

# Module-4

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Inequalities, converges & limit theorem

- ① Markov  $\rightarrow$  upper
- ② Chebyshev  $\rightarrow$  lower
- ③ Law of Large numbers
- ④ Modes of convergence
- ⑤ Central limit theorem
- ⑥ Strong & weak laws of large no.

$$f_Z(z) = \frac{1}{\sqrt{2\pi}\sigma_Z} e^{-(z-\mu_Z)^2/2\sigma_Z^2}$$

$\mu_Z = 0, \sigma_Z^2 = 1$

$$f_Z(z) = \frac{1}{\sqrt{2\pi}} e^{-z^2/2}$$

$$F_Z(z) = \int_{-\infty}^z f_Z(z) dz$$

$$= \frac{1}{\sqrt{2\pi}} \int_{-\infty}^z e^{-t^2/2} dt$$

Q) The order at a restaurant is random. Variance with mean = \$8,  $S.D \sigma = \$2$ .

Estimate the prob that 1st 100 customers spend a total of more than \$840.

Estimate the prob that 1st 100 customers spend a total of between \$780 and \$820.

$$\text{mean} = nm \quad n = 100 \\ = 8 \times 100 = 800$$

$$\text{Variance} = n\sigma^2 = 100 \times 4 = 400$$

$$Z_n = \frac{S_n - nm}{\sigma \sqrt{n}}$$

$$Z_{100} = \frac{S_{100} - 800}{\sqrt{400}}$$

$$P[S_{100} > 840]$$

$$S_{100} = \sum_{i=1}^{100} X_i$$

$$P[20 Z_{100} + 800 > 840]$$

$$P[Z_{100} > \frac{840 - 800}{20}]$$

$$P[X_{100} > 2]$$

$$\approx \Phi(2) = 2.28 \cdot 10^{-2}$$

$$P[780 \leq S_{100} \leq 820]$$

$$\approx P[-1 \leq Z_{100} \leq 1] = 1 - 2\Phi(1) = 0.684$$

## # Convergence Modes

let us consider sequence of RV's  $X_1, X_2, \dots, X_n$   
 then for specific  $\{s\}$  is a sequence of no.  
 that might not converge

→ Convergence everywhere (e)

A sequence of no.  $X_n$  tends to a limit  $x$  if  
 given  $\epsilon > 0$ , we can find a no.  $n_0$  such that  
 $|X_n - x| < \epsilon$  for every  $n > n_0$ .  
 In other words

limit of seq.  $X_n$  is a RV of  $x$   
 $X_n \rightarrow x$  as  $n \rightarrow \infty$

→ Convergence almost everywhere (a.e) - If the  
 set of outcomes ' $s$ ' such that  $\lim X_n(s) = x(s)$   
 as  $n \rightarrow \infty$  exist and its prob equals 1. Then  
 the sequence  $X_n$  converges everywhere (or with prob 1)

$$P\{X_n \rightarrow x\} = 1 \text{ as } n \rightarrow \infty - \textcircled{3}$$

event  $\{X_n \rightarrow x\}$  consist of all outcomes ' $s$ '  
 such that  $X_n \rightarrow x$ .

$X_n$  is a sequence of RV's,  $X$  is a RV

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- 3) Convergence in the mean square sense (M.S.)  
The sequence  $X_n$  tends to RV ' $X'$  in M.S.

$$E \{ |X_n - X|^2 \} \rightarrow 0 \text{ as } n \rightarrow \infty \quad \rightarrow \textcircled{1}$$

if limit in the mean sense is written as  
 $\lim X_n = X \text{ as } n \rightarrow \infty$ .

- 4) Convergence in probability (P).

The probability  $P \{ |X_n - X| \geq \epsilon \}$  of the event  $\{|X_n - X|\}$  is a sequence of numbers depending on  $\epsilon$ , if this sequence tends to 0

$$P \{ |X_n - X| \geq \epsilon \} \rightarrow 0 \text{ as } n \rightarrow \infty \quad \rightarrow \textcircled{2}$$

for any  $\epsilon > 0$ , then the sequence  $X_n$  tends to RV ' $X$ ' in probability. This is also called stochastic convergence.

