Problem Set #3: Deep Learning & Unsupervised Learning

Problem 1 A simple neural network

Let $X = \{x^{(1)}, x^{(2)}, \dots, x^{(m)}\}$ be dataset of m examples with 2 features. That is, $x^{(i)} \in \mathbb{R}^2$. Samples are classified into 2 categorie with labels $y \in \{0, 1\}$, as shown in Figure 1. Want to perform binary classification using a simple neural networks with the architecture shown in Figure 2.

Two features x_1 and x_2 , the three neurons in the hidden layer h_1 , h_2 , h_3 , and the output neuron as o. Weight from x_i to h_j be $w_{i,j}^{[1]}$ for i = 1, 2 and j = 1, 2, 3, and weight from h_j to o be $w_j^{[2]}$. Finally, denote intercept weight for h_j as $w_{0,j}^{[1]}$ and the intercept weight for o as $w_0^{[2]}$. Use average squared loss instead of the usual negative log-likelihood:

$$l = \frac{1}{m} \sum_{i=1}^{m} (o^{(i)} - y^{(i)})^{2}.$$

(a) Suppose we use sigmoid function as activation function for h_1 , h_2 , h_3 , and o. We have

$$h_1 = g(w_1^{[1]}x), \quad h_2 = g(w_2^{[1]}x), \quad h_3 = g(w_3^{[1]}x), \quad o = g(w_3^{[2]}h).$$

Hence,

$$\frac{\partial l}{\partial w_{1,2}^{[1]}} = \frac{1}{m} \sum_{i=1}^{m} 2(o^{(i)} - y^{(i)}) o^{(i)} (1 - o^{(i)}) w_2^{[2]} h_2^{(i)} (1 - h_2^{(i)}) x_1^{(i)},$$

where $h_2^{(i)} = g(w_{0,2}^{[1]} + w_{1,2}^{[1]} x_1^{(i)} + w_{2,2}^{[1]} x_2^{(i)})$ and g is the sigmoid function. Therefore, the gradient descent update to $w_{1,2}^{[1]}$, assuming learning rate α is

$$w_{1,2}^{[1]} := w_{1,2}^{[1]} - \frac{2\alpha}{m} \sum_{i=1}^{m} (o^{(i)} - y^{(i)}) o^{(i)} (1 - o^{(i)}) w_2^{[2]} h_2^{(i)} (1 - h_2^{(i)}) x_1^{(i)}$$

where $h_2^{(i)} = g(w_{0,2}^{[1]} + w_{1,2}^{[1]}x_1^{(i)} + w_{2,2}^{[1]}x_2^{(i)}).$

(b) Now, suppose the activation function for h_1 , h_2 , h_3 , and o is the step function f(x), defined as

$$f(x) = \begin{cases} 1, & (x \ge 0), \\ 0, & (x < 0). \end{cases}$$

Is it possible to have a set of weights that allow the neural network to classify this dataset with 100% accuracy? If so, provide a set of weights by completing optimal_step_weights wihin src/p01_nn.py and explain your reasoning for those weights. If not, please explain the reasoning.

There is a set of weights that allow the neural network to classify this dataset with 100% accuracy. For the step function activation, we have

$$h_{1} = f(w_{1}^{[1]}x) = f(w_{0,1}^{[1]} + w_{1,1}^{[1]}x_{1} + w_{2,1}^{[1]}x_{2})$$

$$h_{2} = f(w_{2}^{[1]}x) = f(w_{0,2}^{[1]} + w_{1,2}^{[1]}x_{1} + w_{2,3}^{[1]}x_{2})$$

$$h_{3} = f(w_{3}^{[1]}x) = f(w_{0,3}^{[1]} + w_{1,3}^{[1]}x_{1} + w_{2,3}^{[1]}x_{2})$$

$$o = f(w_{0,3}^{[2]} + w_{1,3}^{[2]}x_{1} + w_{2,3}^{[2]}x_{2})$$

$$o = f(w_{0,3}^{[2]}x_{1} + w_{1,3}^{[2]}x_{1} + w_{2,3}^{[2]}x_{2})$$

Notice from Figure 1 that the label $y^{(i)} = 0$ if and only if $x^{(i)}$ satisfies

$$\begin{cases} x_2^{(i)} > 0.5, \\ x_1^{(i)} > 0.5, \\ x_1^{(i)} + x_2^{(i)} < 4. \end{cases}$$

