

### Lecture 2

Probability refresher (or introduction)

WILD6900 (Spring 2019)

Processing math: 100%

# Readings:

Hobbs & Hooten 29-70

Warning: The material presented in this lecture is tedious. But the concepts in this lecture are critical to everything that will follow in this course. So push through and try your best to understand these topics. You do not need to be an expert in probability at the end of this lecture - we will reinforce these concepts over and over again throughout the semester - but getting the gist now will help you grasp other topics as we move forward

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- a stochastic model [a | b, c]

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Stochastic processes are different:

- given an input, the model will not always return the same answer
- the output of stochastic processes are uncertain
- Even though stochastic processes are inherently uncertain, they are not unpredictable.

In Bayesian models, all unobserved quantities are treated as a **random variables**, that is they can take on different values due to chance (i.e., stochastic)

Each random variable in our model is governed by a **probabilty distribution**  $^2$ 

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Our goal is to use our data to learn about those distributions

# Probability

Uncertain events are not necessarily unpredictable

# Probability

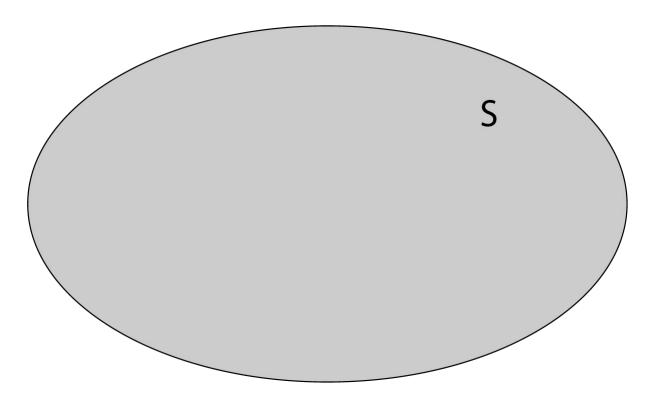
Uncertain events are not necessarily unpredictable

Probability allows us to summarize how likely each possible value of a random variable is to occur

# Sample space

For any given random variable, the sample space S includes all of the possible values the variable can take

For example, for an single-species occupancy model, S would be present or absent. For a model of species abundance, S would be  $0, 1, 2, 3, \ldots, \infty$ .



# Sample space

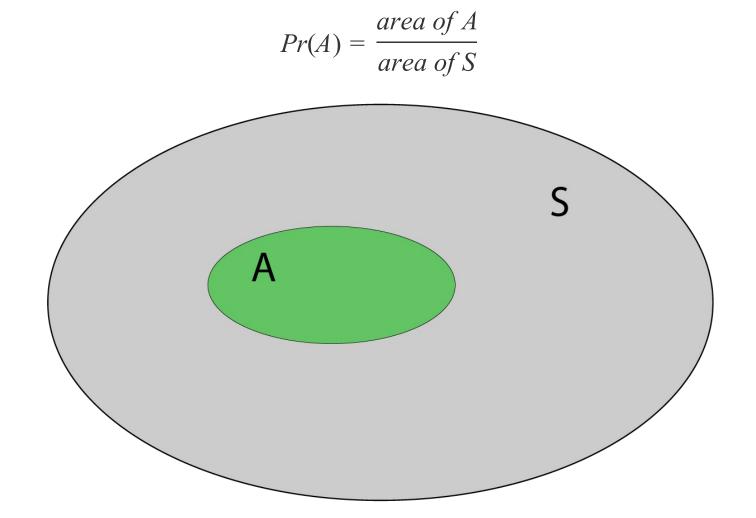
#### **Example**

Imagine an occupancy model in which we want to know if species x is present at a given location

We will denote the occupancy status  $z_{\chi}$  and the sample space includes just two possible values:

$$S_{z_x} = \{0, 1\}$$

The probability that A will occur is the area of A divided by the area of  $S^3$ 



#### **Example**

In our example, let's say that the probability of occupancy for species x is  $Pr(z_x = 1) = 0.4$ 

