

Lecture 07: Evaluation of Point Estimators

Mathematical Statistics II, MATH 60062/70062

Thursday February 10, 2022

Reference: Casella & Berger, 7.3.1-7.3.2

Mean squared error

The **mean squared error (MSE)** of a point estimator $W = W(\mathbf{X})$ of a parameter θ is

$$\begin{aligned}\text{MSE}_{\theta}(W) &= E_{\theta}(W - \theta)^2 \\ &= \text{Var}_{\theta}(W - \theta) + (E_{\theta}(W - \theta))^2 \\ &= \text{Var}_{\theta}(W) + \text{Bias}_{\theta}^2(W),\end{aligned}$$

where $\text{Bias}_{\theta}(W) = E_{\theta}(W) - \theta$ is the **bias** of W , the difference between the expected value of W and θ . If $\text{Bias}_{\theta}(W) = 0$ for all θ , then W is called an **unbiased** estimator.

MSE incorporates two components:

- **Precision**, measured by $\text{Var}_{\theta}(W)$.
- **Accuracy**, measured by $\text{Bias}_{\theta}(W)$.

Normal variance estimators

Suppose that X_1, \dots, X_n are iid $\mathcal{N}(\mu, \sigma^2)$, where $-\infty < \mu < \infty$ and $\sigma^2 > 0$. Set $\boldsymbol{\theta} = (\mu, \sigma^2)$.

- The sample variance

$$S^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2$$

is an unbiased estimator for σ^2 , where $(n-1)S^2/\sigma^2 \sim \chi_{n-1}^2$, $E_{\boldsymbol{\theta}}(S^2) = \sigma^2$ for all σ^2 and $\text{Var}_{\boldsymbol{\theta}}(S^2) = 2\sigma^4/(n-1)$.

- The MLE of σ^2

$$S_{\text{B}}^2 = \frac{1}{n} \sum_{i=1}^n (X_i - \bar{X})^2 = \frac{n-1}{n} S^2$$

is a biased estimator for σ^2 , $E_{\boldsymbol{\theta}}(S_{\text{B}}^2) = \sigma^2(n-1)/n$.

MSE of Normal variance estimators

- MSE of S^2

$$\text{MSE}_{\boldsymbol{\theta}}(S^2) = \text{Var}_{\boldsymbol{\theta}}(S^2) = \frac{2\sigma^4}{n-1}$$

- MSE of S_B^2

$$\begin{aligned}\text{MSE}_{\boldsymbol{\theta}}(S_B^2) &= \text{Var}_{\boldsymbol{\theta}}(S_B^2) + \text{Bias}_{\boldsymbol{\theta}}^2(S_B^2) \\ &= \left(\frac{n-1}{n}\right)^2 \text{Var}_{\boldsymbol{\theta}}(S^2) + \left(\frac{-\sigma^2}{n}\right)^2 \\ &= \left(\frac{2n-1}{n^2}\right) \sigma^4\end{aligned}$$

The biased MLE S_B^2 has smaller MSE than S^2 , since $2/(n-1) > (2n-1)/n^2$ for all $n \geq 2$.

MSE of Bernoulli/Binomial estimators

Suppose that X_1, \dots, X_n are iid $\text{Bern}(\theta)$, where the prior distribution on θ is $\text{Beta}(a, b)$.

- The MLE of θ , $\hat{\theta} = \bar{X}$, is an unbiased estimator of θ , and its MSE is

$$\text{MSE}_{\theta}(\hat{\theta}) = \text{Var}_{\theta}(\bar{X}) = \frac{\theta(1 - \theta)}{n}.$$

- The Bayes estimator is $\hat{\theta}_B = (T + a)/(n + a + b)$, where $T = \sum_{i=1}^n X_i$. The MSE of $\hat{\theta}_B$ is

$$\begin{aligned}\text{MSE}_{\theta}(\hat{\theta}_B) &= \text{Var}_{\theta} \left(\frac{T + a}{n + a + b} \right) + \text{Bias}_{\theta}^2 \left(\frac{T + a}{n + a + b} \right) \\ &= \frac{n\theta(1 - \theta)}{(n + a + b)^2} + \left(\frac{n\theta + a}{n + a + b} - \theta \right)^2.\end{aligned}$$

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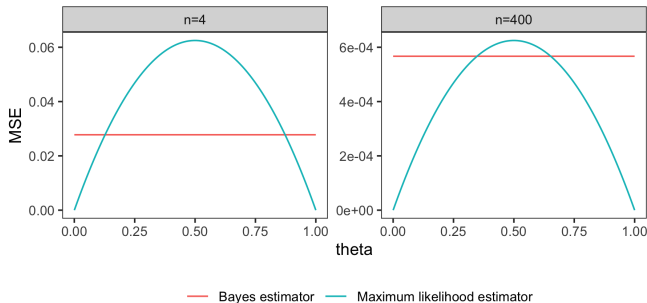
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In the absence of good prior information, we choose $a = b = \sqrt{n/4}$ to make $\text{MSE}_{\theta}(\hat{\theta}_B)$ constant,

$$\hat{\theta}_B = \frac{T + \sqrt{n/4}}{n + \sqrt{n}}, \quad \text{MSE}_{\theta}(\hat{\theta}_B) = \frac{n}{4(n + \sqrt{n})^2}.$$

MSE of Bernoulli/Binomial estimators



- For small n , the Bayes estimator $\hat{\theta}_B$ is a better choice (unless there is a strong belief that θ is near 0 or 1).
- For large n , the MLE $\hat{\theta}$ is a better choice (unless there is a strong belief that θ is close to 1/2).

Best unbiased estimators

In most cases, there is no one “best MSE” estimator. This is because the class of all estimators is too large a class.

Consider a nonsense estimator $\hat{\theta}_N = 0.21$.

- If $\theta = 0.21$, then $\hat{\theta}_N$ is the best in terms of MSE, i.e., $\text{MSE}_\theta(\hat{\theta}_N) = 0$.
- If $\theta \neq 0.21$, then $\hat{\theta}_N$ can be a terrible estimator.

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One way to make the problem of finding a “best” estimator tractable is to limit the class of estimators (e.g., to consider only unbiased estimators).

A best unbiased estimator in terms of MSE is essentially an unbiased estimator whose variance is smaller than any other unbiased estimator.

Uniformly minimum-variance unbiased estimator (UMVUE)

An estimator $W^* = W^*(\mathbf{X})$ is a **uniformly minimum-variance unbiased estimator (UMVUE)** of $\tau(\theta)$ if

- ① $E_{\theta}(W^*) = \tau(\theta)$ for all $\theta \in \Theta$.
- ② $\text{Var}_{\theta}(W^*) \leq \text{Var}_{\theta}(W)$ for all $\theta \in \Theta$, where W is any other unbiased estimator of $\tau(\theta)$.

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Note: UMVUEs may not exist. As we will show later, **if a UMVUE does exist, then it's *unique*.**

Finding UMVUE

Using the definition of UMVUE to find one is difficult. As we have to compare a candidate W^* to all unbiased estimators (there may be infinite of them!).

We will discuss two approaches to find UMVUEs:

- ① (**Cramér-Rao Inequality**) Determine a **lower bound** on the variance of *any* unbiased estimator of $\tau(\theta)$. If we can find an unbiased estimator whose variance attains this lower bound, we have found the UMVUE.
- ② (**Rao-Blackwell Theorem**) Relate the property of UMVUEs with the notation of **sufficiency** and **completeness**.