

Moment Generating Function  $\div$  Let  $X$  be a r.v., then moment generating function (MGF) of  $X$  is defined as

$$M_X(t) = E(e^{tx}) = \sum_{x \in R_X} e^{tx} ; |t| < h \text{ where } h > 0 \text{ (small) and } t \in \mathbb{R}$$

Note: Any order non-central moment can be generated by the help of MGF.

$$\boxed{\frac{d^n M_X(t)}{dt^n} \Big|_{t=0} = E(X^n)} \quad - (1)$$

Q: Find the MGF of Binomial dist<sup>n</sup>?

Sol<sup>n</sup>:  $E(e^{tx}) = \sum_{x \in R_X} e^{tx} \cdot p_X(x)$

$$M_X(t) = \sum_{x=0}^n e^{tx} \cdot \binom{n}{x} p^x q^n$$

$$= \sum_{x=0}^n e^{tx} \cdot \binom{n}{x} p^x q^n$$

$$= \sum_{x=0}^n \binom{n}{x} (pe^t)^x q^n$$

$$\boxed{M_X(t) = (pe^t + q)^n}$$

Q1

Find the mean for Binomial dist<sup>n</sup> from its MGF?

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sol<sup>n</sup>  
 $M_X(t) = (pe^t + q)^n$ ; where  $q = 1-p$

We know that first non-central moment is its mean.

Put  $n=1$  in ①

$$\left. \frac{dM_X(t)}{dt} \right|_{t=0} = E(X)$$

$$\begin{aligned} \Rightarrow E(X) &= \left. \frac{d}{dt} (pe^t + q)^n \right|_{t=0} \\ &= n(pe^t + q)^{n-1} \cdot pe^t \Big|_{t=0} \\ &= n(pe^0 + q)^{n-1} \cdot pe^0 \\ &= n(p \cdot 1 + q)^{n-1} \cdot p \cdot 1 \\ &= n(p+q)^{n-1} \cdot p \\ &= n \cdot 1 \cdot p \end{aligned}$$

$$\boxed{E(X) = np} \quad \#$$



Note:- pmf of  $\text{Bin}(n, p)$  is

$$p_X(x) = \binom{n}{x} p^x q^{n-x}; \quad x=0, 1, 2, \dots, n$$

where  $n \rightarrow$  no of trials.

If take  $n=1$  then  $\text{Bin}(n, p)$  tends to Bernoulli distribution. It is denoted as  $\text{Ben}(p)$  ( $\because n=1$ )

so pmf is given as

$$p_X(x) = p^x q^{1-x}; \quad x=0, 1$$

$$E(X) = p$$

$$(\because np = 1 \cdot p)$$

$$V(X) = pq$$

$$(\because npq = 1 \cdot p \cdot q)$$

$$M_X(t) = (pe^t + q)$$

### Poisson Distribution:

If  $n$  is very large i.e.  $n \rightarrow \infty$   
and  $p$  is very small i.e.  $p \rightarrow 0$  s.t.  $np \rightarrow \lambda$  (constant)

then under these conditions, it is suggested to use Poisson dist<sup>n</sup> instead of  $\text{Bin}(n, p)$ .

Proof:  $P_X(x) = \binom{n}{x} p^x q^{n-x} ; x = 0, 1, 2, 3, \dots, n$

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$$= \frac{n!}{x! (n-x)!} p^x (1-p)^{n-x} ; x = 0, 1, 2, 3, \dots, n$$

$$= \frac{n(n-1)(n-2)\dots(n-(x-1)) \cancel{(n-x)!} p^x (1-p)^{n-x}}{x! \cancel{(n-x)!}}$$

$$= \frac{n(n-1)(n-2)\dots(n-(x-1)) \left(\frac{\lambda}{n}\right)^x \left(1-\frac{\lambda}{n}\right)^{n-x}}{x!} \because np \rightarrow \lambda \Rightarrow p \rightarrow \frac{\lambda}{n}$$

$$= \frac{n \cdot n(1-\frac{1}{n}) \cdot n(1-\frac{2}{n}) \dots n(1-\frac{(x-1)}{n}) \left(\frac{\lambda}{n}\right)^x \left(1-\frac{\lambda}{n}\right)^{n-x}}{x!}$$

