

PS8_12265092

12265092

09/03/2022

```
library(tidyverse)
```

```
## -- Attaching packages ----- tidyverse 1.3.1 --
```

```
## v ggplot2 3.3.5      v purrr  0.3.4
## v tibble  3.1.4      v dplyr  1.0.7
## v tidyr   1.1.3      v stringr 1.4.0
## v readr   2.0.1      v forcats 0.5.1
```

```
## -- Conflicts ----- tidyverse_conflicts() --
```

```
## x dplyr::filter() masks stats::filter()
## x dplyr::lag()     masks stats::lag()
```

```
library(estimatr)
```

```
## Warning: package 'estimatr' was built under R version 4.1.2
```

```
library(Rcpp)
library(readxl)
library(haven)
library(boot)
library(lmtest)
```

```
## Warning: package 'lmtest' was built under R version 4.1.2
```

```
## Loading required package: zoo
```

```
## Warning: package 'zoo' was built under R version 4.1.2
```

```
##
```

```
## Attaching package: 'zoo'
```

```
## The following objects are masked from 'package:base':
```

```
##
```

```
##      as.Date, as.Date.numeric
```

```
library(margins)
```

```
## Warning: package 'margins' was built under R version 4.1.2
```

```
library(mfx)
```

```
## Warning: package 'mfx' was built under R version 4.1.2
```

```
## Loading required package: sandwich
```

```
## Warning: package 'sandwich' was built under R version 4.1.2
```

```
## Loading required package: MASS
```

```
##
```

```
## Attaching package: 'MASS'
```

```
## The following object is masked from 'package:dplyr':
```

```
##
```

```
##      select
```

```
## Loading required package: betareg
```

```
## Warning: package 'betareg' was built under R version 4.1.2
```

```
library(ivreg)
```

```
## Warning: package 'ivreg' was built under R version 4.1.2
```

```
library(aod)
```

```
## Warning: package 'aod' was built under R version 4.1.2
```

```
data <- read_dta("voucher.dta")
```

```
#1a
```

```
table(data$choicetyrs)
```

```
##
```

```
##    0    1    2    3    4
```

```
## 490 211 122 111   56
```

```
# Of 990 students in sample, 56 student attended a choice school for 4 years
```

```
table(data$selecttyrs)
```

```
##
```

```
##    0    1    2    3    4
```

```
## 468 116 148 150 108
```

```
# 108 students had vouchers available for 4 years
# Of 990 students in sample, 468 students never awarded a voucher
```

```
#1b
```

```
reg1b <- lm(formula = choiceyrs ~ selectyrs, data = data)
```

```
summary(reg1b)
```

```
##
## Call:
## lm(formula = choiceyrs ~ selectyrs, data = data)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -3.08725 -0.01992 -0.01992  0.21325  1.21325
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)  0.01992    0.02461   0.809   0.419
## selectyrs    0.76683    0.01259  60.931 <2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.576 on 988 degrees of freedom
## Multiple R-squared:  0.7898, Adjusted R-squared:  0.7896
## F-statistic: 3713 on 1 and 988 DF, p-value: < 2.2e-16
```

```
#Residual standard error: 0.576 on 988 degrees of freedom
#Multiple R-squared:  0.7898,
#Adjusted R-squared:  0.7896
#F-statistic: 3713 on 1 and 988 DF,
#p-value: < 2.2e-16
```

```
#The p-value at 5% significance level is less than 0.05. So we can safely
#reject the null hypothesis that coefficient of 'selectyrs' is zero
#Also the alternate hypothesis cannot be rejected
#The t-statistic calculated says that the 'selectyrs' is statistically
#significant for estimating 'choiceyrs'
```

```
#The coefficient 0.77 shows that selectyrs and choiceyrs are positively
#correlated. This is as expected because the increase in years of voucher will
#increase the number of years student attends choice school
```

```
#Two point criteria for valid IV
```

```
#Correlation : selectyrs, choiceyrs are positively correlated
```

```
#Zero correlation with error term u
```

```
#We can say that selectyrs to be uncorrelated with 'u' because it can be
#understood that the voucher is offered among those who applied, the other
#unmeasurable variables for example student background do not influence
#the variable selectyrs. Thus we can say selectyrs is uncorrelated with the
#error term 'u' and is a valid IV choice for choiceyrs.
```

```

#1c

reg1c <- lm(formula = mnce ~ choiceyrs, data = data)

summary(reg1c)

##
## Call:
## lm(formula = mnce ~ choiceyrs, data = data)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -45.234 -13.234   0.603  12.766  60.114
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)  46.2344     0.8507   54.348 < 2e-16 ***
## choiceyrs    -1.8370     0.5255   -3.495 0.000494 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 20.75 on 988 degrees of freedom
## Multiple R-squared:  0.01222,    Adjusted R-squared:  0.01122
## F-statistic: 12.22 on 1 and 988 DF,  p-value: 0.0004943

#Residual standard error: 20.75 on 988 degrees of freedom
#Multiple R-squared:  0.01222,    Adjusted R-squared:  0.01122
#F-statistic: 12.22 on 1 and 988 DF,
#p-value: 0.0004943

#Coefficient -1.837

#The p-value at 5% significance level and the t-statistic calculated says that
#the variable 'choiceyrs' is statistically significant in estimating 'mnce'
#We expect the percentile score on the math test in 1994 increase with the
#increase in number of years the student attended school in the period
#1991 to 1994
#However, we see that the coefficient is negative, which is in contrary to the
#expectations.

reg1c_1 <- lm (formula = mnce ~ choiceyrs + black + hispanic + female, data = data)

summary(reg1c_1)

##
## Call:
## lm(formula = mnce ~ choiceyrs + black + hispanic + female, data = data)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -56.122 -12.507   0.108  12.156  60.156
##
## Coefficients:

```

```
##           Estimate Std. Error t value Pr(>|t|)
## (Intercept)  57.1219     1.6567  34.479 < 2e-16 ***
## choiceyrs   -0.5652     0.5307  -1.065  0.287
## black       -16.0174     1.7944  -8.926 < 2e-16 ***
## hispanic    -13.4029     2.3168  -5.785 9.73e-09 ***
## female       1.3527     1.2758   1.060  0.289
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 19.99 on 985 degrees of freedom
## Multiple R-squared:  0.08677,    Adjusted R-squared:  0.08307
## F-statistic: 23.4 on 4 and 985 DF,  p-value: < 2.2e-16
```

#Coefficients:

```
#           Estimate Std. Error t value Pr(>|t|)
#(Intercept)  57.1219     1.6567  34.479 < 2e-16 ***
#choiceyrs   -0.5652     0.5307  -1.065  0.287
#black       -16.0174     1.7944  -8.926 < 2e-16 ***
#hispanic    -13.4029     2.3168  -5.785 9.73e-09 ***
#female       1.3527     1.2758   1.060  0.289
```

#p-value: 0.0004943

*#We see that the from the p-value and t statistic calculate above, we see that
#choiceyrs variable is now not statistically significant soon as we added other
#variables black, hispanic, female to our model. Among the new variables,
#only black and hispanic are statistically significant where as female is not
#statistically significant*

#1d

#Why is choiceyrs a endogenous in the equation with choiceyrs, black, hispanic, female

```
regid <- lm(formula = choiceyrs ~ black + hispanic + female, data = data)
summary(regid)
```

```
##
## Call:
## lm(formula = choiceyrs ~ black + hispanic + female, data = data)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -1.4183 -1.0909 -0.1929  0.8071  3.8337
##
## Coefficients:
##           Estimate Std. Error t value Pr(>|t|)
## (Intercept)  0.16633    0.09927   1.676  0.0941 .
## black        0.92456    0.10357   8.927 <2e-16 ***
## hispanic     1.15003    0.13411   8.575 <2e-16 ***
## female       0.10198    0.07648   1.333  0.1827
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
```

```
## Residual standard error: 1.199 on 986 degrees of freedom
## Multiple R-squared:  0.09068,    Adjusted R-squared:  0.08792
## F-statistic: 32.78 on 3 and 986 DF,  p-value: < 2.2e-16
```

