## Lecture 16: P-values

Mathematical Statistics II, MATH 60062/70062

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Reference: Casella & Berger, 8.3.4

# Recap: Hypothesis testing

Given  $X = (X_1, \dots, X_n) \sim f_X(x \mid \theta)$ , where  $\theta \in \Theta$ , consider testing

$$H_0: \theta \in \Theta_0$$
 versus  $H_1: \theta \in \Theta_0^c$ .

A **hypothesis test** finds the rejection region.

After a hypothesis test is done, conclusions may be reported in terms of

- Size  $\alpha$  of the test
- Decision to reject  $H_0$  or accept  $H_0$ .

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After a hypothesis test is done, conclusions may be reported in terms of

- Size  $\alpha$  of the test
- Decision to reject  $H_0$  or accept  $H_0$ .

Another way of reporting the results is to report the value of a certain kind of test statistic called a **p-value**.

#### P-values

A **p-value**  $p(\boldsymbol{X})$  is a test statistic satisfying  $0 \le p(\boldsymbol{x}) \le 1$  for every sample point  $\boldsymbol{x}$ . Small values of  $p(\boldsymbol{X})$  give evidence against  $H_0$ . A p-value is **valid** if, for every  $\theta \in \Theta_0$  and every  $0 \le \alpha \le 1$ ,

$$P_{\theta}(p(\boldsymbol{X}) \leq \alpha) \leq \alpha.$$

- If p(X) is a valid p-value, it is easy to construct a level  $\alpha$  test based on p(X).
- The test that rejects  $H_0$  if and only if  $p(\boldsymbol{X}) \leq \alpha$  is a level  $\alpha$  test. That is,  $\phi(\boldsymbol{x}) = I(p(\boldsymbol{x}) \leq \alpha)$  is a level  $\alpha$  test,

$$\sup_{\theta \in \Theta_0} E_{\theta}[\phi(\boldsymbol{X})] = \sup_{\theta \in \Theta_0} P_{\theta}(p(\boldsymbol{X}) \le \alpha) \le \alpha.$$

A p-value reports the results of a test on a more continuous scale.

#### P-value and test statistic

Let  $W=W(\boldsymbol{X})$  be a test statistic such that large values of W give evidence against  $H_0$ . For each sample point  $\boldsymbol{x}\in\mathcal{X}$ , define

$$p(\boldsymbol{x}) = \sup_{\theta \in \Theta_0} P_{\theta}(W(\boldsymbol{X}) \ge W(\boldsymbol{x})).$$

Then p(X) is a valid p-value.

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Then p(X) is a valid p-value.

Useful result: If X have CDF  $F_X(x)$ , then

$$Y = F_X(X) \sim \text{Unif}(0,1),$$

because

$$P(Y \le y) = P(F_X(X) \le y) = P(X \le F_X^{-1}(y)) = F_X(F_X^{-1}(y)) = y.$$

Fix  $\theta \in \Theta_0$ . Let  $F_{-W}(w \mid \theta)$  denote the CDF of  $-W = -W(\boldsymbol{X})$ . Consider a test rejecting  $H_0$  for large values of W, and define

$$p_{\theta}(\mathbf{x}) = P_{\theta}(W(\mathbf{X}) \ge W(\mathbf{x}))$$
  
=  $P_{\theta}(-W(\mathbf{X}) \le -W(\mathbf{x})) = F_{-W}(-W(\mathbf{x}) \mid \theta).$ 

Thus,  $p_{\theta}(\boldsymbol{X}) = F_{-W}(-W(\boldsymbol{X}) \mid \theta)$  is a  $\mathrm{Unif}(0,1)$  random variable. That is, for every  $0 \le \alpha \le 1$ ,

$$P_{\theta}(p_{\theta}(\boldsymbol{x}) \leq \alpha) \leq \alpha.$$

Now, note that

$$p(\boldsymbol{x}) = \sup_{\theta' \in \Theta_0} P_{\theta'}(W(\boldsymbol{X}) \ge W(\boldsymbol{x})) \ge P_{\theta}(W(\boldsymbol{X}) \ge W(\boldsymbol{x})) = p_{\theta}(\boldsymbol{x})$$

Therefore.

$$P_{\theta}(p(X) \le \alpha) \le P_{\theta}(p_{\theta}(X) \le \alpha) \le \alpha.$$

This is true for every  $\theta \in \Theta_0$  and for every  $0 \le \alpha \le 1$ ; p(X) is a valid p-value.

