

Lab Week 5

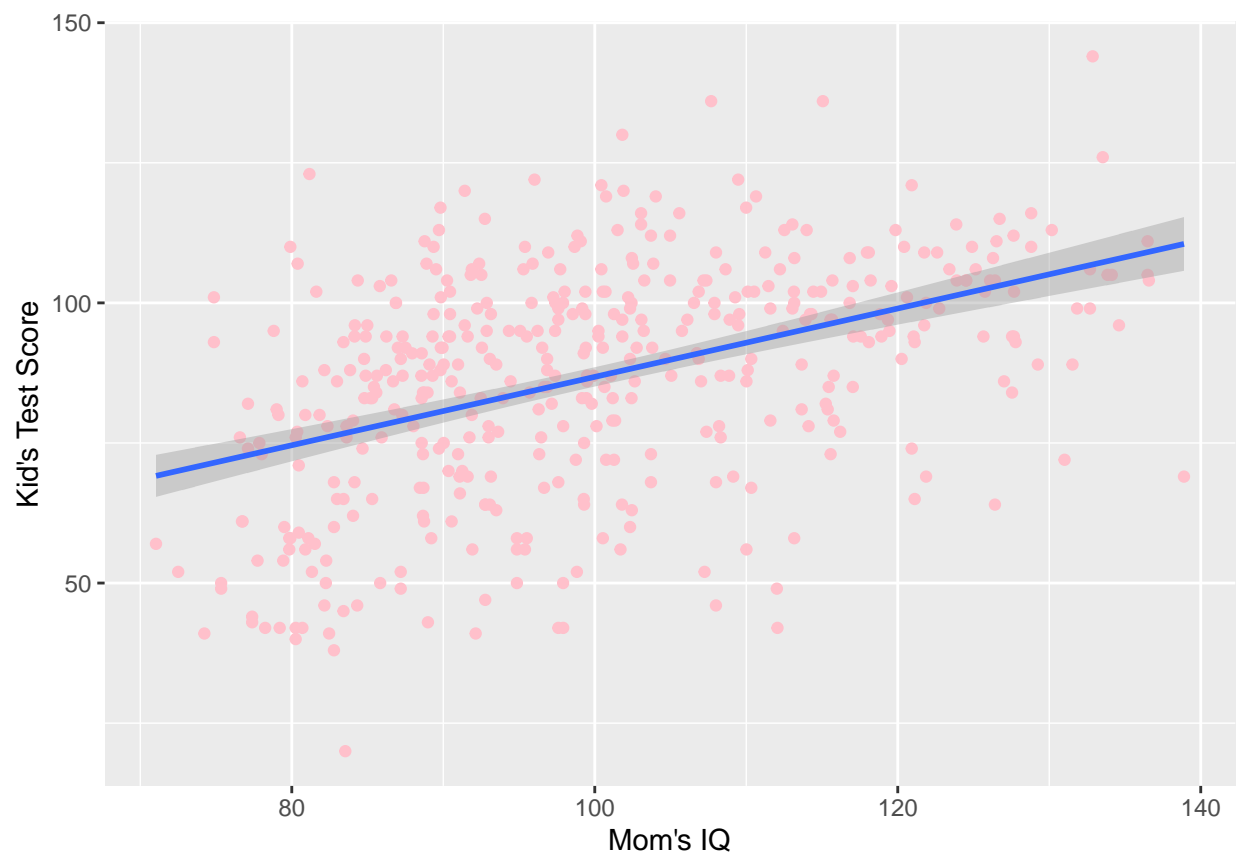
Rosa Fallahpour

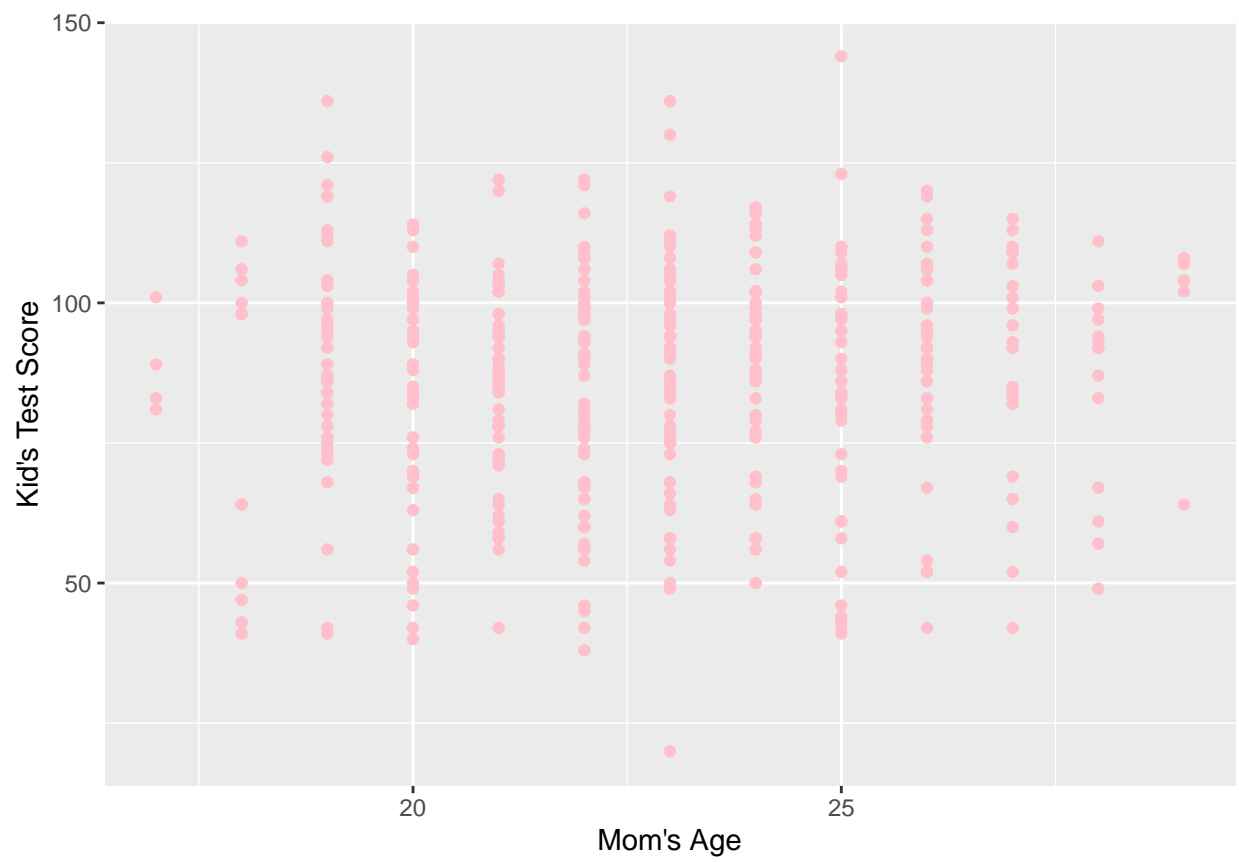
Question 1

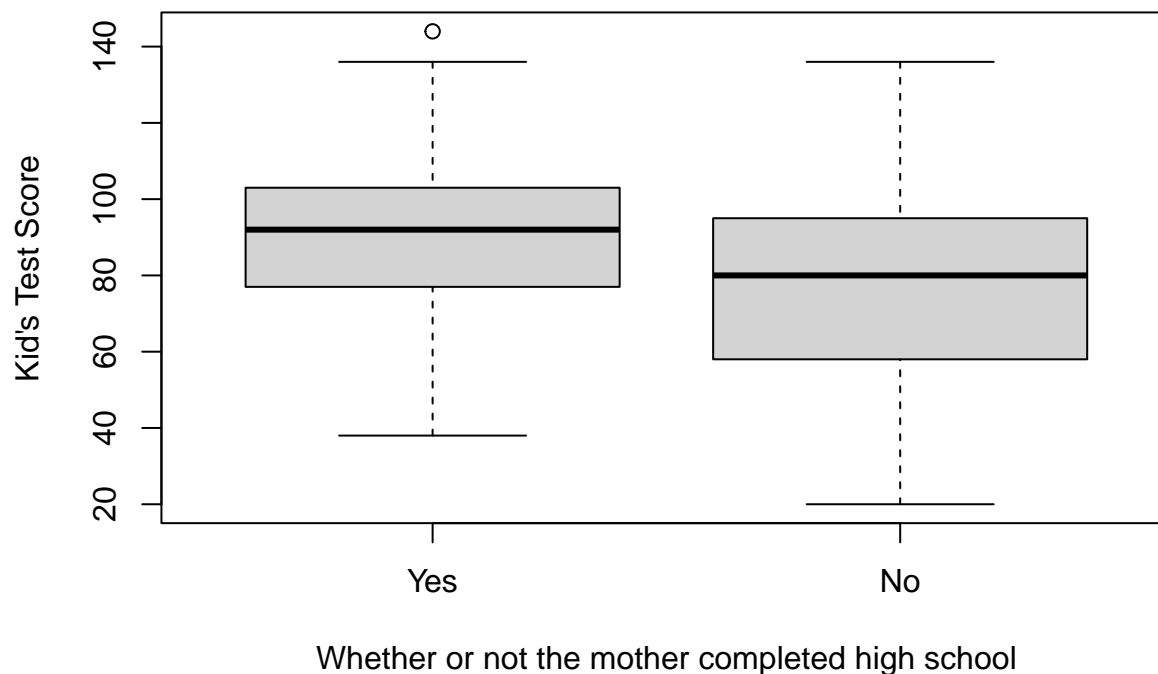
Three following graphs represent the kid's test score with respect to mom's IQ, mom's age and whether or not the mother completed high school, respectively. The first plot shows that the kid's score has the positive relationship with mom's IQ, as we can see that the increase in mom's IQ results in increasing kid's score. The second graph displays the relationship between kid's score and mom's age. It is interesting to see that not a significant change is observed in kid's score in different ages of mothers. Kid's scores are almost in the same range for mothers with different ages. The last plot shows the relationship between kid's score and mom's high school completion. It suggests that the average kid's score for the group whose moms completed their high school is higher than those kid's whose mothers did not complete the high school.

```
kidiq <- read_rds(here("kidiq.RDS"))
kidiq
```

```
## # A tibble: 434 x 4
##   kid_score mom_hs mom_iq mom_age
##   <int>    <dbl> <dbl>   <int>
## 1      65      1  121.     27
## 2      98      1   89.4     25
## 3      85      1  115.     27
## 4      83      1   99.4     25
## 5     115      1   92.7     27
## 6      98      0  108.     18
## 7      69      1  139.     20
## 8     106      1  125.     23
## 9     102      1   81.6     24
## 10     95      1   95.1     19
## # ... with 424 more rows
```







