

# Lab Week 3

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- (1) Assuming  $Y|\theta \sim \text{Bin}(n, \theta)$ , where  $y$  is the number of women who reported to be happy out of the sample of  $n$  women. We know that 118 out of 119 women reported to be happy. So, in order to find out the maximum likelihood estimation of our parameter of interest which is  $\theta$ , we have following procedure:

$$f(y_i; \theta) = \binom{129}{118} (1 - \theta)^{11} \theta^{118}$$

We have the likelihood function as below:

$$L(\theta; y_i) = \prod_{i=1}^{129} \binom{129}{118} (1 - \theta)^{11} \theta^{118},$$

taking log likelihood we will have:

$$l(\theta; y_i) = 129 \left[ \ln \binom{129}{118} + 11 \ln(1 - \theta) + 118 \ln \theta \right],$$

taking derivative with respect to  $\theta$  and set it to zero, we will have MLE of  $\theta$  as  $\hat{\theta} \approx 0.91$ . To calculate the 0.95 confidence interval from a normal distribution we have:

```
n <- 129
se <- sqrt(0.91*0.08/n)
CI <- c(0.91-qnorm(0.975)*se, 0.91+qnorm(0.975)*se)
CI
```

```
## [1] 0.8634394 0.9565606
```

- (2) According to the lecture given  $Y|\theta \sim \text{Bin}(n, \theta)$  and  $\theta \sim \text{Beta}(1, 1)$  ( $p(\theta)$ ), we will have the posterior distribution as  $\theta|y \sim \text{Beta}(y + 1, n - y + 1)$ , thus:

$$\theta|y \sim \text{Beta}(119, 12)$$

We have the posterior mean for  $\hat{\theta}$  and 95% credible interval as following:

```
Posterior_Mean <- 119/(119+12)
Posterior_Mean
```

```
## [1] 0.9083969
```

```
cred_int <- c(qbeta(0.025, 119, 12), qbeta(0.975, 119, 12))
cred_int
```

```
## [1] 0.8536434 0.9513891
```

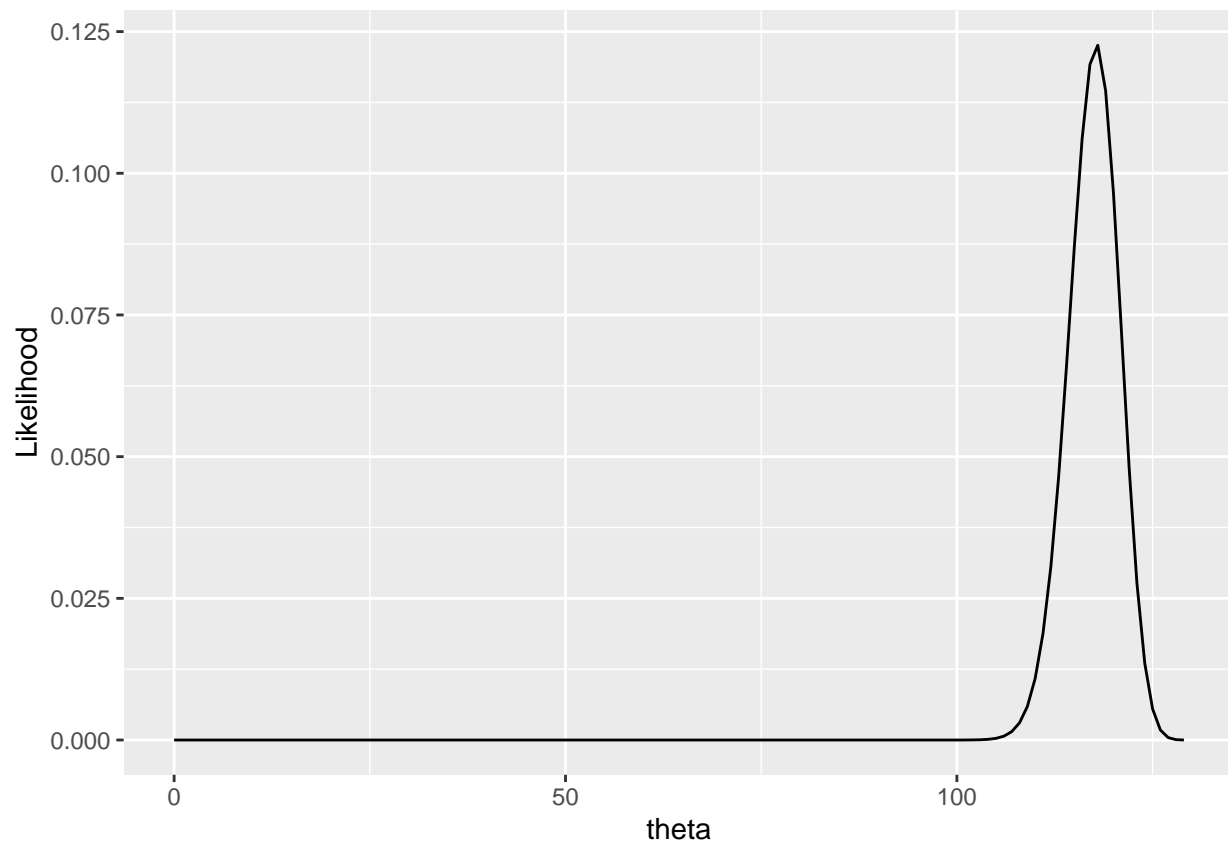
- (3) By considering the prior distribution  $Beta(10,10)$  We are assuming more information about  $\theta$ . Since, the expected proportion of women 65+ being happy ( $\frac{10}{10+10} = 0.5$ ), gains more weight in  $Beta(10,10)$  than  $Beta(1,1)$ , which does not assume any preferences of the range for  $\theta$ . We also can see that the expected prior value is centered in the range of  $[Q1, Q3] = [0.44, 0.58]$ , which is tighter than  $Beta(1,1)$  distribution in question 2.

```
summary(rbeta(100, 10, 10))
```