## Two-sided Normal p-value

Suppose  $X_1,\ldots,X_n$  are iid  $\mathcal{N}(\mu,\sigma^2)$ , where  $-\infty<\mu<\infty$  and  $\sigma^2>0$ . Both parameters are unknown. Set  $\boldsymbol{\theta}=(\mu,\sigma^2)$ . Consider testing

$$H_0: \mu = \mu_0$$
 versus  $H_1: \mu \neq \mu_0$ .

In Lecture 12, we showed that the LRT rejects  ${\cal H}_0$  for large values of

$$W = W(\boldsymbol{X}) = \left| \frac{\bar{X} - \mu_0}{S/\sqrt{n}} \right|.$$

This is known as **one-sample two-sided** t **test**.

The null parameter space is

$$\Theta_0 = \{ \boldsymbol{\theta} = (\mu, \sigma^2) : \mu = \mu_0, \sigma^2 > 0 \}.$$

### With observed value w = W(x), the p-value for the test is

$$\begin{split} p(\boldsymbol{x}) &= \sup_{\boldsymbol{\theta} \in \Theta_0} P_{\boldsymbol{\theta}}(W(\boldsymbol{X}) \geq w) \quad \underset{\text{see wikipedia p-value.}}{\text{W(X): the distribution of hypothesis, } w: \text{ observation see wikipedia p-value.} \\ &= \sup_{\boldsymbol{\theta} \in \Theta_0} P_{\boldsymbol{\theta}} \left( \left| \frac{\bar{X} - \mu_0}{S/\sqrt{n}} \right| \geq w \right) \\ &= \sup_{\boldsymbol{\theta} \in \Theta_0} P_{\boldsymbol{\theta}} \left( \frac{\bar{X} - \mu_0}{S/\sqrt{n}} \geq w \text{ or } \frac{\bar{X} - \mu_0}{S/\sqrt{n}} \leq -w \right) \\ &= \sup_{\boldsymbol{\theta} \in \Theta_0} P_{\boldsymbol{\theta}} \left( \frac{\bar{X} - \mu}{S/\sqrt{n}} \geq w + \frac{\mu_0 - \mu}{S/\sqrt{n}} \text{ or } \frac{\bar{X} - \mu}{S/\sqrt{n}} \leq -w + \frac{\mu_0 - \mu}{S/\sqrt{n}} \right) \\ &= P\left( T_{n-1} \geq w \text{ or } T_{n-1} \leq -w \right) \\ &= 2P(T_{n-1} \geq |w|). \end{split}$$

# One-sided Normal p-value

Suppose  $X_1, \ldots, X_n$  are iid  $\mathcal{N}(\mu, \sigma^2)$ , where  $-\infty < \mu < \infty$  and  $\sigma^2 > 0$ . Both parameters are unknown. Set  $\theta = (\mu, \sigma^2)$ . Consider testing

$$H_0: \mu \leq \mu_0$$
 versus  $H_1: \mu > \mu_0$ .

The LRT rejects  $H_0$  for large values of

$$W = W(\mathbf{X}) = \frac{X - \mu_0}{S/\sqrt{n}}.$$

The null parameter space is

$$\Theta_0 = \{ \boldsymbol{\theta} = (\mu, \sigma^2) : \mu \le \mu_0, \sigma^2 > 0 \}.$$

With observed value w = W(x), the p-value for the test is

$$\begin{split} p(\boldsymbol{x}) &= \sup_{\boldsymbol{\theta} \in \Theta_0} P_{\boldsymbol{\theta}}(W(\boldsymbol{X}) \geq w) \\ &= \sup_{\boldsymbol{\theta} \in \Theta_0} P_{\boldsymbol{\theta}} \left( \frac{\bar{X} - \mu_0}{S/\sqrt{n}} \geq w \right) \\ &= \sup_{\boldsymbol{\theta} \in \Theta_0} P_{\boldsymbol{\theta}} \left( \frac{\bar{X} - \mu}{S/\sqrt{n}} \geq w + \frac{\mu_0 - \mu}{S/\sqrt{n}} \right) \\ &= \sup_{\mu \leq \mu_0} P_{\boldsymbol{\theta}} \left( T_{n-1} \geq w + \frac{\mu_0 - \mu}{S/\sqrt{n}} \right) \\ &= P(T_{n-1} \geq w). \end{split}$$