

$$X \sim \frac{1}{2} e^{-|x-\theta|} \quad \underline{\text{Q1}} \quad \underline{\text{MLE}} \quad \text{(AS 12)}$$

Likelihood:

$$L(x_1, \dots, x_n) = \left(\frac{1}{2}\right)^n e^{-\sum_{i=1}^n |x_i - \theta|}$$

Log-likelihood:

$$n \ln(1/2) - \sum_{i=1}^n |x_i - \theta|$$

⇒ The optimization becomes

$$\arg \max_{\theta} \left[ n \ln(1/2) - \sum_{i=1}^n |x_i - \theta| \right]$$

$$= \arg \min_{\theta} \left( \sum_{i=1}^n |x_i - \theta| \right)$$

We know that,

$$\frac{\partial |x|}{\partial x} = \frac{\partial \sqrt{x^2}}{\partial x} = x(x^2)^{-1/2} = \frac{x}{|x|} = \text{sgn}(x)$$

⇒ Optimizing with respect to  $\theta$  by taking derivative with respect to  $\theta$  and equating with zero, we get

$$\sum_{i=1}^n \text{sgn}(x_i - \theta) = 0$$

↳ condition



Case 1 :  $n$  is odd

If we choose  $\hat{\theta} = \text{median}(x_1, \dots, x_n)$ , then we get  ~~$\frac{n-1}{2}$~~   $\frac{n-1}{2}$  cases where  $x_i < \theta$  & other  $\frac{n-1}{2}$  cases where  $x_i > \theta$ .

$\Rightarrow \hat{\theta}$  satisfies the condition for MLE estimator of  $\theta$ .

Case 2 :  $n$  is even

Now we cannot choose one  $x_i$  to satisfy the condition for MLE estimator. We can still perform the minimization by assuming that  $x_1, \dots, x_n$  are ordered ( $x_1 \leq x_2 \leq \dots \leq x_n$ ) and then choosing either  $x_{n/2}$  or  $x_{\frac{(n+1)}{2}}$ .

$\Rightarrow$  Median is the MLE estimator for the given distribution.



AS 12 Q2 For L1 loss,  $L(\theta, \hat{\theta}) = |\theta - \hat{\theta}|$   
we choose  $\hat{\theta}$  to minimize

$$\begin{aligned} \eta(\hat{\theta}|x) &= \int |\theta - \hat{\theta}| f(\theta|x) d\theta \\ &= \int_{-\infty}^{\hat{\theta}} (\hat{\theta} - \theta) f(\theta|x) d\theta + \int_{\hat{\theta}}^{\infty} (\theta - \hat{\theta}) f(\theta|x) d\theta \quad \text{--- (1)} \end{aligned}$$

Taking derivative of  $\eta(\hat{\theta}|x)$  with respect to  $\hat{\theta}$  & setting it equal to zero:

$$\begin{aligned} \int_{-\infty}^{\hat{\theta}} f(\theta|x) d\theta &= \int_{\hat{\theta}}^{\infty} f(\theta|x) d\theta \\ \Rightarrow 2 \int_{-\infty}^{\hat{\theta}} f(\theta|x) d\theta &= \int_{-\infty}^{\infty} f(\theta|x) d\theta = 1 \\ \Rightarrow \int_{-\infty}^{\hat{\theta}} f(\theta|x) d\theta &= \frac{1}{2} \end{aligned}$$

&  $\hat{\theta}$  is the median of  $f(\theta|x)$ .

The second derivative of (1) with respect to  $\hat{\theta}$  gives  $2 f(\hat{\theta}|x) \geq 0$ .

$\Rightarrow$  The original function is convex & thus the median definitely corresponds to the minimum.

Hence, proved.