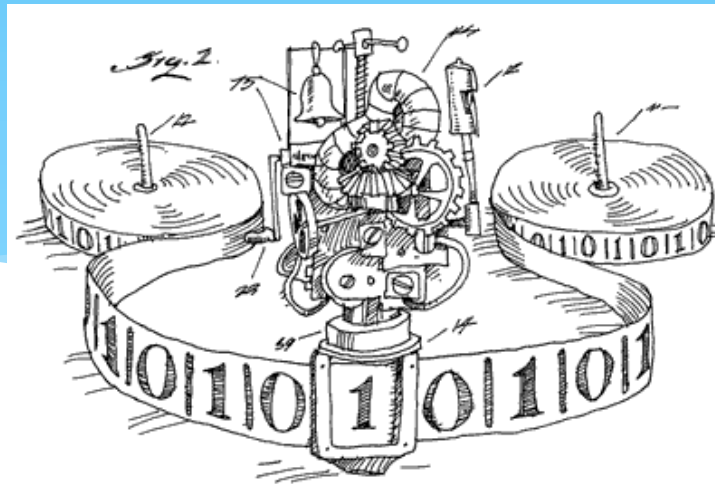


# EECS 376: Foundations of Computer Science

Ali Movaghar

Lecture 22



# Agenda

- \* **Last week:** Quicksort, Skip Lists
  - \* Randomized algorithms that perform well in expectation
  - \* But what's the probability of getting a good result?
- \* **Today:** Concentration Bounds
  - \* Variance and Chebyshev's inequality
  - \* Chernoff-Hoeffding bounds
  - \* Examples: Flipping coins and Polling

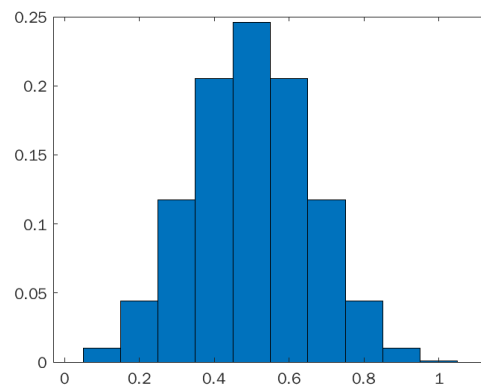
several randomized algorithm analyses use Chernoff bound + Union bound  
EECS 572: Randomness and Computation for a “real” application

# How many heads?

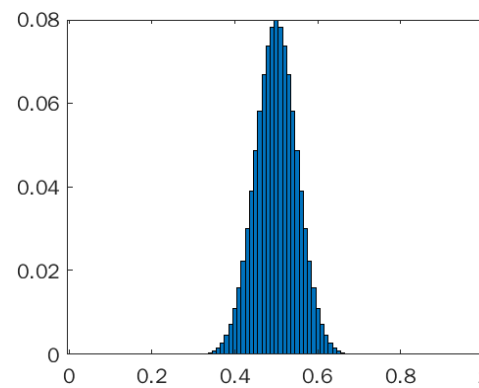
- \* We want to determine if a coin is fair or not.
- \* **Q:** How suspicious would we be if we flip it  $n$  times and see  $k$  heads, for the following values of  $n$  and  $k$ ?
  - \*  $n = 100, k = 51$
  - \*  $n = 10,000, k = 5,100$
  - \*  $n = 1,000,000, k = 510,000$
- \* Need to estimate  $\Pr[X \geq k]$ , where  $X$  is the number of heads after flipping a fair coin  $n$  times!

# Law of Large Numbers

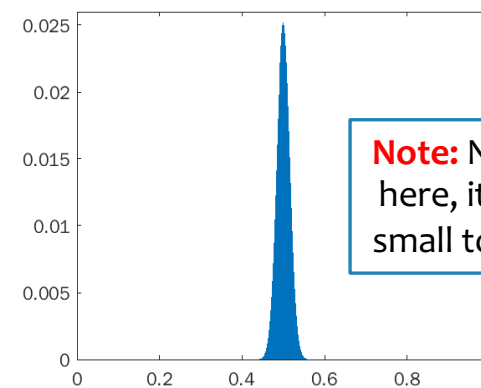
- \* **(Informal)** If  $X_1, X_2, \dots$  are *independent, identically distributed (i.i.d.)* RVs w/ expectation  $\mu$ , then  $\frac{1}{n} \sum_{i=1}^n X_i$  converges to  $\mu$  (a constant) as  $n \rightarrow \infty$ .
- \* **Example:** The fraction of heads obtained when flipping a fair coin  $n$  times converges to  $1/2$  as  $n \rightarrow \infty$ .  
The graphs plot  $\Pr \left[ \frac{1}{n} \sum_{i=1}^n X_i = a \right]$ ; bars are for possible values of  $a$ .



$n = 10$



$n = 100$



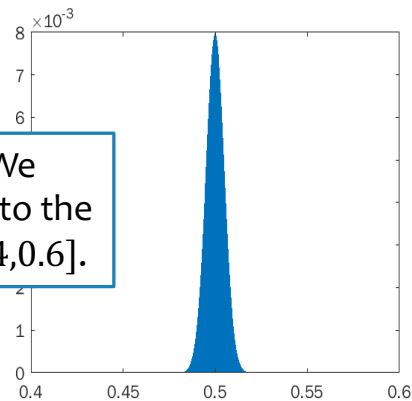
$n = 1000$

**Note:** Nothing's 0 here, it's just too small to show up.

# Limitations of LLN

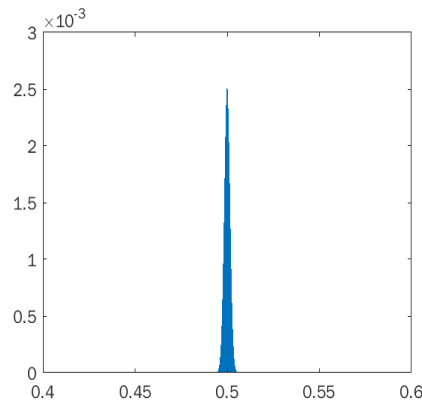
**(Informal)** If  $X_1, X_2, \dots$  are *independent, identically distributed (i.i.d.)* RVs w/ expectation  $\mu$ , then  $\frac{1}{n} \sum_{i=1}^n X_i$  converges to  $\mu$  (a constant) as  $n \rightarrow \infty$ .