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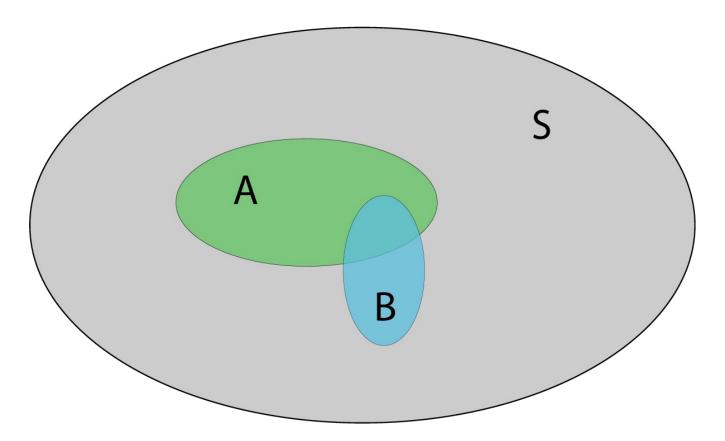
In our example, let's say that the probability of occupancy for species x is  $Pr(z_x = 1) = 0.4$ 

This means that the probability that the site in not occupied is  $Pr(z_x = 0) = 1 - 0.4 = 0.6$ 

Often, we are not interested in the probability of a single event happening but instead of more than one events

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The joint probability refers to the probability that two or more events occur and is usually denoted Pr(A, B)<sup>4</sup>



#### **Example**

To extend our simple example, let's imagine we are interested in the occupancy status of two species - x and y. Our sample space is now:

$$S_{z_x,z_y} = \{(0,0), (0,1), (1,0), (1,1)\}$$

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The answer to that question depends on the relationship between  $Pr(z_{\chi})$  and  $Pr(z_{\chi})$ 

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You can see this very clearly in this awesome visualization

 $Pr(z_y|z_x)$  is the area shared by the two events divided by the area of  $z_x$  (not S!) <sup>5</sup>

$$Pr(z_y|z_x) = \frac{area\ shared\ by\ z_x\ and\ z_y}{area\ of\ z_x} = \frac{Pr(z_x\ \cap\ z_y)}{Pr(z_x)}$$

likewise,

$$Pr(z_x | z_y) = \frac{Pr(z_x \cap z_y)}{Pr(z_y)}$$

For conditional events, the joint probability is:

$$Pr(z_y, z_x) = Pr(z_y | z_x) Pr(z_x) = Pr(z_x | z_y) Pr(z_y)$$

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Thus,

$$Pr(z_x, z_y) = Pr(z_x)Pr(z_y)$$

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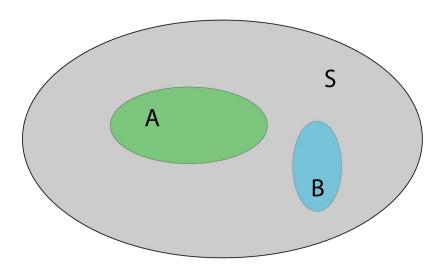
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In this case, knowing that  $z_{\chi}=1$  means that  $z_{\chi}=0$ . In other words,

$$Pr(z_y | z_x) = Pr(z_x | z_y) = 0$$



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This is the area in  $z_x$  and  $z_y$  not including the area of overlap:

$$Pr(z_x \cup z_y) = Pr(z_x) + Pr(z_y) - Pr(z_x, z_y)$$

When  $z_x$  and  $z_y$  are independent,

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If they are disjoint,

$$Pr(z_x \cup z_y) = Pr(z_x) + Pr(z_y)$$

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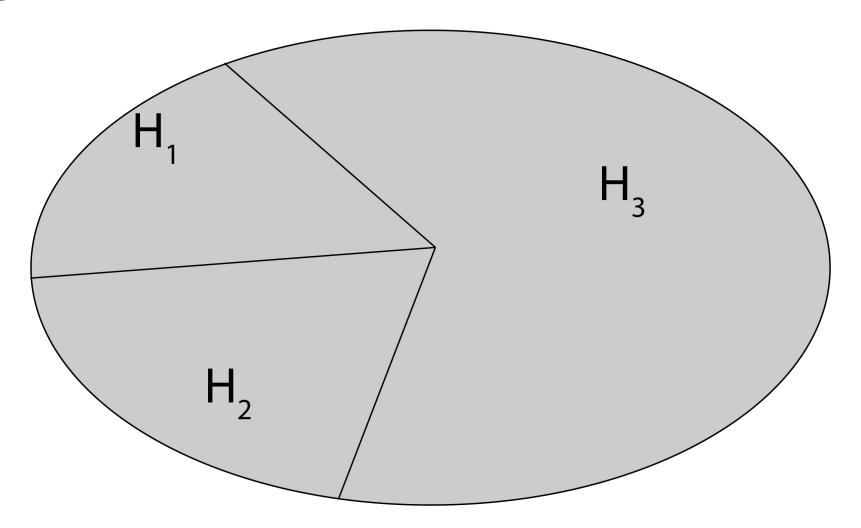
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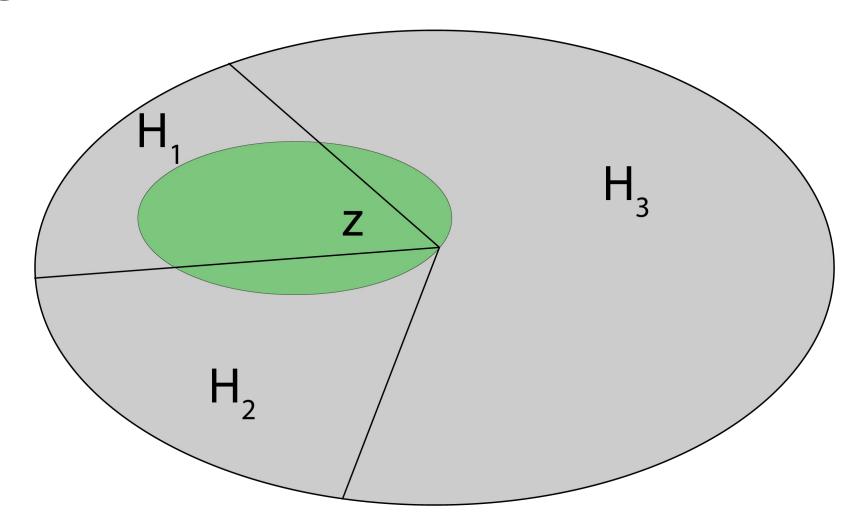
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$$Pr(z_x) = \sum_{i=1}^{3} Pr(z_x | H_i) Pr(H_i)$$

Think of this as a weighted average - the probability that  $z_x = 1$  in each habitat type weighted by the probability that each habitat type occurs.





	H1	H2	НЗ	Total
Occupied	60	10	10	80
Unoccupied	20	70	250	340
Total	80	80	260	420

Many of the models you will work with as an ecologist will contain multiple random variables

$$\left[z, \theta_{p}, \theta_{o}, \sigma_{p}^{2}, \sigma_{s}^{2}, \sigma_{o}^{2}, u_{i} \, \middle| \, y_{i}\right] \propto [y_{i} \, \middle| \, d(\Theta_{o}, u_{i}), \sigma_{o}^{2}][u_{i} \, \middle| \, z, \sigma_{s}^{2}][z \, \middle| \, g(\theta_{p}, x), \sigma_{p}^{2}][\theta_{p}][\theta_{o}][\sigma_{p}^{2}][\sigma_{s}^{2}][\sigma_{o}$$

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Determining the dependences between parameters in the models is aided by **Bayesian network** models

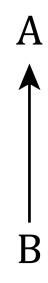
Bayesian networks graphically display the dependence among random variables