Estimating mean, no covariates

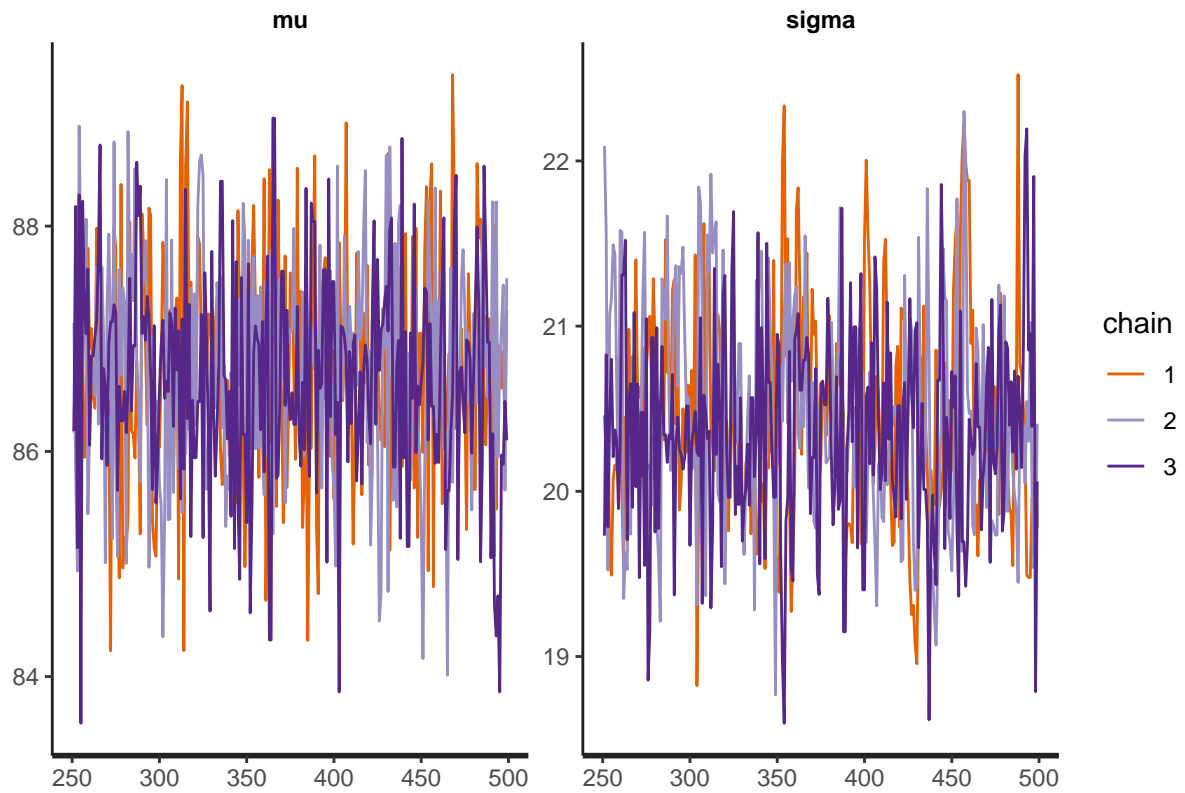
here is the summary:

```
fit
```

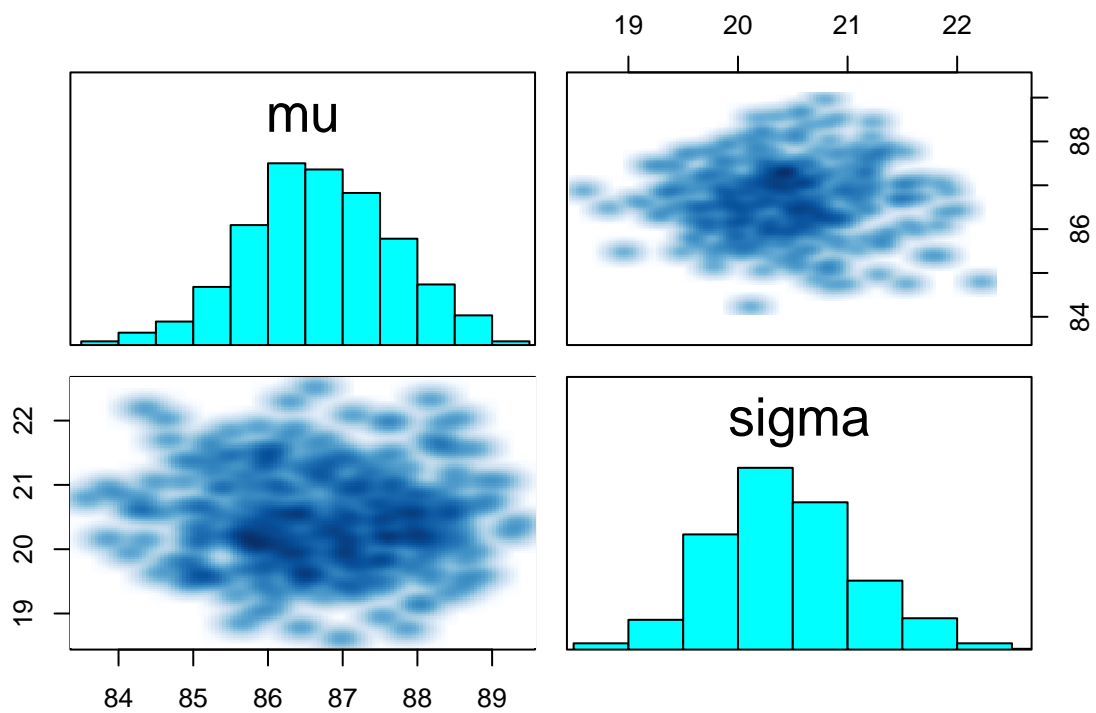
```
## Inference for Stan model: anon_model.
## 3 chains, each with iter=500; warmup=250; thin=1;
## post-warmup draws per chain=250, total post-warmup draws=750.
##
##           mean se_mean   sd    2.5%    25%    50%    75%    97.5%  n_eff
## mu         86.70    0.03 0.98    84.75    86.03    86.68    87.34    88.56   803
## sigma      20.43    0.03 0.65    19.29    19.98    20.40    20.82    21.83   361
## lp__     -1525.70    0.06 0.94 -1528.19 -1526.08 -1525.39 -1525.01 -1524.78   259
##           Rhat
## mu         1.00
## sigma      1.01
## lp__       1.00
##
## Samples were drawn using NUTS(diag_e) at Sun Feb 12 19:50:12 2023.
## For each parameter, n_eff is a crude measure of effective sample size,
## and Rhat is the potential scale reduction factor on split chains (at
## convergence, Rhat=1).
```