- \* LLN says distribution of sum is “concentrated” around its expectation as  $n \rightarrow \infty$ . (However, it doesn’t say how quickly it happens or what the distribution looks like.)
- \* **Example:** The probability of seeing at least 0.51-fraction of heads when flipping a *fair* coin  $n$  times goes to zero as  $n \rightarrow \infty$ . How fast? Not clear.

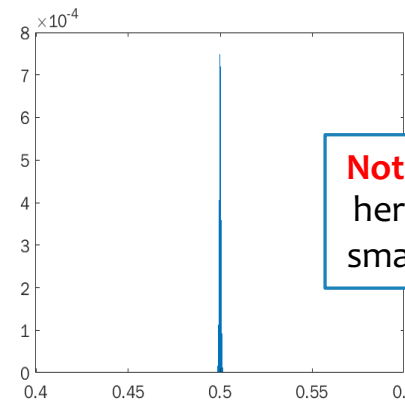


**Note:** We zoomed into the region  $[0.4, 0.6]$ .

$n = 10,000$



$n = 100,000$



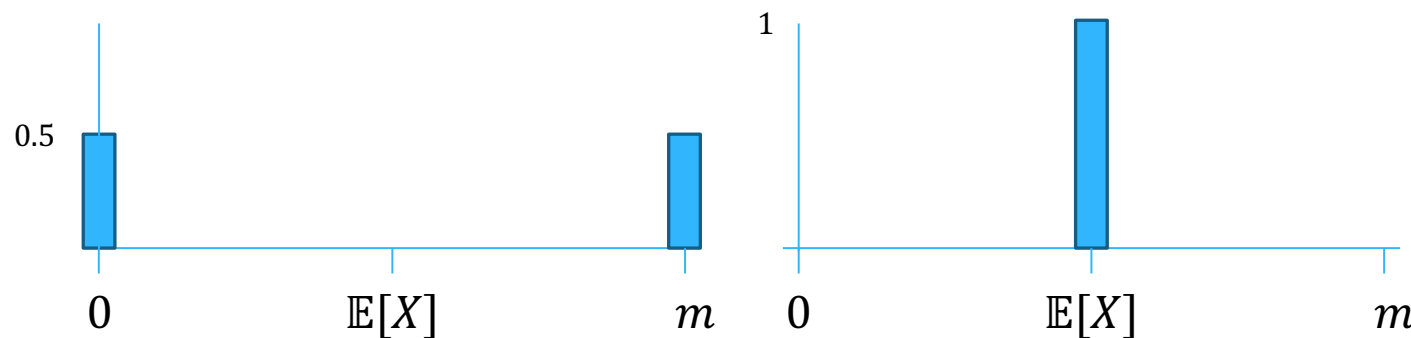
$n = 1,000,000$

**Note:** Nothing's 0 here, it's just too small to show up.

# Variance

- \* The **variance** of a random variable  $X$  is the average squared-distance of  $X$  from its mean, i.e.,  

$$\text{Var}(X) = \mathbb{E}[(X - \mathbb{E}[X])^2] = \mathbb{E}[X^2] - \mathbb{E}[X]^2$$
- \* The **standard deviation** is  $\text{SD}(X) = \sqrt{\text{Var}(X)}$  (it's an upper bound on the average distance of  $X$  from  $\mathbb{E}[X]$ ).



# Example

$$\text{Var}(X) = \mathbb{E}[X^2] - \mathbb{E}[X]^2$$

- \* You have a biased coin that has probability  $p$  of heads,  $1 - p$  of tails. For a single flip of the coin:
- \* Let  $X_i = \begin{cases} 1 & \text{if flip } i \text{ is Heads} \\ 0 & \text{if flip } i \text{ is Tails} \end{cases}$
- \* Then  $\mathbb{E}[X_i] = \Pr[X_i = 1] = p$ ;
- \*  $\mathbb{E}[X_i^2] = 0 \cdot \Pr[X_i^2 = 0] + 1 \cdot \Pr[X_i^2 = 1] = p$
- \*  $\text{Var}(X_i) = \mathbb{E}[X_i^2] - \mathbb{E}[X_i]^2 = p - p^2 = p(1 - p)$
- \* What about for  $n$  flips?