$$= \frac{\underbrace{n \cdot n \cdot n \dots n}_{x \text{ times}} \cdot \frac{(1-\frac{1}{n})(1-\frac{2}{n}) \dots (1-\frac{(x-1)}{n})}{x!} \cdot \frac{\lambda^x}{n^x} \left(1-\frac{\lambda}{n}\right)^{n-x}}{x!}$$

$$P_X(x) = \frac{n^x \cdot \frac{(1-\frac{1}{n})(1-\frac{2}{n}) \dots (1-\frac{(x-1)}{n})}{x!} \cdot \frac{\lambda^x}{n^x} \left(1-\frac{\lambda}{n}\right)^{n-x}}{x!}$$

$$\lim_{n \rightarrow \infty} P_X(x) = \lim_{n \rightarrow \infty} \left[ \frac{(1-0)(1-0) \dots (1-0) \cdot \frac{\lambda^x}{x!} \left(1-\frac{\lambda}{n}\right)^{n-x}}{x!} \right] \left( \because n \rightarrow \infty \right)$$

$$= \frac{\lambda^x}{x!} \lim_{n \rightarrow \infty} \left(1-\frac{\lambda}{n}\right)^{n-x} = e^{-\lambda} \frac{\lambda^x}{x!}$$

$$\text{Let } y = \left(1 - \frac{\lambda}{n}\right)^{n-x}$$

Take log both sides,

$$\ln y = (n-x) \ln\left(1 - \frac{\lambda}{n}\right) \quad \because \ln(1-x) = -x - \frac{x^2}{2} - \frac{x^3}{3} \dots$$

$$= (n-x) \left[ -\left(\frac{\lambda}{n}\right) - \left(\frac{\lambda}{n}\right)^2 \cdot \frac{1}{2} - \left(\frac{\lambda}{n}\right)^3 \cdot \frac{1}{3} \dots \right]$$

$$\ln y = -(n-x) \frac{\lambda}{n} - \frac{\lambda^2 (n-x)}{n^2 \cdot 2} - \frac{\lambda^3 (n-x)}{n^3 \cdot 3} \dots$$

$$\Rightarrow \lim_{n \rightarrow \infty} \ln y = \lim_{n \rightarrow \infty} \left( -\frac{\lambda \cdot n}{n} + \frac{x \lambda}{n} \right) - \frac{\lambda^2 (n-x)}{n^2 \cdot 2} - \frac{\lambda^3 (n-x)}{n^3 \cdot 3} \dots$$

$$\Rightarrow \lim_{n \rightarrow \infty} \ln y = -\lambda + 0 - 0 - 0 \dots$$

$$\Rightarrow y = e^{-\lambda}$$

Thus

$$p_X(x) = \frac{e^{-\lambda} \lambda^x}{x!}, \quad x = 0, 1, 2, 3, \dots$$

This is p.m.f of Poisson dist<sup>n</sup> and denoted by  $P(\lambda)$ .



Q. ① Verify that it is proper p.m.f

② find mean and Variance for poisson dist<sup>n</sup>.

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③ Find mgf of poisson dist<sup>n</sup>.

Sol<sup>n</sup> (1)  $\therefore p_X(x) = \frac{e^{-\lambda} \lambda^x}{x!}; x=0, 1, 2, \dots$

Claim:  $\sum_{x=0}^{\infty} p_X(x) = 1$

$$\begin{aligned} \text{So, } \sum_{x=0}^{\infty} \frac{e^{-\lambda} \lambda^x}{x!} &= e^{-\lambda} \left[ \sum_{x=0}^{\infty} \frac{\lambda^x}{x!} \right] \\ &= e^{-\lambda} \left[ \frac{\lambda^0}{0!} + \frac{\lambda^1}{1!} + \frac{\lambda^2}{2!} + \frac{\lambda^3}{3!} + \dots \right] \\ &= e^{-\lambda} \left[ 1 + \frac{\lambda}{1!} + \frac{\lambda^2}{2!} + \frac{\lambda^3}{3!} + \dots \right] \\ &= e^{-\lambda} \cdot e^{\lambda} \quad \left( \because 1 + \frac{x}{1!} + \frac{x^2}{2!} + \dots = e^x \right) \\ &= e^{-\lambda + \lambda} \\ &= e^0 \\ &= 1 \quad \# \end{aligned}$$