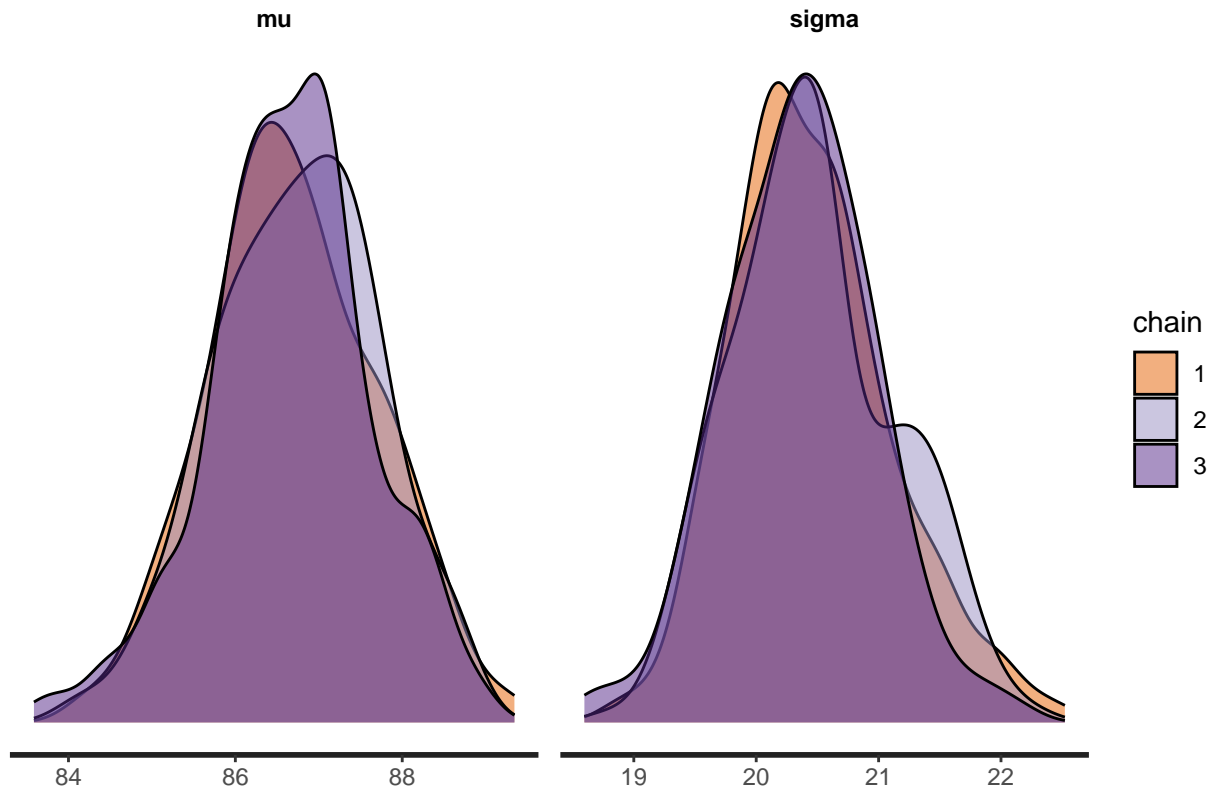
Traceplot:

```
traceplot(fit)
```



```
pairs(fit, pars = c("mu", "sigma"))
```





Understanding output

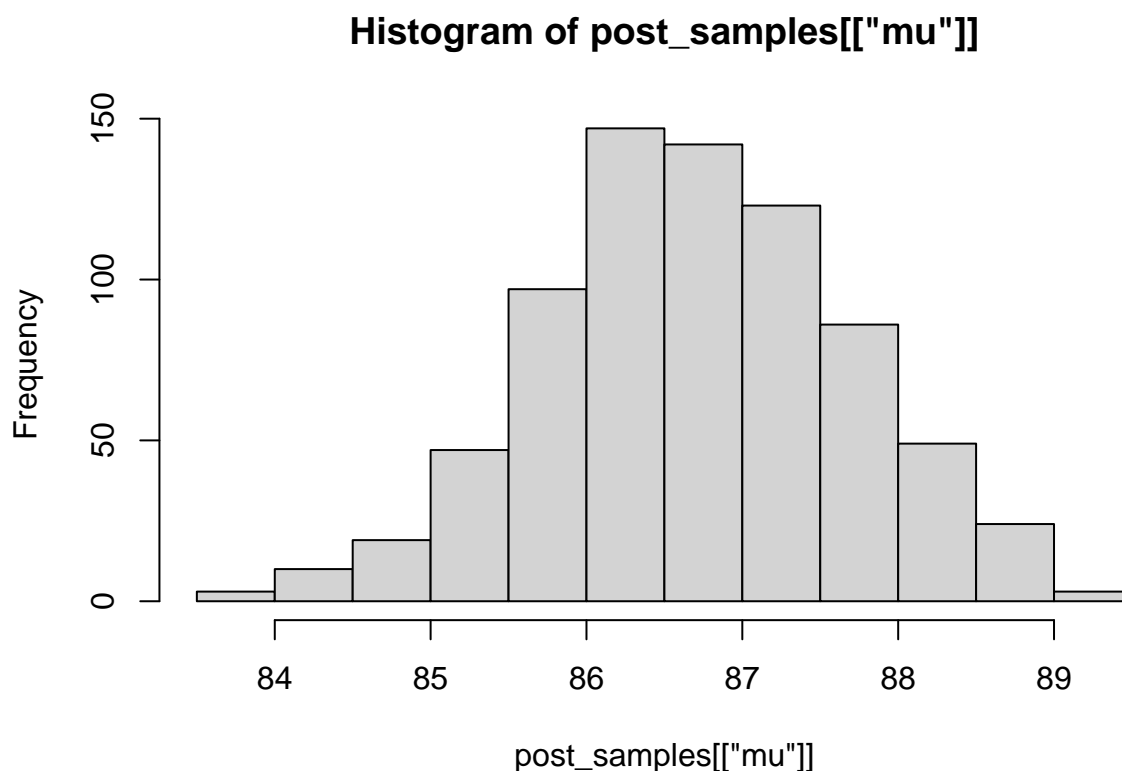
Samples from the posteriors:

```
post_samples <- rstan::extract(fit)
head(post_samples[["mu"]])
```

```
## [1] 87.15510 87.98089 85.67446 86.59396 86.99718 86.63150
```

Histogram of mu:

```
hist(post_samples[["mu"]])
```



```
median(post_samples[["mu"]])
```

```
## [1] 86.68244
```

```
quantile(post_samples[["mu"]], 0.025)
```

```
##      2.5%  
## 84.74647
```

```
quantile(post_samples[["mu"]], 0.975)
```

```
##      97.5%  
## 88.55876
```