# Variance of Sum of Independent RVs

$$\mathbf{Var}(X) = \mathbb{E}[X^2] - \mathbb{E}[X]^2$$

- \* **Fact:** For a sum of independent RVs  $X = X_1 + \dots + X_n$ , we have  $\mathbf{Var}(X) = \mathbf{Var}(\sum X_i) = \sum \mathbf{Var}(X_i)$ .
- \* For one flip of a coin with probability  $p$  of heads, we saw that:
  - \*  $\mathbb{E}[X_i] = p$
  - \*  $\mathbf{Var}(X_i) = p(1 - p)$
- \* For  $n$  flips, we have:
  - \*  $\mathbb{E}[X] = \sum \mathbb{E}[X_i] = np$
  - \*  $\mathbf{Var}(X) = \sum \mathbf{Var}(X_i) = np(1 - p)$



# Chebyshev's Inequality

$$\text{Var}(X) = \mathbb{E}[X^2] - \mathbb{E}[X]^2$$

- \* **(Recall) Markov's Inequality:** For a non-negative RV  $X$  and  $a > 0$ :

$$\Pr[X \geq a] \leq \mathbb{E}[X]/a$$

- \* **Chebyshev's Inequality:** For any RV  $X$  and  $a > 0$ :

$$\Pr[|X - \mathbb{E}[X]| \geq a] \leq \text{Var}(X)/a^2$$

- \* **Proof:** square both sides and apply Markov's ineq.

$$\Pr[|X - \mathbb{E}[X]| \geq a] = \Pr[(X - \mathbb{E}[X])^2 \geq a^2] \quad (\text{sq. both sides})$$

$$\leq \frac{\mathbb{E}[(X - \mathbb{E}[X])^2]}{a^2} \quad (\text{apply Markov})$$

$$\leq \frac{\text{Var}(X)}{a^2} \quad (\text{defn. of Variance})$$

# Chebyshev's Inequality

$$\mathbf{Var}(X) = \mathbb{E}[X^2] - \mathbb{E}[X]^2$$

- \* **Chebyshev's Inequality:** For any RV  $X$  and  $a > 0$ :  
$$\Pr[|X - \mathbb{E}[X]| \geq a] \leq \mathbf{Var}(X)/a^2$$
- \* **Example:** What's the probability of getting  $\leq 49\%$  or  $\geq 51\%$  heads in  $n$  tosses of a fair coin?
- \* 
$$\Pr\left[\left|X - \frac{n}{2}\right| \geq 0.01n\right] \leq \frac{\mathbf{Var}(X)}{(0.01n)^2} = 10,000 \cdot \frac{n \cdot 1/4}{n^2} = \frac{2,500}{n}$$
  - \*  $n = 10,000 \Rightarrow \Pr[\text{deviating by } 1\%] \leq 1/4$
  - \*  $n = 1,000,000 \Rightarrow \Pr[\text{deviating by } 1\%] \leq 1/400$

# Chebyshev's Inequality

For any RV  $X$  and  $a > 0$ :

$$\Pr[|X - \mathbb{E}[X]| \geq a] \leq \mathbf{Var}(X)/a^2$$

$$\mathbf{Var}(X) = \mathbb{E}[X^2] - \mathbb{E}[X]^2$$

- \* For a sum of **i.i.d.** RVs  $X = X_1 + \dots + X_n$ , we are often interested in  $\frac{1}{n}X$  rather than  $X$  itself, since  $\mathbb{E}\left[\frac{1}{n}X\right] = \mathbb{E}[X_i]$  does not depend on  $n$  (unlike  $\mathbb{E}[X] = n\mathbb{E}[X_i]$ ).
- \* **Fact:** For a constant  $c$ ,  $\mathbf{Var}(cX) = c^2\mathbf{Var}(X)$ .
- \* **Chebyshev (Alternative):** For a sum of **i.i.d.**  $X = X_1 + \dots + X_n$ :

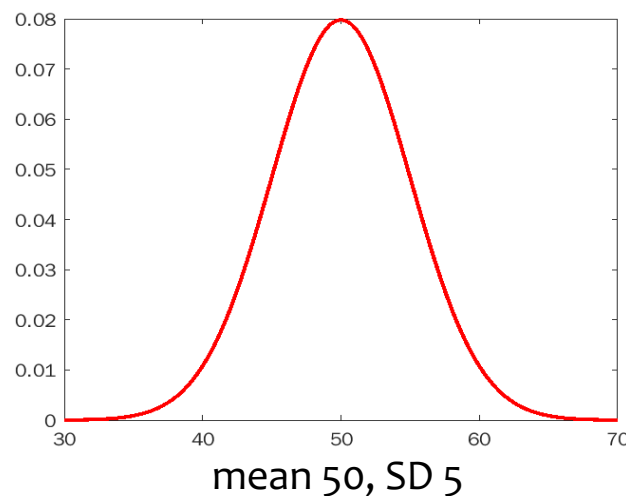
$$\Pr\left[\left|\frac{1}{n}X - \mathbb{E}[X_i]\right| \geq \varepsilon\right] \leq \frac{\mathbf{Var}(X_i)}{\varepsilon^2 n}$$

- \* **Example:** What's the probability of getting  $\leq 49\%$  or  $\geq 51\%$  heads in  $n$  tosses of a fair coin?

$$\Pr\left[\left|\frac{1}{n}X - \frac{1}{2}\right| \geq 0.01\right] \leq \frac{\mathbf{Var}(X_i)}{0.01^2 n} = 10,000 \frac{1/4}{n} = \frac{2,500}{n}$$

# Normal Distribution

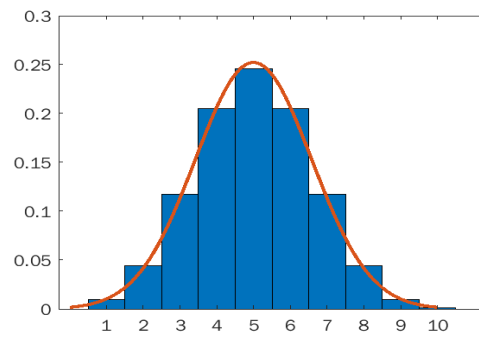
- \* A **normal distribution** has a *bell-curve* shape and is characterized by two parameters, *mean* and *standard deviation*.
  - \* **Examples:** Height, exam scores, measurement error, are “normal-like”...
- \* **66-95-99.7 rule:**  $\approx 66 / 95 / 99.7\%$  of the area under the curve (i.e., probability) is within 1 / 2 / 3 SD from the mean, respectively



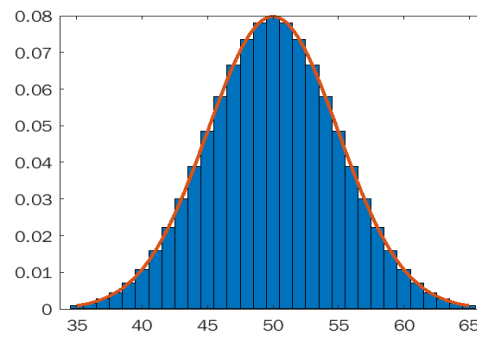
**Note:** Distribution is from  $-\infty$  to  $\infty$  and nothing's 0 here.