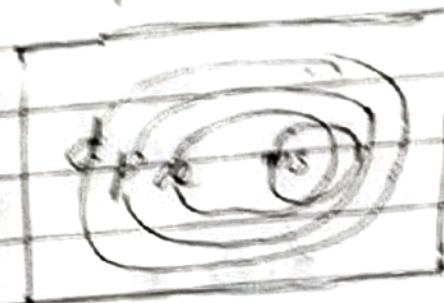
- 5) Convergence in distribution (d) - Let  $F_n(x)$  and  $F(x)$  are the distribution of  $X_n$  &  $X$  respectively if

$$F_n(x) \rightarrow F(x) \text{ as } n \rightarrow \infty \quad \rightarrow \textcircled{3}$$

for every point  $x$  of continuity of  $F(x)$ , then the sequence  $X_n$  tends to RV ' $X$ ' in distribution.

In this case, sequence  $x_n(3)$  need not converge for any  $\epsilon$ .

### Convergence in measure



Each point in vector space represent a random sequence

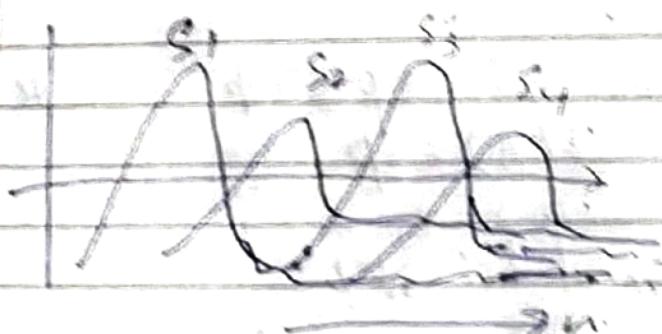
All the sequence converges in the stated mode written on each vector

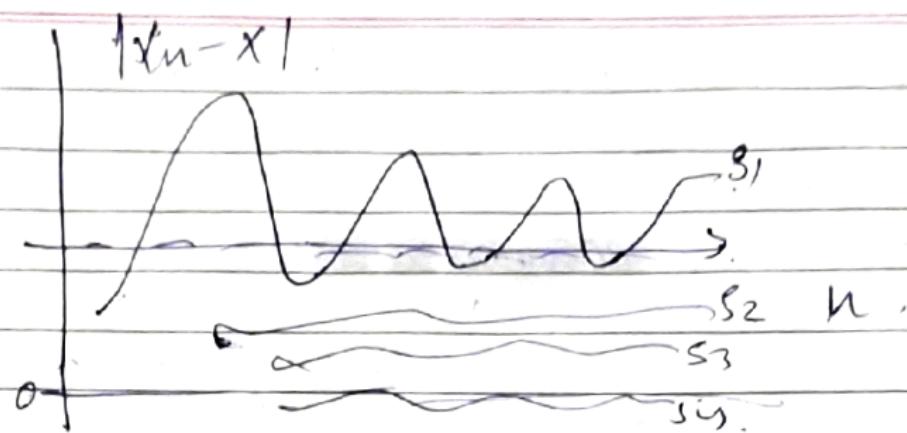
→ To prove & show that if a sequence converges at all then it also converges in distribution

→ If a sequence converges in RS then it also converges in prob. that means

$$\{P\{|x_n - x| \geq \epsilon\} \leq \frac{\epsilon}{c^2}\}$$

→ As convergence in eqn (3)  $\Rightarrow$  also convergence in p. the converge is not necessarily true.





In this fig sequences drawn as curves and each curve represent a particular sequence  $|X_n(s) - X(s)|$

means

→ Convergence in  $P_{\omega,n}$  for specific  $n > n_0$   
only small percent of this curve will have  
ordinates that exceeds  $\epsilon$  (fig 1-6)

→ It is also possible that not of these  
less than  $\epsilon$  for every

Convergence (a.e)  $\Rightarrow$  demands that most  
curve will be below  $\epsilon$  for every  $n > n_0$   
fig (1-6)

## # The Law of Large Numbers

Event A, occur with prob  $p$  in  $k$  times in  
 $n$  trials  $P(A) = p$

then  $\boxed{k \approx np}$

The approximation  $R \approx np$ , means the value  $\frac{R}{n}$  closer to  $p$ .

