

# Measuring Severity in Statistical Inference

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## 1 Introduction: *modus tollens* in statistical inference.

Statistical inference mimics the logical form

$$x \text{ (data)} \rightarrow \Theta \text{ (claim)}.$$

The degree of warrant enjoyed by the claim depends on the data and the strength of the “ $\rightarrow$ ” relation.

**Severity** [3, 2] measures the strength of its contrapositive:

$$\neg\Theta \rightarrow \neg x.$$

Large severity affirms the support of  $x$  for  $\Theta$ .

## 2 Quantifying severity in statistical inference.

Let  $x \in \mathcal{X}$  be the observable data,  $\mathbf{C} = \{C_1, C_0\}$  be a pair of *critical regions* that partition  $\mathcal{X}$ , and  $\{\Theta_0, \Theta_1\}$  be a pair of *inferential conclusions* that partition  $\Theta$ .

**Definition 2.1 (inferential claim).** The function  $T(\cdot; \mathbf{C}) : \mathcal{X} \rightarrow \{\Theta_0, \Theta_1\}$  is a binary *inferential claim* based on  $x$ , where

$$T(x; \mathbf{C}) = \begin{cases} \Theta_1 & \text{if } x \in C_1 \\ \Theta_0 & \text{if } x \in C_0. \end{cases} \quad (2.1)$$

**Definition 2.2 (inferential strategy).** Suppose for every  $a \in \mathcal{A}$ ,  $\mathbf{C}_a = \{C_1(a), C_0(a)\}$  is a partition of  $\mathcal{X}$  such that for  $a, a' \in \mathcal{A}$  where  $a < a'$ ,  $C_1(a) \subseteq C_1(a')$ . Then, the collection of inferential claims

$$\mathcal{T} = \{T(\cdot; \mathbf{C}_a) : \mathcal{X} \rightarrow \{\Theta_0, \Theta_1\}, a \in \mathcal{A}\} \quad (2.2)$$

is called an *inferential strategy*.

An inferential strategy provides a sense of *informativeness* conveyed by the data. A data value is more informative about a conclusion if more inferential claims from the same strategy arrive at that conclusion with that data.

**Definition 2.3 (data informativeness).** Let  $x, x' \in \mathcal{X}$ , and  $\mathcal{T}$  be an inferential strategy whose claims are indexed by  $\mathcal{A}$ . Say that  $x$  is *more informative* than  $x'$  about  $\Theta_i$  with respect to  $\mathcal{T}$ , if there exists some  $a \in \mathcal{A}$  such that  $T_a(x) = \Theta_i$  and  $T_a(x') = \bar{\Theta}_i$ , for  $i = 0, 1$ .

**Corollary 2.1.** If  $x$  is more informative than  $x'$  about  $\Theta_i$  w.r.t  $\mathcal{T}$ , then  $x'$  is more informative than  $x$  about  $\bar{\Theta}_i$ .

**Definition 2.4 (severity).** The *severity* of an assertion  $\Theta \subseteq \Theta$ , as substantiated by an inferential claim  $T$  (Def. 2.1) derived from the inferential strategy  $\mathcal{T}$  (Def. 2.2) and based on data  $x$ , is

$$S(\Theta; T = \Theta_i) = \inf_{\theta \notin \Theta} P_\theta(X \notin C_i^x), \quad i = 0, 1, \quad (2.3)$$

where  $X \sim P_\theta$ ,  $\theta \in \Theta$ , and

$$C_1^x = \cap_{a: T(x, \mathbf{C}_a) = \Theta_1} C_1(a), \quad (2.4)$$

$$C_0^x = \bar{C}_1^x \quad (2.5)$$

are the respective *attained* critical regions.

**Remark 2.1.** *Modus tollens*:

- $X \notin C_i^x$ : hypothetical or future realization of  $X$  results in an inferential claim that is the opposite of the one drawn from the current evidence  $x$ ;
- $\theta \notin \Theta$ : as the parameter varies in the range that is complement to  $\Theta$ .

**Remark 2.2.**  $X \sim P_\theta$ :

- (Frequentist) sampling distribution indexed by  $\theta \in \Theta$ ;
- (Bayes) conditional prior or posterior predictive distribution marginalized over nuisance parameters.

### 3 Severity: properties.

A measure of evidential support should satisfy (see [4]):

- *Coherence*: the larger the hypothesis is, the more support there is;
- *Informativeness*: the farther into the hypothesis the data are, the more support there is.