# Central Limit Theorem

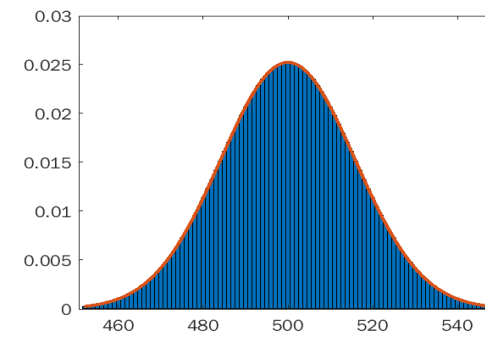
- \* **(Informal)** For large  $n$ , the sum  $X = X_1 + \dots + X_n$  of  $n$  **i.i.d.** RVs is “close” to a normal distribution with mean  $= \mathbb{E}[X]$  and SD  $= \mathbf{SD}(X)$
- \* **Example:** The number of heads after flipping a fair coin  $n$  times, is “close” to a normal distribution with mean  $n/2$  and standard deviation  $\sqrt{n}/2$ .
- \* **Q:** How suspicious are we if we see 510,000 heads after 1,000,000 tosses?



$n = 10$   
mean 5, SD 1.58



$n = 100$   
mean 50, SD 5



$n = 1000$   
mean 500, SD 15.8

# Chernoff-Hoeffding Bounds

(Usually tighter than Chebyshev)

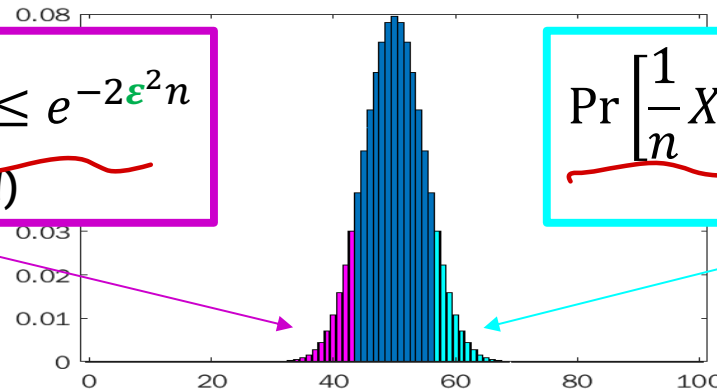
- \* If  $X = X_1 + X_2 + \dots + X_n$  is the sum of  $n$  *i.i.d.* RVs with each  $X_i \in [0, 1]$ , then, for any  $\epsilon > 0$ :

$$\Pr \left[ \frac{1}{n} X \leq \mathbb{E}[X_i] - \epsilon \right] \leq e^{-2\epsilon^2 n}$$

(lower tail bound)

$$\Pr \left[ \frac{1}{n} X \geq \mathbb{E}[X_i] + \epsilon \right] \leq e^{-2\epsilon^2 n}$$

(upper tail bound)



**Example:** If we flip a fair coin  $n = 1,000,000$  times, the probability we see  $\geq (50 + 1)\%$  heads is at most  $e^{-2(0.01)^2 \cdot 1,000,000} = e^{-200}$ !

# Chernoff-Hoeffding Bounds

## Proof Sketch

- \*  $X = \sum_{i=1}^n X_i$  is the sum of  $n$  independent indicators with  $E[X_i] = p$ .
- \*  $\Pr(X - \mathbb{E}(X) \geq t) = \Pr(e^{s(X - \mathbb{E}(X))} \geq e^{st})$  for any  $s > 0$  we wish.

$$\leq \frac{\mathbb{E}(e^{s(X - \mathbb{E}(X))})}{e^{st}} \quad (\text{Markov})$$

$$= \frac{\mathbb{E}(e^{s(X_1 + \dots + X_n - \mathbb{E}(X_1) - \dots - \mathbb{E}(X_n))})}{e^{st}} \quad (\text{Defn of } X)$$

$$= \frac{\mathbb{E}(\prod_{i=1}^n e^{s(X_i - \mathbb{E}(X_i))})}{e^{st}}$$

"Hoeffding's Lemma"

↓

$$= \frac{\prod_{i=1}^n \mathbb{E}(e^{s(X_i - \mathbb{E}(X_i))})}{e^{st}} \leq \frac{\left(\frac{s^2}{8}\right)^n}{e^{st}} = \exp(-st + s^2 n/8)$$

$$= \exp(-2t^2/n). \quad (\text{Choose } s = 4t/n.)$$

# Hoeffding's Lemma

Let  $X$  be a real random variable such that  $X \in [0, 1]$  almost surely. Then for any real number of  $s$ , we have:

$$\mathbb{E}[e^{s(X-E[X])}] \leq \exp\left(\frac{1}{8}s^2\right)$$

The above inequality is proved using the convexity of exponential functions and the arithmetic and geometric means (AM-GM) inequality.



