

Mixed Effects Models Module

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For patient i ($i = 1, \dots, 234$) at time t ($t = 1$ (4 wks), 2 (12 wks), 3 (24 wks), 4 (52 wks))

- **Notation 1:** model equation at the level of statistical units it

$$\text{VISUAL}_{it} = \beta_{0t} + \beta_1 \cdot \text{VISUAL0}_i + \beta_{2t} \cdot \text{TREAT}_i + \epsilon_{it} \quad \text{with} \quad \boxed{\epsilon_{it} \sim \mathcal{N}(0, \sigma^2)}$$

- **Notation 2:** model equation at the patient level i

$$\underline{\text{VISUAL}}_i = \mathbb{X}_i \underline{\beta} + \underline{\epsilon}_i$$

$$\begin{bmatrix} \text{VISUAL}_{i1} \\ \text{VISUAL}_{i2} \\ \text{VISUAL}_{i3} \\ \text{VISUAL}_{i4} \end{bmatrix} \equiv \underline{\text{VISUAL}}_i \quad \text{and} \quad \begin{bmatrix} \epsilon_{i1} \\ \epsilon_{i2} \\ \epsilon_{i3} \\ \epsilon_{i4} \end{bmatrix} \equiv \underline{\epsilon}_i \sim \mathcal{N}(\underline{0}, \mathbf{R}_i) \quad \text{where} \quad \mathbf{R}_i = \sigma^2 \mathbf{R} = \sigma^2 \mathbf{A}_i \mathbf{C}_i \mathbf{A}_i$$

$$\mathbf{A}_i = \begin{bmatrix} 1 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 \\ 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 1 \end{bmatrix} \quad \text{and} \quad \mathbf{C}_i = \begin{bmatrix} 1 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 \\ 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 1 \end{bmatrix}$$

$$\mathbb{X}_i \underline{\beta} = \begin{bmatrix} 1 & 0 & 0 & 0 & \text{VISUAL0}_i & \text{TREAT}_i & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 & \text{VISUAL0}_i & 0 & \text{TREAT}_i & 0 & 0 \\ 0 & 0 & 1 & 0 & \text{VISUAL0}_i & 0 & 0 & \text{TREAT}_i & 0 \\ 0 & 0 & 0 & 1 & \text{VISUAL0}_i & 0 & 0 & 0 & \text{TREAT}_i \end{bmatrix} \cdot \begin{bmatrix} \beta_{0 \text{ 4wks}} \\ \beta_{0 \text{ 12wks}} \\ \beta_{0 \text{ 24wks}} \\ \beta_{0 \text{ 52wks}} \\ \beta_1 \\ \beta_{2 \text{ 4wks}} \\ \beta_{2 \text{ 12wks}} \\ \beta_{2 \text{ 24wks}} \\ \beta_{2 \text{ 52wks}} \end{bmatrix}$$

```
# formula
lm1.form <- visual ~ -1 + time.f + visual0 + treat.f:time.f
```

1 Standard Model: $\epsilon_{it} \sim \mathcal{N}(0, \sigma^2)$

```
# Model 6.1
fm6.1 <- lm(lm1.form, data = armd)
```

In Sections 2 and 3 we relax the variance homogeneity assumption: we want to model Λ_i , (while \mathcal{C}_i is still diagonal). We mainly employ **Notation 1**.

2 Known Variance Weights: $\epsilon_{it} \sim \mathcal{N}(0, \sigma_{it}^2)$

where, for instance, we assume $\sigma_{it} = \sigma_t = \sigma \cdot \sqrt{\text{TIME}_t}$, i.e.,

$$\sigma_t = \begin{cases} \sigma_1 & \\ \sigma_2 & \\ \sigma_3 & \\ \sigma_4 & \end{cases} = \begin{cases} \sigma \cdot \sqrt{4} & \text{if } t = 1 \text{ (4 weeks);} \\ \sigma \cdot \sqrt{12} & \text{if } t = 2 \text{ (12 weeks);} \\ \sigma \cdot \sqrt{24} & \text{if } t = 3 \text{ (24 weeks);} \\ \sigma \cdot \sqrt{52} & \text{if } t = 4 \text{ (52 weeks).} \end{cases}$$

```
# Model 9.0
weights = varFixed(value = ~time)
# the variance covariate needs to be continuous: time.f doesn't work!
fm9.0 <- gls(lm1.form, weights = weights, data = armd)
```

