Assignment 11: Bayes Prior Distributions

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2020-11-18

```
suppressWarnings({
   library(actuar)
   library(tidyverse)
   wholeest<-function(vec0,funcname,prior=c()){</pre>
     n <-length(vec0)</pre>
     meanv<<-mean(vec0)
     snv=sum((vec0-meanv)^2)
     sumv<<-sum(vec0)</pre>
     sqrv=sum(vec0^2)
     # Bernoulli Distribution
     if (funcname=="Bernoulli"){
       cat(sprintf("\n ---- Bernoulli ----\n"))
       funct<-function(x)(x^sumv)*((1-x)^(n-sumv))</pre>
       phat<-sumv/n
       print(paste ("phat (MOM)=",phat))
       pval=optimize(funct,c(0,1),tol=0.0001,maximum = TRUE)
       print(paste ("phat (MLE):", pval$maximum))
       p = seq(0,1, length=n)
       hat=ks.test(vec0, "pbinom", 1, pval maximum) statistic
       print(paste("kstest result:",hat))
       plot(p, dbeta(p, prior[1]+sumv, n-sumv+prior[2]), ylab="density",
            type ="1", main="For Bernoulli Distribution")
     # Geometric Distribution
     if (funcname=="Geometric"){
       cat(sprintf("\n ---- Geometric ----\n"))
       funct < -function(x)(n*log2(x))+(sumv*log2(1-x))
       pval=optimize(funct,c(0,1),tol=0.0001,maximum = TRUE)
       print(paste ("phat (MLE):", pval$maximum))
       phat=1/((sumv/n)+1)
       print (paste("phat(MOM)=",phat))
       hat=ks.test(vec0, "pgeom", pval$maximum)$statistic
       print(paste("kstest result:",hat))
       p = seq(0,1, length=n)
      plot(p, dbeta(p, prior[1]+n, sumv+prior[2]-1), ylab="density",
           type ="1", main="For Geometric Distribution")
```

```
# Normal Distribution
if (funcname=="Normal"){
  cat(sprintf("\n ---- Normal ----\n"))
 mu=sumv/n
 print (paste("mu (MLE):",mu))
 samu=sum((vec0-mu)^2)
 funct < -function(x) - (n/2) * (log(2*pi*x^2)) + (-1/(2*x^2)) * samu
 pval=optimize(funct,c(0,1000),tol=0.001,maximum = TRUE)
 print(paste ("sigma (MLE):", pval$maximum))
 mu=sumv/n
 sig2=snv/n
 print(paste("mu(MOM)=",mu,"sigmasquare(MOM)=",sig2))
 hat=ks.test(vec0, "pnorm", mu, pval$maximum)$statistic
 print(paste("kstest result:",hat))
  \# x = seq(-5, 10, length=n)
  # odsesq=1/(pval$maximum)^2
  # w=odsesq/(odsesq+(1/prior[2]^2))
  # plot(x, dnorm(x,w*sumv/n+(1-w)*prior[1],1/(odsesq+(1/prior[2]^2))),
  # ylab="density", type ="l")
  # Consider a normal sample of 1000 data points
  # sample <- rnorm(n = 1000, mean = 14, sd = 20)
  sample <- vec0
 r <- 1
 tau <- 6
 mu <- 4
 mean_sample <- mean(sample)</pre>
  # Let's assume alpha and beta for the prior distribution to be 1 and 3
 prior_alpha <- 1
 prior_beta <- 3</pre>
  # Now compute the posterior distribution parameters
  conditional_distribution_mean <- (tau*mu + n*mean_sample) / (tau + n)</pre>
  conditional_distribution_precision <- (tau + n) * r</pre>
  print(paste("The posterior normal distribution parameters (mean, precision): ",
      c(conditional_distribution_mean, conditional_distribution_precision)))
  marginal_distribution_alpha <- prior_alpha + n/2</pre>
 marginal_distribution_beta <- 1 / ((1/prior_beta)</pre>
                                    + 1/2*(sum((sample - mean_sample)**2)))
                               + tau*n*((mean_sample - mu)**2)/2*(tau + n)
 print(paste("The marginal posterior gamma distribution parameters are (alpha, beta): ",
            c(marginal_distribution_alpha, marginal_distribution_beta)))
  # Now using the posterior parameters, let's plot the density of 10K data points
  conditional_joint_distribution <- rnorm(n = 10000,</pre>
```

```
mean = conditional_distribution_mean,
                                       1/sqrt(conditional_distribution_precision))
 marginal joint distribution <- rgamma(n = 10000,
                                     marginal distribution alpha,
                                     marginal distribution beta)
 plot(density(conditional_joint_distribution),
    main = "Conditional Joint Probability Distribution
    for Mean M for Normal Distribution")
 plot(density(marginal_joint_distribution),
    main = "Marginal Joint Probability Distribution
    for Precision R for Normal Distribution")
# Binomial Distribution
if (funcname=="Binomial"){
 cat(sprintf("\n ---- Binomial ----\n"))
 phat=1-(snv/sumv)
 nhat=sumv/(n*phat)
 print(paste("p(MOM)=",phat,"n(MOM)=",nhat))
 nval = (1.0/p_true)*mean(vec0)
 pval=sumv / (length(vec0)*nval)
 print(paste ("p (MLE):", pval))
 theta_hat_func <- function(data) {</pre>
   n1 <- length(data)</pre>
   estimated_p <- (1 / n1) * (sum(data)/n1)
   return(estimated_p)
 }
 theta_hat <- theta_hat_func(vec0)</pre>
 print(paste("kstest result:",theta_hat))
 # Consider a binomial sample of 1000 data points
 sample <- vec0
 # Let's assume alpha and beta for the prior distribution to be 1
 prior_alpha <- 1</pre>
 prior_beta <- 1</pre>
 r < -1
 # Now compute the posterior distribution parameters
 posterior_alpha <- prior_alpha + sum(sample)</pre>
 posterior_beta <- prior_beta + r * n - sum(sample)</pre>
 print(paste("The posterior beta distribution parameters are: ",
             c(posterior_alpha, posterior_beta)))
  # Now using this posterior_alpha and posterior_beta plot the density
  # of 10K data points
```