In the sense that, for any  $\epsilon > 0$ , prob that

$$|\left(\frac{R}{n}\right) - p| < \epsilon \rightarrow 1 \text{ as } n \rightarrow \infty$$

i.e.  $\left[ \left\{ \left| \frac{R}{n} - p \right| < \epsilon \right\} \rightarrow 1 \text{ as } n \rightarrow \infty \right]$

To reestablish this result as a limit of sequence of random variables

R.V's  $X_i = \begin{cases} 1, & \text{if } A \text{ occurs at } i\text{th trial} \\ 0, & \text{otherwise.} \end{cases}$

Sample mean is

$$\bar{X}_n = \frac{X_1 + X_2 + \dots + X_n}{n} \text{ of the RV's prob as}$$

proof  $E\{X_i\} = E\{\bar{X}_n\} = p$

$$\sigma_{X_i}^2 = pq \quad \sigma_{\bar{X}_n}^2 = pq/n.$$

further more

$$pq = p(1-p) \leq 1/n$$

Hence

$$P \left\{ |\bar{X}_n - p| < \epsilon \right\} \geq 1 - \frac{1}{n\epsilon^2} \text{ as } n \rightarrow \infty$$

This establishes eq "②" because  $\bar{X}_n$  as  $K$  if  $a$  occurs  $K$  time

## # Probability

## # Strong law of large no.

$X_1, X_2, \dots, X_n$  sequence of RV's with finite mean  $E[X] = \text{mean}$  &  $\text{Var}[X] = \sigma^2$

then  $P[\lim_{n \rightarrow \infty} M_n = m] = 1$

Sample mean  $M_n = \frac{1}{n} \sum_{i=1}^n X_i$

## # Weak law of large no.

$X_1, X_2, \dots, X_n$  sequence of RV's with  $E[X]$ ,  
for  $\epsilon > 0$   $\lim_{n \rightarrow \infty} P[|M_n - m| < \epsilon] = 1$

$P[\lim_{n \rightarrow \infty} M_n = m] = 1$

for a mean finite value of sample mean  $M$  close to true mean.

with Prob 1 every sequence, Sample mean approaches to  $E[X] = m$

e.g. ① Tossing a fair dice  $S = \{1, 2, 3, 4, 5, 6\}$

②  $X_n$  represent the temp  $T = \{1, 2, 3, 4, \dots\}$  discrete  
 $S = \{\text{any temp}\}$ ,  $T = \{1, 2, \dots, 24\}$  at the end of the nth lecture

③  $X(t)$  represent no. of calls received  $\rightarrow 0, 1, \dots$   
 $T = \text{anything}$

④  $X(t)$  represent man<sup>m</sup> temp<sup>r</sup>  $(0, t) \rightarrow \text{any th}$

## Mod-5 Random Process

**Random Variable**  $\Rightarrow$  "a func" that assigns a real no. to every outcome of a random experiment.

**Random process**  $\Rightarrow$  It is a collection or ensemble of random variables  $\{X(s, t)\}$  that are functions of real variable namely time 't' and  $s$ , where  $s \in S$ ,  $t \in T$

Based on  $X(s, t)$  random process can be defined as

- ①  $X(s, t)$  is a RV if it is constant &  $s$  is variable.
- ②  $X(s, t)$  is a random process if both  $s$  and  $t$  are variable.
- ③  $X(s, t)$  is time func if  $s$  is fixed.
- ④  $X(s, t)$  is a number if  $s$  &  $t$  both are fixed

### # Classification of Random process.

Based on continuous or discrete nature of  $S$  and  $T$  as RP is classified into 4 types

- ① If both  $T$  &  $S$  are discrete then the RP is a discrete Random sequence
- ② If  $T$  is discrete &  $S$  is continuous, continuous random
- ③ If  $T$  is continuous &  $S$  is discrete  $\rightarrow$  discrete RP
- ④ If  $T$  &  $S$  are continuous  $\rightarrow$  continuous RP

