

IO-Problem-Sets (/github/RussellMiles310/IO-Problem-Sets/tree/main)  
 / ps3\_submission\_hirvonen\_miles (/github/RussellMiles310/IO-Problem-Sets/tree/main/ps3\_submission\_hirvonen\_miles)  
 / Code\_ACF\_GNR (/github/RussellMiles310/IO-Problem-Sets/tree/main/ps3\_submission\_hirvonen\_miles/Code\_ACF\_GNR)  
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 jupyter\_notebooks\_for\_exposition (/github/RussellMiles310/IO-Problem-Sets/tree/main/ps3\_submission\_hirvonen\_miles/Code\_ACF\_GNR/jupyter\_notebooks\_for\_exposition)

## GNR -- Gandhi, Navarro, Rivers

### Summary of the GNR method:

Derivation of the model:

The firm's profit maximization problem with respect to intermediate inputs is:

$$\max_{M_{jt}} P_t \mathbb{E} [F(k_{jt}, l_{jt}, m_{jt}) e^{\omega_{jt} + \epsilon_{jt}}] - \rho_t M_{jt}.$$

and the first-order condition is:

$$P_t \frac{\partial}{\partial M_{jt}} F(k_{jt}, l_{jt}, m_{jt}) e^{\omega_{jt}} \mathcal{E} = \rho_t.$$

with  $\mathcal{E} \equiv e^{\epsilon_{jt}}$ . Take Logs:

$$\ln P_t + \ln \left( \frac{\partial}{\partial M_{jt}} F(k_{jt}, l_{jt}, m_{jt}) \right) + \omega_{jt} + \ln \mathcal{E} = \ln \rho_t.$$

Now subtract  $\ln Y_{jt} = \ln F(k_{jt}, l_{jt}, m_{jt}) + \omega_{jt} + \epsilon_{jt}$  from both sides

$$\ln \mathcal{E} + \ln \left( \frac{\partial}{\partial M_{jt}} F(k_{jt}, l_{jt}, m_{jt}) \right) - \epsilon_{jt} = \ln \rho_t - \ln P_t - \ln Y_{jt} + \ln F(k_{jt}, l_{jt}, m_{jt}) \equiv \ln \left( \frac{\rho_t F(k_{jt}, l_{jt}, m_{jt})}{P_t Y_{jt}} \right).$$

Using the properties of log derivatives (where lowercase = logs)

$$\ln \left( \frac{\partial}{\partial M_{jt}} F(k_{jt}, l_{jt}, m_{jt}) \right) \equiv \ln \left( \frac{\partial}{\partial m_{jt}} f(k_{jt}, l_{jt}, m_{jt}) \cdot \frac{F(k_{jt}, l_{jt}, m_{jt})}{M_t} \right)$$

Then moving the  $\frac{F(k_{jt}, l_{jt}, m_{jt})}{M_t}$  to the other side we get

$$\underbrace{\ln \mathcal{E} + \ln \left( \frac{\partial}{\partial m_{jt}} f(k_{jt}, l_{jt}, m_{jt}) \right) - \epsilon_{jt}}_{\ln D^{\mathcal{E}}} = \ln \left( \frac{\rho_t M_{jt}}{P_t Y_{jt}} \right),$$

$$\ln D^{\mathcal{E}}(k_{jt}, l_{jt}, m_{jt}) - \epsilon_{jt} = \ln \left( \frac{\rho_t M_{jt}}{P_t Y_{jt}} \right) \equiv s_{jt},$$

where  $s_{jt}$  is the log intermediate share of output.

### Estimation Procedure

First, we want to estimate the elasticity  $D^{\mathcal{E}}(k_{jt}, l_{jt}, m_{jt}) = \text{PolynomialFit}(k_{jt}, l_{jt}, m_{jt})$ . To find the coefficients of the polynomials, we run the estimator

$$\min_{\gamma'} \sum_{j,t} \left\{ s_{jt} - \ln \left( \underbrace{\frac{\gamma'_0 + \gamma'_k k_{jt} + \gamma'_l l_{jt} + \gamma'_m m_{jt} + \gamma'_{kk} k_{jt}^2 + \gamma'_{ll} l_{jt}^2}{\gamma'_{mm} m_{jt}^2 + \gamma'_{kl} k_{jt} l_{jt} + \gamma'_{km} k_{jt} m_{jt} + \gamma'_{lm} l_{jt} m_{jt}}}_{D^{\mathcal{E}}} \right) \right\}^2$$

where  $s_{jt} \equiv \ln \left( \frac{\rho_t M_{jt}}{P_t Y_{jt}} \right)$  is the log intermediate share of input. Then using the gamma coefficients, we get

$$\hat{D}_{jt} \equiv D^{\mathcal{E}}(k_{jt}, l_{jt}, m_{jt}) = \text{Polynomial}(k_{jt}, l_{jt}, m_{jt} | \gamma')$$

