

## Lecture 20: Evaluation of Interval Estimators

Mathematical Statistics II, MATH 60062/70062

Thursday April 14, 2022

Reference: Casella & Berger, 9.3.1

# Size and coverage probability

When evaluating a confidence set, we want the set to have small **size** and large **coverage probability**.

- Coverage probability is typically measured by the **confidence coefficient**. If the coverage probability is not equal to  $1 - \alpha$  for all  $\theta \in \Theta$ , we would like it to be as close as possible to the nominal  $1 - \alpha$  level.
- If the set is an interval, the size is usually measured by the **interval length**. Shorter intervals are more informative.

These two goals are often in opposition to one another. (Clearly, we can have a large coverage probability by increasing the size.)

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- If the set is an interval, the size is usually measured by the **interval length**. **Shorter intervals are more informative.**

These two goals are often in opposition to one another. (Clearly, we can have a large coverage probability by increasing the size.)

*Finding an optimal interval using these criteria is often a constrained minimization problem, where **for a specified coverage the goal is to find the confidence interval with the shortest length.***

# Optimizing length

Suppose  $X_1, \dots, X_n$  are iid  $\mathcal{N}(\mu, \sigma^2)$ , where  $-\infty < \mu < \infty$  and  $\sigma^2 > 0$ . Both parameters are unknown. Let  $\boldsymbol{\theta} = (\mu, \sigma^2)$ . Based on the pivot

$$Q = Q(\mathbf{X}, \boldsymbol{\theta}) = \frac{\bar{X} - \mu}{S/\sqrt{n}},$$

a  $1 - \alpha$  confidence interval for  $\mu$  can be obtained with

$$1 - \alpha = P_{\boldsymbol{\theta}} \left( a \leq \frac{\bar{X} - \mu}{S/\sqrt{n}} \leq b \right).$$

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Suppose  $n = 25$  and  $1 - \alpha = 0.95$ . Some values of  $a$  and  $b$  that satisfy the condition are

- $a = t_{24,0.96} = -1.828$  and  $b = t_{24,0.01} = 2.492$
- $a = t_{24,0.97} = -1.974$  and  $b = t_{24,0.02} = 2.172$
- $a = t_{24,1} = -\infty$  and  $b = t_{24,0.05} = 1.171$
- $a = t_{24,0.95} = -1.171$  and  $b = t_{24,0} = \infty$

An infinite number of intervals satisfy this condition.

# Optimal length with specified coverage probability

Suppose  $Q = Q(\mathbf{X}, \theta)$  is a pivotal quantity and  $P_\theta(a \leq Q \leq b) = 1 - \alpha$ , where  $a$  and  $b$  are constants. Let  $f_Q(q)$  be the unimodal PDF of  $Q$ . If

①  $\int_a^b f_Q(q) dq = 1 - \alpha$

②  $f_Q(a) = f_Q(b) > 0$

③  $a \leq q^* \leq b$ , where  $q^*$  is the mode of  $f_Q(q)$ ,

then  $b - a$  is minimized relative to  $Q$ .

Proof: Let  $(a', b')$  be any interval with  $b' - a' < b - a$ . We will show that this implies

$$\int_{a'}^{b'} f_Q(q) dq < 1 - \alpha.$$

The result will be proved for  $a' \leq a$ , where two cases need to be considered,  $b' \leq a$  and  $b' > a$ . The proof for  $a' > a$  is similar.

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If  $b' \leq a$ , then  $a' \leq b' \leq a \leq q^*$  and

$$\begin{aligned} \int_{a'}^{b'} f_Q(q) dq &\leq f_Q(b')(b' - a') \\ &\leq f_Q(a)(b' - a') \\ &< f_Q(a)(b - a) \\ &\leq \int_a^b f_Q(q) dq = 1 - \alpha, \end{aligned}$$

which completes the proof in the first case.



If  $b' > a$ , then  $a' \leq a < b' < b$ , since if  $b' \geq b$ ,  $b' - a' \geq b - a$ . In this case, we write

$$\begin{aligned}\int_{a'}^{b'} f_Q(q) dq &= \int_a^b f_Q(q) dq + \left[ \int_{a'}^a f_Q(q) dq - \int_{b'}^b f_Q(q) dq \right] \\ &= (1 - \alpha) + \left[ \int_{a'}^a f_Q(q) dq - \int_{b'}^b f_Q(q) dq \right].\end{aligned}$$

Using the unimodality of  $f$ , the ordering  $a' \leq a < b' < b$ , and  $a \leq q^* \leq b$ , we have

$$\int_{a'}^a f_Q(q) dq \leq f_Q(a)(a - a') \quad \text{and} \quad \int_{b'}^b f_Q(q) dq \geq f_Q(b)(b - b').$$

Since  $f_Q(a) = f_Q(b)$ , the expression in square brackets is

$$\begin{aligned}\int_{a'}^a f_Q(q) dq - \int_{b'}^b f_Q(q) dq &\leq f_Q(a)(a - a') - f_Q(b)(b - b') \\ &= f_Q(a)[(a - a') - (b - b')] \\ &= f_Q(a)[(b' - a') - (b - a)],\end{aligned}$$

which is negative. This implies  $\int_{a'}^{b'} f_Q(q) dq < 1 - \alpha$  for  $b' > a$ , completing the proof in the second case.

# Optimizing length

Suppose  $X_1, \dots, X_n$  are iid  $\mathcal{N}(\mu, \sigma^2)$ , where  $-\infty < \mu < \infty$  and  $\sigma^2 > 0$ . Both parameters are unknown. Let  $\boldsymbol{\theta} = (\mu, \sigma^2)$ . Based on the pivot

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If we choose  $a = -t_{n-1, \alpha/2}$  and  $b = t_{n-1, \alpha/2}$ , then the conditions for optimal length are satisfied. Therefore,

$$\left( \bar{X} - t_{n-1, \alpha/2} \frac{S}{\sqrt{n}}, \bar{X} + t_{n-1, \alpha/2} \frac{S}{\sqrt{n}} \right)$$

has the shortest expected length among all  $1 - \alpha$  confidence intervals based on  $Q$ .

# Optimal length with specified coverage probability

## Remark:

- The method works well for location families and location-scale families because the interval's length is proportional to  $b - a$ .
- When the interval's length is not proportional to  $b - a$  (e.g., in scale families), then the method is not directly applicable. However, a modified version of the method might be applicable (e.g., see CB pp. 443-444).