# Some Useful Inequalities

- \* Let  $X$  be a random variable. Then:
  - \*  $\Pr[X \geq a] \geq \Pr[X > a]$  and  $\Pr[X \leq a] \geq \Pr[X < a]$ 
    - \* Why?  $\Pr[X \geq a] = \Pr[X > a] + \Pr[X = a]$
  - \*  $\Pr[X \geq a] = 1 - \Pr[X < a] \geq 1 - \Pr[X \leq a]$
  - \*  $\Pr[X \leq a] = 1 - \Pr[X > a] \geq 1 - \Pr[X \geq a]$

- \* Let  $a \leq b$ . Then:



- \*  $\Pr[X \geq a] \geq \Pr[X \geq b]$  and  $\Pr[X \leq a] \leq \Pr[X \leq b]$ 
  - \* Why?  $\Pr[X \geq a] = \Pr[X \geq b] + \Pr[b > X \geq a]$

# Union Bound

$$\Pr\left[\frac{1}{n}X \leq \mathbb{E}[X_i] - \epsilon\right] \leq e^{-2\epsilon^2 n}$$

(lower tail bound)

$$\Pr\left[\frac{1}{n}X \geq \mathbb{E}[X_i] + \epsilon\right] \leq e^{-2\epsilon^2 n}$$

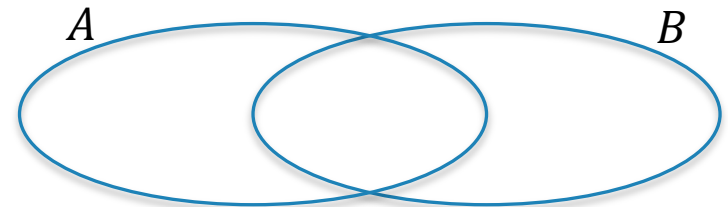
(upper tail bound)

- \* For any two events  $A, B$ :

- \*  $\Pr[A \cup B] \leq \Pr[A] + \Pr[B]$

- \* This also implies:

- \*  $\Pr[A \cap B] \geq 1 - (\Pr[\bar{A}] + \Pr[\bar{B}])$



- \* **Example:** Combined Chernoff-Hoeffding bound:

- \*  $\Pr\left[\left|\frac{1}{n}X - \mathbb{E}[X_i]\right| \geq \epsilon\right] \leq 2e^{-2\epsilon^2 n}$

Since  $\Pr\left[\left|\frac{1}{n}X - \mathbb{E}[X_i]\right| \geq \epsilon\right] = \Pr\left[\left(\frac{1}{n}X \leq \mathbb{E}[X_i] - \epsilon\right) \cup \left(\frac{1}{n}X \geq \mathbb{E}[X_i] + \epsilon\right)\right]$

# Polling

- \* There are  $m$  candidates for president. How can we estimate their relative support without asking the entire population?
- \* **A:** Sample people at random and compute the relative frequencies.
- \* Two types of “accuracy”:
  1. The probability that we indeed obtain a “good” estimate
  2. The extent to which our estimate approximates reality
- \* **Fine print:** “This poll has been conducted with a *confidence level* of 95% and *statistical error* of  $\pm 2\%$ ”

# Polling

- \* **Algorithm for one candidate (approval rating):**
  - \* Sample at random  $n$  people (ask: “Do you support?” Yes/No)
  - \* Let  $X$  be the number of supporters
  - \* Return  $X/n$  as an estimate
- \* Let  $0 \leq p \leq 1$  be the true level of support. How large does  $n$  have to be so that we get good “accuracy” with high “confidence”?
- \* **Fine print:** “This poll has been conducted with a *confidence level* of 95% and *statistical error* of  $\pm 2\%$ ”
- \* Thus, we want  $\Pr \left[ \left| \frac{1}{n}X - p \right| \leq 0.02 \right] \geq 0.95$ .

# Combined Chernoff-Hoeffding Bound

Combined Chernoff-Hoeffding:

$$\Pr \left[ \left| \frac{1}{n} X - \mathbb{E}[X_i] \right| \geq \varepsilon \right] \leq 2e^{-2\varepsilon^2 n}$$

Compare to Chebyshev:

$$\Pr \left[ \left| \frac{1}{n} X - \mathbb{E}[X_i] \right| \geq \varepsilon \right] \leq \frac{\text{Var}(X_i)}{\varepsilon^2 n}$$

- \* **Goal:** Find  $n$  such that  $\Pr \left[ \left| \frac{1}{n} X - p \right| \leq 0.02 \right] \geq 0.95$ .
- \* Define indicators for  $i = 1..n$ :
 
$$X_i = \begin{cases} 1, & \text{person } i \text{ supports the candidate} \\ 0, & \text{otherwise} \end{cases}$$
- \* Then  $\mathbb{E}[X_i] = \Pr[X_i = 1] = p$  and  $X = X_1 + X_2 + \dots + X_n$ .
- \* **Q:** What should the value of  $n$  be to satisfy the fine print?
- \* **Equivalently:** We want  $\Pr \left[ \left| \frac{1}{n} X - p \right| > 0.02 \right] \leq 0.05$ .
- \* By the combined CH bound:

$$\Pr \left[ \left| \frac{1}{n} X - p \right| > 0.02 \right] \leq \Pr \left[ \left| \frac{1}{n} X - p \right| \geq 0.02 \right] \leq 2e^{-2 \cdot 0.02^2 n}$$

# Polling Analysis

Combined Chernoff-Hoeffding:

$$\Pr \left[ \left| \frac{1}{n} X - \mathbb{E}[X_i] \right| \geq \varepsilon \right] \leq 2e^{-2\varepsilon^2 n}$$

- \* By the combined CH bound:

$$\Pr \left[ \left| \frac{1}{n} X - p \right| > 0.02 \right] \leq \Pr \left[ \left| \frac{1}{n} X - p \right| \geq 0.02 \right] \leq 2e^{-2 \cdot 0.02^2 n}$$