From there, get the residuals

$$\hat{\epsilon}_{jt} = \ln \hat{D}_{jt} - s_{jt}$$

Next, estimate  $\hat{\mathcal{E}} = \frac{1}{JT} \sum_{j,t} e^{\hat{\epsilon}_{jt}}$ . Then using

$$\ln \hat{D}_{jt}^{\mathcal{E}} = \ln \hat{\mathcal{E}} + \ln \left( \frac{\partial}{\partial m_{jt}} f_{jt} \right)$$

$$\hat{D}_{jt}^{\mathcal{E}} \cdot \frac{1}{\hat{\mathcal{E}}} = \frac{\partial}{\partial m_{jt}} f_{jt}$$

$$\hat{X}_{\text{poly}} \cdot \frac{\hat{\gamma}'}{\hat{\mathcal{E}}} = \frac{\partial}{\partial m_{jt}} f_{jt}$$

That is, we can use the same polynomial design matrix used to estimate  $\hat{D}_{jt}^{\mathcal{E}}$ , but just divide the gamma coefficients by  $\hat{\mathcal{E}}$ , to get the  $\frac{\partial}{\partial m_{jt}} f_{jt}$ .

## Integral:

Next we want to recover the constant of integration in

$$\int \frac{\partial}{\partial m_{jt}} f(k_{jt}, l_{jt}, m_{jt}) dm_{jt} = f(k_{jt}, l_{jt}, m_{jt}) + \mathcal{C}(k_{jt}, l_{jt})$$

This amounts to integrating the polynomial using  $\gamma = \gamma' / \hat{\mathcal{E}}$ .

$$\text{integral} = \mathcal{D}(k_{jt}, l_{jt}, m_{jt}) = \left( \gamma_0 + \gamma_k k_{jt} + \gamma_l l_{jt} + \frac{\gamma_m}{2} m_{jt} + \gamma_{kk} k_{jt}^2 + \gamma_{ll} l_{jt}^2 + \frac{\gamma_{mm}}{3} m_{jt}^2 + \gamma_{kl} k_{jt} l_{jt} + \frac{\gamma_{km}}{2} k_{jt} m_{jt} + \frac{\gamma_{lm}}{2} l_{jt} m_{jt} \right) m_{jt}$$

Using the above  $\mathcal{D}$ , we can calculate  $\mathcal{Y}_{jt} \equiv y_{jt} - \epsilon_{jt} - \mathcal{D}_{jt} = -\mathcal{C}(k_{jt}, l_{jt}) + \omega_{jt}$

$$\hat{\mathcal{Y}}_{jt} = \ln \left( \frac{Y_{jt}}{e^{\hat{\epsilon}_{jt}} e^{\mathcal{D}_{jt}}} \right)$$

## Constant of Integration:

We approximate  $\omega$  with a function  $h(\cdot)$ :

$$\mathcal{Y}_{jt} = -\mathcal{C}(k_{jt}, l_{jt}) + h(\mathcal{Y}_{jt-1} + \mathcal{C}(k_{jt-1}, l_{jt-1})) + \eta_{jt}.$$

In this case we use

$$\begin{aligned} \mathcal{C}(k_{jt}, l_{jt}) &= \text{PolynomialFit}(k_{jt}, l_{jt}, \text{degree}_C) \\ h_A(\omega_{jt-1}) &= \text{PolynomialFit}(\omega_{jt-1}, \text{degree}_\omega) = \text{PolynomialFit}(\mathcal{Y}_{jt-1} + \mathcal{C}(k_{jt-1}, l_{jt-1}), \text{degree}_\omega) + \eta_{jt} \end{aligned}$$

All together,

$$\hat{\mathcal{Y}}_{jt} = - \sum_{0 < \tau_k + \tau_l \leq \tau} \alpha_{\tau_k, \tau_l} k_{jt}^{\tau_k} l_{jt}^{\tau_l} + \sum_{0 \leq a \leq A} \delta_a \left( \hat{\mathcal{Y}}_{jt-1} + \sum_{0 < \tau_k + \tau_l \leq \tau} \alpha_{\tau_k, \tau_l} k_{jt-1}^{\tau_k} l_{jt-1}^{\tau_l} \right)^a + \eta_{jt}$$

The moments to estimate are:

$$\begin{aligned} (i) \quad E \left[ \varepsilon_{jt} \frac{\partial \ln D_r(k_{jt}, l_{jt}, m_{jt})}{\partial \gamma} \right] &= 0 \\ (ii) \quad E \left[ \eta_{jt} k_{jt}^{\tau_k} l_{jt}^{\tau_l} \right] &= 0, \\ (iii) \quad E \left[ \eta_{jt} \mathcal{Y}_{jt-1}^a \right] &= 0, \end{aligned}$$

## Moving to the second step of GNR

The moments to estimate are:

$$\begin{aligned} (i) \quad E \left[ \varepsilon_{jt} \frac{\partial \ln D_r(k_{jt}, l_{jt}, m_{jt})}{\partial \gamma} \right] &= 0 \\ (ii) \quad E \left[ \eta_{jt} k_{jt}^{\tau_k} l_{jt}^{\tau_l} \right] &= 0, \\ (iii) \quad E \left[ \eta_{jt} \mathcal{Y}_{jt-1}^a \right] &= 0, \end{aligned}$$

We can "concentrate out" (i) by running the nonlinear least squares regression of of the share equation  $s_{jt}$  on  $\ln D_{jt}$ .