- Random variables are nodes
- Arrows point from parents to children



Bayesian networks graphically display the dependence among random variables  $^9\,$ 

- Children nodes are on the LHS of conditioning symbols
- Parent nodes are on the RHS of conditioning symbols
- Nodes without a parent are expressed unconditionally

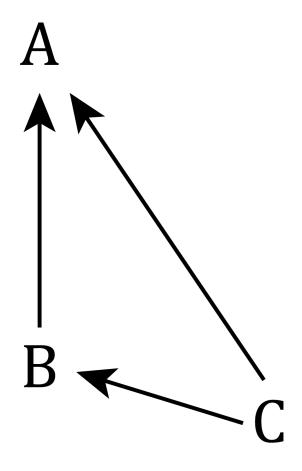


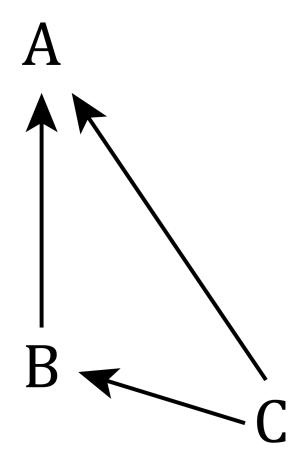
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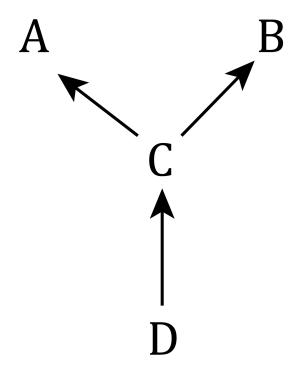


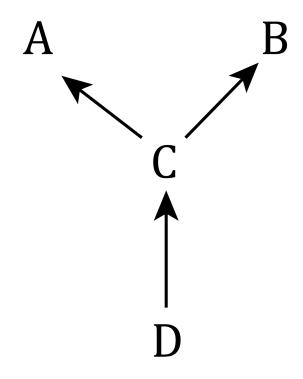
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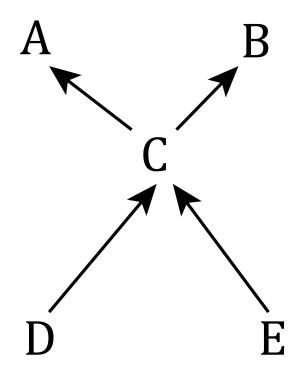


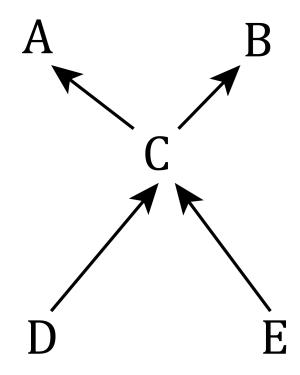
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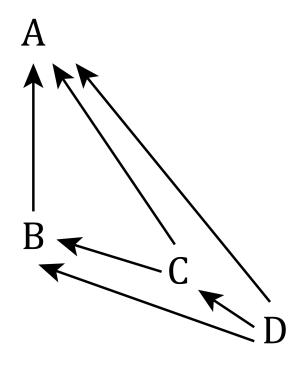


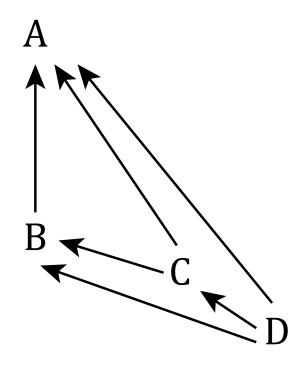
 $Pr(A, B, C, D) = Pr(A \mid C)Pr(B \mid C)Pr(C \mid D)Pr(D)$ 