**Theorem 3.1 (coherence).** For a pair of assertions  $\Theta, \tilde{\Theta}$  satisfying  $\Theta \subseteq \tilde{\Theta}$ ,

$$S(\Theta; T = \Theta_i) \leq S(\tilde{\Theta}; T = \Theta_i) \quad (3.1)$$

for all  $\Theta \in \Theta$  and  $i \in \{0, 1\}$ .

**Theorem 3.2 (informativeness).** If the data  $x$  is more informative than  $x'$  about  $\Theta_i$  with respect to  $\mathcal{T}$  for some  $i \in \{0, 1\}$ , then for all  $\Theta \in \Theta$ ,

$$S(\Theta; T(x) = \Theta_i) \geq S(\Theta; T(x') = \Theta_i).$$

The severity enjoyed by an assertion  $\Theta$  depends on the design of the inferential strategy  $\mathcal{T}$ , rather than the specific inferential claim derived from it.

**Theorem 3.3 (strategic equivalence).** For fixed data  $x \in \mathcal{X}$ , let  $T_a$  and  $T_{a'}$  be two inferential claims derived from the inferential strategy  $\mathcal{T}$  such that  $T_a(x) = T_{a'}(x) = \Theta_i$ , where  $i \in \{0, 1\}$ . Then for all  $\Theta \in \Theta$ ,

$$S(\Theta; T_a(x) = \Theta_i) = S(\Theta; T_{a'}(x) = \Theta_i). \quad (3.2)$$

**Theorem 3.4 (subadditivity).** For fixed data  $x \in \mathcal{X}$ , let  $T_a$  and  $T_{a'}$  be two inferential claims derived from the inferential strategy  $\mathcal{T}$  such that  $T_a(x) = \Theta_i$  and  $T_{a'}(x) = \bar{\Theta}_i$ , where  $i \in \{0, 1\}$ . Then for all  $\Theta \in \Theta$ ,

$$S(\Theta; T_a(x) = \Theta_i) + S(\Theta; T_{a'}(x) = \bar{\Theta}_i) \leq 1. \quad (3.3)$$

The equality attains if  $\Theta$  is a singleton.

### 4 Severity in binary classification.

A classifier  $T$  aims to discern between the following two assertions:

$\Theta_-$ : patient  $x$  is healthy, vs.  $\Theta_+$ : patient  $x$  is ill.

The severity of the ill/positive diagnosis is ( $\mathcal{X} = \Theta = \{-, +\}$  in this case):

$$S(\Theta_+; T(x) = \Theta_+) = \inf_{x \in \Theta_-} P_x(X \in \Theta_-) = \Pr(X \in \Theta_- | x \in \Theta_-) = \text{Specificity}(T).$$

On the other hand, the severity of the healthy/negative diagnosis is

$$S(\Theta_-; T(x) = \Theta_-) = \inf_{x \in \Theta_+} P_x(X \in \Theta_+) = \Pr(X \in \Theta_+ | x \in \Theta_+) = \text{Sensitivity}(T).$$

### 5 Severity in frequentist inference.

Consider a family of tests for the pair of null and alternative hypotheses  $\Theta_0$  and  $\Theta_1$ , characterized by the collection of rejection regions  $\{R(\alpha)\}$ ,  $\alpha \in (0, 1)$ . The attained power function is

$$\beta(\theta; x) = P_\theta(X \in R^x) \quad (5.1)$$

where  $R^x = \cap_{\alpha: x \in R(\alpha)} R(\alpha)$  is the attained rejection region, and  $p_x = \sup_{\theta \in \Theta_0} \beta(\theta; x)$  is the  $p$ -value of the test.

**Corollary 5.1 (severity and  $p$ -value).** The inferential strategy  $\mathcal{T}$ , characterized by critical regions that are rejection regions of a family of hypothesis tests:  $C_1(\alpha) = R(\alpha)$ , satisfies that for every  $T \in \mathcal{T}$ ,

$$S(\Theta_1; T(x) = \Theta_1) = 1 - p_x,$$

Furthermore, if the attained power function  $\beta(\theta; x)$  satisfies  $\sup_{\theta \in \Theta_0} \beta(\theta; x) = \inf_{\theta \in \Theta_1} \beta(\theta; x)$  for  $x \in \mathcal{X}$ , then

$$S(\Theta_0; T(x) = \Theta_0) = p_x.$$

**Corollary 5.2 (severity and the attained power function).** Suppose  $\Theta = \mathbb{R}$ , and consider one-sided assertions of interest  $\tilde{\Theta}_0 = (-\infty, \tilde{\theta}]$  versus  $\tilde{\Theta}_1 = (\tilde{\theta}, \infty)$ . If the attained power function  $\beta(\theta; x)$  is continuous and monotone increasing in  $\theta$  for every  $x$ , we have that

$$S(\tilde{\Theta}_1; T = \Theta_1) = 1 - \beta(\tilde{\theta}; x),$$

and

$$S(\tilde{\Theta}_0; T = \Theta_0) = \beta(\tilde{\theta}; x).$$

**Remark 5.1.** Under the assumption of Corollary 5.2, the attained power function  $\beta(\theta; x)$  is a **significance function** [1] (also called the  $p$ -value function), which is also a **confidence distribution** [5].