#Coefficients

```
#black      0.92456    0.10357    8.927    <2e-16 ***
#hispanic   1.15003    0.13411    8.575    <2e-16 ***
#female     0.10198    0.07648    1.333    0.1827
```

```
#p-value: < 2.2e-16
```

```
#We see from the regression results that the variables black and hispanic
#are statistically significant in estimating choiceyrs and variable female is
#not. This means that the number of years a student attended choice schools
#in period 1991-1994 is dependent on race. This is like multicollinearity.
#Thus the choiceyrs is a endogenous variable in the equation with the
#variables choiceyrs, black, hispanic, female
```

```
#1e
```

```
#use selectyrs as IV for choiceyrs
```

```
ivregle <- ivreg(formula = mnce ~ choiceyrs + black + hispanic + female |
                  selectyrs + black + hispanic + female, data = data)
```

```
summary(ivregle)
```

```
##
```

```
## Call:
```

```
## ivreg(formula = mnce ~ choiceyrs + black + hispanic + female |
##       selectyrs + black + hispanic + female, data = data)
```

```
##
```

```
## Residuals:
```

```
##      Min       1Q   Median       3Q      Max
## -56.0680 -12.5098  -0.0476  12.0769  59.2142
```

```
##
```

```
## Coefficients:
```

```
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)  57.0680     1.6577  34.426 < 2e-16 ***
## choiceyrs    -0.2413     0.6053  -0.399   0.690
## black       -16.3169     1.8148  -8.991 < 2e-16 ***
## hispanic    -13.7754     2.3412  -5.884 5.49e-09 ***
## female        1.3197     1.2763   1.034   0.301
```

```
##
```

```
## Diagnostic tests:
```

```
##              df1 df2 statistic p-value
## Weak instruments    1 985  3281.836 <2e-16 ***
## Wu-Hausman          1 984    1.241   0.265
## Sargan              0 NA         NA      NA
```

```
## ---
```

```
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
##
```

```
## Residual standard error: 19.99 on 985 degrees of freedom
```

```
## Multiple R-Squared: 0.08643, Adjusted R-squared: 0.08272
## Wald test: 23.15 on 4 and 985 DF, p-value: < 2.2e-16
```

```
#Coefficients:
```

```
#           Estimate Std. Error t value Pr(>|t|)
##(Intercept)  57.0680      1.6577  34.426 < 2e-16 ***
##choiceyrs   -0.2413      0.6053  -0.399  0.690
##black       -16.3169      1.8148  -8.991 < 2e-16 ***
##hispanic    -13.7754      2.3412  -5.884 5.49e-09 ***
##female       1.3197      1.2763   1.034  0.301
```

```
#p-value: < 2.2e-16
```

```
#selectyrs as IV for 'choiceyrs' doesnt produce a positive effect on 'mnce'
```

```
#The variable black and hispanic are statistically significant at the
#5% level as seen from the t value calculated -8.991(Pr 2e-16) and -5.884
#(Pr 5.49e-09). The variable female is not statistically significant
#as seen from the t-statistic calculated 1.034 (Pr 0.301) at 5% significance
#level. Thus we can say that mnce is more dependent on the race (black or
#hispanic) of the student.
```

```
#1f
```

```
reg1f <- lm(formula = mnce ~ choiceyrs + black + hispanic +
             female + mnce90, data = data)
```

```
summary(reg1f)
```

```
##
## Call:
## lm(formula = mnce ~ choiceyrs + black + hispanic + female + mnce90,
##     data = data)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -47.921 -11.669   0.773  10.686  50.838
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)  22.1529      3.6204   6.119 2.73e-09 ***
## choiceyrs     0.4106      0.7359   0.558 0.57726
## black        -8.3052      2.5461  -3.262 0.00123 **
## hispanic     -4.1050      3.3624  -1.221 0.22303
## female       -0.8829      1.7760  -0.497 0.61945
## mnce90        0.6204      0.0484  12.817 < 2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 16.03 on 322 degrees of freedom
## (662 observations deleted due to missingness)
## Multiple R-squared:  0.4237, Adjusted R-squared:  0.4147
## F-statistic: 47.34 on 5 and 322 DF, p-value: < 2.2e-16
```