Plot estimates

Get the posterior samples for mu and sigma in long format:

```
library(tidybayes)  
dsamples <- fit |>  
  gather_draws(mu, sigma) # gather = long format  
dsamples
```



```
## # A tibble: 1,500 x 5
## # Groups:   .variable [2]
##   .chain .iteration .draw .variable .value
##   <int>      <int> <int> <chr>      <dbl>
## 1      1          1     1 1 mu        86.2
## 2      1          2     2 2 mu        87.2
## 3      1          3     3 3 mu        85.9
## 4      1          4     4 4 mu        86.4
## 5      1          5     5 5 mu        86.0
## 6      1          6     6 6 mu        86.9
## 7      1          7     7 7 mu        85.9
## 8      1          8     8 8 mu        86.2
## 9      1          9     9 9 mu        87.8
## 10     1         10    10 10 mu        86.9
## # ... with 1,490 more rows
```

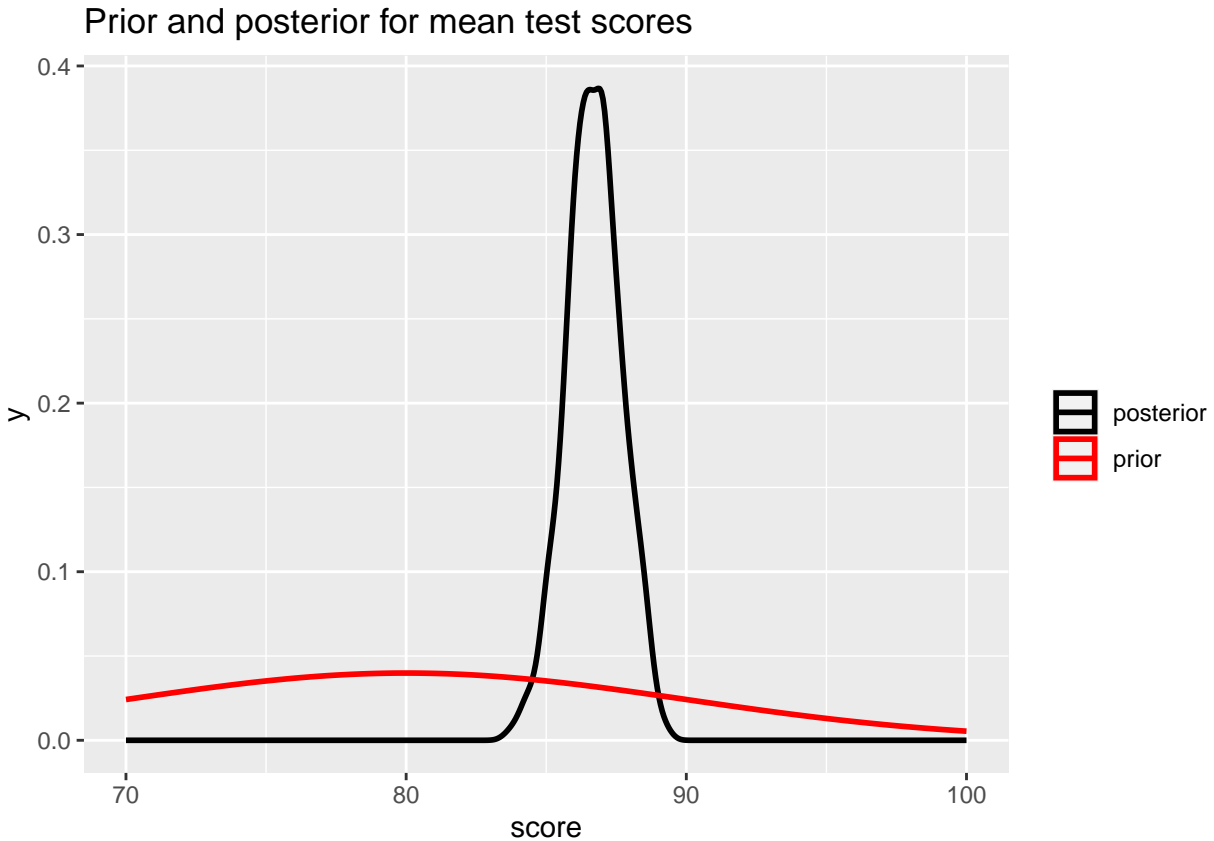
```
# wide format
fit |> spread_draws(mu, sigma)
```

```
## # A tibble: 750 x 5
##   .chain .iteration .draw    mu sigma
##   <int>      <int> <int> <dbl> <dbl>
## 1      1          1     1 86.2 20.5
## 2      1          2     2 87.2 20.2
## 3      1          3     3 85.9 20.5
## 4      1          4     4 86.4 19.6
## 5      1          5     5 86.0 19.5
## 6      1          6     6 86.9 20.1
## 7      1          7     7 85.9 20.2
## 8      1          8     8 86.2 20.1
## 9      1          9     9 87.8 20.7
## 10     1         10    10 86.9 20.9
## # ... with 740 more rows
```

```
# quickly calculate the quantiles using
dsamples |>
  median_qi(.width = 0.8)
```

```
## # A tibble: 2 x 7
##   .variable .value .lower .upper .width .point .interval
##   <chr>      <dbl> <dbl> <dbl> <dbl> <chr> <chr>
## 1 mu        86.7  85.5  88.0   0.8 median qi
## 2 sigma     20.4  19.6  21.3   0.8 median qi
```

Let's plot the density of the posterior samples for mu and add in the prior distribution:



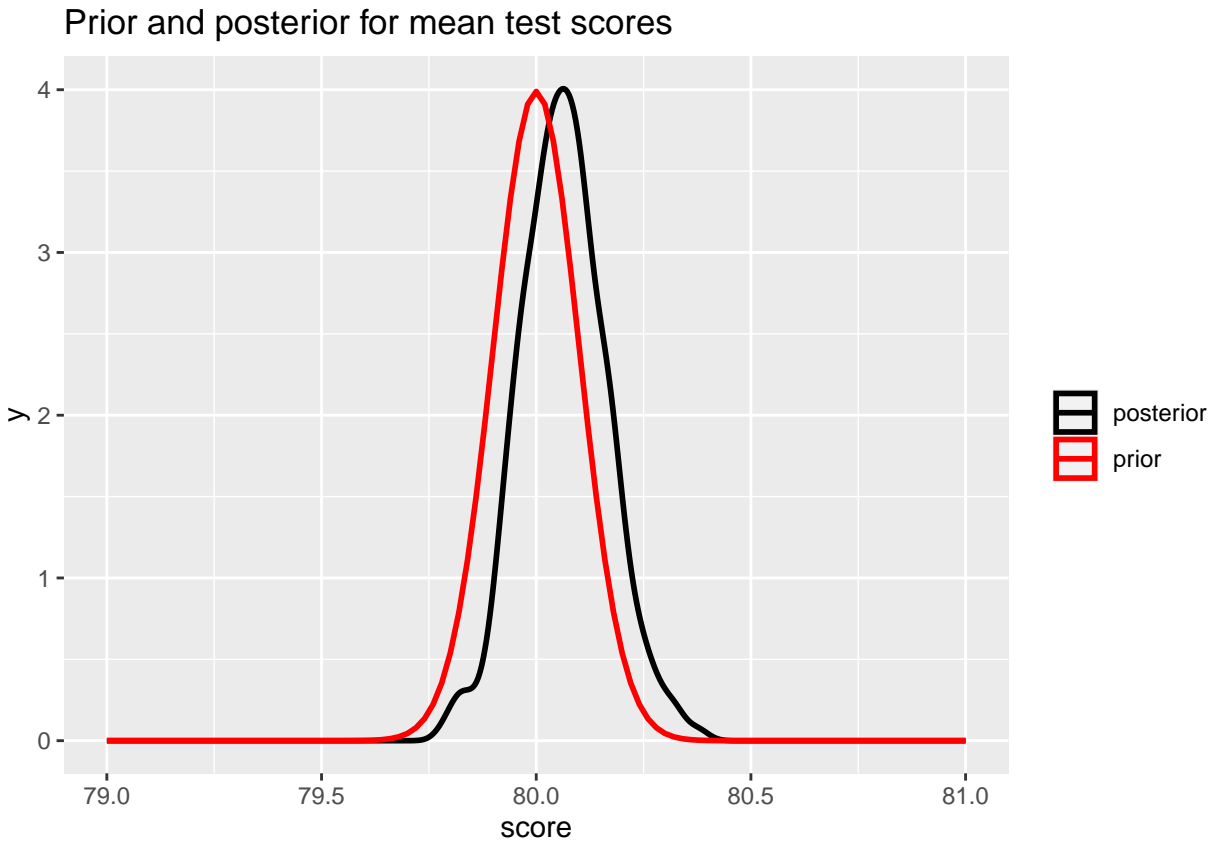
Question 2

In the model with more informative prior, μ estimate and its standard error has decreased. However, it shows a slight increase in the σ value in the more informative model.

```
fit_inform
```

```
## Inference for Stan model: anon_model.
## 3 chains, each with iter=500; warmup=250; thin=1;
## post-warmup draws per chain=250, total post-warmup draws=750.
##
##           mean se_mean  sd    2.5%    25%    50%    75%   97.5% n_eff
## mu       80.06   0.00 0.10   79.87   80.00   80.06   80.13   80.27   542
## sigma    21.40   0.03 0.73   19.94   20.88   21.38   21.89   22.78   648
## lp__    -1548.38  0.05 0.97 -1550.99 -1548.77 -1548.09 -1547.69 -1547.39  446
##      Rhat
## mu      1
## sigma   1
## lp__    1
##
## Samples were drawn using NUTS(diag_e) at Sun Feb 12 19:50:14 2023.
## For each parameter, n_eff is a crude measure of effective sample size,
## and Rhat is the potential scale reduction factor on split chains (at
## convergence, Rhat=1).
```

Plotting the prior and posterior densities:



Adding Covariates

fit2

```
## Inference for Stan model: anon_model.
## 4 chains, each with iter=1000; warmup=500; thin=1;
## post-warmup draws per chain=500, total post-warmup draws=2000.
##
##           mean se_mean  sd    2.5%    25%    50%    75%    97.5%
## alpha      78.07   0.07 1.98   74.07   76.70   78.12   79.41   81.97
## beta[1]    11.10   0.08 2.21    6.77    9.63   11.10   12.58   15.41
## sigma      19.81   0.02 0.65   18.61   19.35   19.79   20.23   21.13
## lp__     -1514.33   0.04 1.20 -1517.40 -1514.83 -1514.02 -1513.47 -1512.97
##           n_eff Rhat
## alpha      870 1.00
## beta[1]    869 1.00
## sigma      830 1.01
## lp__       800 1.00
##
## Samples were drawn using NUTS(diag_e) at Sun Feb 12 19:50:50 2023.
## For each parameter, n_eff is a crude measure of effective sample size,
## and Rhat is the potential scale reduction factor on split chains (at
## convergence, Rhat=1).
```

Question 3

- (a) As we can see in the following summaries which are related to lm model and model fit2, the estimates are very close to each other.

```
model_lm <- lm(kid_score~mom_hs, data=kidiq)
summary(model_lm)$`coefficient`
```

