

Lecture 19: Sufficiency

Mathematical Statistics I, MATH 60061/70061

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Reference: Casella & Berger, 6.1-6.2.1

Factorization Theorem

So far, we've used the definition of sufficiency directly by showing that the conditional distribution of X given T is free of θ . What if we need to find a sufficient statistic?

Factorization Theorem: A statistic $T = T(\mathbf{X})$ is **sufficient** for θ if and only if there exists functions $g(t \mid \theta)$ and $h(\mathbf{x})$ such that

$$f_{\mathbf{X}}(\mathbf{x} \mid \theta) = g(t \mid \theta)h(\mathbf{x}),$$

for all sample points $\mathbf{x} \in \mathcal{X}$ and all $\theta \in \Theta$.

Normal sufficient statistic

Suppose X_1, \dots, X_n are iid from $\mathcal{N}(\mu, \sigma^2)$, with parameter $\theta = (\mu, \sigma^2)$. The joint PDF is

$$\begin{aligned}f_X(x \mid \theta) &= \prod_{i=1}^n \frac{1}{\sqrt{2\pi\sigma^2}} e^{-(x_i - \mu)^2 / 2\sigma^2} \\&= \left(\frac{1}{\sqrt{2\pi\sigma^2}} \right)^n \exp \left(- \sum_{i=1}^n \frac{(x_i - \mu)^2}{2\sigma^2} \right) \\&= \left(\frac{1}{\sqrt{2\pi\sigma^2}} \right)^n \exp \left(- \sum_{i=1}^n \frac{(x_i - \bar{x})^2}{2\sigma^2} - \frac{n(\bar{x} - \mu)^2}{2\sigma^2} \right) \\&= \left(\frac{1}{\sqrt{2\pi\sigma^2}} \right)^n \exp \left(- \frac{(n-1)s^2}{2\sigma^2} - \frac{n(\bar{x} - \mu)^2}{2\sigma^2} \right)\end{aligned}$$

where s^2 is the realization of the sample variance
 $S^2 = (n-1)^{-1} \sum_{i=1}^n (X_i - \bar{X})^2$.

$$f_X(x | \theta) = \left(\frac{1}{\sqrt{2\pi\sigma^2}} \right)^n \exp \left(-\frac{(n-1)s^2}{2\sigma^2} - \frac{n(\bar{x} - \mu)^2}{2\sigma^2} \right)$$

By the Factorization Theorem, (\bar{X}, S^2) is a sufficient statistic for $\theta = (\mu, \sigma^2)$.

- If σ^2 is known, then \bar{X} is sufficient for μ .
- If μ is known, then S^2 is sufficient for σ^2 .
- If both μ and σ^2 are unknown, we cannot say that \bar{X} is sufficient for μ (or S^2 is sufficient for σ^2); the correct statement is that \bar{X} and S^2 together is sufficient for μ and σ^2 .
- We can also say that (\bar{X}, S^2) is sufficient for μ (or σ^2).

Sufficient statistics in the Exponential family

Suppose X_1, \dots, X_n are iid from the **Exponential family**

$$f_X(x \mid \boldsymbol{\theta}) = h(x)c(\boldsymbol{\theta}) \exp \left(\sum_{j=1}^k w_j(\boldsymbol{\theta}) t_j(x) \right),$$

where $\boldsymbol{\theta} = (\theta_1, \dots, \theta_d)$, $d \leq k$. Then

$$\boldsymbol{T} = \boldsymbol{T}(\boldsymbol{X}) = \left(\sum_{i=1}^n t_1(X_i), \sum_{i=1}^n t_2(X_i), \dots, \sum_{i=1}^n t_k(X_i) \right)$$

is sufficient for $\boldsymbol{\theta}$.

Binomial sufficient statistic

Suppose X_1, \dots, X_n are iid $\text{Bern}(\theta)$ with parameter $0 < \theta < 1$. Then $T(\mathbf{X}) = X_1 + \dots + X_n$ is a sufficient statistic for θ .

The PMF of \mathbf{X} is given by

$$f_{\mathbf{X}}(\mathbf{x} \mid \theta) = \prod_{i=1}^n \theta^{x_i} (1 - \theta)^{1-x_i}.$$

Note that $T(\mathbf{X})$ counts the number of X_i 's that equal 1, so $T(\mathbf{X})$ has a $\text{Bin}(n, \theta)$ distribution,

$$f_T(t \mid \theta) = \binom{n}{t} \theta^t (1 - \theta)^{n-t}.$$

With $t = \sum_{i=1}^n x_i$, the conditional distribution

$$f_{\mathbf{X}|T}(\mathbf{x} \mid t) = \frac{f_{\mathbf{X}}(\mathbf{x} \mid \theta)}{f_T(t \mid \theta)} = \frac{\theta^t (1 - \theta)^{n-t}}{\binom{n}{t} \theta^t (1 - \theta)^{n-t}} = \frac{1}{\binom{n}{\sum x_i}},$$

which is free of θ . Therefore, $T(\mathbf{X}) = \sum_{i=1}^n X_i$ is a sufficient statistic.

Suppose X_1, \dots, X_n are iid $\text{Bern}(\theta)$ with parameter $0 < \theta < 1$. For $x = 0, 1$, the PMF of X is

$$\begin{aligned} f_X(x \mid \theta) &= \theta^x (1 - \theta)^{1-x} \\ &= (1 - \theta) \left(\frac{\theta}{1 - \theta} \right)^x \\ &= (1 - \theta) \exp \left(\log \left(\frac{\theta}{1 - \theta} \right) x \right) \\ &= h(x) c(\theta) \exp(w_1(\theta) t_1(x)), \end{aligned}$$

where $h(x) = 1$, $c(\theta) = 1 - \theta$, $w_1(\theta) = \log(\theta/(1 - \theta))$, and $t_1(x) = x$.

Therefore,

$$T = T(\mathbf{X}) = \sum_{i=1}^n t_1(X_i) = \sum_{i=1}^n X_i$$

is sufficient.