# First order cdf  $F(x, t) = P[X(t) \leq x]$

Pdf  $f(x, t) = \frac{d}{dx} F(x, t)$  Regd. No. \_\_\_\_\_  
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# Description of Random process

Cdf of a R.P.  $\rightarrow X(t)$  or  $X(s, t)$  is defined  
as  $F_{X,T}(x, t) = P\{X(t) \leq x\}$

$F(x, t) = P\{X(t) \leq x\} \rightarrow$  1st order distribution

$F(x_1, x_2, t_1, t_2) = P\{X(t_1) \leq x_1; X(t_2) \leq x_2\}$

$f(x_1, x_2, \dots, x_n; t_1, t_2, \dots, t_n) \rightarrow$  2nd order  
 $= \frac{\partial^n}{\partial x_1 \partial x_2 \dots \partial x_n} F(x_1, x_2, \dots, x_n; t_1, t_2, \dots, t_n)$

The pdf of RP  $X(t)$  is

$f(x, t) = \frac{\partial^n}{\partial x^n} F(x, t)$

$f(x_1, x_2; t_1, t_2) = \frac{\partial^2}{\partial x_1 \partial x_2} F(x_1, x_2; t_1, t_2)$

# Stationary process  $\rightarrow$  if certain probability distributions or averages do not depend on time  
then RP  $\{X(t)\}$  is called stationary process

① Strict sense stationary process (SSS)

② Wide sense stationary process (WSS).

A random process is called strongly stationary process or strict sense process if all its finite dimensional distributions are invariant under the translation of time.

① If  $f(x, t) = f(x, t+t_0)$  where  $t_0 > 0$   
and it is possible if  $f(x, t)$  is independent of  $t$ .

② It is possible if  $f(x, t)$  is independent of  $t$ .

③ Ass  $E\{X(t)\} = m = \text{a constant}$

④ Auto correlation of  $R(t_1, t_2)$  or  $K_c$

of this  $R(t_1, t_2) = E[X(t_1)X(t_2)]$  is also  
a function of  $t_1 - t_2$ .  
 $t_1 = t_2$

⑤ Wide sense stationary process (WSS)

⑥ if  $E\{X(t)\} = m = \text{a constant}$ .

⑦  $\text{Var}\{X(t)\} = \sigma^2 = \text{a constant}$ ,

⑧  $R(t_1, t_2) = E[X(t_1)X(t_2)]$

= function of  $(t_1 - t_2)$

# Mean  $\Rightarrow t = \text{continuous } X(t)$

$$\mu(t) = E\{X(t)\}$$

# Autocorrelation

$$R_{XX}(t_1, t_2) = R_X(t_1, t_2) = R(t_1, t_2) \\ = E\{X(t_1)X(t_2)\}$$

$$C_{XX}(t_1, t_2) \text{ or } C_X(t_1, t_2) \text{ or } C(t_1, t_2) = R(t_1, t_2) - E\{X(t_1)X(t_2)\}$$

## \* Stationary process .

(1) Show that random process  $X(t) = A \cos(\omega_0 t + \theta)$  is WSS process - if  $A$  and  $\omega_0$  are constants and  $\theta$  is uniformly distributed RV over  $(0, 2\pi)$

$$f_\theta(\omega) = \frac{1}{2\pi}, \text{ for } (0, 2\pi).$$

$$\textcircled{1} \text{ Mean } E[X(t)] = \int_0^{2\pi} X(t) f_\theta(\omega) d\omega.$$

$$= \int_0^{2\pi} A \cos(\omega_0 t + \theta) \cdot \frac{1}{2\pi} d\omega$$

$$= \frac{A}{2\pi} \int (\cos \omega_0 \cos \theta - \sin \omega_0 \sin \theta) d\omega$$

$$= 0$$

= a const.