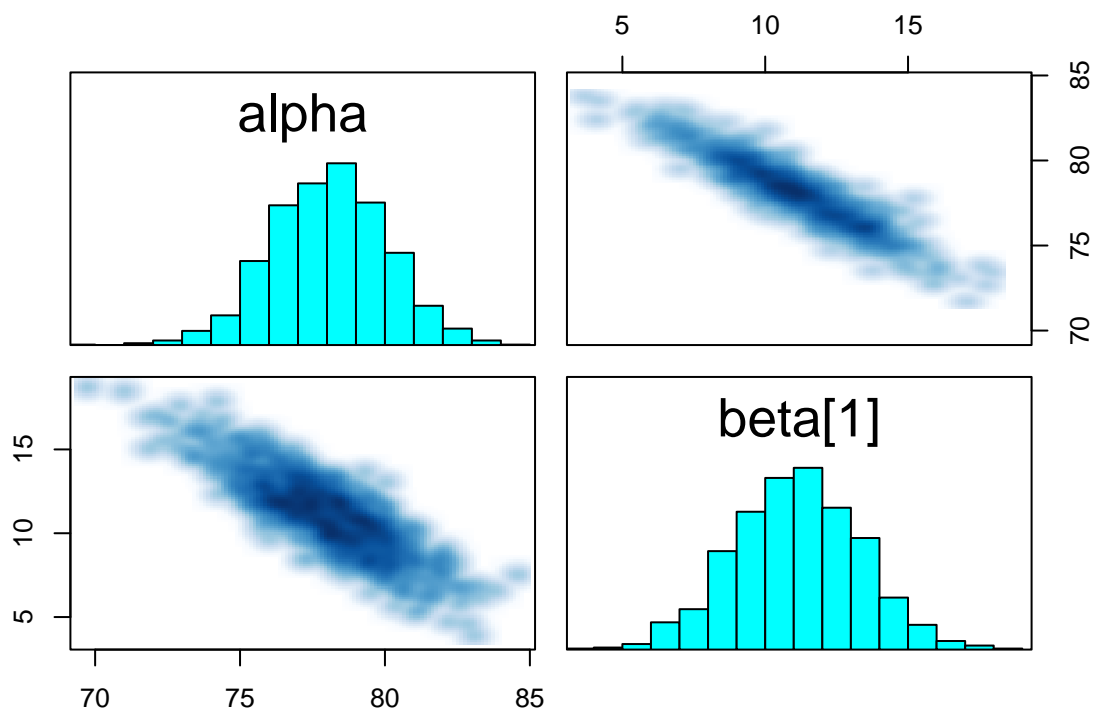
```
##           Estimate Std. Error   t value    Pr(>|t|)
## (Intercept) 77.54839    2.058612  37.670231 1.392224e-138
## mom_hs      11.77126    2.322427   5.068516 5.956524e-07
```

```
summary(fit2)$summary[c("alpha", "beta[1]"), ]
```

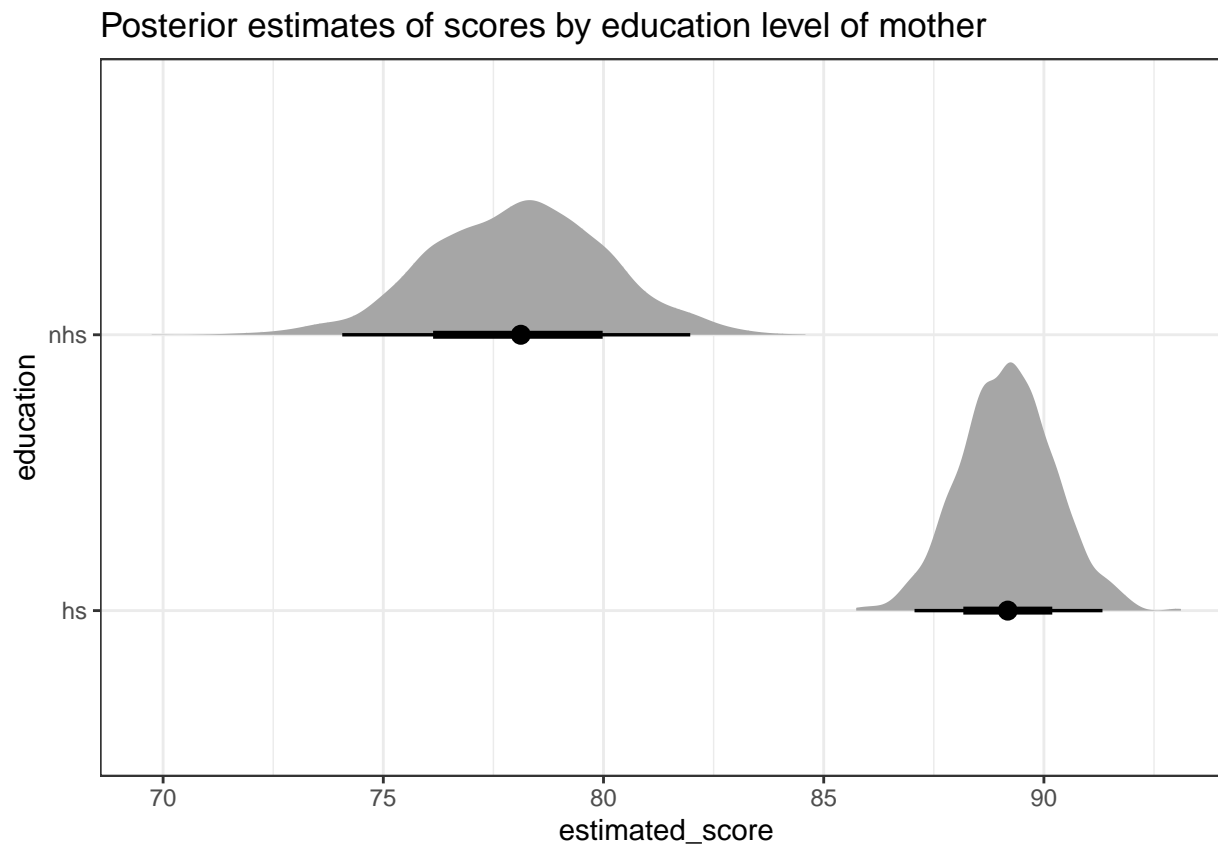
```
##           mean    se_mean      sd    2.5%    25%    50%    75%
## alpha      78.06743 0.06719963 1.981812 74.067805 76.703994 78.12197 79.40693
## beta[1]    11.10351 0.07504760 2.212830  6.766445  9.631034 11.09916 12.58276
##           97.5%    n_eff      Rhat
## alpha      81.96982 869.7431 1.002156
## beta[1]    15.41314 869.4061 1.001961
```

- (b) As the figure shows, the slope variation includes the opposite variation of the intercept. Thus, the intercept interpretation and sampling would be harder.

```
pairs(fit2, pars = c("alpha", "beta[1]"))
```



