

Random Variable

A Random Variable X is a symbol that represents the outcome of an experiment.

Example - consider an experiment of tossing a pair of identical coins and let the random variable X be the number of heads obtained. The sample space S is follows.

Outcome	TT	HT	TH	HH
Value of X	0	1	1	2

$$P(X=0) = \frac{1}{4}, \quad P(X=1) = \frac{2}{4} = \frac{1}{2} \quad \text{and} \quad P(X=2) = \frac{1}{4}$$

Random Variable

Discrete

continuous

→ Discrete Random Variable

A discrete variable is a variable which can only take a countable numbers of values.

Eg. $(0, 1, 2, 3)$,

→ Continuous Random Variable

A continuous random variable is random variable where the data can take infinitely many values

Example $[0, 1]$, $[a, b]$

Distribution function

The cumulative distribution function (CDF also cumulative density function) of a real valued random variable X or just distribution function of X evaluated at x is the probability that X will take a value less than or equal to x .

$$F_X(x) = P_X(X \leq x) \quad \forall x$$

Remark Domain of Distribution function is $(-\infty, \infty)$ and Range is $[0, 1]$

→ Discrete Probability Distribution

The set of values x_i with their probability p_i constitute a discrete probability distribution of the discrete variate X

$P(x) = P(X = x)$ and if x_1, x_2, \dots are the possible values of X .

where (i) $P(x_i) \geq 0 \quad \forall i$

(ii) $\sum P(x_i) = 1$

→ Continuous Probability Distribution

$$\text{if } F(x) = P(X \leq x) = \int_{-\infty}^x f(x) dx$$

$$\text{or } P\left(x - \frac{1}{2} dx \leq X \leq x + \frac{1}{2} dx\right) = f(x) dx$$

where $f(x)$ is defined as the cumulative distribution function or simply the distribution function of the continuous variate X .

It is the probability that the value of the variate X will be $\leq x$.

→ The distribution function $f(x)$ has the following Properties

(i) $f'(x) = f(x) \geq 0$

So that $f(x)$ is a non-decreasing function.

(ii) $f(-\infty) = 0$

(iii) $f(\infty) = 1$

(iv)
$$P(a \leq x \leq b) = \int_a^b f(x) dx = \int_{-\infty}^b f(x) dx - \int_{-\infty}^a f(x) dx$$
$$= f(b) - f(a)$$

Example 26.28. A die is tossed thrice. A success is 'getting 1 or 6' on a toss. Find the mean and variance of the number of successes.
(V.T.U., 2011 S; Rohtak, 2004)

Solution. Probability of a success = $\frac{2}{6} = \frac{1}{3}$, Probability of failures = $1 - \frac{1}{3} = \frac{2}{3}$.

$$\therefore \text{prob. of no success} = \text{Prob. of all 3 failures} = \frac{2}{3} \times \frac{2}{3} \times \frac{2}{3} = \frac{8}{27}$$

$$\text{Probability of one successes and 2 failures} = {}^3C_1 \times \frac{1}{3} \times \frac{2}{3} \times \frac{2}{3} = \frac{4}{9}$$

$$\text{Probability of two successes and one failure} = {}^3C_2 \times \frac{1}{3} \times \frac{1}{3} \times \frac{2}{3} = \frac{2}{9}$$

$$\text{Probability of three successes} = \frac{1}{3} \times \frac{1}{3} \times \frac{1}{3} = \frac{1}{27}$$

Now	$x_i = 0$	1	2	3
	$p_i = 8/27$	4/9	2/9	1/27

$$\therefore \text{mean} \quad \mu = \sum p_i x_i = 0 + \frac{4}{9} + \frac{4}{9} + \frac{1}{9} = 1.$$

$$\text{Also} \quad \sum p_i x_i^2 = 0 + \frac{4}{9} + \frac{8}{9} + \frac{9}{27} = \frac{5}{3}$$

$$\therefore \text{variance} \quad \sigma^2 = \sum p_i x_i^2 - \mu^2 = \frac{5}{3} - 1 = \frac{2}{3}.$$

Example 26.29. The probability density function of a variate X is

$X :$	0	1	2	3	4	5	6
$p(X) :$	k	$3k$	$5k$	$7k$	$9k$	$11k$	$13k$

(i) Find $P(X < 4)$, $P(X \geq 5)$, $P(3 < X \leq 6)$.

(V.T.U., 2013)

(ii) What will be the minimum value of k so that $P(X \leq 2) > 3$.