- \* Therefore, we need  $2e^{-2 \cdot 0.02^2 n} \leq 0.05$

$$\Rightarrow 40 \leq e^{2 \cdot 0.02^2 n} \Rightarrow \ln 40 \leq 2 \cdot 0.02^2 n \Rightarrow n \geq 4612$$

- \* **Remark:**  $n$  does not depend on the population size!
- \* **Q:** How large should  $n$  be if we want error  $\varepsilon$  with probability  $\leq \delta$ ?

$$* \quad 2e^{-2\varepsilon^2 n} \leq \delta \Leftrightarrow n \geq \frac{\ln(2/\delta)}{2\varepsilon^2}.$$

# Polling General Case

- \* **Algorithm for  $m$  candidates:**

- \* Sample at random  $n$  people (ask: “Who do you support?”)
- \* Let  $X^{(j)}$  be the number of supporters of candidate  $j$
- \* For each  $j$ : Return  $X^{(j)}/n$

- \* **Fine print:** “This poll has been conducted with a **confidence level** of  $1 - \delta$  and **statistical error** of  $\pm \varepsilon$ ”

- \* **Formally:** Let  $p_1, \dots, p_m$  be the support levels of the candidates.

- \* **We want:**  $\Pr \left[ \text{for every } j = 1..m: \left| \frac{1}{n} X^{(j)} - p_j \right| \leq \varepsilon \right] \geq 1 - \delta.$

# Polling General Case

- \* **We want:**  $\Pr \left[ \text{for every } j = 1..m: \left| \frac{1}{n} X^{(j)} - p_j \right| \leq \varepsilon \right] \geq 1 - \delta.$
- \* How many samples do we need now?
- \* When  $m = 1$ , we need  $n \geq \frac{1}{2\varepsilon^2} \ln \left( \frac{2}{\delta} \right)$
- \* **Wrong answer:** for  $m$  candidates we need  $n \geq m \cdot \frac{1}{2\varepsilon^2} \ln \left( \frac{2}{\delta} \right)$
- \* **Sampling Theorem:** If  $n \geq \frac{1}{2\varepsilon^2} \ln \left( \frac{2m}{\delta} \right)$  then we can assert that our estimates satisfy the fine print. (proof via union bound)
- \* **Conclusion:** The dependence on  $m$  is logarithmic!

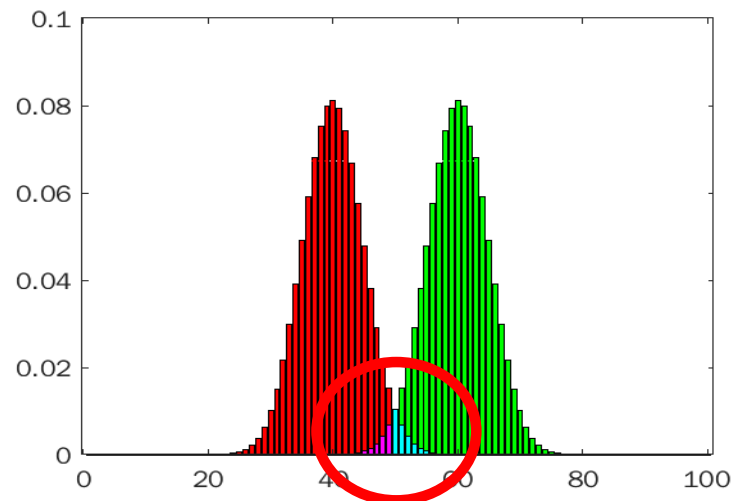


# Distinguishing Biased Coins

- \* You're given a coin that is  $\epsilon$ -biased to either heads or tails.
  - \* i.e.,  $\Pr[H] = \frac{1}{2} + \epsilon$  and  $\Pr[T] = \frac{1}{2} - \epsilon$  or  $\Pr[H] = \frac{1}{2} - \epsilon$  and  $\Pr[T] = \frac{1}{2} + \epsilon$
- \* To determine if it's biased towards heads, you flip the coin  $n$  times.
  - \* If you see at least  $\frac{1}{2}n$  heads, you guess "yes"
  - \* Otherwise, you guess "no".

**Note:** We have *two-sided* error; *false positives* and *false negatives* are possible!

**Q:** How large should  $n$  be to guarantee an error probability of  $\delta$ ?



# Probability of False Negatives

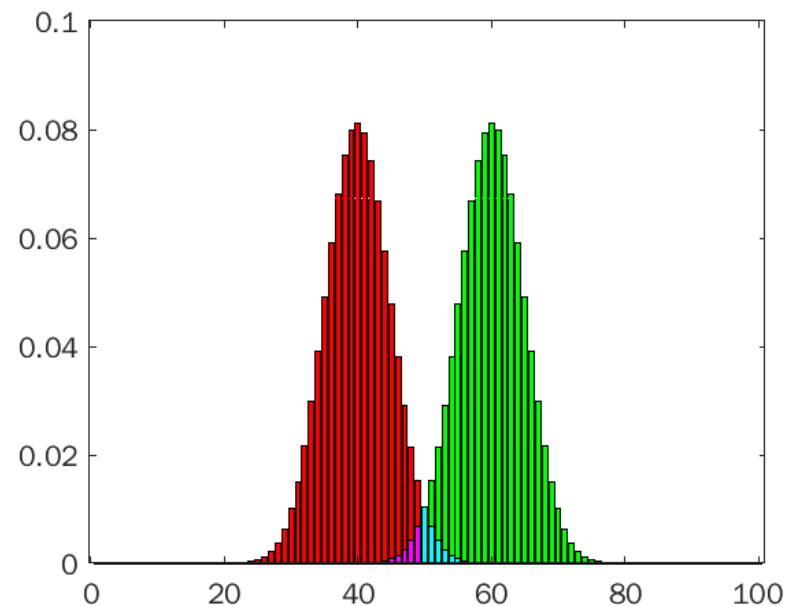
If  $X = X_1 + X_2 + \dots + X_n$  is the sum of  $n$  *i.i.d.* RVs with **each**  $X_i \in [0, 1]$ , then, for any  $\varepsilon > 0$ :

$$\Pr \left[ \frac{1}{n} X \leq \mathbb{E}[X_i] - \varepsilon \right] \leq e^{-2\varepsilon^2 n}$$