## 6 Severity in Bayesian inference: an illustration.

Suppose the sampling distribution is  $X \sim \text{Bin}(n, \theta)$ . Consider

$$\Theta_0: \theta \sim \delta_{(\theta=\frac{1}{2})} \quad \text{versus} \quad \Theta_1: \theta \sim \text{Unif}(0,1).$$

Notice that both  $\Theta_0$  and  $\Theta_1$  are *singleton* sets. This is the setting known as the Lindley's paradox.

The marginal likelihood for the data under the two respective assertions are

$$P(X=x|\Theta_0) = \binom{n}{x} \left(\frac{1}{2}\right)^n \quad \text{and} \quad P(X=x|\Theta_1) = \int \binom{n}{x} \theta^x (1-\theta)^{n-x} d\theta = \frac{1}{n+1}.$$

Under equal prior probabilities,  $P(\Theta_0) = P(\Theta_1) = 0.5$ , the Bayes factor

$$BF(x;n) = \frac{P(x|\Theta_0)}{P(x|\Theta_1)} = (n+1) \binom{n}{x} \left(\frac{1}{2}\right)^n. \quad (6.1)$$

**The Bayesian inferential strategy.** The Bayesian inferential strategy is to endorse  $\Theta_1$  if the attained Bayes factor  $BF(x;n)$  falls below a certain pre-determined threshold. This creates critical regions of the form

$$C_1(a) = \left\{ x \in [n] : \left| x - \frac{n}{2} \right| > a \right\} \quad (6.2)$$

for some  $a \geq 0$ , and  $C_0(a) = \bar{C}_1(a)$ . For illustration, suppose  $n=100$  and  $x=35$ . The attained critical region for the inferential strategy is

$$C_1^{35} = \{x < 36 \text{ or } x > 64\},$$

and  $C_0^{35} = \bar{C}_1^{35}$ .

**Case 1: A ten-fold Bayes factor.** Adopt a specific inferential claim that a Bayes factor less than **0.1** is considered strong evidence towards  $\Theta_1$ . This effectively sets  $a = 14$  in (6.2). The attained Bayes factor  $BF(x=35;n=100) = 8.72 \times 10^{-2}$ , and the inferential claim results in the conclusion  $T_{14}(35) = \Theta_1$ , with severity

$$S(\Theta_1; T_{14}(35) = \Theta_1) = \inf_{\theta \notin \Theta_1} P_\theta(X \notin C_1^{35}) = \Pr\left(36 \leq X \leq 64 \mid \theta = \frac{1}{2}\right) \doteq 99.65\%.$$

On the other hand, the severity for the unendorsed claim  $\Theta_0$  is

$$S(\Theta_0; T_{14}(35) = \Theta_1) = \inf_{\theta \notin \Theta_0} P_\theta(X \notin C_1^{35}) = \sum_{x=36}^{64} \int \binom{n}{x} \theta^x (1-\theta)^{n-x} d\theta = \frac{29}{101}.$$

**Case 2: A hundred-fold Bayes factor.** Adopt a more stringent inferential claim, that only a Bayes factor less than **0.01** would be considered strong evidence towards  $\Theta_1$ . This is equivalent to setting  $a = 18$ . The inferential claim concludes  $T_{18}(x=35) = \Theta_0$ , with severity

$$S(\Theta_0; T_{18}(35) = \Theta_0) = \inf_{\theta \notin \Theta_0} P_\theta(X \notin C_0^{35}) = 1 - \sum_{x=36}^{64} \int \binom{n}{x} \theta^x (1-\theta)^{n-x} d\theta = \frac{82}{101},$$

whereas the severity for the unendorsed claim  $\Theta_1$  is

$$S(\Theta_1; T_{18}(35) = \Theta_0) = \inf_{\theta \notin \Theta_1} P_\theta(X \notin C_0^{35}) = 1 - P\left(36 \leq X \leq 64 \mid \theta = \frac{1}{2}\right) \doteq 0.35\%.$$

## References

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