We can concentrate out (iii) with the following procedure:

- Guess  $\alpha$
- Form  $\omega_{jt-1}(\alpha) = \hat{\mathcal{Y}}_{jt-1} + \text{Poly}_{\text{no intercept}}^\alpha(k_{jt-1}, l_{jt-1})$
- Calculate  $\hat{\omega}_{jt} = \hat{\mathcal{Y}}_{jt} + \text{Poly}_{\text{no intercept}}^\alpha(k_{jt-1}, l_{jt-1})$ , plugging in the guessed regression coefficients  $\alpha$  and multiplying them with the polynomial design matrix to get the predicted omega.
- Polynomial-regress  $\hat{\omega}_{jt} \sim \hat{\omega}_{jt-1}$  to get the innovation  $\hat{\eta}_{jt}$
- Use the moments  $\mathbb{E}[\hat{\eta}_{jt} X_{jt}] = 0$  for all  $X \in [k, l, k^2, l^2, kl, \dots, \text{higher order polynomial terms}]$ .

So, we will have as many moments as terms in the polynomial regression. The regression to approximate the constant of integration has no intercept.

Now we have  $\alpha$ , which lets us recover the constant of integration, and we have  $\delta$  (which is just the regression coefficients from regression  $\omega$  on the polynomial of lagged  $\omega$ , which lets us recover the productivity  $\omega_{jt}$ . The production function is nonparametrically identified by

$$f(k_{jt}, l_{jt}, m_{jt}) = \mathcal{D}_{jt} - \mathcal{C}_{jt}$$

that is, it's equal to the "integral" minus the "constant of integration."

## Load in the data

```
In [4]: import autograd.numpy as np
        from autograd import grad, jacobian, hessian
        #import numpy as np
        import pandas as pd
        import scipy.optimize as opt
        import matplotlib.pyplot as plt
        from mpl_toolkits.mplot3d import Axes3D
        import math
        from itertools import combinations_with_replacement, chain
```

```
In [5]: filename = "../PS3_data_changedtoxlsx.xlsx"
        df0 = pd.read_excel(filename)
        #Remove missing materials columns
        df = df0[['year', 'firm_id', 'X03', 'X04', 'X05', 'X16', 'X40', 'X43', 'X44', 'X45', 'X49']]
        #new_names = ["year", "firm_id", "obs", "ly", "s01", "s02", "lc", "ll", "lm"]
        new_names = ["t", "firm_id", "y_gross", "s01", "s02", "s13", "k", "l", "m", 'py', 'pm']
        df.columns = new_names
        #Drop missing materials data
        df=df[df['m']!=0]
        #Keep industry 1 only
        df=df[df['s13']==1]

        #Creating value-added y
        df['y'] = df['y_gross']
        #Creating the intermeidate share of output cost variable: emember everything is already in Logs
        #df['s'] = np.log(np.exp(df['pm'])*np.exp(df['m']))/(np.exp(df['py'])*np.exp(df['y']))
        df['s'] = df['pm']+df['m'] - df['py'] - df['y']
        #Creating Lagged variables
        df = df.sort_values(by=['firm_id', 't'])
        df['kprev'] = df.groupby('firm_id')['k'].shift(1)
        df['lprev'] = df.groupby('firm_id')['l'].shift(1)
        df['mprev'] = df.groupby('firm_id')['m'].shift(1)
```

```
In [6]: #Creates an iterator of tuples, useful for constructing polynomial regression design matrices.
def poly_terms(n_features, degree):
    #This thing creates an iterator structure of tuples, used to create polynomial interaction terms.
    #It looks something like this: (0,), (1,), (2,), (0, 0), (0, 1)
    polynomial_terms = chain(
        *(combinations_with_replacement(range(n_features), d) for d in range(1, degree+1))
    )
    return(polynomial_terms)

#Constructs a polynomial regression design matrix.
def poly_design_matrix(xvars, degree):
    if xvars.ndim == 1:
        xvars = xvars.reshape(1, -1)
    # Get the number of samples (n) and number of features (m) from X
    n_samples, n_features = xvars.shape
    # Start with a column of ones for the intercept term
    X_poly = np.ones((n_samples, 1))
    #Create iterator used to construct polynomial terms
    polynomial_terms = poly_terms(n_features, degree)
    # Generate polynomial terms and interaction terms up to 4th degree
    for terms in polynomial_terms: # For degrees 1 to 4
        #print(terms)
        X_poly = np.hstack((X_poly, np.prod(xvars[:, terms], axis=1).reshape(-1, 1)))
    # Compute the coefficients using the normal equation: beta = (X.T * X)^(-1) * X.T * y
    return X_poly

#Runs a regression
def regress(y, X):
    beta = np.linalg.solve(X.T@X, X.T@y)
    yhat = X@beta
    resids = y-yhat
    return beta, yhat, resids

#Function for finding the nonlinear Least squares objective function
def nlls_share_obj(gamma, X_poly, s):
    #gamma is the vector of coefficients
    #X_poly is the design matrix containing all of the polynomial coefficients
    #Check of X_poly@gamma is negative, if so, make it small
    #epsilon = 1e-16
    Dhat = X_poly@gamma

    #Autograd does not like in-place modification. Use np.where to set negative elements to a small number
    #Dhat = np.where(Dhat <= epsilon, epsilon, Dhat)

    #print(min(Dhat))
    #Evaluate the objective
    obj = np.sum((s.flatten() - np.log(Dhat))**2) #/(X_poly.shape[0])
    return obj

def nlls_residuals(gamma, X_poly, s):
    #gamma is the vector of coefficients
    #X_poly is the design matrix containing all of the polynomial coefficients
    #Check of X_poly@gamma is negative, if so, make it small
    Dhat = X_poly@gamma

    #Autograd does not like in-place modification. Use np.where to set negative elements to a small number
    #Dhat = np.where(Dhat <= epsilon, epsilon, Dhat)

    #print(min(Dhat))
    #Evaluate the objective
    resids = (s.flatten() - np.log(Dhat)) #/(X_poly.shape[0])
    return resids
```