3 Variance Functions: $\epsilon_{it} \sim \mathcal{N}(0, \sigma_{it}^2)$

3.1 varIdent(.) - $\langle \delta \rangle$ -group

$$\epsilon_{it} \sim \mathcal{N}(0, \sigma_t^2), \quad \text{s.t.} \quad \sigma_t = \begin{cases} \sigma_1 & \\ \sigma_2 & \\ \sigma_3 & \\ \sigma_4 & \end{cases} = \begin{cases} \sigma \cdot 1 & \text{if } t = 1 \text{ (4 weeks);} \\ \sigma \cdot \delta_2 & \text{if } t = 2 \text{ (12 weeks);} \\ \sigma \cdot \delta_3 & \text{if } t = 3 \text{ (24 weeks);} \\ \sigma \cdot \delta_4 & \text{if } t = 4 \text{ (52 weeks).} \end{cases}$$

we get: $\delta_2 = \frac{\sigma_2}{\sigma_1}$; $\delta_3 = \frac{\sigma_3}{\sigma_1}$; $\delta_4 = \frac{\sigma_4}{\sigma_1}$

```
# Model 9.1
weights = varIdent(form = ~1|time.f)
fm9.1 <- gls(lm1.form, weights = weights, data = armd)
```

3.2 varPower(.) - < δ >-group

- We do not include any stratification in the model

$$\begin{aligned}\epsilon_{it} &\sim \mathcal{N}(0, \sigma_{it}^2) \quad \text{s.t.} \quad \sigma_{it} = \sigma \cdot \lambda_{it} \\ &= \sigma \cdot \lambda(\delta, \text{TIME}_{it}) \\ &= \sigma \cdot |\text{TIME}_{it}|^\delta \quad \text{since } \lambda \text{ is } \text{varPower}(\cdot)\end{aligned}$$

Following Notation 2, this could be written as $\underline{\epsilon}_i \sim \mathcal{N}(\underline{0}, \sigma^2(\mathbf{\Lambda}_i \mathbf{C}_i \mathbf{\Lambda}_i))$ s.t.

$$\mathbf{\Lambda}_i = \begin{bmatrix} |\text{TIME}_{i1}|^\delta & 0 & 0 & 0 \\ 0 & |\text{TIME}_{i2}|^\delta & 0 & 0 \\ 0 & 0 & |\text{TIME}_{i3}|^\delta & 0 \\ 0 & 0 & 0 & |\text{TIME}_{i4}|^\delta \end{bmatrix} \quad \text{and} \quad \mathbf{C}_i = \begin{bmatrix} 1 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 \\ 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 1 \end{bmatrix}$$

```
# Model 9.2
weights = varPower(form = ~time)
fm9.2 <- gls(lm1.form, weights = weights, data = armd)
```

- We include stratification in the model

$$\begin{aligned}\epsilon_{it} &\sim \mathcal{N}(0, \sigma_{it}^2) \quad \text{s.t.} \quad \sigma_{it} = \sigma \cdot \lambda_{it} \\ &= \sigma \cdot \lambda(\underline{\delta}, \text{TIME}_{it}) \\ &= \sigma \cdot \lambda([\delta_1, \delta_2]', \text{TIME}_{it}) \\ &= \begin{cases} \sigma \cdot |\text{TIME}_{it}|^{\delta_1} & \text{if active} \\ \sigma \cdot |\text{TIME}_{it}|^{\delta_2} & \text{if placebo} \end{cases}\end{aligned}$$

```
# Model 9.3
weights = varPower(form = ~time|treat.f)
fm9.3 <- gls(lm1.form, weights = weights, data = armd)
```

In Section 4 we also relax the independence assumption; we mainly employ **Notation 2**. Specifically, we want to modify \mathbf{C}_i , allowing the visual acuity measurements for the same individual to be correlated.