Plotting the results



Question 4

fit3

```
## Inference for Stan model: anon_model.
## 4 chains, each with iter=1000; warmup=500; thin=1;
## post-warmup draws per chain=500, total post-warmup draws=2000.
##
##           mean se_mean  sd    2.5%    25%    50%    75%    97.5%
## alpha      82.37    0.06 1.89    78.62    81.11    82.39    83.63    86.06
## beta[1]     5.62    0.07 2.12     1.41     4.17     5.59     6.99     9.81
## beta[2]     0.57    0.00 0.06     0.45     0.53     0.57     0.61     0.68
## sigma      18.11    0.02 0.61    16.97    17.71    18.09    18.51    19.37
## lp__      -1474.40    0.06 1.42   -1477.90 -1475.12 -1474.08 -1473.35 -1472.63
##           n_eff Rhat
## alpha      937    1
## beta[1]   1004    1
## beta[2]   1378    1
## sigma    1446    1
## lp__      583    1
##
```

```
## Samples were drawn using NUTS(diag_e) at Sun Feb 12 19:50:51 2023.
## For each parameter, n_eff is a crude measure of effective sample size,
## and Rhat is the potential scale reduction factor on split chains (at
## convergence, Rhat=1).
```

For a given mother's high school completion, one unit increase in centered mom's IQ score, results in the posterior mean of the kid's score to increase by 0.57.

Question 5

As we can see the estimates of two models are comparable.

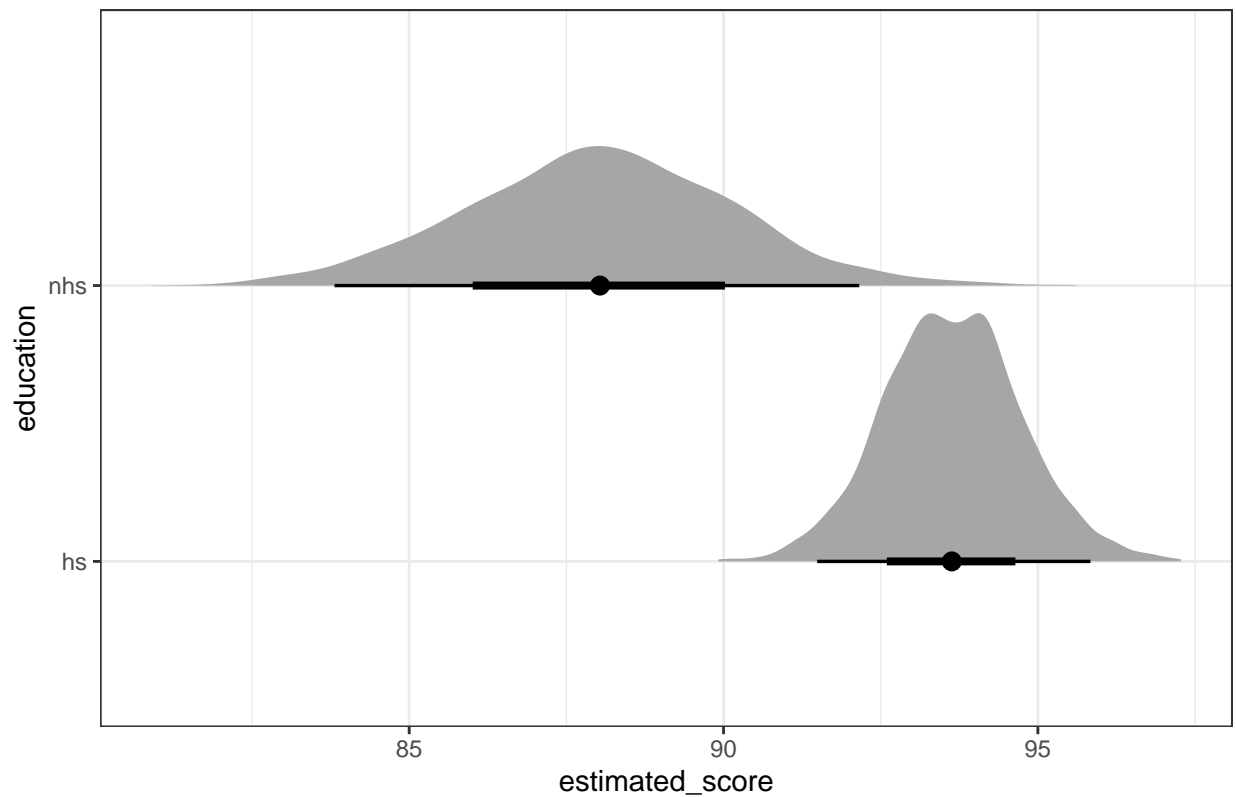
```
momiq_2 <- kidiq$mom_iq-mean(kidiq$mom_iq)
model2_lm <- lm(kid_score~mom_hs+momiq_2, data=kidiq)
summary(model2_lm)$`coefficient`
```

```
##              Estimate Std. Error  t value      Pr(>|t|)
## (Intercept) 82.122143 1.94370047 42.250411 2.435765e-155
## mom_hs      5.950117 2.21181218  2.690155 7.419327e-03
## momiq_2     0.563906 0.06057408  9.309362 6.609618e-19
```

Question 6

```
data <- as.data.frame(fit3 %>% spread_draws(alpha, beta[condition], sigma))
data %>%
  reshape(
    idvar = c(".iteration", ".draw", ".chain"),
    timevar = "condition", v.names = "beta", direction = "wide"
  ) %>% mutate(nhs = alpha + beta.2 * 10, hs = alpha + beta.1 + beta.2 * 10) %>%
  pivot_longer(nhs:hs, names_to = "education", values_to = "estimated_score") %>%
  ggplot(aes(y = education, x = estimated_score)) +
  stat_halfeye() +
  theme_bw() +
  ggtitle("Posterior estimates of scores by education level of mother")
```

Posterior estimates of scores by education level of mother



Question 7

```
postsample <- rstan ::extract(fit3)
alpha <- postsample[["alpha"]]
beta1 <- postsample[["beta"]][,1]
beta2 <- postsample[["beta"]][,2]
x_new_2 <- 95-mean(kidiq$mom_iq)
lin_pred <- alpha + beta1*1+beta2*-5
sigma <- postsample[["sigma"]]
y_new <- rnorm(n= length(sigma),mean = lin_pred, sd=sigma)
hist(y_new, xlab = "Kid's Score", main="Posterior Predictive Distribution", col="pink")
```

Posterior Predictive Distribution