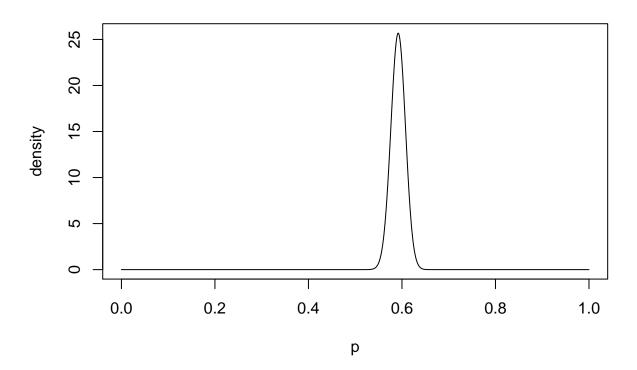
```
posterior_distribution_sample <- rbeta(n = 10000,</pre>
                                          shape1 = posterior_alpha,
                                          shape2 = posterior_beta)
  plot(density(posterior_distribution_sample),
       main = "Posterior Beta Distribution for Binomial Distribution")
# Uniform Distribution
if (funcname=="Uniform"){
  cat(sprintf("\n ---- Uniform ----\n"))
  a=sumv/n-(sqrt(3*snv/n))
 b=sumv/n+(sqrt(3*snv/n))
  print(paste("a (MOM)=",a,"b (MOM)=",b))
  theta_hat_func <- function(data) {</pre>
    estimated_a <- min(data)</pre>
    estimated_b <- max(data)</pre>
    return (c(estimated_a, estimated_b))
  }
  theta_hat <- theta_hat_func(vec0)</pre>
  print(paste("a (MLE)=",theta_hat[1],"b (MLE)=",theta_hat[2]))
 nboot <- 1000
  q_{\text{hat}} \leftarrow q_{\text{unif}}(c(1:n)/(n+1), \text{ theta_hat}[1], \text{ theta_hat}[2])
  DO <- ks.test(vec0, q_hat)$statistic
  D_vec<-NULL
  for(i in 1:nboot){
    x_star <- runif(n, theta_hat[1], theta_hat[2])</pre>
    theta_hat_star <- theta_hat_func(x_star)</pre>
    q_hat_star <- qunif(c(1:n)/(n+1), theta_hat_star[1], theta_hat_star[2])
    D_star <- ks.test(x_star, q_hat_star)$statistic</pre>
    D_vec <- c(D_vec, D_star)</pre>
  p_value <- sum(D_vec > D0)/nboot
  print(paste("kstest result: The p-value is",p_value))
  # Consider a uniform sample of 1000 data points
  sample <- vec0</pre>
  \# Let's assume alpha and WO for the prior distribution to be 1
  prior_w0 <- 1</pre>
 prior_alpha <- 1</pre>
  # Now compute the posterior distribution parameters
  posterior_w0 <- max(c(prior_w0, sample))</pre>
```

```
posterior_alpha <- prior_alpha + n</pre>
 print(paste("The posterior pareto distribution parameters are: ",
              c(posterior_w0, posterior_alpha)))
  # Now using the posterior_wO and posterior_alpha, plot the density
  # of 10K data points
 posterior_distribution_sample <- rpareto(n = 10000,</pre>
                                           posterior_w0,
                                           posterior_alpha)
 plot(density(posterior_distribution_sample),
      main = "Posterior Pareto Distribution for Uniform Distribution")
# Poisson Distribution
if (funcname=="Poisson"){
 cat(sprintf("\n ---- Poisson ----\n"))
 lambdahat=sumv/n
 print(paste("lambda (MOM)=",lambdahat))
 theta_hat = meanv
 print(paste("lambda (MLE)=",theta_hat))
 q_hat \leftarrow qpois(c(1:n)/(n+1), theta_hat)
 D0 <- ks.test(vec0, q_hat)$statistic
 D_vec<-NULL
 nboot = 1000
 for(i in 1 : nboot){
    x_star <- rpois(n, theta_hat)</pre>
   theta_hat_star <- mean(x_star)</pre>
   q_hat_star <- qpois(c(1:n)/(n+1), theta_hat_star)</pre>
   D_star <- ks.test(x_star, q_hat_star)$statistic</pre>
   D_vec <- c(D_vec, D_star)</pre>
 p_value <- sum(D_vec > D0)/nboot
 print(paste("kstest result: The p-value is",p_value))
 sample <- vec0
  # Let's assume alpha and beta for the prior distribution to be 1
 prior_alpha <- 1</pre>
 prior_beta <- 1</pre>
  # Now compute the posterior distribution parameters
 posterior_alpha <- prior_alpha + sum(sample)</pre>
 posterior_beta <- 1/(1/prior_beta + n)</pre>
```

```
print(paste("The posterior gamma distribution parameters are: ",
              c(posterior_alpha, posterior_beta)))
  # Now using the posterior_alpha and posterior_beta, plot the density
  # of 10K data points
 posterior_distribution_sample <- rgamma(n = 10000, posterior_alpha, posterior_beta)</pre>
 plot(density(posterior distribution sample),
       main = "Posterior Gamma Distribution for Poisson Distribution")
# Exponential Distribution
if (funcname=="Exponential"){
  cat(sprintf("\n ---- Exponential ----\n"))
 beta=sumv/n
 print(paste("beta (MOM)=",beta))
 theta_hat_func <- function(data) {</pre>
    estimated_theta <- mean(data)</pre>
   return(estimated_theta)
 theta_hat <- theta_hat_func(vec0)</pre>
 print(paste("beta (MLE)=", theta_hat))
 nboot <- 1000
  q_{n+1} \leftarrow q_{n+1}(c(1:n) / (n+1), beta)
 D_0 <- ks.test(vec0, q_hat)$statistic</pre>
 D_vec<-NULL
 for(i in 1:nboot){
    x_star <- rexp(n, beta)</pre>
    theta_hat_star <- theta_hat_func(x_star)</pre>
   q_hat_star <- qexp(c(1:n)/(n+1), theta_hat_star)</pre>
   D_star <- ks.test(x_star, q_hat_star)$statistic</pre>
   D_vec <- c(D_vec, D_star)</pre>
 p_value <- sum(D_vec > D_0)/nboot
 print(paste("kstest result: The p-value is",p_value))
  # Consider a normal sample of 1000 data points
 sample <- vec0</pre>
  # Let's assume alpha and beta for the prior distribution to be 1
 prior_alpha <- 1</pre>
 prior_beta <- 1</pre>
  # Now compute the posterior distribution parameters
 posterior_alpha <- prior_alpha + n</pre>
```