$$\textcircled{2} \text{ Auto-correlation } R(t_1, t_2) = E[X(t_1) X(t_2)]$$

$$= E[A^2 \cos(\omega_0 t_1 + \theta) \cdot \cos(\omega_0 t_2 + \theta)]$$

$$= A^2 \cdot E[\cos(\omega_0 t_1 + \theta) \cos(\omega_0 t_2 + \theta)]$$

$$= \frac{A^2}{2} E[2 \cos(\omega_0 t_1 + \theta) \cos(\omega_0 t_2 + \theta)]$$

$$= \frac{A^2}{2} E[(\cos(t_1 + t_2)\omega_0 + \lambda\theta) + \cos(\omega_0(t_1 - t_2))]$$

$$= \frac{A^2}{2} \int_0^{2\pi} \frac{1}{\alpha^2} x'(t) d\theta.$$

$$\Rightarrow \frac{A^2}{4\pi} \int_0^{2\pi} \left\{ \cos[(t_1 + t_2)\omega_0 + \phi_0] + \cos[\omega_0(t_1 - t_2)] d\theta \right\}$$

$$= \frac{A^2}{4\pi} \frac{A^2}{2} \cos[\omega_0(t_1 - t_2)]$$

= f<sup>n</sup> of  $t_1 - t_2$  i.e. time dependent.

To do Q) Verify that the process  $X(t) = Y(\cos t)$  is S.S where  $Y$  is uniformly distributed b/w 0 to 1.

$$\Rightarrow F(x, t) = F(x, t + \Delta t)$$

$$f(x, t) = f(x, t + \Delta t)$$

$$X(t) = Y \cos \omega t$$

$$F(x, t) = P[X(t) \leq x] = P[Y \cos \omega t \leq x]$$

$$\Rightarrow P\left\{Y \leq \frac{x}{\cos \omega t}\right\}, \text{ if } \cos \omega t > 0$$

$$\left\{ P\left\{Y \geq \frac{x}{\cos \omega t}\right\}, \text{ if } \cos \omega t < 0\right.$$

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$$\text{i.e } F_{X(t)}(x, t) = \begin{cases} F_Y\left(\frac{x}{\cos \omega t}\right), & \text{if } \cos \omega t \\ 1 - F_Y\left(\frac{x}{\cos \omega t}\right), & \text{if } \cos \omega t \end{cases}$$

$$\therefore f_{X(t)}(x, t) = \frac{1}{|\cos \omega t|} \times f_Y\left(\frac{x}{\cos \omega t}\right)$$

= "fun" of ' $\omega t$ '.

Hence not SSS.

① Mean of the process.  $\{x(t)\}$

$$m(t) = E\{x(t)\}$$

② Autocorrelation

$$R(t_1, t_2) = E\{x(t_1) \cdot x(t_2)\}$$

③ Autocovariance

$$\begin{aligned} C(t_1, t_2) &= E\{[x(t_1) - m(t_1)] \cdot \\ &\quad [x(t_2) - m(t_2)]\} \\ &= R(t_1, t_2) - m(t_1)m(t_2) \end{aligned}$$

④ Correlation coefficient

$$\rho(t_1, t_2) = \frac{C(t_1, t_2)}{\sqrt{C(t_1, t_1) \cdot C(t_2, t_2)}}$$

$$R_{XY}(t_1, t_2) = E\{x(t_1) \cdot Y(t_2)\}$$

## # Autocorrelation.

$$R(t_1, t_2) = E[X(t_1) X(t_2)]$$

$$= E[X(t) X(t-\epsilon)]$$

$$\therefore \epsilon = t_1 - t_2$$

where  $X(t)$  is a random process.

## # Properties of Autocorrelation f"

$\rightarrow R(\epsilon)$  is even f" of  $\epsilon$

$$\text{i.e } R(-\epsilon) = R(\epsilon)$$

$\rightarrow R(\epsilon)$  is maximum at  $\epsilon = 0$ .

$$|R(\epsilon)| \leq R(0)$$

3) If ACF of real stationary process  $\{X(t)\}$  is continuous at  $\epsilon = 0$ ; then it is continuous at every point.