Solution. (i) If X is a random variable, then

$$\sum_{i=0}^6 p(x_i) = 1 \text{ i.e., } k + 3k + 5k + 7k + 9k + 11k + 13k = 1 \text{ or } k = 1/49.$$

$$\therefore P(X < 4) = k + 3k + 5k + 7k = 16k = 16/49.$$

$$P(X \geq 5) = 11k + 13k = 24k = 24/49.$$

$$P(3 < X \leq 6) = 9k + 11k + 13k = 33k = 33/49.$$

$$(ii) \quad P(X \leq 2) = k + 3k + 5k = 9k > 0.3 \text{ or } k > 1/30$$

Thus minimum value of $k = 1/30$.

Example 26.31. (i) Is the function defined as follows a density function?

$$f(x) = e^{-x}, \quad x \geq 0$$

$$= 0, \quad x < 0,$$

if yes, determine the probability that the variate having this density will fall in the interval $(1, 2)$ and also find the cumulative probability function $F(x)$?

Solution. (i) $f(x)$ is clearly > 0 for every x in $(1, 2)$ and

$$\int_{-\infty}^{\infty} f(x) dx = \int_{-\infty}^0 0 \cdot dx + \int_0^{\infty} e^{-x} dx = 1$$

874

Hence the function $f(x)$ satisfies the requirements for a density function.

(ii) Required probability $= P(1 \leq x \leq 2) = \int_1^2 e^{-x} dx = e^{-1} - e^{-2} = 0.368 - 0.135 = 0.233$.

This probability is equal to the shaded area in Fig. 26.3 (a).

(iii) Cumulative probability function $F(x)$

$$\int_{-\infty}^2 f(x) dx = \int_{-\infty}^0 0 \cdot dx + \int_0^2 e^{-x} dx = 1 - e^{-2} = 1 - 0.135 = 0.865$$

which is shown in Fig. 26.3 (b).

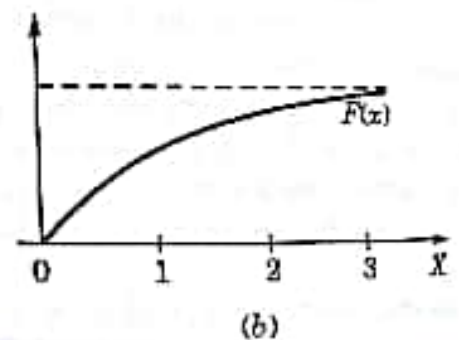
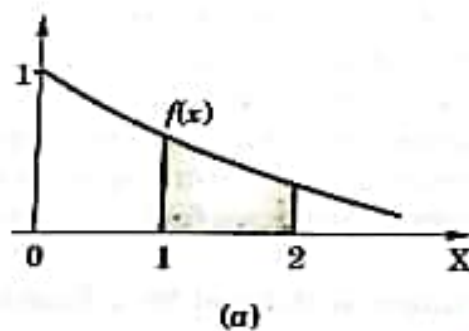


Fig. 26.3

26.10 (1) EXPECTATION

The mean value (μ) of the probability distribution of a variate X is commonly known as its expectation and is denoted by $E(X)$. If $f(x)$ is the probability density function of the variate X , then

$$\sum_i x_i f(x_i)$$

(discrete distribution)

or

$$E(X) = \int_{-\infty}^{\infty} x f(x) dx$$

(continuous distribution)

In general, expectation of any function $\phi(x)$ is given by

$$E[\phi(x)] = \sum_i \phi(x_i) f(x_i)$$

(discrete distribution)

or

$$E[\phi(x)] = \int_{-\infty}^{\infty} \phi(x) f(x) dx$$

(continuous distribution)

(2) Variance of a distribution is given by

$$\sigma^2 = \sum_i (x_i - \mu)^2 f(x_i)$$

(discrete distribution)

or

$$\sigma^2 = \int_{-\infty}^{\infty} (x - \mu)^2 f(x) dx$$

(continuous distribution)

where σ is the standard deviation of the distribution.