(lower tail bound)

- \* Let  $X_i$  be the indicator RV for whether  $i$ 'th coin flip was  $H$ .
- \* Suppose the coin we had was  $\varepsilon$ -biased towards **heads**.
  - \* Then  $\mathbb{E}[X_i] = \frac{1}{2} + \varepsilon$ .
- \* **Q:** When do we get an error (**false negative**) in this case?
- \* **A:** When  $\frac{1}{n} X < \frac{1}{2} = \mathbb{E}[X_i] - \varepsilon$
- \* Therefore:

$$\begin{aligned} & \Pr[\text{error} | H \text{ bias}] \\ &= \Pr \left[ \frac{1}{n} X < \mathbb{E}[X_i] - \varepsilon \right] \leq e^{-2\varepsilon^2 n} \end{aligned}$$



To determine if it's biased towards heads, you flip the coin  $n$  times.  
 If you see at least  $\frac{1}{2}n$  heads, you guess "yes".  
 Otherwise, you guess "no".

# Probability of False Positives

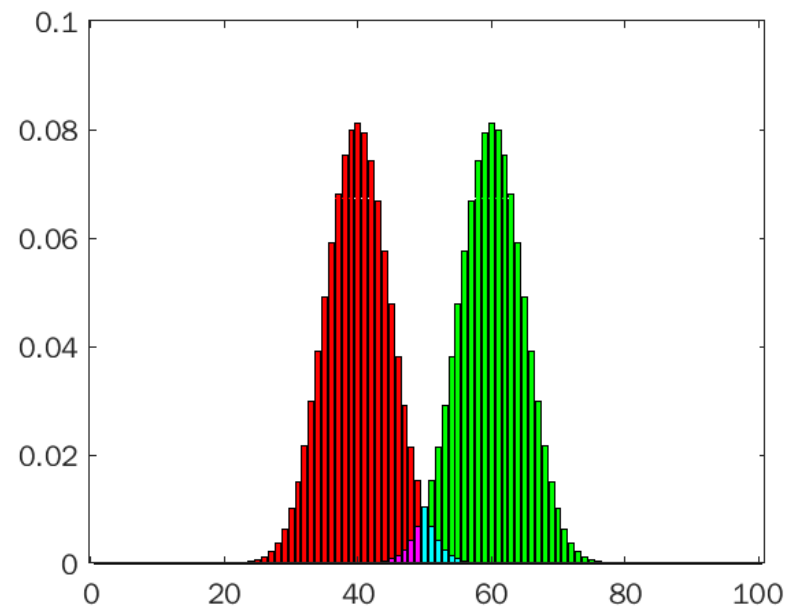
If  $X = X_1 + X_2 + \dots + X_n$  is the sum of  $n$  *i.i.d.* RVs with **each**  $X_i \in [0, 1]$ , then, for any  $\varepsilon > 0$ :

$$\Pr \left[ \frac{1}{n} X \geq \mathbb{E}[X_i] + \varepsilon \right] \leq e^{-2\varepsilon^2 n}$$

(lower tail bound)

- \* Let  $X_i$  be the indicator RV for whether  $i$ 'th coin flip was  $H$ .
- \* Suppose the coin we had was  $\varepsilon$ -biased towards **tails**.
  - \* Then  $\mathbb{E}[X_i] = \frac{1}{2} - \varepsilon$ .
- \* **Q:** When do we get an error (**false positive**) in this case?
- \* **A:** When  $\frac{1}{n} X \geq \frac{1}{2} = \mathbb{E}[X_i] + \varepsilon$
- \* Therefore:

$$\begin{aligned} & \Pr[\text{error} | T \text{ bias}] \\ &= \Pr \left[ \frac{1}{n} X \geq \mathbb{E}[X_i] + \varepsilon \right] \leq e^{-2\varepsilon^2 n} \end{aligned}$$



To determine if it's biased towards heads, you flip the coin  $n$  times.  
 If you see at least  $\frac{1}{2}n$  heads, you guess "yes".  
 Otherwise, you guess "no".

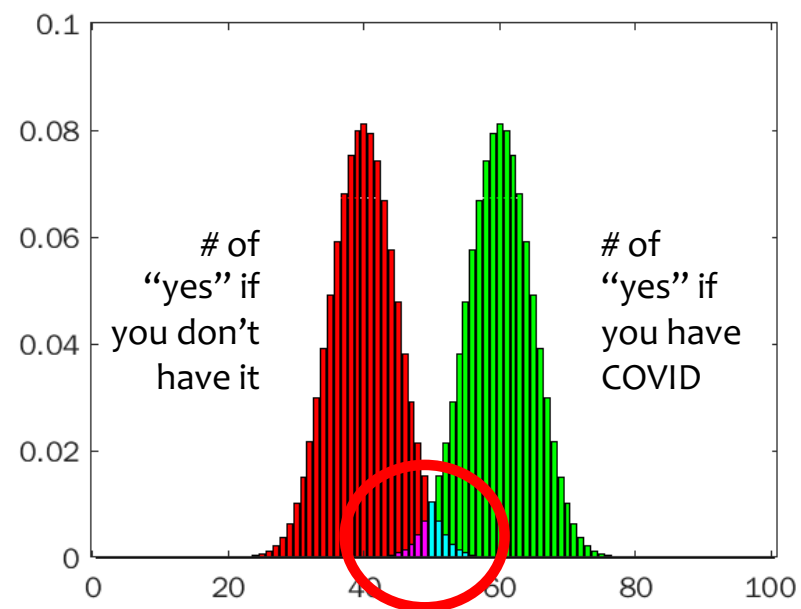
# How large should $n$ be?