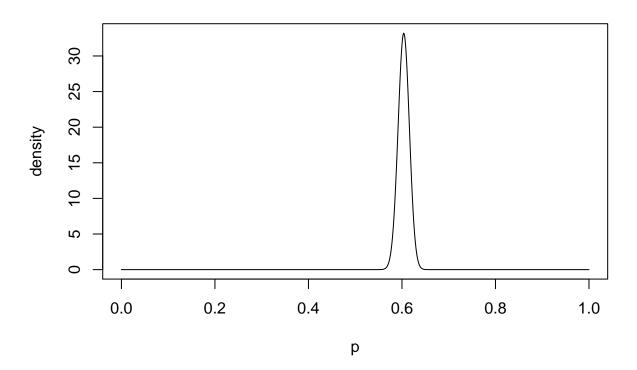
```
posterior_beta <- 1/(1/prior_beta + sum(sample))</pre>
       print(paste("The posterior gamma distribution parameters are:",
                   c(posterior_alpha, posterior_beta)))
       # Now using the posterior parameters, let's plot the density of 10K data points
       posterior_distribution_sample <- rgamma(n = 10000, posterior_alpha, posterior_beta)</pre>
       plot(density(posterior distribution sample),
            main = "Posterior Gamma Distribution for Exponential Distribution")
     }
   }
   wholeest(rbinom(1000,1,0.6), "Bernoulli", c(1,1))
   wholeest(rgeom(1000,0.6), "Geometric", c(1,1))
   wholeest(rnorm(1000,3,5),"Normal",c(5,15))
   n_true<-1000
   p_true<-0.76
   wholeest(rbinom(n_true, 1, p_true), "Binomial")
   wholeest(runif(1000, 0, 15), "Uniform")
   wholeest(rpois(1000,4),"Poisson")
   wholeest(rexp(1000,8), "Exponential")
})
## Attaching package: 'actuar'
## The following object is masked from 'package:grDevices':
##
##
      cm
## -- Attaching packages ------ tidyverse 1.3.0 --
## v ggplot2 3.3.2 v purrr 0.3.4
## v tibble 3.0.3 v dplyr 1.0.2
## v tidyr 1.1.2 v stringr 1.4.0
## v readr 1.3.1
                    v forcats 0.5.0
## -- Conflicts ----- tidyverse_conflicts() --
## x dplyr::filter() masks stats::filter()
## x dplyr::lag()
                   masks stats::lag()
##
## ---- Bernoulli ----
## [1] "phat (MOM) = 0.592"
## [1] "phat (MLE): 0.591994411491554"
## [1] "kstest result: 0.592"
```