4 Correlation Structure: $\boxed{\underline{\epsilon}_i \sim \mathcal{N}(\underline{0}, \sigma^2 \mathbf{\Lambda}_i \mathbf{C}_i \mathbf{\Lambda}_i)}$

4.1 corCompSymm(.) (& varPower(.))

Compound Symmetry Correlation Structure

$$\mathbf{C}_i = \begin{bmatrix} 1 & \rho & \rho & \rho \\ \rho & 1 & \rho & \rho \\ \rho & \rho & 1 & \rho \\ \rho & \rho & \rho & 1 \end{bmatrix} \quad \text{and} \quad \mathbf{\Lambda}_i = \begin{bmatrix} |\text{TIME}_{i1}|^\delta & 0 & 0 & 0 \\ 0 & |\text{TIME}_{i2}|^\delta & 0 & 0 \\ 0 & 0 & |\text{TIME}_{i3}|^\delta & 0 \\ 0 & 0 & 0 & |\text{TIME}_{i4}|^\delta \end{bmatrix}$$

```
# Model 12.1
correlation = corCompSymm(form = ~1|subject)
weights = varPower(form = ~time)
fm12.1 <- gls(lm1.form,
              weights = weights,
              correlation = correlation,
              data = armd)
```

4.2 corAR1(.) (& varPower(.))

Heteroscedastic Autoregressive Residual Errors

$$\mathbf{C}_i = \begin{bmatrix} 1 & \rho & \rho^2 & \rho^3 \\ \rho & 1 & \rho & \rho^2 \\ \rho^2 & \rho & 1 & \rho \\ \rho^3 & \rho^2 & \rho & 1 \end{bmatrix} \quad \text{and} \quad \mathbf{\Lambda}_i = \begin{bmatrix} |\text{TIME}_{i1}|^\delta & 0 & 0 & 0 \\ 0 & |\text{TIME}_{i2}|^\delta & 0 & 0 \\ 0 & 0 & |\text{TIME}_{i3}|^\delta & 0 \\ 0 & 0 & 0 & |\text{TIME}_{i4}|^\delta \end{bmatrix}$$

```
# Model 12.2
correlation = corAR1(form = ~tp|subject)
# using (form = ~1 | subject) would be a mistake
weights = varPower(form = ~time)
fm12.2 <- gls(lm1.form,
              weights = weights,
              correlation = correlation,
              data = armd)
```

4.3 corSymm(.) (& varPower(.))

$$\mathbf{C}_i = \begin{bmatrix} 1 & \rho_{12} & \rho_{13} & \rho_{14} \\ \rho_{12} & 1 & \rho_{23} & \rho_{24} \\ \rho_{13} & \rho_{23} & 1 & \rho_{34} \\ \rho_{14} & \rho_{24} & \rho_{34} & 1 \end{bmatrix} \quad \text{and} \quad \mathbf{\Lambda}_i = \begin{bmatrix} |\text{TIME}_{i1}|^\delta & 0 & 0 & 0 \\ 0 & |\text{TIME}_{i2}|^\delta & 0 & 0 \\ 0 & 0 & |\text{TIME}_{i3}|^\delta & 0 \\ 0 & 0 & 0 & |\text{TIME}_{i4}|^\delta \end{bmatrix}$$

```
# Model 12.3
correlation = corSymm(form = ~tp|subject)
weights = varPower(form = ~time)
fm12.3 <- gls(lm1.form,
              weights = weights,
              correlation = correlation,
              data = armd)
```

4.4 corSymm(.) (& varIdent(.))