## First step of coding: Fit $D$ .

$$\min_{\gamma'} \sum_{j,t} \left\{ s_{jt} - \ln \left( \underbrace{\gamma'_0 + \gamma'_k k_{jt} + \gamma'_l l_{jt} + \gamma'_m m_{jt} + \gamma'_{kk} k_{jt}^2 + \gamma'_{ll} l_{jt}^2}_{D^{\mathcal{E}}} + \gamma'_{mm} m_{jt}^2 + \gamma'_{kl} k_{jt} l_{jt} + \gamma'_{km} k_{jt} m_{jt} + \gamma'_{lm} l_{jt} m_{jt} \right) \right\}^2$$

```
In [23]: degree= 2
#Make the polynomial design matrix
xvars = df[['k', 'l', 'm']].to_numpy()
s = df[['s']].to_numpy()
X_poly_D = poly_design_matrix(xvars, degree)
#calculate the gradient of the objective function using AutoGrad
autogradient_nlls = grad(nlls_share_obj)
autohessian_nlls = hessian(nlls_share_obj)
#initial guess
gamma_prime0 = np.ones(X_poly_D.shape[1])
```

## Scipy minimize seems to fit the data well and finds a minimum (gradient close to 0)

```
In [25]: #minimize to fit the coefficients gamma_prime
#Enforce that X@gamma is nonnegative, otherwise we get negative values in the log
nonnegative_b = {'type': 'ineq', 'fun': lambda b: (X_poly_D@b)}

gamma_prime_results = opt.minimize(nlls_share_obj, gamma_prime0, args=(X_poly_D, s),
                                   constraints = [nonnegative_b],
                                   tol=1e-12, jac=autogradient_nlls, hess = autohessian_nlls, method='trust-constr'
                                   )

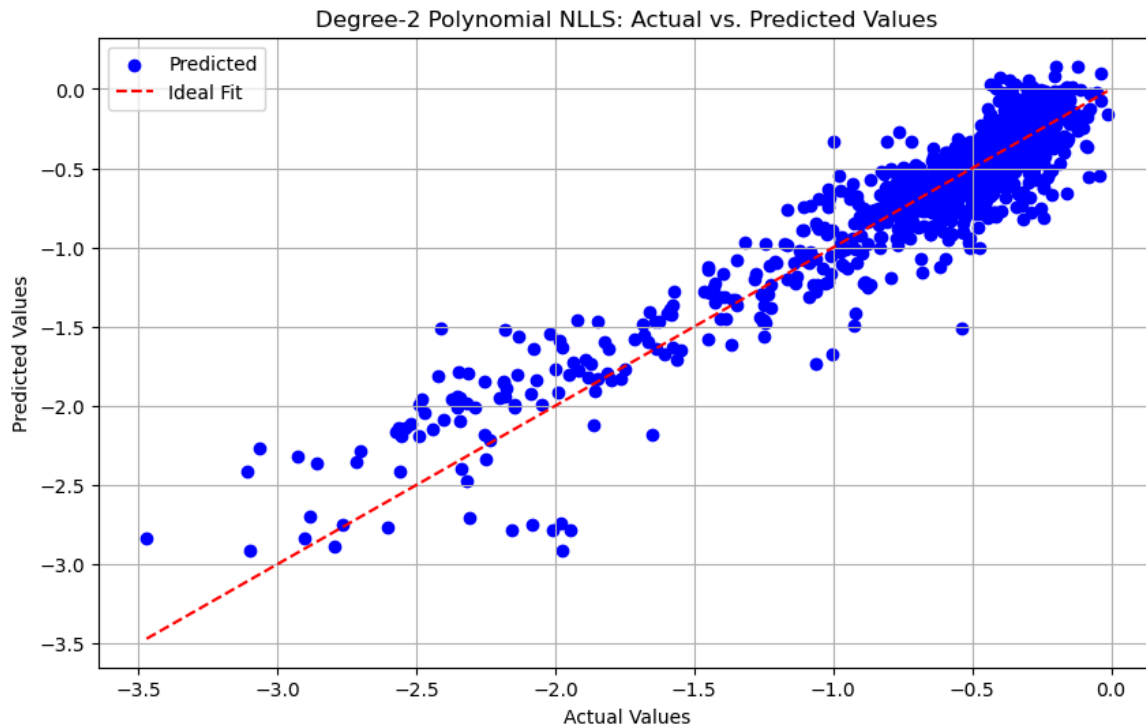
print("The error is:", gamma_prime_results.fun)
print("The gradient is:", gamma_prime_results.grad)
print("The coefficients in the degree-1 fit are:", gamma_prime_results.x)

shat = np.log(X_poly_D@gamma_prime_results.x)

The error is: 45.751893555022164
The gradient is: [4.09656451e-08 4.67529361e-07 1.98948108e-07 5.06100143e-07
 5.42993808e-06 2.32990450e-06 5.83383169e-06 1.02908564e-06
 2.51373481e-06 6.32519689e-06]
The coefficients in the degree-1 fit are: [-0.06492208  0.07132593 -0.04254507 -0.05276808 -0.00885781  0.02737997
 -0.00261827  0.00630371 -0.03944874  0.01855989]
```

## Actual by predicted plot for my polynomial approximation of $\Phi$ , run on the entire sample

```