(3) The **rth moment** about the mean (denoted by μ_r) is defined by

$$\mu_r = \sum_i (x_i - \mu)^r f(x_i)$$

(discrete distribution)

or

$$\mu_r = \int_{-\infty}^{\infty} (x - \mu)^r f(x) dx$$

(continuous distribution)

(4) Mean deviation from the mean is given by

$$\sum_i |x_i - \mu| f(x_i)$$

(discrete distribution)

or by

$$\int_{-\infty}^{\infty} |x - \mu| f(x) dx$$

(continuous distribution)

Example 26.32. In a lottery, m tickets are drawn at a time out of n tickets numbered from 1 to n . Find the expected value of the sum of the numbers on the tickets drawn.
(Rohtak, 2011)

Solution. Let x_1, x_2, \dots, x_n be the variables representing the numbers on the first, second, ..., n th ticket. The probability of drawing a ticket out of n tickets being in each case $1/n$, we have

$$E(x_i) = 1 \cdot \frac{1}{n} + 2 \cdot \frac{1}{n} + 3 \cdot \frac{1}{n} + \dots + n \cdot \frac{1}{n} = \frac{1}{2} (n+1)$$

$$\begin{aligned} \therefore \text{expected value of the sum of the numbers on the tickets drawn} \\ &= E(x_1 + x_2 + \dots + x_m) = E(x_1) + E(x_2) + \dots + E(x_m) \\ &= mE(x_i) = \frac{1}{2} m (n+1). \end{aligned}$$

Example 26.33. X is a continuous random variable with probability density function given by

$$\begin{aligned} f(x) &= kx \quad (0 \leq x < 2) \\ &= 2k \quad (2 \leq x < 4) \\ &= -kx + 6k \quad (4 \leq x < 6) \end{aligned}$$

(J.N.T.U., 2003)

Find k and mean value of X .

Solution. Since the total probability is unity

$$\int_0^6 f(x) dx = 1$$

$$\int_0^2 kx dx + \int_2^4 2k dx + \int_4^6 (-kx + 6k) dx = 1$$

$$\begin{aligned} k \left[\frac{x^2}{2} \right]_0^2 + 2k \left[x \right]_2^4 + \left(-\frac{kx^2}{2} + 6kx \right)_4^6 &= 1 \\ 2k + 4k + (-10k + 12k) &= 1 \text{ i.e., } k = 1/8. \end{aligned}$$

$$\text{Mean of } X = \int_0^6 x f(x) dx$$

$$= \int_0^2 kx^2 dx + \int_2^4 2kx dx + \int_4^6 x(-kx + 6k) dx$$

$$= k \left[\frac{x^3}{3} \right]_0^2 + 2k \left[\frac{x^2}{2} \right]_2^4 + \left(-k \left[\frac{x^3}{3} \right]_4^6 + 6k \left[\frac{x^2}{2} \right]_4^6 \right)$$

$$= k(8/3) + k(12) - k(152/3) + 3k(20) = \frac{1}{8}(24) = 3.$$

Example 26.34. A variate X has the probability distribution

x	:	-3	6	9
$P(X=x)$:	1/6	1/2	1/3

Find $E(X)$ and $E(X^2)$. Hence evaluate $E(2X+1)^2$.

Solution.

$$E(X) = -3 \times \frac{1}{6} + 6 \times \frac{1}{2} + 9 \times \frac{1}{3} = 11/2.$$

$$E(X^2) = 9 \times \frac{1}{6} + 36 \times \frac{1}{2} + 81 \times \frac{1}{3} = 93/2$$

$$\begin{aligned} E(2X+1)^2 &= E(4X^2 + 4X + 1) = 4E(X^2) + 4E(X) + 1 \\ &= 4(93/2) + 4(11/2) + 1 = 209. \end{aligned}$$

Example 26.35. The frequency distribution of a measurable characteristic varying between 0 and 2 is as under

$$\begin{aligned} f(x) &= x^3, & 0 \leq x \leq 1 \\ &= (2-x)^3, & 1 \leq x \leq 2. \end{aligned}$$

Calculate the standard deviation and also the mean deviation about the mean.

Solution. Total frequency $N = \int_0^1 x^3 dx + \int_1^2 (2-x)^3 dx = \frac{1}{4} + \frac{1}{4} = \frac{1}{2}$

$$\begin{aligned} \therefore \mu_1' \text{ (about the origin)} &= \frac{1}{N} \left[\int_0^1 x \cdot x^3 dx + \int_1^2 x(2-x)^3 dx \right] \\ &= 2 \left\{ \left[\frac{x^5}{5} \right]_0^1 + \left[-x \cdot \frac{(2-x)^4}{4} \right]_1^2 - \left[\frac{(2-x)^5}{20} \right]_1^2 \right\} = 2 \left(\frac{1}{5} + \frac{1}{4} + \frac{1}{20} \right) = 1 \end{aligned}$$

$$\begin{aligned} \mu_2' \text{ (about the origin)} &= \frac{1}{N} \left[\int_0^1 x^2 \cdot x^3 dx + \int_1^2 x^2 (2-x)^3 dx \right] \\ &= 2 \left\{ \left[\frac{x^6}{6} \right]_0^1 + \left[-x^2 \frac{(2-x)^4}{4} \right]_1^2 + \frac{1}{2} \int_1^2 x(2-x)^4 dx \right\} \\ &= 2 \left\{ \frac{1}{6} + \frac{1}{4} + \frac{1}{2} \left[\frac{1}{5} + \frac{1}{30} \right] \right\} = \frac{16}{15} \end{aligned}$$