#Coefficient of choiceyrs = 0.4106

```
ivreg1f <- ivreg(mnce ~ choiceyrs + black + hispanic + female + mnce90 |
                selectyrs + black + hispanic +
                female + mnce90, data = data)

summary(ivreg1f)
```

```
##
## Call:
## ivreg(formula = mnce ~ choiceyrs + black + hispanic + female +
##       mnce90 | selectyrs + black + hispanic + female + mnce90,
##       data = data)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -51.171 -11.160   1.067  10.883  50.802
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)  21.53886    3.64548   5.908 8.79e-09 ***
## choiceyrs     1.79938    0.86019   2.092 0.037236 *
## black        -9.06711    2.57142  -3.526 0.000483 ***
## hispanic     -5.00373    3.39279  -1.475 0.141240
## female       -1.02048    1.78630  -0.571 0.568205
## mnce90        0.62881    0.04874  12.901 < 2e-16 ***
##
## Diagnostic tests:
##              df1 df2 statistic p-value
## Weak instruments    1 322    916.06 < 2e-16 ***
## Wu-Hausman          1 321     10.43 0.00137 **
## Sargan              0 NA         NA      NA
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 16.12 on 322 degrees of freedom
## Multiple R-Squared: 0.4173, Adjusted R-squared: 0.4082
## Wald test: 47.64 on 5 and 322 DF, p-value: < 2.2e-16
```

#Coefficient of choiceyrs = 1.79938

*#The coefficient of choiceyrs in IV is 1.79938 which is greater than the
#coefficient of choiceyrs in OLS in 0.4106. The variable is statistically
#significant in IV*

*#The p-value 0.037 with t-statistic 2.092 for choiceyrs shows that the
#statistically significance. This is a large effect that each year in a choice
#school is worth 1.7993 times in crease in the math percentile score.*


```
#1g
```

```
summary(data$mnce90)
```

```
##      Min. 1st Qu.  Median    Mean 3rd Qu.    Max.     NA's  
##      1.00   30.50   43.00   43.93   57.25   99.00     662
```

```
nobs(ivreg1e)
```

```
## [1] 990
```

```
nobs(ivreg1f)
```

```
## [1] 328
```

*#The stark difference in the number of observations in both the iv regressions
990 and 328 can be accounted for low precision in the coefficients
#estimated in 1f which can be a reason for why the analysis 1f to be unconvincing*

```
#1h
```

```
ivreg1h <- ivreg(formula = mnce ~ choiceyrs1 + choiceyrs2 + choiceyrs3 +  
                  choiceyrs4 + black + hispanic + female + mnce90 |  
                  selectyrs1 + selectyrs2 + selectyrs3 + selectyrs4 +  
                  black + hispanic + female + mnce90, data = data)
```

```
summary(ivreg1h)
```

```
##  
## Call:  
## ivreg(formula = mnce ~ choiceyrs1 + choiceyrs2 + choiceyrs3 +  
##       choiceyrs4 + black + hispanic + female + mnce90 | selectyrs1 +  
##       selectyrs2 + selectyrs3 + selectyrs4 + black + hispanic +  
##       female + mnce90, data = data)  
##  
## Residuals:  
##      Min      1Q   Median      3Q      Max  
## -58.4487 -11.2629   0.9051  10.4675  54.5467  
##  
## Coefficients:  
##              Estimate Std. Error t value Pr(>|t|)  
## (Intercept)  20.91132    3.76450   5.555 5.86e-08 ***  
## choiceyrs1   -2.15904    5.16369  -0.418  0.67614  
## choiceyrs2    1.49306    4.62755   0.323  0.74717  
## choiceyrs3    1.08156    7.14421   0.151  0.87976  
## choiceyrs4   13.92571    6.35406   2.192  0.02913 *  
## black        -8.34689    2.66087  -3.137  0.00187 **  
## hispanic     -3.58953    3.61733  -0.992  0.32180  
## female       -1.03110    1.81870  -0.567  0.57115  
## mnce90        0.63574    0.05028  12.645 < 2e-16 ***
```

```
##
## Diagnostic tests:
##
##              df1 df2 statistic p-value
## Weak instruments (choiceyrs1)    4 319    74.480 <2e-16 ***
## Weak instruments (choiceyrs2)    4 319   101.952 <2e-16 ***
## Weak instruments (choiceyrs3)    4 319    37.555 <2e-16 ***
## Weak instruments (choiceyrs4)    4 319    62.387 <2e-16 ***
## Wu-Hausman                      4 315     3.126 0.0153 *
## Sargan                          0 NA         NA      NA
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 16.34 on 319 degrees of freedom
## Multiple R-Squared: 0.4067, Adjusted R-squared: 0.3918
## Wald test: 29.22 on 8 and 319 DF, p-value: < 2.2e-16
```

```
#From the t-statistic and p value observed above, we can say that the variables
#choiceyrs4 and black are statistically significant.
#From the coefficeints
#The coefficient of variable black -3.137 indicates the negative correlation
#between black race and the math percentile score
#The coefficient of variable choiceyrs4 13.925 indicates that 4 additional
#years at the school of choice increases the math percentile score in 1994
#mnce by 13.925 times
#This makes sense as the coefficient found here is in line
#with the coefficients in parts d,e and f
```

