



Lindeberg-Feller CLT and Lyapunov's Condition

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1 Lindeberg-Feller Central Limit Theorem.

1.1 About the Theorem.

Sufficiency is proved by Lindeberg in 1922 and necessity by Feller in 1935. *Lindeberg-Feller CLT* is one of the most far-reaching results in probability theory. Nearly all generalizations of various types of central limit theorems spin from Lindeberg-Feller CLT.

The insights of the Lindeberg condition are that the “wild” values of the random variables, compared with s_n , the standard deviation of S_n as the normalizing constant, are insignificant and can be truncated off without affecting the general behavior of the partial sum S_n .

1.2 Lindeberg - Feller CLT.

Suppose X_1, X_2, \dots, X_n are independent r.v.s with mean 0 and variance σ_n^2 . Let $s_n^2 = \sum_{i=1}^n \sigma_j^2$ denote the variance of partial sum $S_n = X_1 + X_2 + \dots + X_n$. If, for every $\epsilon > 0$,

$$\frac{1}{s_n^2} \sum_{j=1}^n \mathbb{E} (X_j^2 \mathbf{1}_{\{|X_j| > \epsilon s_n\}}) \rightarrow 0,$$

then $\frac{S_n}{s_n} \xrightarrow{\mathcal{D}} \mathcal{N}(0, 1)$.

Conversely, if $\max_{1 \leq j \leq n} \frac{\sigma_j^2}{s_n^2} \rightarrow 0$ and $\frac{S_n}{s_n} \xrightarrow{\mathcal{D}} \mathcal{N}(0, 1)$, then Lindeberg Condition holds.

1.3 Prerequisites

1.

$$\forall x > 0, |e^{-x} - 1 + x| \leq x^2/2$$

2. For complex z_j and w_j with $|z_j| \leq 1$ and $|w_j| \leq 1$, $\left| \prod_{j=1}^n z_j - \prod_{j=1}^n w_j \right| \leq \sum_{j=1}^n |z_j - w_j|$

3.

$$\forall x \in \mathbb{R} \quad \cos x - 1 \geq -\frac{x^2}{2}$$

4.

$$|\log(x+1) - x| \leq x^2 \quad \text{for } x > 0$$

5. $\forall k \geq 1$

$$\left| \varphi_X(\lambda) - 1 - \sum_{j=1}^k \frac{(i\lambda)^j}{j!} \mathbb{E}(X^j) \right| \leq \min \left\{ \frac{2|\lambda|^k \mathbb{E}|X|^k}{k!}, \frac{|\lambda|^{k+1} \mathbb{E}|X|^{k+1}}{(k+1)!} \right\} \quad \forall \lambda \in \mathbb{R}$$

2 If part

2.1 Proof of Lindeberg-Feller CLT contd.

“ \Leftarrow ” The Lindeberg Condition implies,

$$\begin{aligned}
 \sigma_j^2 &= \mathbb{E} (X_j^2 \mathbf{1}_{\{|X_j| > \epsilon s_n\}}) + \mathbb{E} (X_j^2 \mathbf{1}_{\{|X_j| \leq \epsilon s_n\}}) \\
 &\leq \mathbb{E} (X_j^2 \mathbf{1}_{\{|X_j| > \epsilon s_n\}}) + \epsilon^2 s_n^2 \quad \forall j \\
 &\leq \sum_{i=1}^n \mathbb{E} (X_j^2 \mathbf{1}_{\{|X_j| > \epsilon s_n\}}) + \epsilon^2 s_n^2 \quad \forall j \\
 \Rightarrow \frac{\sigma_j^2}{s_n^2} &\leq \frac{1}{s_n^2} \sum_{i=1}^n \mathbb{E} (X_j^2 \mathbf{1}_{\{|X_j| > \epsilon s_n\}}) + \epsilon^2 \quad \forall j \\
 \therefore \max_{1 \leq j \leq n} \left(\frac{\sigma_j^2}{s_n^2} \right) &\leq \underbrace{\frac{1}{s_n^2} \sum_{i=1}^n \mathbb{E} (X_j^2 \mathbf{1}_{\{|X_j| > \epsilon s_n\}})}_{\rightarrow 0} + \epsilon^2
 \end{aligned}$$