$R(\epsilon)$  is continuous at  $\epsilon = 0$   $\lim_{\epsilon \rightarrow 0} R(\epsilon) = R(0)$

$R(\epsilon)$  is continuous for all  $\epsilon$ .

$$\lim_{h \rightarrow 0} \{R(\epsilon + h)\} = R(\epsilon)$$

$$\text{mean } m_x = \overline{\lim_{\epsilon \rightarrow \infty} R(\epsilon)}$$

$$R(\epsilon) = E[X(t)Y(t-\epsilon)]$$

cross correlat"

6) Cross covariance of two processes.

$\{X(t)\}$  &  $\{Y(t)\}$  is.

$$c_{xy}(t_1, t_2) = R_{xy}(t_1, t_2) - M_x(t_1) \cdot M_y(t_2)$$

7) Cross correlation coeff. of two processes.

$\{X(t)\}$  &  $\{y(t)\}$  is

$$f_{xy}(t_1, t_2) = \frac{c_{xy}(t_1, t_2)}{\sqrt{c_{xx}(t_1, t_1)c_{yy}(t_2, t_2)}}$$

$$x(t) = A \cos(\omega_0 t + \phi)$$

$$y(t) = A \sin(\omega_0 t)$$

# Auto correlation of  $x^n$ .

## # Ergodicity

A process is said to be ergodic if its ensemble mean is equal to time avg

for exp  $\Rightarrow$  ensemble mean of a discrete random process  $\{X(t)\}$  is

and time avg of a single sample fun<sup>n</sup>

$$X_T = \frac{1}{2T} \int_{-T}^T X(t) dt.$$

const  
current

$$\boxed{m_x = \bar{X}_T}$$

## # Correlation Ergodic

The stationary process  $\{X(t)\}$  is correlation ergodic if process  $\{\gamma(t)\}$  is mean ergodic, where  $\gamma(t) = x(t)x(t+\varepsilon)$

$$\text{i.e } \bar{\gamma}_T = \frac{1}{2T} \int_{-T}^T x(t)x(t+\varepsilon) dt$$

tends to  $E\{x(t+\varepsilon)^T x(t)\} = \kappa(\varepsilon)$  as  $T \rightarrow \infty$

- Q) If WSS process. Is  $x(t) = 10 \cos(100t + \theta)$   
 where  $\theta$  is uniformly distributed over  $(-\pi, \pi)$   
 Prove that  $x(t)$  is correlation ergodic.

$$V_T = \frac{1}{dT} \int_{-\infty}^{\infty} x(t) x(t + \varepsilon) d\varepsilon$$

$$R(\varepsilon) = \lim_{dT \rightarrow 0} \left\{ x(t + \varepsilon) x(t) \right\}$$

# Power spectral density (PSD)

if  $x(t)$  is a stationary process

$$\Rightarrow S(\omega) = \int_{-\infty}^{\infty} R(\varepsilon) e^{-j\omega\varepsilon} d\varepsilon$$

$$= F.T [R(\varepsilon)]$$

$$\omega = 2\pi f$$

$$\text{Then } s(f) = \int_{-\infty}^{\infty} R(\varepsilon) e^{-j2\pi f \varepsilon} d\varepsilon$$

(q) Given that PSD, the ACF  $R(\varepsilon)$  is -

$$R(\varepsilon) = \frac{1}{2\pi} \int_{-\infty}^{\infty} |S(\omega)| e^{j\omega\varepsilon} d\omega$$

i.e

$$R(\varepsilon) = \int_{-\infty}^{\infty} s(f) e^{j2\pi f \varepsilon} df$$

$$= FT \{S(\omega)\}$$

# Properties ①  $S(0) = \int_{-\infty}^{\infty} R(\epsilon) d\epsilon$   
 at  $\epsilon = 0$   $\int_{-\infty}^{\infty}$  Area under  $R(\epsilon)$ .

②  $E\{x^2(t)\} = R(0) = \int_{-\infty}^{\infty} S(f) df$

Mean square  
of value  
of  $x(t)$

Total area under  
curve  
PSD.