In [28]: # Example data (replace these with your actual y and yhat values)
# Create a scatter plot for actual vs. predicted
plt.figure(figsize=(10, 6))
plt.scatter(s, shat, color='blue', label='Predicted', marker='o')
plt.plot([min(s), max(s)], [min(s), max(s)], color='red', linestyle='--', label='Ideal Fit')
# Add labels and title
plt.xlabel('Actual Values')
plt.ylabel('Predicted Values')
plt.title('Degree-2 Polynomial NLLS: Actual vs. Predicted Values')
plt.legend()
plt.grid()
# Show the plot
plt.show()
```



Now we have  $\hat{D}_{jt}$ .

```
In [30]: gammaprime = gammaprime_results.x
#Get Dhat, the elasticities
df['Dhat'] = X_poly_D@gammaprime
#Back out the residuals, epsilons
df['epsilonhat'] = np.log(df['Dhat']) - df['s']
# mean of epsilon is 1e-12 --- good sign
mean_eps = np.mean((df['epsilonhat']))
var_eps = np.var((df['epsilonhat']))
#From here, estimate curlyE which is the sample average of exp(epsilons)
#Turns out to be 1.02. the mean of a lognormal variable of mean theta is e^(sigma^2/2). That implies the variance of the epsilons is vey l
CurlyEhat = (np.mean(np.exp(df['epsilonhat'])))
lognormal_guess_curlyEhat = np.exp(var_eps/2)
#The theoretical guess for CurlyEhat given epsilon ~ N(theta, sigma^2) is very close to the actual curlyEhat,
#suggesting the epsilons are approximately normally distributed.
#It follows from the math above that ...
gamma = gammaprime/CurlyEhat
df['df_dm'] = X_poly_D@gamma
```

Next we want to recover the integral:

$$\int \frac{\partial}{\partial m_{jt}} f(k_{jt}, l_{jt}, m_{jt}) dm_{jt} = f(k_{jt}, l_{jt}, m_{jt}) + \mathcal{C}(k_{jt}, l_{jt})$$

This amounts to integrating the polynomial using  $\gamma = \gamma' / \hat{\mathcal{E}}$ .

$$\text{integral} = \mathcal{D}(k_{jt}, l_{jt}, m_{jt}) = \left( \gamma_0 + \gamma_k k_{jt} + \gamma_l l_{jt} + \frac{\gamma_m}{2} m_{jt} + \gamma_{kk} k_{jt}^2 + \gamma_{ll} l_{jt}^2 + \frac{\gamma_{mm}}{3} m_{jt}^2 + \gamma_{kl} k_{jt} l_{jt} + \frac{\gamma_{km}}{2} k_{jt} m_{jt} + \frac{\gamma_{lm}}{2} l_{jt} m_{jt} \right) m_{jt}$$