$$\mathbf{C}_i = \begin{bmatrix} 1 & \rho_{12} & \rho_{13} & \rho_{14} \\ \rho_{12} & 1 & \rho_{23} & \rho_{24} \\ \rho_{13} & \rho_{23} & 1 & \rho_{34} \\ \rho_{14} & \rho_{24} & \rho_{34} & 1 \end{bmatrix} \quad \text{and} \quad \mathbf{\Lambda}_i = \begin{bmatrix} 1 & 0 & 0 & 0 \\ 0 & \delta_2 & 0 & 0 \\ 0 & 0 & \delta_3 & 0 \\ 0 & 0 & 0 & \delta_4 \end{bmatrix}$$

```
# Model 12.3.b
correlation = corSymm(form = ~tp|subject)
weights = varIdent(form = ~1|time.f)
fm12.3.b <- gls(lm1.form,
               weights = weights,
               correlation = correlation,
               data = armd)
```

5 Linear Mixed Models

For patient i ($i = 1, \dots, 234$) at time t ($t = 1$ (4 wks), 2 (12 wks), 3 (24 wks), 4 (52 wks))

- **Notation 1:** We define our new model and **add a random intercept b_{0i}**

$$\text{VISUAL}_{it} = \beta_0 + \beta_1 \cdot \text{VISUAL0}_i + \beta_2 \cdot \text{TIME}_{it} + \beta_3 \cdot \text{TREAT}_i + \beta_4 \cdot \text{TREAT}_i \cdot \text{TIME}_{it} + b_{0i} + \epsilon_{it}$$

with $\epsilon_{it} \sim \mathcal{N}(0, \sigma_{it}^2)$ and $b_{0i} \sim \mathcal{N}(0, \sigma^2 d_{11})$

- **Notation 2:** Model equation at the patient level i

$$\underline{\text{VISUAL}}_i = \mathbb{X}_i \underline{\beta} + \underline{1}_i b_{0i} + \underline{\epsilon}_i$$

with $\underline{\epsilon}_i \sim \mathcal{N}(\underline{0}, \mathbf{R}_i)$ where $\mathbf{R}_i = \sigma^2 \mathbf{R}_i = \mathbf{\Lambda}_i \mathbf{C}_i \mathbf{\Lambda}_i$ and $b_{0i} \sim \mathcal{N}(0, \sigma^2 d_{11})$

$$\begin{bmatrix} \text{VISUAL}_{i1} \\ \text{VISUAL}_{i2} \\ \text{VISUAL}_{i3} \\ \text{VISUAL}_{i4} \end{bmatrix} \equiv \underline{\text{VISUAL}}_i \quad \text{and} \quad \begin{bmatrix} \epsilon_{i1} \\ \epsilon_{i2} \\ \epsilon_{i3} \\ \epsilon_{i4} \end{bmatrix} \equiv \underline{\epsilon}_i \quad \text{and} \quad \begin{bmatrix} 1 \\ 1 \\ 1 \\ 1 \end{bmatrix} \equiv \underline{1}_i$$

$$\mathbb{X}_i \underline{\beta} = \begin{bmatrix} 1 & \text{VISUAL0}_i & 4 & \text{TREAT}_i & 4 \cdot \text{TREAT}_i \\ 1 & \text{VISUAL0}_i & 12 & \text{TREAT}_i & 12 \cdot \text{TREAT}_i \\ 1 & \text{VISUAL0}_i & 24 & \text{TREAT}_i & 24 \cdot \text{TREAT}_i \\ 1 & \text{VISUAL0}_i & 52 & \text{TREAT}_i & 52 \cdot \text{TREAT}_i \end{bmatrix} \cdot \begin{bmatrix} \beta_0 \\ \beta_1 \\ \beta_2 \\ \beta_3 \\ \beta_4 \end{bmatrix}$$