For Bernoulli Distribution



```
##
## ---- Geometric ----
## [1] "phat (MLE): 0.603499416969091"
## [1] "phat(MOM) = 0.603500301750151"
## [1] "kstest result: 0.603499416969091"
```

For Geometric Distribution



##

##

---- Normal ----

```
## [1] "mu (MLE): 3.14207941950593"

## [1] "sigma (MLE): 5.06722078024272"

## [1] "mu(MOM)= 3.14207941950593 sigmasquare(MOM)= 25.6763684876148"

## [1] "kstest result: 0.0225458129442648"

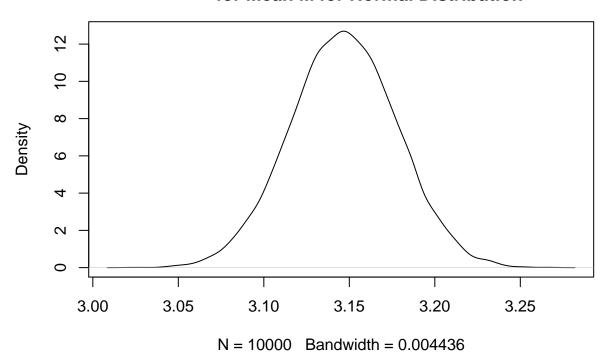
## [1] "The posterior normal distribution parameters (mean, precision): 3.1471962420536"

## [2] "The posterior normal distribution parameters (mean, precision): 1006"

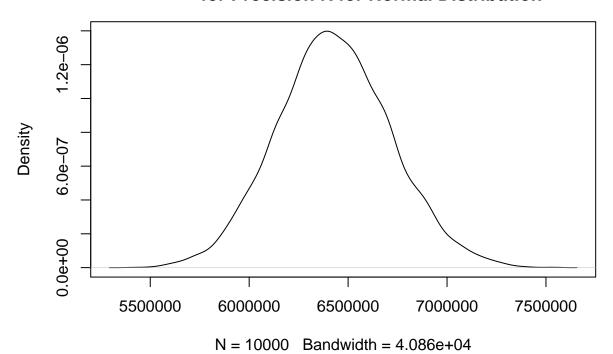
## [1] "The marginal posterior gamma distribution parameters are (alpha, beta): 501"

## [2] "The marginal posterior gamma distribution parameters are (alpha, beta): 7.78906126810561e-05"
```

