof. Kim used the following generalized Hamilton model for the log of real GNP (Lam; 1990) in his paper and book.

$$\begin{aligned}\ln(GNP_t) &= n_t + x_t \\ n_t &= n_{t-1} + \mu_0 + \mu_1 s_t \\ x_t &= \phi_1 x_{t-1} + \phi_2 x_{t-2} + u_t \\ u_t &\sim N(0, \sigma^2) \\ s_t &= 0, 1 \quad P_{tj} = \begin{bmatrix} p_{00} & p_{01} \\ p_{10} & p_{11} \end{bmatrix}\end{aligned}$$

where $\ln(GNP_t)$ is a real GNP level. n_t is a deterministic series with a regimeswitching growth rate and x_t is stationary AR(2) cycle process. Since $\ln(GNP_t)$ is the log level variable, the difference of it, $y_t = \ln(GNP_t) - \ln(GNP_{t-1})$, can be represented as a state space model in the following way.

$$\begin{aligned}y_t &= \mu_0 + \mu_1 s_t + x_t - x_{t-1} \\ x_t &= \phi_1 x_{t-1} + \phi_2 x_{t-2} + u_t\end{aligned}$$

It is a typical approach that a state space model is represented as a vector-matrix form for using Kalman filter as follows.

$$\begin{aligned}
y_t &= \mu_{s_t} + \begin{bmatrix} 1 & -1 \end{bmatrix} \begin{bmatrix} x_t \\ x_{t-1} \end{bmatrix} \\
\begin{bmatrix} x_t \\ x_{t-1} \end{bmatrix} &= \begin{bmatrix} \phi_1 & \phi_2 \\ 1 & 0 \end{bmatrix} \begin{bmatrix} x_{t-1} \\ x_{t-2} \end{bmatrix} + \begin{bmatrix} u_t \\ 0 \end{bmatrix} \\
u_t &\sim N(0, \sigma^2) \\
\mu_{s_t} &= \mu_0 + \mu_1 s_t, \quad \mu_1 > 0 \\
s_t &= 0, 1 \quad P_{tj} = \begin{bmatrix} p_{00} & p_{01} \\ p_{10} & p_{11} \end{bmatrix} \\
&\Downarrow \\
y_t &= \mu_{s_t} + F \mathbf{x}_t \\
\mathbf{x}_t &= A \mathbf{x}_{t-1} + v_t
\end{aligned}$$

For the sake of notational simplicity, we use x_t instead of \mathbf{x}_t .

1.2 Kalman Filtering

Kalman filter with regime switching is used to get state estimates from a state space model taking regime transition into account and has the following recursion.

$$\begin{aligned}
x_{t|t-1}^{ij} &= A x_{t-1|t-1}^i \\
P_{t|t-1}^{ij} &= A P_{t-1|t-1}^i A^T + Q \\
\eta_{t|t-1}^{ij} &= y_t - \mu_j - F x_{t|t-1}^{ij} \\
H_{t|t-1}^{ij} &= F P_{t|t-1}^{ij} F^T + R \\
K^{ij} &= P_{t|t-1}^{ij} F^T \left[H_{t|t-1}^{ij} \right]^{-1} \\
x_{t|t}^{ij} &= x_{t|t-1}^{ij} + K^{ij} \eta_{t|t-1}^{ij} \\
P_{t|t}^{ij} &= (I - K^{ij} F) P_{t|t-1}^{ij}
\end{aligned}$$

In regime-dependent Kalman filter, all the notations are augmented with superscript $\{ij\}$ except $x_{t-1|t-1}^i$ and $P_{t-1|t-1}^i$ since these two estimates are in i-state (two-state) but other estimates must reflect state transitions from i to j (four-state). For example,

$x_{t|t-1}$ and $x_{t|t-1}^{ij}$ are different in terms of conditioning information.

$$\begin{aligned} x_{t|t-1} &= E[X_t | \psi_{t-1}] \\ x_{t|t-1}^{ij} &= E[X_t | \psi_{t-1}, S_t = j, S_{t-1} = i] \end{aligned}$$

In contrast to the single regime, however, in the multiple regimes, $x_{t|t}^{ij}$ and $P_{t|t}^{ij}$ cannot be used the next state prediction due simply to the mismatch both 1) between $x_{t|t}^{ij}$ and $x_{t-1|t-1}^i$ and 2) between $P_{t|t}^{ij}$ and $P_{t-1|t-1}^i$. To resolve this mismatch problem, Kim (1994) developed a dimension collapsing algorithm.

1.3 Kim(1994)' s Collapsing procedure

Kim (1994) introduces a collapsing procedure (approximation) to reduce the $(M \times M)$ posteriors $(x_{t|t}^{ij}$ and $P_{t|t}^{ij})$ into M to complete the above Kalman filter recursion.

$$\begin{aligned} x_{t|t}^j &= \frac{\sum_{i=1}^M P[S_{t-1} = i, S_t = j | \psi_t] x_{t|t}^{ij}}{P[S_t = j | \psi_t]} \\ P_{t|t}^j &= \frac{\sum_{i=1}^M P[S_{t-1} = i, S_t = j | \psi_t] \left[P_{t|t}^{ij} + (x_{t|t}^j - x_{t|t}^{ij})(x_{t|t}^j - x_{t|t}^{ij})^T \right]}{P[S_t = j | \psi_t]} \end{aligned}$$

To calculate the above approximation, when we calculate $P[S_{t-1} = i, S_t = j | \psi_t]$, $P[S_t = j | \psi_t]$ is easily obtained by summing its M branches from each i states.