③ PSD of real random process is an even.  
 $\Rightarrow S(w) = S(-w)$

④ The power spectral density of  $x(t)$  is  
 real & even function of  $w$  and non-

$$S^*(w) = S(w)$$

⑤ The PSD & ACF of a real WSS process  
 from a Fourier cosine transform pair

⑥  $R(\epsilon) = \text{FICF of } \left[ \frac{1}{2} S(w) \right]$

Proof  $\rightarrow$  Wiener - Kacchine Theorem.

#

Markov's  $\rightarrow$  upper bound.

Mean  $\rightarrow$  markov's

Let  $X$  be a positive random variable and also real no. other

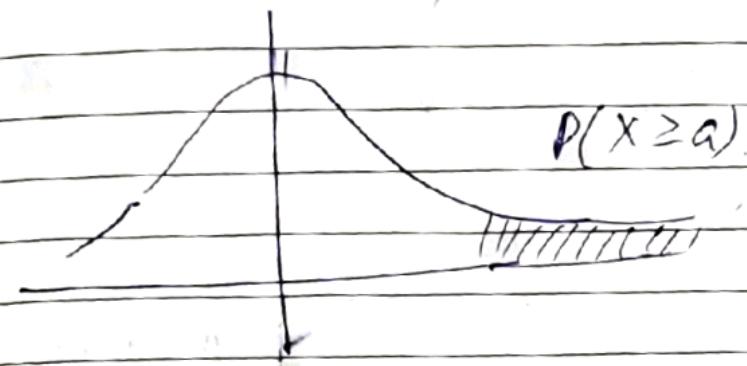
$$P(X \geq a) \leq \frac{E(X)}{a} \quad E(X) = \text{mean}$$

$$X = R \cdot V$$

$a \geq a \rightarrow$  upper bound

$a \leq a \rightarrow$  lower bound

$a = \text{upper bound}$



Let  $X$  be a discrete random variable with pmf (summation)  $p(x)$ .

$$\Rightarrow E(X) = \sum x p(x) \quad \begin{matrix} 1+2=3 \\ 1 \geq 3 \end{matrix}$$

$$= \sum_{x < a} x p(x) + \sum_{x \geq a} x p(x)$$

$$\geq \sum_{x \geq a} x p(x) \quad ; \quad x \geq a$$

$$\geq \sum_{x \geq a} a p(x) \quad x = a$$

$$\geq a \sum_{x \geq a} p(x)$$

$$\boxed{P(X \geq a) \leq \frac{E(X)}{a}}$$

A positive continuous R.V. with pdf  $f(x)$

$$E(X) = \int_{-\infty}^{\infty} xf(x) dx, \quad a > 0.$$

$$= \int_0^{\infty} xf(x) dx.$$

$$= \int_0^a xf(x) dx + \int_a^{\infty} xf(x) dx$$

$$= \int_a^{\infty} xf(x) dx$$

$$\geq a \int f(x) dx$$

$$\geq a P(X \geq a)$$

$$P(X \geq a) \leq \frac{E(X)}{a}$$

\* Only have to know expected value (not say var.)

→ It does not depend on any property of probability distribution of the R.V.

(Q)  $G(x) = 70\%$

$$a = 90\%$$

$$P(X \geq 0.9) \leq \frac{E(X)}{0.9}, \leq \frac{7}{9}$$

$$= 77.8\%$$

## Binomial.

Q)  $n = 20 \rightarrow$  terms  $\rightarrow$  Binomial

$$E(X) = np \quad p = \frac{20}{100} = 1$$

$$\text{Varia} = npq$$

$$E(X) = np = \frac{20 \times 1}{100} = 4$$

$$P(X \geq 16) \leq \frac{E(X)}{16} \leq \frac{4}{16} = \frac{1}{4}$$

Q)	$x$	0	1
$p(x)$		$\frac{24}{25}$	$\frac{1}{25}$
		24	1

$$E(X) = 0 \times \frac{24}{25} + 1 \times \frac{1}{25} \quad P(X \geq 5)$$

$$\geq \frac{1}{25}$$

$$P(X \geq 5) \leq \frac{E(X)}{5} = \frac{\frac{1}{25}}{5} \leq \frac{1}{125}$$

# Chebyshev's

Chelyshew's

$$\# P(|X-\mu|) \geq k\sigma \leq \frac{1}{k^2}$$

$\mu$  = mean  
 $\sigma$  = variance.