Using the above  $\mathcal{D}$ , we can calculate  $\mathcal{Y}_{jt} \equiv y_{jt} - \epsilon_{jt} - \mathcal{D}_{jt} = -\mathcal{C}(k_{jt}, l_{jt}) + \omega_{jt}$

$$\hat{\mathcal{Y}}_{jt} = \ln \left( \frac{Y_{jt}}{e^{\hat{\epsilon}_{jt}} e^{\mathcal{D}_{jt}}} \right)$$

but I don't see any reason to take exps and logs. Seems easier to evaluate the log and use

$$\hat{\mathcal{Y}}_{jt} \equiv y_{jt} - \hat{\epsilon}_{jt} - \mathcal{D}_{jt}$$

```
In [32]: #by default, integrate with respect to xvars[2] = m
def poly_integral_design_matrix(xvars, degree, w_r_t = 2):
    #Get number of observations (n) and number of independent variables (k)
    if xvars.ndim == 1:
        xvars = xvars.reshape(1, -1)
    # Get the number of samples (n) and number of features (m) from X
    n_samples, n_features = xvars.shape
    # Start with a column of ones for the intercept term
    X_poly_integral0 = np.ones((n_samples, 1))
    #Create iterator used to construct polynomial terms
    polynomial_terms = poly_terms(n_features, degree)
    # Generate polynomial terms and interaction terms up to 4th degree
    for terms in polynomial_terms: # For degrees 1 to 4
        integration_divisor = terms.count(w_r_t) + 1 #count the number of xvars[2] (i.e. m) appearing in the term and add 1.
        #Divide the column by that term to compute the "integration scalar"
        xcolumn = np.prod(xvars[:, terms], axis=1).reshape(-1, 1) / integration_divisor
        X_poly_integral0 = np.hstack((X_poly_integral0, xcolumn))
    #Elementwise-multiply all columns in the resulting matrix by m
    X_poly_integral = X_poly_integral0 * xvars[:,w_r_t].reshape(xvars.shape[0],1)
    return X_poly_integral
```

```
In [33]: #Get the design matrix associated with the integral of the polynomial
X_poly_D_integral = poly_integral_design_matrix(xvars, degree, w_r_t = 2)
#Evaluate it to get curlyD, which is the integral of the log elasticities
df['CurlyD'] = X_poly_D_integral@gamma
#from here, get CurlyY
df['CurlyY'] = df['y'] - df['epsilonhat'] - df['CurlyD']
```

## Moving to the second step of GNR

The moments to estimate are:

$$\begin{aligned} (i) \quad E \left[ \varepsilon_{jt} \frac{\partial \ln D_r(k_{jt}, l_{jt}, m_{jt})}{\partial \gamma} \right] &= 0 \\ (ii) \quad E \left[ \eta_{jt} k_{jt}^{\tau_k} l_{jt}^{\tau_l} \right] &= 0, \\ (iii) \quad E \left[ \eta_{jt} \mathcal{Y}_{jt-1}^{\alpha} \right] &= 0, \end{aligned}$$

We can "concentrate out" (i) by running the nonlinear least squares regression of of the share equation  $s_{jt}$  on  $\ln D_{jt}$ .

We can concentrate out (iii) with the following procedure:

- Guess  $\alpha$
- Form  $\omega_{jt-1}(\alpha) = \hat{\mathcal{Y}}_{jt-1} + \text{Poly}_{\text{no intercept}}^{\alpha}(k_{jt-1}, l_{jt-1})$
- Calculate  $\hat{\omega}_{jt} = \hat{\mathcal{Y}}_{jt} + \text{Poly}_{\text{no intercept}}^{\alpha}(k_{jt-1}, l_{jt-1})$ , plugging in the guessed regression coefficients  $\alpha$  and multiplying them with the polynomial design matrix to get the predicted omega.
- Polynomial-regress  $\hat{\omega}_{jt} \sim \hat{\omega}_{jt-1}$  to get the innovation  $\hat{\eta}_{jt}$
- Use the moments  $\mathbb{E}[\hat{\eta}_{jt} X_{jt}] = 0$  for all  $X \in [k, l, k^2, l^2, kl, \dots, \text{higher order polynomial terms}]$ .

Note that there is NO INTERCEPT in the approximation of the constant of integration. Even if we put an intercept in, it would not be identified.