More in general: **with both a random intercept and slopes $\underline{b}_i = [b_{0i} \ b_{1i} \ \dots]'$**

$$\underline{\text{VISUAL}}_i = \mathbb{X}_i \underline{\beta} + \mathbb{Z}_i \underline{b}_i + \underline{\epsilon}_i$$

with $\underline{\epsilon}_i \sim \mathcal{N}(\underline{0}, \mathbf{R}_i)$ where $\mathbf{R}_i = \sigma^2 \mathbf{R}_i$ and $\underline{b}_i \sim \mathcal{N}(\underline{0}, \mathbf{D})$ where $\mathbf{D} = \sigma^2 \mathbf{D}$

for instance, we can choose TIME to be our slope:

$$\mathbb{Z}_i = \begin{bmatrix} 1 & 4 \\ 1 & 12 \\ 1 & 24 \\ 1 & 52 \end{bmatrix} \quad \text{and} \quad \underline{b}_i = \begin{bmatrix} b_{0i} \\ b_{1i} \end{bmatrix}$$

```
# formula
lm2.form <- visual ~ visual0 + time + treat.f + treat.f:time
```

we know that $\mathbf{V}_i = \sigma^2 \mathbf{V}_i = \sigma^2 [\mathbb{Z}_i \mathbf{D} \mathbb{Z}_i' + \mathbf{R}_i]$

- with `'getVarCov(model, type = 'conditional')` we extract $\sigma^2 \mathbf{R}_i$;
- with `'getVarCov(model, type = 'marginal')` we extract $\sigma^2 \mathbf{V}_i$;
- with `'VarCorr(model)` we extract $\sigma^2 \mathbf{D}$ (also from the summary).

5.1 Homoscedastic residuals

5.1.1 Random intercept only

$$D = [d_{11}]$$

$$R_i = \Lambda_i \mathcal{C}_i \Lambda_i = \begin{bmatrix} 1 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 \\ 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 1 \end{bmatrix}$$

$$\mathbf{V}_i = \mathbb{Z}_i D \mathbb{Z}_i' + R_i = \begin{bmatrix} 1 \\ 1 \\ 1 \\ 1 \end{bmatrix} [d_{11}] \begin{bmatrix} 1 & 1 & 1 & 1 \end{bmatrix} + \begin{bmatrix} 1 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 \\ 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 1 \end{bmatrix}$$

$$\Rightarrow \mathbf{V}_i = \begin{bmatrix} 1 + d_{11} & d_{11} & d_{11} & d_{11} \\ d_{11} & 1 + d_{11} & d_{11} & d_{11} \\ d_{11} & d_{11} & 1 + d_{11} & d_{11} \\ d_{11} & d_{11} & d_{11} & 1 + d_{11} \end{bmatrix}$$

Note that the **implied marginal variance-covariance structure** is that of *compound symmetry* with a common correlation equal to $\rho = d_{11}/(1 + d_{11}) > 0$ since $d_{11} > 0$.

$$Var(\text{VISUAL}_{it}) = \sigma^2(d_{11} + 1)$$

```
# Model 16.1
# with lme4 library
fm16.1mer <- lmer(visual ~ visual0 + time * treat.f + (1|subject),
  data = armd)
# with nlme library
fm16.1 <- lme(lm2.form, random = ~1|subject, data = armd)
```

5.1.2 Random intercept & slope

- General D

$$D = \begin{bmatrix} d_{11} & d_{12} \\ d_{21} & d_{22} \end{bmatrix}$$

$$R_i = \Lambda_i \mathcal{C}_i \Lambda_i = \begin{bmatrix} 1 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 \\ 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 1 \end{bmatrix}$$

$$\mathbf{V}_i = \mathbb{Z}_i \mathbf{D} \mathbb{Z}_i' + \mathbf{R}_i = \begin{bmatrix} 1 & 4 \\ 1 & 12 \\ 1 & 24 \\ 1 & 52 \end{bmatrix} \begin{bmatrix} d_{11} & d_{12} \\ d_{21} & d_{22} \end{bmatrix} \begin{bmatrix} 1 & 1 & 1 & 1 \\ 4 & 12 & 24 & 52 \end{bmatrix} + \begin{bmatrix} 1 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 \\ 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 1 \end{bmatrix}$$