Conditional Joint Probability Distribution for Mean M for Normal Distribution



Marginal Joint Probability Distribution for Precision R for Normal Distribution

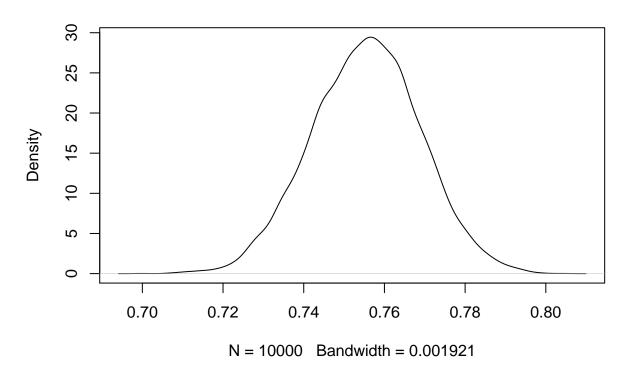


```
##
## ---- Binomial ----
## [1] "p(MOM)= 0.756 n(MOM)= 1"
## [1] "p (MLE): 0.76"
## [1] "kstest result: 0.000756"
## [1] "The posterior beta distribution parameters are: 757"
```

[2] "The posterior beta distribution parameters are:

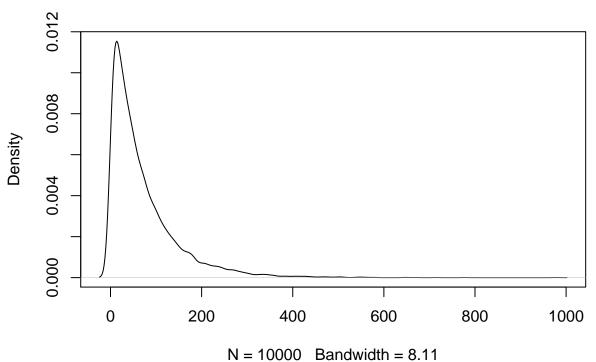
245"

Posterior Beta Distribution for Binomial Distribution



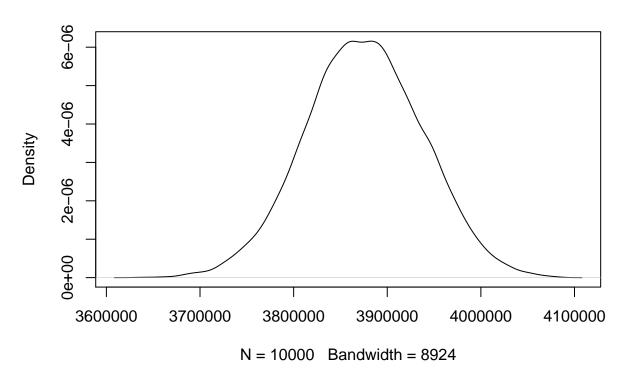
```
##
## ---- Uniform ----
## [1] "a (MOM)= 0.362370456091329 b (MOM)= 15.0829414918765"
## [1] "a (MLE)= 0.0477080466225743 b (MLE)= 14.9909275991376"
## [1] "kstest result: The p-value is 0.087"
## [1] "The posterior pareto distribution parameters are: 14.9909275991376"
## [2] "The posterior pareto distribution parameters are: 1001"
```

Posterior Pareto Distribution for Uniform Distribution



```
##
## ---- Poisson ----
## [1] "lambda (MOM)= 3.871"
## [1] "lambda (MLE)= 3.871"
## [1] "kstest result: The p-value is 0.649"
## [1] "The posterior gamma distribution parameters are: 3872"
## [2] "The posterior gamma distribution parameters are: 0.000999000999000999"
```

Posterior Gamma Distribution for Poisson Distribution



```
##
## ---- Exponential ----
## [1] "beta (MOM)= 0.130028004109076"
## [1] "beta (MLE)= 0.130028004109076"
## [1] "kstest result: The p-value is 0.399"
## [1] "The posterior gamma distribution parameters are: 1001"
## [2] "The posterior gamma distribution parameters are: 0.00763195628903526"
```

Posterior Gamma Distribution for Exponential Distribution