So, we will have as many moments as terms in the polynomial regression.

```
In [35]: #First, drop all NaNs
df['CurlyYprev'] = df.groupby('firm_id')['CurlyY'].shift(1)

df_nonans = df.dropna().copy()

degree_omega = 2
xvars_omega = df_nonans[["k", "l"]].to_numpy()
xvars_prev_omega = df_nonans[["kprev", "lprev"]].to_numpy()

#This polynomial fit has NO INTERCEPT. Even if we wanted an intercept it would not be identified because we end up taking first differences
X_poly_omega = poly_design_matrix(xvars_omega, degree_omega)[: , 1:]
Xprev_poly_omega = poly_design_matrix(xvars_prev_omega, degree_omega)[: , 1:]

#Previous CurlyY
CurlyY = df_nonans['CurlyY'].to_numpy()
CurlyYprev = df_nonans['CurlyYprev'].to_numpy()
```

```
In [36]: #Moment error function
def gmm_stage2_error_GNR(alpha, X_poly_omega, Xprev_poly_omega, CurlyY, CurlyYprev):
    #Given alpha, the previous omega is Curly Y + the integration constant curlyC, which is a polynomial fit on lagged k and L
    #Note that there's NO INTERCEPT in this polynomial fit
    #Even if we included an intercept,
    omegaprev = CurlyYprev + Xprev_poly_omega@alpha
    #Then, calculate current omega = curlyY
    omega = CurlyY + X_poly_omega@alpha
    #Regress omega on omegaprev
    degree_omega = 2
    #xvars
    Xo = poly_design_matrix(omegaprev.reshape(-1, 1), degree_omega)
    #Fit the regression omegaprev ~ polynomial(omega)
    #delta is the coefficients, eta are the residuals
    delta, _, eta = regress(omega, Xo)
    #calculate the moment errors
    #The moments are simply the polynomial terms in poly(alpha) used to approximate the constant of integration.
    moment_error = (X_poly_omega.T @ eta)/len(eta)
    return moment_error, delta, eta

def gmm_obj_fcn_GNR(alpha, X_poly_omega, Xprev_poly_omega, CurlyY, CurlyYprev, W):
    #Get the moment errors
    moment_error = gmm_stage2_error_GNR(alpha, X_poly_omega, Xprev_poly_omega, CurlyY, CurlyYprev)[0]
    #calucLate GMM error objective function
    obj = moment_error.T@W@moment_error
    return obj

#Define autogradient
autogradient_GNR = grad(gmm_obj_fcn_GNR)
```

```
In [37]: #initial guess for alpha, the polynomial fit for omega
alpha0 = np.ones(X_poly_omega.shape[1])/2
W0 = np.eye(len(alpha0))

args_GNR = (X_poly_omega, Xprev_poly_omega, CurlyY, CurlyYprev, W0)

gmm_results_GNR = opt.minimize(gmm_obj_fcn_GNR, alpha0, args=args_GNR,
                               tol=1e-24, jac=autogradient_GNR, method='L-BFGS-B'
                               )

alpha = gmm_results_GNR.x
delta, eta = gmm_stage2_error_GNR(alpha, X_poly_omega, Xprev_poly_omega, CurlyY, CurlyYprev)[1:3]

print("The error is:", gmm_results_GNR.fun)
print("The gradient is:", gmm_results_GNR.jac)
print("The coefficients for the integration constant [alpha] are:", gmm_results_GNR.x)
print("The coefficients for productivity omega [delta] are:", delta)

The error is: 4.0982459107368714e-25
The gradient is: [-2.57379564e-11 -1.50131709e-11 -6.16007307e-10 -3.14746311e-10
-1.55119473e-10]
The coefficients for the integration constant [alpha] are: [ 0.22962852 -2.6976986 -0.08259352  0.26112497  0.05193076]
The coefficients for productivity omega [delta] are: [1.97089299 0.17895714 0.08014284]
```

```
In [38]: #Calculate the omegas
df_nonans['ConstantC'] = X_poly_omega@alpha
df_nonans['omega'] = df_nonans['ConstantC'] + df_nonans['CurlyY']

Eomega = np.mean(df_nonans['omega'])
Edf_dm = np.mean(df_nonans['df_dm'])
print("the average productivity [omega] is:", Eomega)
print("the average elasticity [df/dm] is:", Edf_dm)
```

the average productivity [omega] is: 3.8396919943168815  
the average elasticity [df/dm] is: 0.5696607852776345

The production function is nonparametrically identified by

$$f(k_{jt}, l_{jt}, m_{jt}) = \mathcal{D}_{jt} - \mathcal{C}_{jt}$$

that is, it's equal to the "integral" minus the "constant of integration."

For comparability with ACF and panel methods, assume Cobb-Douglas form.