$$\text{Var}(\text{VISUAL}_{it}) = \sigma^2(d_{11} + 2d_{12}\text{TIME}_{it} + d_{22}\text{TIME}_{it}^2 + 1)$$

```
# Model 16.2A
# with lme4 library
fm16.2mer <- lmer(visual ~ visual0 + time * treat.f + (1+time|subject),
                  data = armd, control=lmerControl(optimizer="bobyqa",
                                                    optCtrl=list(maxfun=2e5)))

# with nlme library
fm16.2A <- lme(lm2.form, random = ~1 + time | subject, data = armd)
```

• **Diagonal \mathbf{D}**

$$\mathbf{D} = \begin{bmatrix} d_{11} & 0 \\ 0 & d_{22} \end{bmatrix}$$

$$\mathbf{R}_i = \mathbf{\Lambda}_i \mathbf{C}_i \mathbf{\Lambda}_i = \begin{bmatrix} 1 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 \\ 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 1 \end{bmatrix}$$

$$\mathbf{V}_i = \mathbb{Z}_i \mathbf{D} \mathbb{Z}_i' + \mathbf{R}_i = \begin{bmatrix} 1 & 4 \\ 1 & 12 \\ 1 & 24 \\ 1 & 52 \end{bmatrix} \begin{bmatrix} d_{11} & 0 \\ 0 & d_{22} \end{bmatrix} \begin{bmatrix} 1 & 1 & 1 & 1 \\ 4 & 12 & 24 & 52 \end{bmatrix} + \begin{bmatrix} 1 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 \\ 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 1 \end{bmatrix}$$

$$\text{Var}(\text{VISUAL}_{it}) = \sigma^2(d_{11} + d_{22}\text{TIME}_{it}^2 + 1)$$

```
# Model 16.2B
# with lme4 library
fm16.2dmer <- lmer(visual ~ visual0 + time * treat.f +
                  (1|subject) + (0 + time|subject),
                  data = armd, control=lmerControl(optimizer="bobyqa",
                                                    optCtrl=list(maxfun=2e5)))

# with nlme library
fm16.2B <- lme(lm2.form, random = list(subject = pdDiag(~time)), data = armd)
```


5.2 Heteroscedastic residuals: varPower(·)

5.2.1 Random intercept only

$$\mathbf{D} = [d_{11}]$$

$$\mathbf{R}_i = \mathbf{\Lambda}_i \mathbf{C}_i \mathbf{\Lambda}_i = \begin{bmatrix} |\text{TIME}_{i1}|^{2\delta} & 0 & 0 & 0 \\ 0 & |\text{TIME}_{i2}|^{2\delta} & 0 & 0 \\ 0 & 0 & |\text{TIME}_{i3}|^{2\delta} & 0 \\ 0 & 0 & 0 & |\text{TIME}_{i4}|^{2\delta} \end{bmatrix}$$

$$\mathbf{V}_i = \mathbb{Z}_i \mathbf{D} \mathbb{Z}_i' + \mathbf{R}_i = \begin{bmatrix} 1 \\ 1 \\ 1 \\ 1 \end{bmatrix} [d_{11}] \begin{bmatrix} 1 & 1 & 1 & 1 \end{bmatrix} + \begin{bmatrix} |\text{TIME}_{i1}|^{2\delta} & 0 & 0 & 0 \\ 0 & |\text{TIME}_{i2}|^{2\delta} & 0 & 0 \\ 0 & 0 & |\text{TIME}_{i3}|^{2\delta} & 0 \\ 0 & 0 & 0 & |\text{TIME}_{i4}|^{2\delta} \end{bmatrix}$$