by letting $n \rightarrow \infty$ and then $\epsilon \downarrow 0$,

$$\max_{1 \leq j \leq n} \left(\frac{\sigma_j^2}{s_n^2} \right) \rightarrow 0$$

For any $\epsilon > 0$, then $\forall j$

$$\begin{aligned}
 &\left| \mathbb{E} \left(e^{itX_j/s_n} \right) - e^{-t^2 \sigma_j^2 / 2s_n^2} \right| \\
 &= \left| \mathbb{E} \left(e^{itX_j/s_n} \right) - \mathbb{E} \left(1 + it \frac{X_j}{s_n} - \frac{t^2}{2} \frac{X_j^2}{s_n^2} \right) + 1 + it \mathbb{E} \left(\frac{X_j}{s_n} \right) - \frac{t^2}{2} \mathbb{E} \left(\frac{X_j^2}{s_n^2} \right) - e^{-t^2 \sigma_j^2 / 2s_n^2} + 1 - \frac{t^2 \sigma_j^2}{2s_n^2} - 1 + \frac{t^2 \sigma_j^2}{2s_n^2} \right| \\
 &\leq \left| \mathbb{E} \left(e^{itX_j/s_n} \right) - \mathbb{E} \left(1 + it \frac{X_j}{s_n} - \frac{t^2}{2} \frac{X_j^2}{s_n^2} \right) \right| + \left| e^{-t^2 \sigma_j^2 / 2s_n^2} - 1 + \frac{t^2 \sigma_j^2}{2s_n^2} \right| \\
 &\leq \mathbb{E} \left[\min \left(\frac{t^2 X_j^2}{s_n^2}, \frac{|tX_j|^3}{6s_n^3} \right) \right] + \frac{t^4 \sigma_j^4}{8s_n^4} \\
 &\leq \mathbb{E} \left[\min \left(\frac{t^2 X_j^2}{s_n^2}, \frac{|tX_j|^3}{6s_n^3} \right) \mathbf{1}_{\{|X_j| > \epsilon s_n\}} \right] + \mathbb{E} \left[\min \left(\frac{t^2 X_j^2}{s_n^2}, \frac{|tX_j|^3}{6s_n^3} \right) \mathbf{1}_{\{|X_j| \leq \epsilon s_n\}} \right] + \frac{t^4 \sigma_j^4}{8s_n^4} \\
 &\leq \mathbb{E} \left[\frac{t^2 X_j^2}{s_n^2} \mathbf{1}_{\{|X_j| > \epsilon s_n\}} \right] + \mathbb{E} \left[\frac{|tX_j|^3}{6s_n^3} \mathbf{1}_{\{|X_j| \leq \epsilon s_n\}} \right] + \frac{t^4 \sigma_j^4}{8s_n^4} \\
 &\leq \frac{t^2}{s_n^2} \mathbb{E} [X_j^2 \mathbf{1}_{\{|X_j| > \epsilon s_n\}}] + \frac{|t|^3 \epsilon}{6s_n^2} \mathbb{E} [X_j^2] + \frac{t^4 \sigma_j^2}{s_n^2} \max_{1 \leq k \leq n} \left(\frac{\sigma_k^2}{s_n^2} \right)
 \end{aligned}$$

2.2 Proof of Lindeberg-Feller CLT contd.

Then for any fixed t ,

$$\begin{aligned}
& \left| \mathbb{E} \left(e^{itS_n/s_n} \right) - e^{-t^2/2} \right| \\
&= \left| \prod_{j=1}^n \mathbb{E} \left(e^{itX_j/s_n} \right) - \prod_{j=1}^n \left(e^{-t^2\sigma_j^2/2s_n^2} \right) \right| \\
&\leq \sum_{j=1}^n \left| \mathbb{E} \left(e^{itX_j/s_n} \right) - e^{-t^2\sigma_j^2/2s_n^2} \right| \\
&\leq \sum_{j=1}^n \left(\frac{t^2}{s_n^2} \mathbb{E} [X_j^2 \mathbf{1}_{\{|X_j| > \epsilon s_n\}}] + \frac{|t|^3 \epsilon}{6s_n^2} \mathbb{E} [X_j^2] + \frac{t^4 \sigma_j^2}{s_n^2} \max_{1 \leq k \leq n} \left(\frac{\sigma_k^2}{s_n^2} \right) \right) \\
&\leq \left(\frac{t^2}{s_n^2} \sum_{j=1}^n \mathbb{E} [X_j^2 \mathbf{1}_{\{|X_j| > \epsilon s_n\}}] + \frac{|t|^3 \epsilon}{6} + t^4 \max_{1 \leq k \leq n} \left(\frac{\sigma_k^2}{s_n^2} \right) \right) \\
&\rightarrow \frac{\epsilon |t|^3}{6} \text{ as } n \rightarrow \infty
\end{aligned}$$