Now, let

$$w_1^{[1]} = \begin{bmatrix} 0.5 \\ 0 \\ -1 \end{bmatrix}, \quad w_2^{[1]} = \begin{bmatrix} 0.5 \\ -1 \\ 0 \end{bmatrix}, \quad w_3^{[1]} = \begin{bmatrix} -4 \\ 1 \\ 1 \end{bmatrix}, \quad w_1^{[1]} = \begin{bmatrix} -0.5 \\ 1 \\ 1 \\ 1 \end{bmatrix}.$$

This set of weights will capture all the conditions and allow the nerual network to classify this dataset with 100% accuracy.

(c) Let the activation function for h_1 , h_2 , h_3 , and o is the linear function f(x) = x, and the activation function for o be the same step function as before. Is it possible to have a set of weights that allow the neural network to classify this dataset with 100% accuracy? If so, provide a set of weights by completing optimal_linear_weights within src/p01_nn.py and explain your reasoning for those weights. If not, please explain the reasoning.

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Problem 2 KL divergence and maximum likelihood

Kullback-Leibler (KL) divergence is a measure of how much one probability distribution is different from a second one. The KL divergence between two discrete-valued distribution P(X), Q(X) over the outcome space \mathcal{X} is defined as follows:

$$D_{\mathrm{KL}}(P \parallel Q) = \sum_{x \in \mathcal{X}} P(x) \log \frac{P(x)}{Q(x)}.$$

Assuem P(x) > 0 for all x. (One other standard thing to do is adopt the convention that $0 \log 0 = 0$.) Sometimes, we also write the KL divergence more explicitly as $D_{\text{KL}}(P \parallel Q) = D_{\text{KL}}(P(X) \parallel Q(X))$.

Background on Information Theory

The *entropy* of a probability distribution P(X), defined as

$$H(P) = -\sum_{x \in \mathcal{X}} P(x) \log P(x).$$

measures how dispersed a probability distribution is. Notably, $\mathcal{N}(\mu, \sigma^2)$ has the highest entropy among all possible continuous distribution that has mean μ and variance σ^2 . The entropy H(P) is the best possible long term average bits per message (optimal) that can be achieved under probability distribution P(X).

The *cross entropy* is defined as

$$H(P,Q) = -\sum_{x \in \mathcal{X}} P(x) \log Q(x).$$

The cross entropy H(P,Q) is the long term average bits per message (suboptimal) that results under a distribution P(X), by reusing an encoding scheme designed to be optimal for a scenario with probability distribution Q(X).

Notice that

$$D_{\mathrm{KL}}(P \parallel Q) = \sum_{x \in \mathcal{X}} P(x) \log P(x) - \sum_{x \in \mathcal{X}} P(x) \log Q(x) = H(P, Q) - H(P).$$

If H(P,Q)=0, then it necessarily means P=Q. In ML, it is common task to find distribution Q that is close to another distribution P. To achieve this, we optimize $D_{\mathrm{KL}}(P\parallel Q)$. Later we will see that Maximum Likelihood Estimation turns out to be equivalent minimizing KL divergence between the training data and the model.

(a) **Nonnegativity.** Prove that

$$D_{\mathrm{KL}}(P \parallel Q) \ge 0$$

and $D_{KL}(P \parallel Q) = 0$ if an only if P = Q.

Hint: Use Jensen's inequality.

Proof. By definition,

$$D_{\mathrm{KL}}(P \parallel Q) = \sum_{x \in \mathcal{X}} P(x) \log \frac{P(x)}{Q(x)} = -\sum_{x \in \mathcal{X}} P(x) \log \frac{Q(x)}{P(x)}.$$

Since $-\log x$ is strictly convex, by Jensen's inequality, we have

$$D_{\mathrm{KL}}(P \parallel Q) = -\sum_{x \in \mathcal{X}} P(x) \log \frac{Q(x)}{P(x)} \ge -\log \sum_{x \in \mathcal{X}} P(x) \frac{Q(x)}{P(x)} = 0.$$

When the equality holds,

$$\log \frac{Q(x)}{P(x)} = 0$$

with probability 1. That is, Q = P with probability 1. This completes the proof.