$$\Rightarrow \mathbf{V}_i = \begin{bmatrix} |\text{TIME}_{i1}|^{2\delta} + d_{11} & d_{11} & d_{11} & d_{11} \\ d_{11} & |\text{TIME}_{i2}|^{2\delta} + d_{11} & d_{11} & d_{11} \\ d_{11} & d_{11} & |\text{TIME}_{i3}|^{2\delta} + d_{11} & d_{11} \\ d_{11} & d_{11} & d_{11} & |\text{TIME}_{i4}|^{2\delta} + d_{11} \end{bmatrix}$$

$$\text{Var}(\text{VISUAL}_{it}) = \sigma^2(d_{11} + |\text{TIME}_{it}|^{2\delta})$$

```
# Model 16.3 - nlme library only
fm16.3 <- lme(lm2.form, random = ~1|subject,
              weights = varPower(form = ~ time), data = armd)
```

5.2.2 Random intercept & slope

- General \mathbf{D}

$$\mathbf{D} = \begin{bmatrix} d_{11} & d_{12} \\ d_{21} & d_{22} \end{bmatrix}$$

$$\mathbf{R}_i = \mathbf{\Lambda}_i \mathbf{C}_i \mathbf{\Lambda}_i = \begin{bmatrix} |\text{TIME}_{i1}|^{2\delta} & 0 & 0 & 0 \\ 0 & |\text{TIME}_{i2}|^{2\delta} & 0 & 0 \\ 0 & 0 & |\text{TIME}_{i3}|^{2\delta} & 0 \\ 0 & 0 & 0 & |\text{TIME}_{i4}|^{2\delta} \end{bmatrix}$$

$$\mathbf{V}_i = \mathbb{Z}_i \mathbf{D} \mathbb{Z}_i' + \mathbf{R}_i = \begin{bmatrix} 1 & 4 \\ 1 & 12 \\ 1 & 24 \\ 1 & 52 \end{bmatrix} \begin{bmatrix} d_{11} & d_{12} \\ d_{21} & d_{22} \end{bmatrix} \begin{bmatrix} 1 & 1 & 1 & 1 \\ 4 & 12 & 24 & 52 \end{bmatrix} + \mathbf{R}_i$$

$$\text{Var}(\text{VISUAL}_{it}) = \sigma^2(d_{11} + 2d_{12}\text{TIME}_{it} + d_{22}\text{TIME}_{it}^2 + |\text{TIME}_{it}|^{2\delta})$$

```
# Model 16.4A - nlme library only
fm16.4A <- update(fm16.3,
  random = ~1 + time | subject,
  data = armd)
```

- Diagonal D

$$D = \begin{bmatrix} d_{11} & 0 \\ 0 & d_{22} \end{bmatrix}$$

$$R_i = \Lambda_i C_i \Lambda_i = \begin{bmatrix} |\text{TIME}_{i1}|^{2\delta} & 0 & 0 & 0 \\ 0 & |\text{TIME}_{i2}|^{2\delta} & 0 & 0 \\ 0 & 0 & |\text{TIME}_{i3}|^{2\delta} & 0 \\ 0 & 0 & 0 & |\text{TIME}_{i4}|^{2\delta} \end{bmatrix}$$

$$V_i = \mathbb{Z}_i D \mathbb{Z}_i' + R_i = \begin{bmatrix} 1 & 4 \\ 1 & 12 \\ 1 & 24 \\ 1 & 52 \end{bmatrix} \begin{bmatrix} d_{11} & 0 \\ 0 & d_{22} \end{bmatrix} \begin{bmatrix} 1 & 1 & 1 & 1 \\ 4 & 12 & 24 & 52 \end{bmatrix} + R_i$$

$$\text{Var}(\text{VISUAL}_{it}) = \sigma^2(d_{11} + d_{22}\text{TIME}_{it}^2 + |\text{TIME}_{it}|^{2\delta})$$

```
# Model 16.4B - nlme library only
fm16.4B <- update(fm16.3,
  random = list(subject = pdDiag(~time)), # Diagonal D
  data = armd)
```