Problem Set #1: Supervised Learning

Problem 1 Linear Classifiers (Logistic Regression and GDA)

Consider two datasets provided in the following files:

- i. data/ds1_{train,valid},csv
- ii. data/ds2_{train,valid},csv

Each file contains m examples, one example per row. The i-th row contains columns $x_0^{(i)} \in \mathbb{R}$, $x_1^{(i)} \in \mathbb{R}$ and $y^{(i)} \in \{0,1\}$. Use logistic regression and GDA to perform binary classification.

(a) Average empirical loss for logistic regression:

$$J(\theta) = -\frac{1}{m} \sum_{i=1}^{m} y^{(i)} \log(h_{\theta}(x^{(i)})) + (1 - y^{(i)}) \log(1 - h_{\theta}(x^{(i)})),$$

where $y^{(i)} \in \{0, 1\}, h_{\theta}(x^{(i)}) = g(\theta^T x)$ and $g(z) = 1/(1 + e^{-z})$.

The gradient of the function

$$\frac{\partial J}{\partial \theta_j} = -\frac{1}{m} \sum_{i=1}^{m} (y^{(i)} - h_{\theta}(x^{(i)})) x_j^{(i)}.$$

It follows that

$$\frac{\partial^2 J}{\partial \theta_k \partial \theta_j} = \frac{1}{m} \sum_{i=1}^m h_{\theta}(x^{(i)}) (1 - h_{\theta}(x^{(i)})) x_k^{(i)} x_j^{(i)}.$$

Hence, The Hessian H of this function is

$$H = \frac{1}{m} \sum_{i=1}^{m} h_{\theta}(x^{(i)}) (1 - h_{\theta}(x^{(i)})) x^{(i)} (x^{(i)})^{T}.$$

Now, for any vector z, using Einstein's summation, we have

$$z^{T}Hz = \frac{1}{m} \sum_{i=1}^{m} h_{\theta}(x^{(i)}) (1 - h_{\theta}(x^{(i)})) z_{k} x_{k}^{(i)} x_{j}^{(i)} z_{j}$$
$$= \frac{1}{m} \sum_{i=1}^{m} h_{\theta}(x^{(i)}) (1 - h_{\theta}(x^{(i)})) (x^{T}z)^{2}$$
$$\geq 0$$

This shows that H is PSD, and J is convex.

- (b) Coding problem.
- (c) To show that GDA results in a classifier that has a linear decision boundary, we want to show

$$p(y = 1 \mid x; \phi, \mu_0, \mu_1, \Sigma) = \frac{1}{1 + \exp(-(\theta^T x + \theta_0))}$$

for some $\theta \in \mathbb{R}^n$ and $\theta_0 \in \mathbb{R}$ as functions of ϕ , Σ , μ_0 , and μ_1 . We have

$$p(y=1 \mid x) = \frac{p(x \mid y=1)p(y=1)}{p(x \mid y=1)p(y=1) + p(x \mid y=0)p(y=0)}$$

$$= \frac{\phi \exp\left(-\frac{1}{2}(x-\mu_1)^T \Sigma^{-1}(x-\mu_1)\right)}{\phi \exp\left(-\frac{1}{2}(x-\mu_1)^T \Sigma^{-1}(x-\mu_1)\right) + (1-\phi) \exp\left(-\frac{1}{2}(x-\mu_0)^T \Sigma^{-1}(x-\mu_0)\right)}$$

$$= \frac{1}{1 + \frac{1-\phi}{\phi} \exp\left(-\frac{1}{2}(x-\mu_0)^T \Sigma^{-1}(x-\mu_0) + \frac{1}{2}(x-\mu_1)^T \Sigma^{-1}(x-\mu_1)\right)}$$

$$= \frac{1}{1 + \frac{1-\phi}{\phi} \exp\left(-((\mu_1 - \mu_0)^T \Sigma^{-1}x + \frac{1}{2}(\mu_0^T \Sigma^{-1}\mu_0 - \mu_1^T \Sigma^{-1}\mu_1))\right)}.$$

This is the desired form, where

$$\theta = \Sigma^{-1}(\mu_1 - \mu_0),$$

$$\theta_0 = \frac{1}{2}(\mu_0^T \Sigma^{-1} \mu_0 - \mu_1^T \Sigma^{-1} \mu_1) - \log \frac{1 - \phi}{\phi}.$$

(d) The log-likelihood of the data is

$$\ell(\phi, \mu_0, \mu_1, \Sigma) = \log \prod_{i=1}^{m} p(x^{(i)} \mid y^{(i)}; \mu_0, \mu_1, \Sigma) p(y^{(i)}; \phi)$$

$$= \sum_{i=1}^{m} 1\{y^{(i)} = 1\} \left(-\frac{1}{2} (x^{(i)} - \mu_1)^T \Sigma^{-1} (x^{(i)} - \mu_1) + \log \phi \right)$$

$$+ \sum_{i=1}^{m} 1\{y^{(i)} = 0\} \left(-\frac{1}{2} (x^{(i)} - \mu_0)^T \Sigma^{-1} (x^{(i)} - \mu_0) + \log(1 - \phi) \right)$$

$$- \frac{m}{2} \log |\Sigma| + C,$$

where C is some constant independent of the parameters.