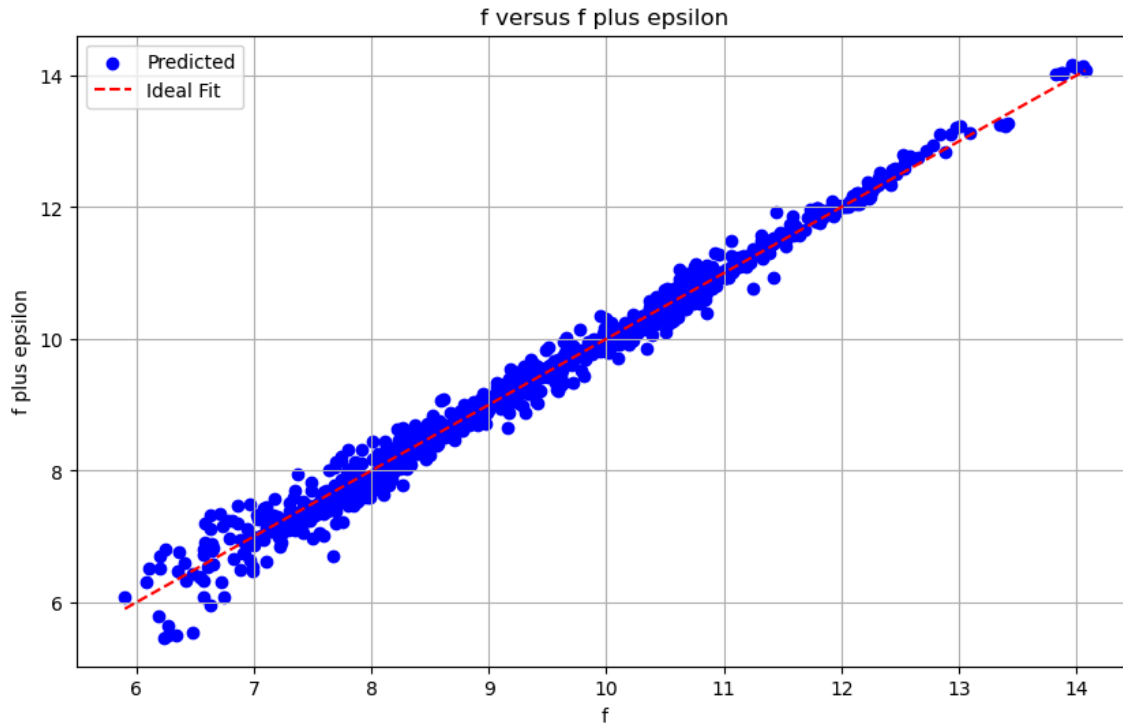
$$f(k_{jt}, l_{jt}, m_{jt}) = \beta_0 + \beta_k k_{jt} + \beta_l l_{jt} + \beta_m m_{jt}$$

Then we can get the elasticities out using ols of  $f(k_{jt}, l_{jt}, m_{jt})$  on 1,  $k_{jt}$ ,  $l_{jt}$ ,  $m_{jt}$

```
In [40]: df_nonans['f'] = df_nonans['CurlyD'] - df_nonans['ConstantC']
df_nonans['f_plus_epsilon'] = df_nonans['y'] - df_nonans['omega']
f = df_nonans['f']
f_plus_epsilon = df_nonans['f_plus_epsilon']
```



```
In [41]: # Example data (replace these with your actual y and yhat values)
# Create a scatter plot for actual vs. predictedZZZZZ
plt.figure(figsize=(10, 6))
plt.scatter(f, f_plus_epsilon, color='blue', label='Predicted', marker='o')
plt.plot([min(f), max(f)], [min(f), max(f)], color='red', linestyle='--', label='Ideal Fit')
# Add Labels and title
plt.xlabel('f')
plt.ylabel('f plus epsilon')
plt.title('f versus f plus epsilon')
plt.legend()
plt.grid()
# Show the plot
plt.show()
```



```
In [42]: klm = df_nonans[['k', 'l', 'm']]
Xklm = np.hstack((np.ones((klm.shape[0],1)), klm.to_numpy()))

fbeta_cobbdouglas, fhat, fresids = regress(f, Xklm)
fbeta_cobbdouglas
results = np.concatenate((fbeta_cobbdouglas.flatten(), [Eomega], [Edf_dm]))

print("Cobb-Douglas coefficients are: [beta_0, beta_k, beta_l, beta_m] = ", fbeta_cobbdouglas)

Cobb-Douglas coefficients are: [beta_0, beta_k, beta_l, beta_m] = [-0.21136415  0.19574309  0.31153059  0.46819569]
```

```
In [43]: # Example data (replace these with your actual y and yhat values)
# Create a scatter plot for actual vs. predictedZZZZZ
plt.figure(figsize=(10, 6))
plt.scatter(f, fhat, color='blue', label='Predicted', marker='o')
plt.plot([min(f), max(f)], [min(f), max(f)], color='red', linestyle='--', label='Ideal Fit')
# Add Labels and title
plt.xlabel('Actual Values')
plt.ylabel('Predicted Values')
plt.title('f versus a linear regression on f on 1, k, l, m')
plt.legend()
plt.grid()
# Show the plot
plt.show()
```