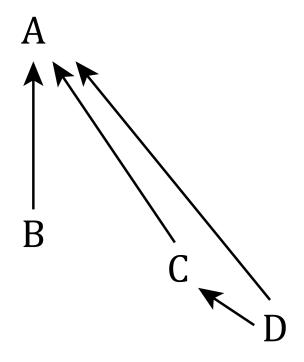


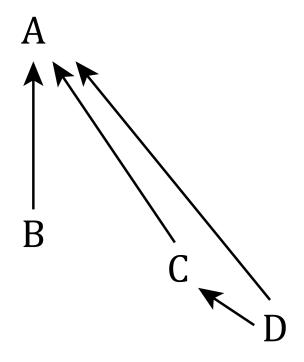
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#### Properties of probability distributions

Because all unobserved quantities are treated as random variables governed by probability distributions, using and understanding Bayesian methods requires understanding probability distributions.

As ecologists, there are a number of very common probability distributions that we encounter and use regularly  $^{10}$ :

- normal
- Poisson
- binomial
- gamma

#### Discrete vs. continuous distributions

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**Continuous** random variables can take on an infinite number of values on a specific interval  $^{11}$ 

- Normal  $(-\infty to \infty)$
- Gamma  $(0 \text{ to } \infty)$
- Beta (0 to 1)
- Uniform (? to ?)

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**Discrete** random variables are those that take on distinct values, usually integers  $^{12}$ 

- Poisson (integers  $\geq 0$ )
- Bernoulli (0 or 1)
- Binomial
- Multinomial

#### Probability functions

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Probability functions differ between continuous and discrete distributions so we will discuss these separately

For discrete random variables, the probability that the variable will take a specific value z is defined by the probability mass function (pmf)

All pmf's share two properties:

$$0 \le [z] \le 1$$

$$\sum_{z \in S} [z] = 1$$

where S is the set of all z for which [z] > 0 (the range of possible values of z).

As an example, let's assume a random variable that follows a Poisson distribution

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The shape of the Poisson distribution is determined by 1 parameter called  $\lambda$ 

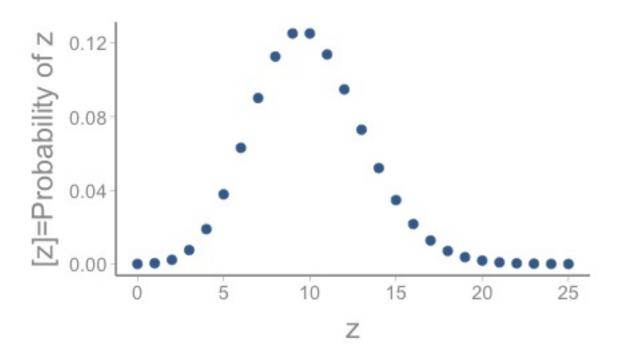
- $\lambda$  is the expected value (the most likely value) of a random variable generated from the Poisson distribution
- larger  $\lambda$  means larger values of the variable

If  $\lambda = 10$ , what is the probability that z will equal 10? Or 8? Or 15?

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In R, probability mass is estimating using the dpois() function (or the equivalent for other discrete distributions)

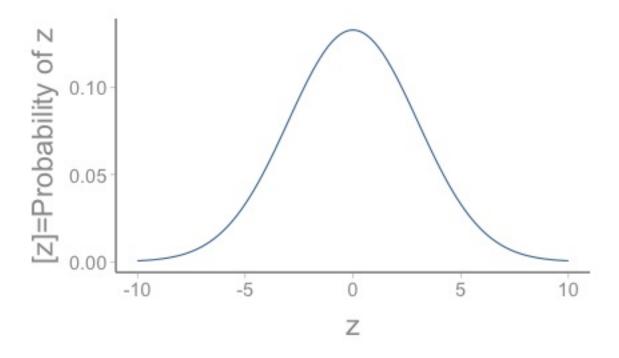
- takes two arguments: the value we are interested in estimating the probability of  $(z)^{14}$  and the expected value of our distribution  $(\lambda)$
- dpois(x = seq(0,25), lambda = 10)



Probability mass functions provide the probability that a discrete random variable takes on a specific value  $\it z$ 