Let $\nabla_{\phi} \ell = 0$, we have

$$\phi = \frac{1}{m} \sum_{i=1}^{m} 1\{y^{(i)} = 1\}.$$

Let $\nabla_{\mu_1} \ell = 0$, we have

$$\sum_{i=1}^{m} 1\{y^{(i)} = 1\} \Sigma^{-1} x^{(i)} = \sum_{i=1}^{m} 1\{y^{(i)} = 1\} \Sigma^{-1} \mu_1,$$

and thus

$$\mu_1 = \frac{\sum_{i=1}^m 1\{y^{(i)} = 1\}x^{(i)}}{\sum_{i=1}^m 1\{y^{(i)} = 1\}}, \quad \mu_0 = \frac{\sum_{i=1}^m 1\{y^{(i)} = 0\}x^{(i)}}{\sum_{i=1}^m 1\{y^{(i)} = 0\}}.$$

To derive Σ , recall that $\nabla_A \log |A| = (A^{-1})^T$, so we have

$$\nabla_{\Sigma^{-1}} \ell = -\frac{m}{2} \Sigma^{-1} + \frac{1}{2} \sum_{i=1}^{m} (x^{(i)} - \mu_{y^{(i)}}) (x^{(i)} - \mu_{y^{(i)}})^{T}.$$

Hence,

$$\Sigma = \frac{1}{m} \sum_{i=1}^{m} (x^{(i)} - \mu_{y^{(i)}}) (x^{(i)} - \mu_{y^{(i)}})^{T}.$$

We conclude that the maximum likelihood estimates of the parameters are given by

$$\phi = \frac{1}{m} \sum_{i=1}^{m} 1\{y^{(i)} = 1\},$$

$$\mu_0 = \frac{\sum_{i=1}^{m} 1\{y^{(i)} = 0\}x^{(i)}}{\sum_{i=1}^{m} 1\{y^{(i)} = 0\}},$$

$$\mu_1 = \frac{\sum_{i=1}^{m} 1\{y^{(i)} = 1\}x^{(i)}}{\sum_{i=1}^{m} 1\{y^{(i)} = 1\}},$$

$$\Sigma = \frac{1}{m} \sum_{i=1}^{m} (x^{(i)} - \mu_{y^{(i)}})(x^{(i)} - \mu_{iy})^T.$$

- (e) Coding problem.
- (f) See jupyter notebook for plots.
- (g) See jupyter notebook for plots. On Dataset 1 GDA perform worse than logistic regression. This might be the case because for Dataset 1, the distribution of features are not quite multivariate normal.
- (h) *** TO-DO ***

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Problem 2 Incomplete, Positive-Only Labels

Dataset without full access to labels. In particular, we have labels only for a subset of positive examples. All negative examples and the rest of positive examples are unlabeled. Assume dataset $\{(x^{(i)}, t^{(i)}, y^{(i)})\}_{i=1}^m$ where $t^{(i)} \in \{0, 1\}$ is true label and where

$$y^{(i)} = \begin{cases} 1 & x^{(i)} \text{ is labeled} \\ 0 & \text{otherwise.} \end{cases}$$

All labeled examples are positive, which is to say $p(t^{(i)} = 1 \mid y^{(i)} = 1) = 1$. Goal is to construct a binary classifier h of true label t which only access to partial labels y. That is, construct h such that $h(x^{(i)}) \approx p(t^{(i)} = 1 \mid x^{(i)})$ as closely as possible, using only x and y.

(a) Suppose each $y^{(i)}$ and $x^{(i)}$ conditionally independent given $t^{(i)}$:

$$p(y^{(i)} = 1 \mid t^{(i)} = 1, x^{(i)}) = p(y^{(i)} = 1 \mid t^{(i)} = 1).$$

That is, labeled examples are selected uniformly at random from positive examples.