#Question 2

```
#Corn - percapita consumption of corn in bushels at country level
#price price per buschel of corn
#income - percapita country income
#rainfall - inches of rainfall during last corn-growing season
```

```
#from the given equations we can say that the exogenous variables are
#rainfall and income and the endogenous variables are corn and price as they are
#dependent on the rainfall and income variables.
```

```
#Equation 1 is the demand equation as the equation demonstrates the consumption
#of corn in form of price of the corn and the income of country. Thus this is
#the demand function
```

```
#Equation 2 is supply function. Supply or availability of corn can be said from
#the rainfall in the last season , usually increase in rainfall leading to
#increase in supply of corn and also depends on the price. Thus 2 is supply eqn.
```

#Question 3

```
#a
```

```
#In the basic model where only poorhealth is estimated on smokes yes or no
#the tax can be a valid instrument if we observe that we observe some correlation
```

```

#between smoker and other health factors or if we see some correlation with
#the error term u. Criteria for the same can be evaluated if below is true
# Cov(tax, u ) = 0. tax is not correlated with error term u
# Cov(tax, smoker) != 0. tax is correlated with smoker

#b
#We need to run the following regressions in order to obtain instrumental
#variables estimator

#smoker = alpha0 + alpha1*tax + u
#smoker^ = alpha0^ + alpha1^*tax

#and

#poorhealth = beta0 + beta1*smoker + u
#poorhealth^ = beta0^ + beta1^*smoker

#c

#Tax as a valid instrument, omitting does not create omitted variable bias while
#estimating poorhealth. However, we can say that it can result in
#multicollinearity, depends on the strenght of tax as a valid instrument on
#regression for variable smoker.
#As tax is a valid instrument, that means the correlation between tax and the
#error term u is zero. And the error term u accounts for the omitted variable
# bias in the regression. So omitting tax will not create omitted variable
#bias in regression. These two statemtns says that variable tax is related to
#poorhealth via smoker. Thus we may not get supportive coefficient to prove
#statistical significance of tax in OLS but can prove multicollinearity with
#a less precision.

```

#Question 4

```
votedata <- read_dta('vote2.dta')
```

```
#a
```

```
#Difference the equation across two years
```

```
#Regression
```

```
reg4a <- lm_robust(cvote ~ clinexp + clchexp + cincshr,
                  data = votedata)
```

```
summary(reg4a)
```

```
##
```

```
## Call:
```

```
## lm_robust(formula = cvote ~ clinexp + clchexp + cincshr, data = votedata)
```

```
##
```

```
## Standard error type: HC2
```

```
##
## Coefficients:
##           Estimate Std. Error t value Pr(>|t|) CI Lower CI Upper DF
## (Intercept)  -2.5559     0.5873  -4.3519 2.456e-05  -3.7162  -1.3957 153
## clinexp      -1.2915     1.3076  -0.9877 3.249e-01  -3.8749   1.2918 153
## clchexp      -0.5985     0.5859  -1.0215 3.086e-01  -1.7561   0.5591 153
## cincshr       0.1559     0.0543   2.8707 4.676e-03   0.0486   0.2631 153
##
## Multiple R-squared:  0.2437 ,    Adjusted R-squared:  0.2288
## F-statistic: 17.27 on 3 and 153 DF,  p-value: 1.037e-09
```

```
#p value 1.037e-09
```

```
#From the values we observed in the regression summary table, at 5% significance
#level we can say that cincshr is statistically significant whereas the
#other two variables clinexp and clchexp are not statistically
#significant in estimating cvote
```

```
#4b
```

```
wald.test(Sigma = vcov(reg4a), b = coef(reg4a), Terms = 2:3)
```

```
## Wald test:
```

```
## -----
```

```
##
```

```
## Chi-squared test:
```

```
## X2 = 3.3, df = 2, P(> X2) = 0.19
```

```
#The p value of 0.19 is more than 0.05, so we cannot reject the null hypothesis
#Thus we can say that the coefficients of the variables delta_ln(inexp) and
#delta_ln(chexp) are zero and these variables are statistically insignificant
#Thus we can say that the ln change in incumbent expenditures and the ln change
#in challenger expenditures do not have joint role in estimating the incumbent
#vote share
```

```
#4c
```

```
reg4c <- lm_robust(cvote ~ cincshr, data = votedata)
```

```
summary(reg4c)
```

```
##
```

```
## Call:
```

```
## lm_robust(formula = cvote ~ cincshr, data = votedata)
```

```
##
```

```
## Standard error type: HC2
```

```
##
```

```
## Coefficients:
```

```
##           Estimate Std. Error t value Pr(>|t|) CI Lower CI Upper DF
## (Intercept)  -2.6811     0.6218  -4.312 2.871e-05  -3.9095  -1.4528 155
## cincshr       0.2176     0.0374   5.818 3.311e-08   0.1437   0.2915 155
##
```

```
## Multiple R-squared:  0.2287 ,    Adjusted R-squared:  0.2237
```

```
## F-statistic: 33.85 on 1 and 155 DF,  p-value: 3.311e-08
```