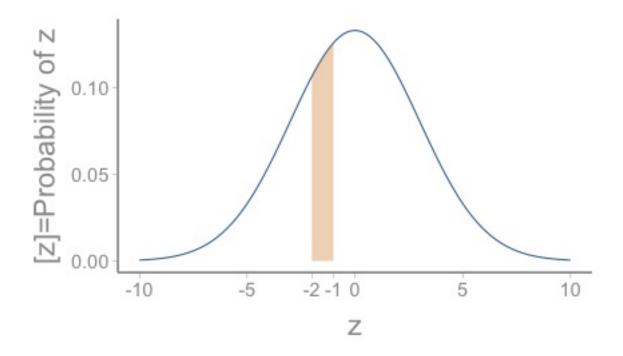
For continuous variables, estimating probabilities is a little trickier because Pr(z) = 0 for any specific value z

Why? Let's look at the probability distribution for a normal random variable with mean = 0 and standard deviation = 3:



The probability density is the area under the curve for an interval between a and b, which we'll call  $\Delta_z = (a - b)$ .

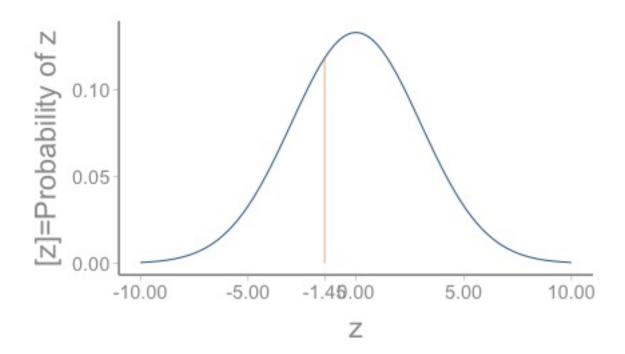
For example, the shaded area below shows the probability density  $Pr(-2 \le z \le -1)$ :



This area can be approximated by multiplying the width times the (average) height of the rectangle:

$$Pr(a \le z \le b) \approx \Delta_z[(a+b)/2]$$

By making the range  $\Delta_z = a - b$  smaller and smaller, we get closer to Pr(z):



At z,  $\Delta_z = 0$ , thus [z] = 0

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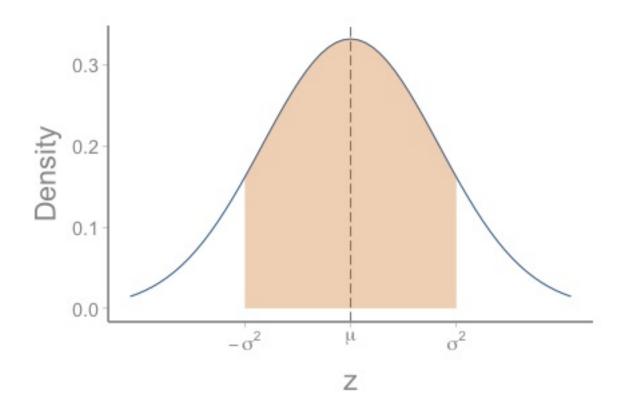
Estimating probability density in R is the same as for discrete variables: dnorm()

## Moments

### Moments

Every probability distribution we will use in the course can be described by its moments

- 1<sup>st</sup> moment is the expected value (i.e., mean)
- $2^{nd}$  moment is the variance



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$$\mu = E(z) = \sum_{z \in S} z[z]$$

For continuous variables, we need to use an integral instead of a sum:

$$\mu = E(z) = \int_{-\infty}^{\infty} z[z]dz$$

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For discrete variables, variance can be estimated as the weighted average of the squared difference (squared to prevent negative values) between each value z and the mean  $\mu$  of the distribution:

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and for continuous variables:

$$\sigma^2 = E((z - \mu)^2) = \int_{-\infty}^{\infty} (z - \mu)^2 [z] dz$$

One way to estimate moments is by simulating a large number of values from a probability distribution and then using these samples to calculate the first and second moments  $^{17}$ 

This approach is very easy to do in R using the r class of functions (e.g., rnorm(), rpois(), etc.)