Since, $\epsilon > 0$ is arbitrary, it follows that, $\mathbb{E} (e^{itS_n/s_n}) \rightarrow e^{-t^2/2}$ for all t . Levy's continuity theorem implies,

$$\frac{S_n}{s_n} \xrightarrow{\mathcal{D}} \mathcal{N}(0, 1)$$

3 Only If Part.

3.1 Proof of the only if part.

" \implies " Let ψ_j be the Characteristic Function of X_j . The asymptotic normality is equivalent to,

$$\prod_{j=1}^n \psi_j \left(\frac{t}{s_n} \right) \rightarrow e^{-\frac{t^2}{2}}$$

Notice that,

$$\left| \psi_j \left(\frac{t}{s_n} \right) - 1 \right| \leq \frac{t^2 \sigma_j^2}{2s_n^2}$$

Write as $n \rightarrow \infty$,

$$\sum_{j=1}^n \left(\psi_j \left(\frac{t}{s_n} \right) - 1 \right) + \frac{t^2}{2}$$

$$\begin{aligned}
&= \sum_{j=1}^n \left(\psi_j \left(\frac{t}{s_n} \right) - 1 - \log \psi_j \left(\frac{t}{s_n} \right) \right) + \underbrace{\sum_{i=1}^n \log \psi_j \left(\frac{t}{s_n} \right) + \frac{t^2}{2}}_{=o(1)} \\
&\leq \sum_{j=1}^n \left| \psi_j \left(\frac{t}{s_n} \right) - 1 \right|^2 + o(1) \\
&\leq \max_{1 \leq j \leq n} \left| \psi_j \left(\frac{t}{s_n} \right) - 1 \right| \times \sum_{j=1}^n \left| \psi_j \left(\frac{t}{s_n} \right) - 1 \right| + o(1) \\
&\leq \max_{1 \leq j \leq n} \frac{t^2 \sigma_j^2}{2s_n^2} \times \sum_{j=1}^n \frac{t^2 \sigma_j^2}{2s_n^2} + o(1) = o(1)
\end{aligned}$$

by the assumption, $\max_{1 \leq k \leq n} \left(\frac{\sigma_k^2}{s_n^2} \right) \rightarrow 0$

3.2 Proof of the only if part contd.

On the other hand, by definition of characteristic function, the above expression is, as $n \rightarrow \infty$,

$$\begin{aligned}
o(1) &= \sum_{j=1}^n \left(\psi_j \left(\frac{t}{s_n} \right) - 1 \right) + \frac{t^2}{2} \\
&= \sum_{j=1}^n \mathbb{E} \left(e^{itX_j/s_n} - 1 \right) + \frac{t^2}{2} \\
&= \sum_{j=1}^n \mathbb{E} \left(\cos \left(\frac{tX_j}{s_n} \right) - 1 \right) + \frac{t^2}{2} + i \sum_{j=1}^n \mathbb{E} \left(\sin \left(\frac{tX_j}{s_n} \right) \right) \\
&= \sum_{j=1}^n \mathbb{E} \left\{ \left(\cos \left(\frac{tX_j}{s_n} \right) - 1 \right) \mathbf{1}_{\{|X_j| > \epsilon s_n\}} \right\} + \sum_{j=1}^n \mathbb{E} \left\{ \left(\cos \left(\frac{tX_j}{s_n} \right) - 1 \right) \mathbf{1}_{\{|X_j| \leq \epsilon s_n\}} \right\} + \frac{t^2}{2} + \text{imaginary part (immaterial)}
\end{aligned}$$

Since, $\cos x - 1 \geq -\frac{x^2}{2}$ for all real x ,

$$\begin{aligned}
&\frac{1}{s_n^2} \sum_{j=1}^n \mathbb{E} \left(X_j^2 \mathbf{1}_{\{|X_j| > \epsilon s_n\}} \right) = 1 - \frac{2}{t^2} \sum_{j=1}^n \mathbb{E} \left(\frac{t^2 X_j^2}{2s_n^2} \mathbf{1}_{\{|X_j| \leq \epsilon s_n\}} \right) \\
&\leq \frac{2}{t^2} \left(\frac{t^2}{2} + \sum_{j=1}^n \mathbb{E} \left\{ \left(\cos \left(\frac{tX_j}{s_n} \right) - 1 \right) \mathbf{1}_{\{|X_j| \leq \epsilon s_n\}} \right\} \right) \\
&\leq \frac{2}{t^2} \left(\left| \sum_{j=1}^n \mathbb{E} \left\{ \left(\cos \left(\frac{tX_j}{s_n} \right) - 1 \right) \mathbf{1}_{\{|X_j| > \epsilon s_n\}} \right\} \right| + o(1) \right)
\end{aligned}$$

$$\begin{aligned}
 &\leq \frac{2}{t^2} \sum_{j=1}^n 2\mathbb{P}(|X_j| > \epsilon s_n) + o(1) \\
 &\leq \frac{4}{t^2} \sum_{j=1}^n \left(\frac{\sigma_j^2}{\epsilon^2 s_n^2} \right) + o(1) \text{ by Chebyshev's Inequality} \\
 &\leq \frac{4}{t^2 \epsilon^2} + o(1).
 \end{aligned}$$

Since t can be chosen arbitrarily large, Lindeberg condition holds.