(b) Chain rule for KL divergence. The KL divergence between 2 conditional distributions $P(X \mid Y)$, $Q(X \mid Y)$ is defined as follows:

$$D_{\mathrm{KL}}(P(X\mid Y)\parallel Q(X\mid Y)) = \sum_{y} P(y) \left(\sum_{x} P(x\mid y) \log \frac{P(x\mid y)}{Q(x\mid y)}\right).$$

This can be thought of as the expected KL divergence between the corresponding conditional distributions on x. That is, between $P(X \mid Y = y)$ and $Q(X \mid Y = y)$, where the expectation is taken over the random y.

Prove the following chain rule for KL divergence:

$$D_{\mathrm{KL}}(P(X,Y) \parallel Q(X,Y)) = D_{\mathrm{KL}}(P(X) \parallel Q(X)) + D_{\mathrm{KL}}(P(Y \mid X) \parallel Q(Y \mid X)).$$

Proof.

LHS =
$$\sum_{x} \sum_{y} P(x, y) \log \frac{P(x, y)}{Q(x, y)}$$

= $\sum_{x} \sum_{y} P(y \mid x) P(x) \left[\log \frac{P(y \mid x)}{Q(y \mid x)} + \log \frac{P(x)}{Q(x)} \right]$
= $\sum_{x} \sum_{y} P(y \mid x) P(x) \log \frac{P(y \mid x)}{Q(y \mid x)} + \sum_{x} P(x) \log \frac{P(x)}{Q(x)} \sum_{y} P(y \mid x)$
= $\sum_{x} \sum_{y} P(y \mid x) P(x) \log \frac{P(y \mid x)}{Q(y \mid x)} + \sum_{x} P(x) \log \frac{P(x)}{Q(x)}$
= $D_{KL}(P(X) \parallel Q(X)) + D_{KL}(P(Y \mid X) \parallel Q(Y \mid X))$
= RHS.

(c) **KL** and maximum likelihood. Consider density estimation problem and suppose we are given training set $\{x^{(i)}\}_{i=1}^m$. Let the empirical distribution be $\hat{P}(x) = \frac{1}{m} \sum_{i=1}^m \mathbb{I}\{x^{(i)} = x\}$. (\hat{P} is just the uniform distribution over the training set; i.e., sampling from the empirical distribution is the same as picking a random example from the training set.)

Suppose we have a family of distributions P_{θ} parametrized by θ . Prove that finding the maximum likelihood estimates for the parameter θ is equivalent to finding P_{θ} with minimal KL divergence from \hat{P} . That is, prove that

$$\underset{\theta}{\operatorname{argmin}} D_{\mathrm{KL}}(\hat{P} \parallel P_{\theta}) = \underset{\theta}{\operatorname{argmax}} \sum_{i=1}^{m} \log P_{\theta}(x^{(i)}).$$

Remark: Consider the relationship between parts (b-c) and multi-variate Bernoulli Naive bayes parameter estimation. In NB model we assumed P_{θ} is the following form: $P_{\theta}(x, y) = p(y) \prod_{i=1}^{n} p(x_i \mid y)$. By the chain rule for KL divergence, we therefore have

$$D_{\mathrm{KL}}(\hat{P} \parallel P_{\theta}) = D_{\mathrm{KL}}(\hat{P}(y) \parallel p(y)) + \sum_{i=1}^{n} D_{\mathrm{KL}}(\hat{P}(x_i \mid y) \parallel p(x_i \mid y)).$$

This shows that finding the maximum likelihood/minimum KL divergence estimates of the parameters decomposes into 2n + 1 independent optimization problems: One for the class priors p(y), and one for each conditional distributions $p(x_i \mid y)$ for each feature x_i given each of the two possible labels for y. Specifically, finding the maximum likelihood estimates for each of these problems individually results in also maximizing the likelihood of the joint distribution. This similarly applies to bayesian networks.