

VI updates for ASIS

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1 Sufficient augmentation

Under the model,

$$\begin{aligned}\theta &\sim \text{Uniform}(\mathbb{R}) \\ \mu|\theta &\sim \mathcal{N}(\theta, V) \\ Y|\mu &\sim \mathcal{N}(\mu, 1)\end{aligned}$$

we have that μ is a sufficient statistic for θ . We seek the posterior distributions for the latent variables μ and θ . Cheating slightly, we can compute by hand the posterior variances of μ and θ ; in particular, we find that they're given by $1 + V$ and $V/(V + 1)$, respectively. Therefore, we make the mean field approximations

$$\begin{aligned}q_{\hat{\theta}}(\theta) &= \mathcal{N}(\theta; \hat{\theta}, V + 1) \\ q_{\hat{\mu}}(\mu) &= \mathcal{N}(\mu; \hat{\mu}, V/(V + 1))\end{aligned}$$

With this mean field approximation, we compute the evidence lower bound:

$$\begin{aligned}\text{elbo} &= E_q[\log p(Y|\mu)] + E_q[\log p(u|\theta)] + E_q[\log p(\theta)] - E_q[q(\theta)] - E_q[q(\mu)] \\ &= \frac{-1}{2}E_q(Y - \mu)^2 + \frac{-1}{2V}E_q(\mu - \theta)^2 + \frac{1}{2(V + 1)}E_q(\theta - \hat{\theta})^2 + \frac{V + 1}{2V}E(\mu - \hat{\mu})^2 + K \\ &= -\frac{1}{2}\left(E_q\mu^2 - 2YE_q\mu\right) + \frac{-1}{2V}\left(E_q\mu^2 - 2E_q[\mu\theta] + E_q\theta^2\right) + K \\ &= -\frac{1}{2}\hat{\mu}^2 + Y\hat{\mu} - \frac{1}{2V}\hat{\mu}^2 + \frac{1}{V}\hat{\mu}\hat{\theta} - \frac{1}{2V}\hat{\theta}^2 + K\end{aligned}$$

To do coordinate ascent, we compute the two partial derivatives:

$$\begin{aligned}\frac{d(\text{elbo})}{d\hat{\theta}} &= \frac{1}{V}\hat{\mu} - \frac{1}{V}\hat{\theta} \\ \frac{d(\text{elbo})}{d\hat{\mu}} &= -\hat{\mu} + Y - \frac{1}{V}\hat{\mu} + \frac{1}{V}\hat{\theta}\end{aligned}$$

Setting these partials derivatives equal to zero, and solving for $\hat{\theta}$ and $\hat{\mu}$, our CAVI updates are given by

$$\begin{aligned}\hat{\mu}^{(t+1)} &= \frac{VY + \hat{\theta}^{(t)}}{1 + V} \\ \hat{\theta}^{(t+1)} &= \hat{\mu}^{(t+1)}\end{aligned}$$

2 Auxillary augmentation

Under this model,

$$\begin{aligned}\theta &\sim \text{Uniform}(\mathbb{R}) \\ \eta &\sim \mathcal{N}(0, V) \\ Y|\theta, \eta &\sim \mathcal{N}(\theta + \eta, 1)\end{aligned}$$

and in this case, η is an auxillary statistic for θ , since its distribution does not depend on θ . We seek the posterior distributions for the latent variables η and θ , and again we compute by hand the posterior variances of η and θ ; they are given by $V/(V+1)$ and 1, respectively. Therefore, we make the mean field approximations

$$\begin{aligned}q_{\hat{\theta}}(\theta) &= \mathcal{N}(\theta; \hat{\theta}, 1) \\ q_{\hat{\mu}}(\eta) &= \mathcal{N}(\eta; \hat{\eta}, V/(V+1))\end{aligned}$$

We compute the evidence lower bound:

$$\begin{aligned}\text{elbo} &= E_q[\log p(Y|\eta, \theta)] + E_q[\log p(\eta)] + E_q[\log p(\theta)] - E_q[q(\theta)] - E_q[q(\eta)] \\ &= \frac{-1}{2} E_q(Y - \theta - \eta)^2 + \frac{-1}{2V} E_q \eta^2 - \frac{1}{2} E_q(\theta - \hat{\theta})^2 - \frac{1+V}{2V} E(\eta - \hat{\eta})^2 + K \\ &= Y E(\theta + \eta) - \frac{1}{2} E(\theta + \eta)^2 - \frac{1}{2V} E_q \eta^2 + K \\ &= Y \hat{\theta} + Y \hat{\eta} - \frac{1}{2} \hat{\theta}^2 - \frac{1}{2} \hat{\eta}^2 + \hat{\eta} \hat{\theta} - \frac{1}{2V} \hat{\eta}^2 + K\end{aligned}$$

Taking partial derivatives,

$$\begin{aligned}\frac{d(\text{elbo})}{d\hat{\theta}} &= Y - \hat{\theta} + \hat{\eta} \\ \frac{d(\text{elbo})}{d\hat{\eta}} &= Y - \hat{\eta} + \hat{\theta} - \frac{1}{V} \hat{\eta}\end{aligned}$$

setting equal to zero and solving, the CAVI updates are given by

$$\begin{aligned}\hat{\theta}^{(t+1)} &= Y + \hat{\eta}^{(t)} \\ \hat{\eta}^{(t+1)} &= \frac{(Y + \hat{\theta}^{(t+1)})V}{1 + V}\end{aligned}$$