$$P[S_t = j | \psi_t] = \sum_{i=1}^M P[S_{t-1} = i, S_t = j | \psi_t]$$

Knowing $P[S_{t-1} = i, S_t = j | \psi_t]$ means that we observe time t data since the last time of information set is t and the state is migrated from i to j . For this data information into account, we can think of the marginal probability of state transition by integrating out data.

$$\begin{aligned} P[S_{t-1} = i, S_t = j | \psi_t] &= \frac{f(y_t, S_{t-1} = i, S_t = j | \psi_{t-1})}{f(y_t | \psi_{t-1})} \\ f(y_t | \psi_{t-1}) &= \sum_{j=1}^M \sum_{i=1}^M f(y_t, S_{t-1} = i, S_t = j | \psi_{t-1}) \end{aligned}$$

As can be seen from the above equations, when we know $f(y_t, S_{t-1} = i, S_t = j \mid \psi_{t-1})$ which is the joint density of data and two states, $P[S_{t-1} = i, S_t = j \mid \psi_t]$ and $f(y_t \mid \psi_{t-1})$ are easily obtained. We get the following the joint density.

$$\begin{aligned} & f(y_t, S_{t-1} = i, S_t = j \mid \psi_{t-1}) \\ &= f(y_t \mid S_{t-1} = i, S_t = j, \psi_{t-1}) \times P(S_{t-1} = i, S_t = j \mid \psi_{t-1}) \end{aligned}$$

Now we need to know two parts. The first part is the forecast error given data. (MVN : probability density function of multivariate normal distribution with zero mean and forecast error variance)

$$\begin{aligned} & f(y_t \mid S_{t-1} = i, S_t = j, \psi_{t-1}) \\ &= MVN(\text{forecast error, its variance}) \end{aligned}$$

The second part is calculated by the multiplication of the transition probability and the summation of its branches.

$$\begin{aligned} & P(S_{t-1} = i, S_t = j \mid \psi_{t-1}) \\ &= P[S_t = j \mid S_{t-1} = i] \times \sum_{k=1}^M f(S_{t-2} = k, S_{t-1} = i \mid \psi_{t-1}) \\ &= P_{ij} \times f(S_{t-1} = i \mid \psi_{t-1}) \end{aligned}$$

In the above equation, as we know, P_{ij} is already known as transition probability matrix and $f(S_{t-2} = k, S_{t-1} = i \mid \psi_{t-1})$ is exactly what we want to find but evaluated at the previous $t-1$ time. Therefore for the iteration, $f(S_{-1} = k, S_0 = i \mid \psi_0)$ calls for initialization with the steady state probabilities. Therefore, we can calculate $P[S_{t-1} = i, S_t = j \mid \psi_t]$ and $P[S_t = j \mid \psi_t]$ through the above equations.

1.4 Kim (1994) Filter for Regime Switching State Space model

Kim filtering procedure is summarized in a sequence of equations. Kalman Filtering

$$\begin{aligned}
x_{t|t-1}^{ij} &= Ax_{t-1|t-1}^i \\
P_{t|t-1}^{ij} &= AP_{t-1|t-1}^i A^T + Q \\
\eta_{t|t-1}^{ij} &= y_t - \mu_j - Fx_{t|t-1}^{ij} \\
H_{t|t-1}^{ij} &= FP_{t|t-1}^{ij} F^T + R \\
K^{ij} &= P_{t|t-1}^{ij} F^T \left[H_{t|t-1}^{ij} \right]^{-1} \\
x_{t|t}^{ij} &= x_{t|t-1}^{ij} + K^{ij} \eta_{t|t-1}^{ij} \\
P_{t|t}^{ij} &= (I - K^{ij} F) P_{t|t-1}^{ij} \\
&\Downarrow
\end{aligned}$$

Hamilton Filtering

$$\begin{aligned}
f(y_t, S_{t-1} = i, S_t = j \mid \psi_{t-1}) &= N\left(\eta_{t|t-1}^{ij}, H_{t|t-1}^{ij}\right) \times P_{ij} \times P(S_{t-1} = i \mid \psi_{t-1}) \\
f(y_t \mid \psi_{t-1}) &= \sum_{j=1}^M \sum_{i=1}^M f(y_t, S_{t-1} = i, S_t = j \mid \psi_{t-1}) \\
P[S_{t-1} = i, S_t = j \mid \psi_t] &= \frac{f(y_t, S_{t-1} = i, S_t = j \mid \psi_{t-1})}{f(y_t \mid \psi_{t-1})} \\
P[S_t = j \mid \psi_t] &= \sum_{i=1}^M P[S_{t-1} = i, S_t = j \mid \psi_t] \\
&\Downarrow
\end{aligned}$$

Kim's Collapsing

$$\begin{aligned}
x_{t|t}^j &= \frac{\sum_{i=1}^M P[S_{t-1} = i, S_t = j \mid \psi_t] x_{t|t-1}^{ij}}{P[S_t = j \mid \psi_t]} \\
P_{t|t}^j &= \frac{\sum_{i=1}^M P[S_{t-1} = i, S_t = j \mid \psi_t] \left[P_{t|t-1}^{ij} + \left(x_{t|t}^j - x_{t|t-1}^{ij} \right) \left(x_{t|t}^j - x_{t|t-1}^{ij} \right)^T \right]}{P[S_t = j \mid \psi_t]}
\end{aligned}$$

In particular, three red colored terms $\left(x_{t-1|t-1}^i, P_{t-1|t-1}^i, \text{ and } P(S_{t-1} = i \mid \psi_{t-1}) \right)$ are initialized for iteration to be started. Once the iterations get started, these red colored terms are replaced with blue colored terms $\left(x_{t|t}^j, P_{t|t}^j, \text{ and } P(S_t = j \mid \psi_t) \right)$ for each iterations.

A R code