# By Markov's inequality  $X$  is RV also and

$$P(|X| \geq a) \leq \frac{E|X|}{a}$$

$$X = (x - \mu)^2 \quad a = K^2 \sigma^2$$

$$\begin{aligned} P(|x-\mu| \geq k\sigma) &\leq \frac{E((x-\mu)^2)}{K^2 \sigma^2} \\ &\leq \frac{\sigma^2}{K^2 \sigma^2} \\ &\leq \frac{1}{K^2} \end{aligned}$$

→ four continuous variable

$$\# \sigma^2 = E((x-\mu)^2) \quad (x-\mu) = K^2 \sigma^2$$

$$= \int_{-\infty}^{\infty} |x-\mu|^2 f(x) dx$$

$$= \int_{|\mu| < K\sigma} |x-\mu|^2 f(x) dx + \int_{|\mu| \geq K\sigma} |x-\mu|^2 f(x) dx$$

$$\geq \int_{|\mu| > K\sigma} |\mu|^2 f(\mu)$$

$$\# \geq \int k^2 \sigma^2 f(x) dx$$

$$|x - \mu| \geq k\sigma$$

$$\text{P} | (x - \mu) | \geq k\sigma \leq k^2$$

$$\sigma^2 \geq \int k^2 \sigma^2 f(x) dx$$

$$\frac{1}{k^2} \geq \int f(x) dx$$

$$\text{P} | (x - \mu) | \geq k\sigma \leq \frac{1}{k^2}$$

$$\# \sigma^2 = E |(x - \mu)|^2 \quad \text{for discrete } X.$$

$$= \sum |x - \mu|^2 f(x)$$

$$= \sum_{|x - \mu| < k\sigma} (x - \mu)^2 f(x) + \sum_{|x - \mu| \geq k\sigma} (x - \mu)^2 f(x)$$

$$\geq \sum_{|x - \mu| \geq k\sigma} |x - \mu|^2 f(x).$$

$$\geq \sum_{|x - \mu| \geq k\sigma} k^2 \sigma^2 f(x)$$

$$\frac{1}{k^2} \geq \sum_{|x - \mu| \geq k\sigma} f(x)$$

$$(P | (x - \mu) | \geq k\sigma) \leq \frac{1}{k^2}$$

Q) Suppose that  $\alpha(t) \xrightarrow{\text{fixed mean}} \mu(t) = 3$ .  
 $R(t_1, t_2) = 4 + 4e^{-0.2|t_1 - t_2|}$ .

Given  $N$ , Cov.  $Z = X(5)$  &  $W = X(8)$

$$E(Z) = 3 \quad E(W) = 3$$

$$\text{Var}(X(s)) = R(s, s) - (\mu(s))^2$$

$$= 13 - 9 = 4$$

$$\text{Var}(X(8)) = R(8, 8) - (\mu(8))^2$$

$$= 13 - 9$$

Covariance  $C(s, 8) = R(s, 8) - \mu(s)\mu(8)$

$$= 9 + 4e^{-0.6} - 9 = 2.19.$$

Q)  $X(t) = A \cos(\omega t + \theta)$  RV in  $(0, 2\pi)$ .

$$E(X) = E(A \cos(\omega t + \theta)) = \text{uniform distn}$$

$$\text{mean} = \frac{1}{b-a} \int_a^b xf(x) dx$$

$$= \frac{1}{2\pi} \int_0^{2\pi} A \cos(\omega t + \theta) \frac{1}{2\pi} d\theta \quad \frac{1}{2\pi-d} \frac{1}{2\pi}$$

$$= \frac{A}{2\pi} \left[ \sin(\omega t + \theta) \right]_0^{2\pi}$$

$$= \frac{A}{2\pi} [\sin(2\pi + \omega t) - \sin \omega t]$$

$$= 0$$

$$0$$