• These functions generate specified number of random draws (r for random) from a given probability distribution.

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In R, we can generate and visualize a large number (e.g., 10000) random draws from the gamma distribution using the following code:

```
n <- 10000 # Sample size
samp <- rgamma(n, shape = 0.5, rate = 2)</pre>
```

Now let's use these sample to estimate the first moment (the mean) and the second moment (the variance) of the distribution

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We estimate the first moment by taking the arithmetic mean of our samples  $(\frac{1}{n}\sum_{i=1}^{n}z_i)$  and the variance as  $(\frac{1}{n}\sum_{i=1}^{n}(z_i-\mu)^2)$ :

```
mu <- sum(samp)/n # mean of the sample
sigma2 <- sum((samp - mu)^2)/n # variance of the sample
```

How close are these values to the true moments? For the gamma distribution:

$$\mu = \frac{\alpha}{\beta}$$

$$\sigma^2 = \frac{\alpha}{\beta}$$

For our samples: 1

mu # Estimated mean

## [1] 0.2542

0.5/2 # True mean

## [1] 0.25

How close are these values to the true moments? For the gamma distribution:

$$\mu = \frac{\alpha}{\beta}$$

$$\sigma^2 = \frac{\alpha}{\beta^1}$$

For our distribution:

sigma2 # Estimated variance

## [1] 0.1315

0.5/2^2

## [1] 0.125

Try this on your own - simulate data from a Poisson distribution and see if the moments you estimate from the sample are close to the true moments

**Hint** - the Poisson distribution has a single parameter  $\lambda$ , which is both the mean and the variance of the distribution

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**Hint** - the Poisson distribution has a single parameter  $\lambda$ , which is both the mean and the variance of the distribution

Change both  $\lambda$  and n. Does varying these values change how well your sample estimates the moments?  $^{20}$ 

What if you know the mean and variance of a distribution and need the parameters?

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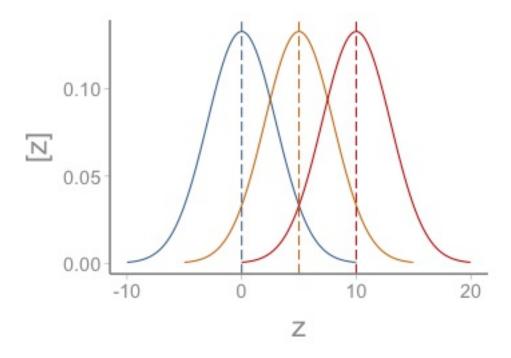
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Rather than using simulation, each distribution has a set of formulas for converting between parameters and moments (called moment matching)

Moment matching is very important because often we have the mean and variance of distributions but need to convert those summaries into the parameters of the underlying distribution  $^{21,22}$ 

For the normal distribution, it is relatively easy to understand moments because the parameters of the distribution (mean and standard deviation) are the first and second moments



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For example, the beta distribution is a continuous distribution with values between 0 and 1  $^{23,24}$  . Its first and second moments are:

$$\mu = \frac{\alpha}{\alpha + \beta}$$

$$\sigma^2 = \frac{\alpha\beta}{(\alpha + \beta)^2(\alpha + \beta + 1)}$$

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$$\alpha = \left(\frac{1-\mu}{\sigma^2} - \frac{1}{\mu}\right)\mu^2$$

$$\beta = \alpha \left(\frac{1}{\mu} - 1\right)$$

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$$\alpha = \left(\frac{1-\mu}{\sigma^2} - \frac{1}{\mu}\right)\mu^2$$

$$\beta = \alpha \left(\frac{1}{\mu} - 1\right)$$

For our model, that means  $^{25}$ :

```
(alpha <- ( (1 - 0.3)/0.025 - (1/0.3) )*0.3^2)
```

## [1] 2.22

## [1] 5.18