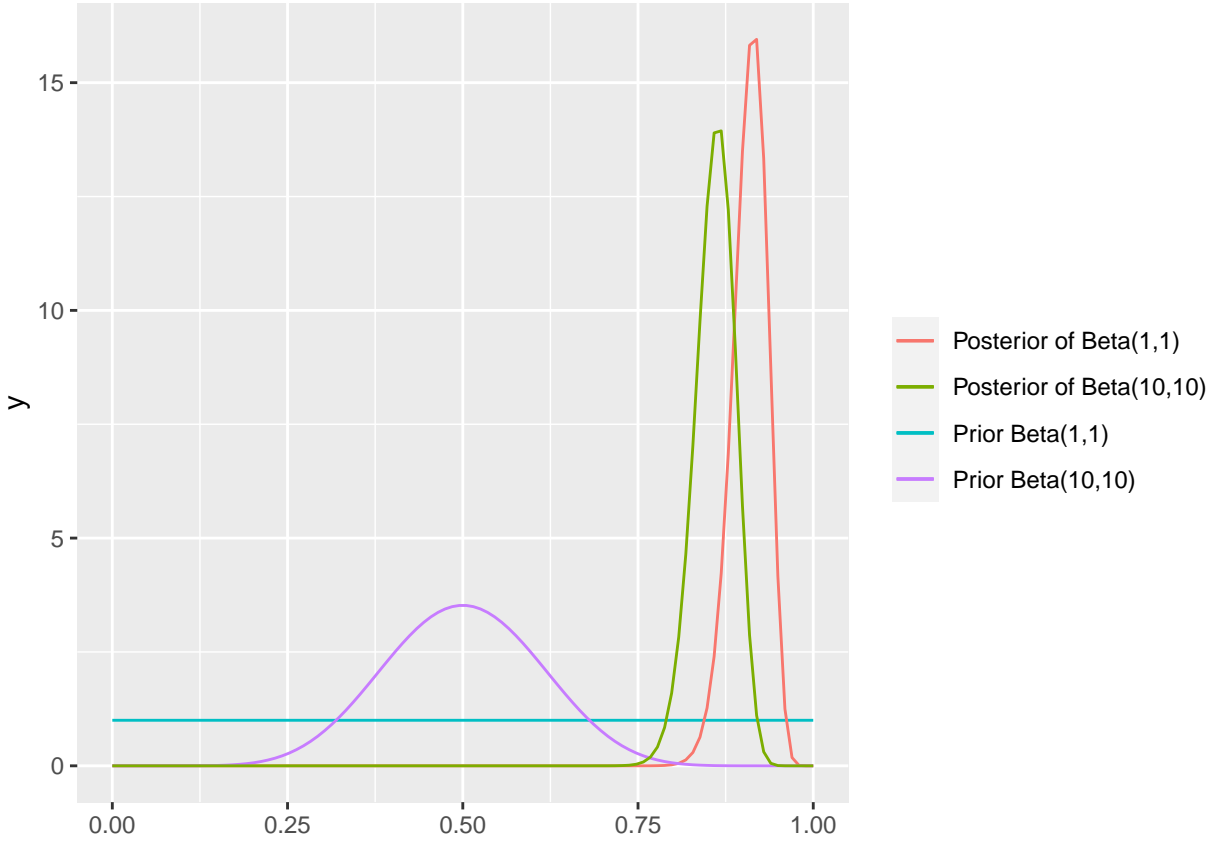
```
##      Min. 1st Qu.  Median    Mean 3rd Qu.    Max.
## 0.2664  0.4000  0.4942  0.4911  0.5645  0.7751
```

- (4) As the following graphs display, the posterior distributions related to priors  $Beta(1,1)$  and  $Beta(10,10)$  are close. However, the posterior of  $Beta(1,1)$  shifts towards 1. In addition, the posterior mean of prior  $Beta(1,1)$  is closer to MLE of  $\theta$ .

```
library(tidyverse)
theta <- seq(0,129)
likefun <- dbinom(theta,129,0.91)
df <- data.frame(x=theta,y=likefun)
df |> ggplot(aes(x=theta,y=likefun))+geom_line()+labs(x="theta",y="Likelihood")
```



```
df |> ggplot(aes()) + stat_function(fun = dbeta, n = 100, args = list(shape1 = 1, shape2 = 1), aes(colour = "Posterior of Beta(1,1)")) +
  stat_function(fun = dbeta, n = 100, args = list(shape1 = 119, shape2 = 12), aes(colour = "Posterior of Beta(119,12)")) +
  stat_function(fun = dbeta, n = 100, args = list(shape1 = 10, shape2 = 10), aes(colour = "Prior Beta(10,10)")) +
  stat_function(fun = dbeta, n = 100, args = list(shape1 = 128, shape2 = 21), aes(colour = "Posterior of Beta(128,21)"))
```



- (5) A noninformative prior distribution for  $\theta$  could be uniform distribution over real line ( $\theta \sim U(-\infty, \infty)$ ), which does not assign any weight on  $\theta$  over this interval, thus does not provide any specific information about it.

Since we have  $\theta$  defined as average improvement in success probability, it takes value between 0 and 1. In addition, we expect an improvement in student's performance after practicing for one month, therefore, I would consider  $Beta(2, 5)$  as the prior distribution for this parameter of interest.