- \* By previous analysis,  $\Pr[\text{error}] \leq e^{-2\varepsilon^2 n}$ .
  - \* We saw the error in either case is at most this.
- \* How large should  $n$  be if we want error to be  $\leq \delta$ ?
  - \*  $e^{-2\varepsilon^2 n} \leq \delta \Leftrightarrow n \geq \frac{\ln(1/\delta)}{2\varepsilon^2}$ .
- \* **Example:** If  $\varepsilon = 0.01$  and  $\delta = 0.0001$  (correct 99.99%), then we need  $n \geq \frac{\ln(0.0001^{-1})}{2 \cdot 0.01^2} \approx 46,052$  flips.

# Extra Practice

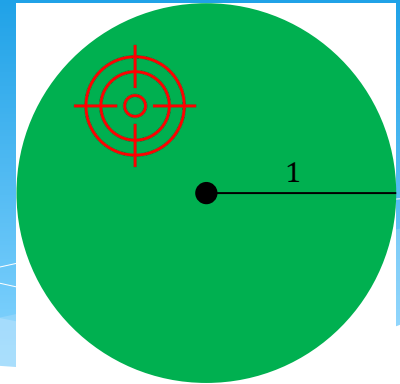
# Decreasing Error

- \* A low-quality COVID test has two-sided error:
  - \* If a person has COVID, it says “yes” w.p.  $2/3$ .
  - \* Otherwise, it says “no” w.p.  $2/3$ .
  - \* Different runs are independent.
- \* You decide to buy and run the test  $n$  times and take the *majority* answer you get.
- \* **Q:** How large should  $n$  be to guarantee that the answer is correct w.p.  $1 - \delta$ ?
  - \* Same as distinguishing  $\varepsilon$ -biased coins with  $\varepsilon = 1/6$ !



**Note:** *false positives* and *false negatives* are possible!

# Estimating $\pi$



- \* Suppose there is a 2x2 square with a unit circle inside
- \* **Q:** If we toss a dart *uniformly at random* towards the square, what's the probability that we hit the circle?
  - \*  $(\text{area of circle})/(\text{area of board}) = \pi/4$
- \* We toss  $n$  darts uniformly at random towards the square
  - \*  $X_i$  = indicator 0/1 RVs for whether we hit circle on  $i$ 'th toss
  - \* **Q:** What is  $\mathbb{E} \left[ \frac{1}{n} \sum_{i=1}^n X_i \right]$ ?
- \* **Q:** How might we estimate  $\pi$  by tossing darts?
  - \* It's roughly 4\*fraction of times we hit circle;  
CH to bound error.

$$\Pr \left[ \left| \frac{1}{n} X - \mathbb{E}[X_i] \right| \geq \varepsilon \right] \leq 2e^{-2\varepsilon^2 n}$$

(combined bound)

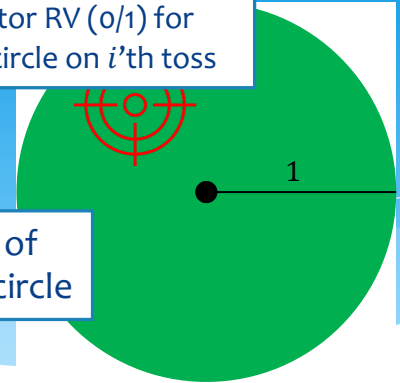
# Math

(to show that this is a bad idea)

We toss  $n$  darts uniformly at random towards the square

Let  $X_i$  = indicator RV (0/1) for whether we hit circle on  $i$ 'th toss

$\pi \approx 4 \times \text{fraction of times we hit the circle}$



\* Let  $X = X_1 + \dots + X_n$  be the number of times we hit the circle.

\*  $\mathbb{E}[X_i] = \frac{\pi}{4}$  so  $\left| \frac{1}{n} X - \frac{\pi}{4} \right| < \varepsilon$  with probability  $\geq 1 - 2e^{-2\varepsilon^2 n}$

\* To estimate  $\pi$  within  $\gamma$ , i.e.  $\left| \frac{4}{n} X - \pi \right| < \gamma$ , set  $\varepsilon = \frac{\gamma}{4}$ .

\*  $\left| \frac{4}{n} X - \pi \right| < \gamma \Leftrightarrow \left| \frac{1}{n} X - \frac{\pi}{4} \right| < \frac{\gamma}{4} = \varepsilon$   
 (with probability  $\geq 1 - 2e^{-2\varepsilon^2 n} = 1 - 2e^{-\gamma^2 n/8}$ )

\* For probability  $\geq 1 - \delta$ , set  $n = 8 \ln(2/\delta) / \gamma^2$ .

\*  $1 - 2e^{-\gamma^2 n/8} \geq 1 - \delta \Leftrightarrow \delta \geq 2e^{-\gamma^2 n/8} \Leftrightarrow \ln \delta/2 \geq -\gamma^2 n/8 \Leftrightarrow n \geq 8 \ln(2/\delta) / \gamma^2$

**Example:** To get our estimate between 3.140 and 3.142 ( $\gamma = 0.001$ ) 99.99% of the time ( $\delta = 0.0001$ ), we should toss  $n \approx 79,227,901$  darts