#The t statistic 5.818 and p 2.817e-05 prove that the variable delta_incs shr is statistically significant. The coefficient estimate is 0.2176 which means that the incumbent's share of vote rises by 2.176 percentage points for an increase of incumbents share of spending by 10 percentage points

```
#vote_rptchall <- vote %>% filter(rptchall == 1)
#dim(vote_rptchall)
```

```
table(votedata$rptchall)
```

```
##
##      0      1
## 150    36
```

```
rptchall1 <- votedata %>% filter(rptchall == 1)
```

```
rptchall1
```

```
## # A tibble: 36 x 26
##   state district democ vote90 vote88 inexp90 chexp90 inexp88 chexp88 prtyst
##   <chr>      <dbl> <dbl> <dbl> <dbl> <dbl> <dbl> <dbl> <dbl> <dbl>
## 1 AZ          2      1     66     73 112373  1445  99607  3065    55
## 2 CA          3      1     60     71 734005  4628 638688  7695    49
## 3 CA         11      1     66     71 620782 97030 269510 95575    58
## 4 CA         23      1     62     64 201404 360389 140486 100956    56
## 5 CA         24      1     69     72 287505  1830 191334 15449    65
## 6 CA         28      1     73     76 113669  3799 114523    NA     73
## 7 CA         32      1     62     67 462503  7579 457410  20608    49
## 8 CA         42      0     59     64 398963 28350 494487  11889    65
## 9 GA          1      1     71     67 399035 19130 337048  40461    39
## 10 GA         5      1     76     78 108118  7755 101540  6047     68
## # ... with 26 more rows, and 16 more variables: rptchall <dbl>, tenure <dbl>,
## #   lawyer <dbl>, linexp90 <dbl>, lchexp90 <dbl>, linexp88 <dbl>,
## #   lchexp88 <dbl>, incshr90 <dbl>, incshr88 <dbl>, cvote <dbl>, clinexp <dbl>,
## #   clchexp <dbl>, cincshr <dbl>, win88 <dbl>, win90 <dbl>, cwin <dbl>
```

```
#rptchall1$incshr9088 <- rptchall1$incshr90 - rptchall1$incshr88
```

```
reg4d <- lm_robust(cvote ~ cincshr, data = rptchall1)
summary(reg4d)
```

```
##
## Call:
## lm_robust(formula = cvote ~ cincshr, data = rptchall1)
##
## Standard error type: HC2
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|) CI Lower CI Upper DF
## (Intercept) -2.24982    0.95077  -2.366   0.0244 -4.18893  -0.3107 31
## cincshr      0.09239    0.06185   1.494   0.1453 -0.03375   0.2185 31
```

```
##  
## Multiple R-squared:  0.03694 ,    Adjusted R-squared:  0.005875  
## F-statistic: 2.232 on 1 and 31 DF,  p-value: 0.1453
```

```
#The t statistic of 1.494 with p-value 0.1453 greater than 0.05 says that the  
#variable difference in increase in share is not statistically significant  
#in estimating the votes.  
#By using the pairs that have repeated challengers,  
#we are eliminating any unmeasurable quality effects of the incumbent. THE  
#assumption that holds true here is that the quality characteristics of the  
#incumbent will not change over the period 1988 - 1990. THis is  
#because we are differencing and the challenger is same in both times making the  
#approach preferable.
```