Hence

$$\sigma^2 = \mu_2 = \mu_2' - (\mu_1')^2 = \frac{1}{15}$$

i.e., standard deviation $\sigma = \frac{1}{\sqrt{15}}$.

Mean deviation about the mean

$$\begin{aligned} &= \frac{1}{N} \left\{ \int_0^1 |x-1| x^3 dx + \int_1^2 |x-1| (2-x)^3 dx \right\} \\ &= 2 \left\{ \int_0^1 (1-x)x^3 dx + \int_1^2 (x-1)(2-x)^3 dx \right\} \\ &= 2 \left\{ \left(\frac{1}{4} - \frac{1}{5} \right) + \left(0 + \frac{1}{20} \right) \right\} = \frac{1}{5} \end{aligned}$$

1008

COVARIANCE AND VARIANCE

Covariance: If $E(X) = \bar{X}$ and $E(Y) = \bar{Y}$, then the covariance of the random variables X and Y , denoted by $\text{cov}(X, Y)$, is given by

$$\begin{aligned}\text{cov}(X, Y) &= E[(X - \bar{X})(Y - \bar{Y})] = E(XY - X\bar{Y} - Y\bar{X} + \bar{X}\bar{Y}) \\ &= E(XY) - \bar{Y}E(X) - \bar{X}E(Y) + \bar{X}\bar{Y} \\ &= E(XY) - \bar{Y}\bar{X} - \bar{X}\bar{Y} + \bar{X}\bar{Y}\end{aligned}$$

or

$$\text{cov}(X, Y) = E(XY) - \bar{X}\bar{Y}.$$

Note that, when X and Y are independent, we have
 $E(XY) = E(X)E(Y)$, then $\text{cov}(X, Y) = 0$.

Also we can easily establish that

1. $\text{cov}(aX, bY) = ab \text{cov}(X, Y)$
2. $\text{cov}(X + a, Y + b) = \text{cov}(X, Y)$
3. $\text{cov}\left(\frac{X-a}{h}, \frac{Y-b}{k}\right) = \frac{1}{hk} \text{cov}(X, Y)$
4. $\text{cov}(X, Y, Z) = \text{cov}(X, Y) + \text{cov}(X, Z)$.

Variance: It is easy to verify the following properties:

1. $\text{Var}(X) = \text{cov}(X, X)$
2. $\text{Var}(ax + b) = a^2 \text{Var}(X)$
3. $\text{Var}(aX + bY) = a^2 \text{Var}(X) + b^2 \text{Var}(Y) + 2ab \text{cov}(X, Y)$.

MOMENT GENERATING FUNCTION

The moment generating function (m.g.f.) of a random variable X about the origin is defined as

$$M_0(t) = E(e^{tX}) = \sum e^{tX} P(X).$$

This function $M_0(t)$ is called moment generating function because all the moments of X can be obtained from $M_0(t)$, as follows:

$$\begin{aligned}M_0(t) &= \left[1 + tX + \frac{t^2}{2!}X^2 + \frac{t^3}{3!}X^3 + \dots + \frac{t^r}{r!}X^r + \dots \right] p(X) \\ &= 1 + tE(X) + \frac{t^2}{2!}E(X^2) + \dots + \frac{t^r}{r!}E(X^r) + \dots \\ &= 1 + t\mu'_1 + \frac{t^2}{2!}\mu'_2 + \dots + \frac{t^r}{r!}\mu'_r + \dots \quad \text{where } \mu'_r = E(X^r).\end{aligned}$$

Obviously, we can write here

$$\mu'_r = \frac{d^r}{dt^r} [M_0(t)].$$

As a general case, the moment generating function of the random variable X about an arbitrary point 'a' which can be \bar{X} (or μ) also, is defined as

$$M_a(t) = E(e^{t(X-a)}) = e^{-at} E(e^{tX}).$$

Moment generating function of the sum of two independent variables is the product of their moment generating functions.