Want to show $p(t^{(i)} = 1 \mid x^{(i)}) = p(y^{(i)} = 1 \mid x^{(i)})/\alpha$ for some $\alpha \in \mathbb{R}$. As $p(\cdot \mid x^{(i)})$ is a conditional measure, we have

$$\begin{split} p(y^{(i)} = 1 \mid x^{(i)}) &= p(y^{(i)} = 1 \mid t^{(i)} = 1, x^{(i)}) p(t^{(i)} = 1 \mid x^{(i)}) \\ &+ p(y^{(i)} = 1 \mid t^{(i)} = 0, x^{(i)}) p(t^{(i)} = 0 \mid x^{(i)}) \\ &= p(y^{(i)} = 1 \mid t^{(i)} = 1, x^{(i)}) p(t^{(i)} = 1 \mid x^{(i)}) \\ &= p(y^{(i)} = 1 \mid t^{(i)} = 1) p(t^{(i)} = 1 \mid x^{(i)}). \end{split}$$

Hence, $p(t^{(i)} = 1 \mid x^{(i)}) = p(y^{(i)} = 1 \mid x^{(i)})/\alpha$ where $\alpha = p(y^{(i)} = 1 \mid t^{(i)} = 1)$.

(b) Estimate α using a trained classifier h and a held-out validation set V. Let $V_+ = \{x^{(i)} \in V \mid y^{(i)} = 1\}$. Assuming $h(x^{(i)}) \approx p(y^{(i)} = 1 \mid x^{(i)})$ for all $x^{(i)}$. Want to show

$$h(x^{(i)}) \approx \alpha \text{ for all } x^{(i)} \in V_+.$$

May assume that $p(t^{(i)} = 1 \mid x^{(i)}) \approx 1$ when $x^{(i)} \in V_+$.

We have

$$h(x^{(i)}) \approx p(y^{(i)} = 1 \mid x^{(i)})$$

= $p(y^{(i)} = 1 \mid t^{(i)} = 1, x^{(i)})p(t^{(i)} = 1 \mid x^{(i)})$
 $\approx \alpha.$

- (c) Coding problem.
- (d) Coding problem.

(e) Coding problem. Estimate the constant α using validation set.

$$\alpha \approx \frac{1}{|V_+|} \sum_{x^{(i)} \in V_+} h(x^{(i)}).$$

To plot the decision boundary, we need to calculate the rescaled θ , write θ_* . The new decision boundary is given by $\frac{1}{\alpha} \frac{1}{1 + \exp(-\theta^T x)} = \frac{1}{2}$. We have

$$\theta^T x + \log\left(\frac{2}{\alpha} - 1\right) = 0.$$

This is equivalent to $\theta_*^T x = 0$. This shows that θ_* and θ differs only in the 0-th index by a constant $\log(\frac{2}{\alpha} - 1)$.

Problem 3 Poisson Regression

(a) The poisson distribution parametrized by λ is

$$p(y;\lambda) = \frac{e^{-\lambda}\lambda^y}{y!}.$$

Therefore, we have

$$p(y; \lambda) = \frac{1}{y!} \exp(-\lambda + y \log \lambda).$$

Compare with $p(y; \eta) = b(y) \exp(\eta^T T(y) - a(\eta))$, we conclude that the poisson distribution is in the exponential family, with

$$b(y) = \frac{1}{y!},$$

$$T(y) = y,$$

$$\eta = \log \lambda,$$

$$a(\eta) = e^{\eta}.$$

(b) The canonical response function for the family

$$\mathbb{E}[T(y); \eta] = \mathbb{E}[T(y); \eta] = \lambda = e^{\eta}.$$

(c) For a general linear model and a training set, the log likelihood

$$\log p(y^{(i)} \mid x^{(i)}; \eta) = \log b(y) \exp(\eta^T T(y) - a(\eta))$$

= \log b(y) + \eta^T T(y) - a(\eta).

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For our model with poisson responses y, we have

$$\ell = \log p(y^{(i)} \mid x^{(i)}; \theta) = -\log y! + (\theta^T x^{(i)}) y^{(i)} - \exp(\theta^T x^{(i)}).$$

Taking the derivative with respect to θ_j , we have

$$\frac{\partial \ell}{\partial \theta_j} = (y^{(i)} - \exp(\theta^T x^{(i)})) x_j^{(i)}$$

Hence, the stochastic gradient ascent update rule for learning using a GLM model with poisson response y is

$$\theta_j := \theta_j + \alpha \frac{\partial \ell}{\partial \theta_j}$$

$$:= \theta_j + \alpha (y^{(i)} - \exp(\theta^T x^{(i)})) x_j^{(i)}.$$

(d) Coding problem. To predict the dataset, recall that the hypothese function for our model with poisson response y is

$$h_{\theta}(x) = \mathbb{E}[y \mid x] = e^{\eta} = e^{\theta^T x}.$$

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