To establish it, consider two independent random variables X and Y whose m.g.fs. are $M_X(t)$ and $M_Y(t)$ respectively, then m.g.f. of $X + Y$ is given by

$$\begin{aligned} M_{X+Y}(t) &= E[e^{t(X+Y)}] = E(e^{tX} \cdot e^{tY}) \\ &= E(e^{tX}) E(e^{tY}) \end{aligned} \quad (\text{since } X \text{ and } Y \text{ are independent})$$

Thus, $M_{X+Y}(t) = M_X(t) M_Y(t)$.

Further, it is easy to verify that

1. $M_{cX}(t) = M_X(ct)$ where c is some constant

2. If $U = \frac{X-a}{h}$ then $M_U(t) = e^{at/h} M_X\left(\frac{t}{h}\right)$.

EXAMPLE 31.5.

Find the probability distribution of the number of green balls drawn when three balls are drawn one by one without replacement from a bag containing three green and five white balls.

SOLUTION: Let X be the random variable which is the number of green balls drawn when three balls are drawn without replacement.

$$\therefore P(X=0) = P(\text{no green ball is drawn}) = P(W, W, W)$$

$$= \frac{5}{8} \cdot \frac{4}{7} \cdot \frac{3}{6}$$

$$P(X=1) = P(\text{one green ball is drawn})$$

$$= P(G, W, W) + P(W, G, W) + P(W, W, G)$$

$$= \frac{3}{8} \cdot \frac{5}{7} \cdot \frac{4}{6} + \frac{5}{8} \cdot \frac{3}{7} \cdot \frac{4}{6} + \frac{5}{8} \cdot \frac{4}{7} \cdot \frac{3}{6}$$

$$P(X=2) = P(\text{two green balls are drawn})$$

$$= P(G, G, W) + P(G, W, G) + P(W, G, G)$$

$$= \frac{3}{8} \cdot \frac{2}{7} \cdot \frac{5}{6} + \frac{3}{8} \cdot \frac{5}{7} \cdot \frac{2}{6} + \frac{5}{8} \cdot \frac{3}{7} \cdot \frac{3}{6}$$

$$P(X=3) = P(G, G, G) = \frac{3}{8} \cdot \frac{2}{7} \cdot \frac{1}{6}$$

Probability distribution is

X	0	1	2	3
$P(X)$	$\frac{5}{28}$	$\frac{15}{28}$	$\frac{15}{28}$	$\frac{1}{56}$

EXAMPLE 31.7.

A random variate X has the following probability distribution

x	0	1	2	3	4
$p(x)$	k	$2k$	$2k$	k^2	$5k^2$

Find the value of k and obtain $P(X < 3)$ and $P(0 < X < 4)$.

Determine the distribution function of X , also mean and variance of X .

SOLUTION: Since $\sum p(x) = 1$ we have $k + 2k + 2k + k^2 + 5k^2 = 1$ or $6k^2 + 4k + k - 1 = 0$ which gives $k = -1$ and $k = 1/6$. Therefore $k = 1/6$ as k cannot be negative. Ans.

Next, $P(X < 3) = p(0) + p(1) + p(2) = \frac{1}{6} + \frac{2}{6} + \frac{2}{6} = \frac{5}{6}$.

and $P(0 < X < 4) = p(1) + p(2) + p(3) = \frac{2}{6} + \frac{2}{6} + \frac{1}{36} = \frac{25}{36}$

Now, the probability distribution and the distribution function is as follows:

x	0	1	2	3	4
$p(x)$	$\frac{1}{6}$	$\frac{2}{6}$	$\frac{2}{6}$	$\frac{1}{36}$	$\frac{5}{36}$
$F(x)$	$\frac{1}{6}$	$\frac{3}{6}$	$\frac{5}{6}$	$\frac{31}{36}$	1

Then $\text{mean} = \mu_x = \sum x_i p_i = 0 + \frac{2}{6} + \frac{4}{6} + \frac{3}{36} + \frac{20}{36} = \frac{59}{36}$

$$\text{Variance} = \sigma_x^2 = \left[0\left(\frac{1}{6}\right) + 1\left(\frac{2}{6}\right) + 4\left(\frac{2}{6}\right) + 9\left(\frac{1}{36}\right) + 16\left(\frac{5}{36}\right) \right] - \left(\frac{59}{36}\right)^2$$

Two most important discrete probability distributions are – Binomial Distribution and Poisson distribution.