4 Lyapunov's Condition.

4.1 Lyapunov's Condition.

Suppose X_1, X_2, \dots, X_n are independent r.v.s with mean 0 and variance σ_n^2 . Let $s_n^2 = \sum_{i=1}^n \sigma_i^2$ denote the variance of partial sum $S_n = X_1 + X_2 + \dots + X_n$ and $\mathbb{E}(|X_j|^{2+\delta}) < \infty \forall j$. Then,

$$\frac{1}{s_n^{2+\delta}} \sum_{j=1}^n \mathbb{E}(|X_j|^{2+\delta}) \longrightarrow 0 \text{ as } n \rightarrow \infty$$

for some $\delta > 0$.

4.2 Lyapunov's Condition \implies Lindeberg's Condition.

Note that,

$$\begin{aligned}
 \frac{1}{s_n^2} \sum_{j=1}^n \mathbb{E}(X_j^2 \mathbf{1}_{\{|X_j| > \epsilon s_n\}}) &= \frac{1}{s_n^2} \sum_{j=1}^n \mathbb{E}\left(X_j^2 \mathbf{1}_{\left\{\left|\frac{X_j}{\epsilon s_n}\right| > 1\right\}}\right) \\
 &= \frac{1}{s_n^2} \sum_{j=1}^n \mathbb{E}\left(X_j^2 \mathbf{1}_{\left\{\frac{|X_j|^\delta}{\epsilon^\delta s_n^\delta} > 1\right\}}\right) \\
 &\leq \frac{1}{s_n^2} \sum_{j=1}^n \mathbb{E}\left(X_j^2 \frac{|X_j|^\delta}{\epsilon^\delta s_n^\delta}\right) \leq \frac{1}{\epsilon^\delta s_n^\delta} \sum_{j=1}^n \mathbb{E}(|X_j|^{2+\delta}) \longrightarrow 0
 \end{aligned}$$

if Lyapunov's Condition holds.

5 Multivariate Version.

5.1 Multivariate Lindeberg-Feller CLT.

Let $\mathbf{X}_1, \mathbf{X}_2, \dots, \mathbf{X}_n$ be independent random vectors with $\mathbb{E}(\mathbf{X}_j) = \mathbf{0}$ and $\text{Var}(\mathbf{X}_j) = \Sigma_j$. Suppose that, $\frac{1}{n}(\Sigma_1 + \Sigma_2 + \dots + \Sigma_n) \rightarrow \Sigma$ as $n \rightarrow \infty$ and,

$$\forall \epsilon > 0 \quad \frac{1}{n} \sum_{j=1}^n \mathbb{E} \left(\|\mathbf{X}_j\|^2 \mathbf{1}_{\{\|\mathbf{X}_j\| > \epsilon \sqrt{n}\}} \right) \rightarrow 0 \text{ as } n \rightarrow \infty$$

Then,

$$\frac{\mathbf{S}_n}{\sqrt{n}} \xrightarrow{\mathcal{D}} \mathcal{N}_d(\mathbf{0}, \Sigma) \text{ as } n \rightarrow \infty$$

where $\mathbf{S}_n = \mathbf{X}_1 + \mathbf{X}_2 + \dots + \mathbf{X}_n$.

5.2 Proof of Multivariate Lindeberg-Feller CLT.

We have $\mathbf{S}_n = \sum_{j=1}^n \mathbf{X}_j$, where $\mathbf{X}_j \sim (\mathbf{0}, \Sigma_j)$ independently. We shall use *Cramer-Wold Device* to establish the theorem. We have to show, $\forall \mathbf{a} \in \mathbb{R}^d \setminus \{\mathbf{0}\}$, $(0, \mathbf{a}'\Sigma\mathbf{a})$.

$$\frac{\mathbf{a}'\mathbf{S}_n}{\sqrt{n}} \xrightarrow{\mathcal{D}} \mathcal{N}(0, \mathbf{a}'\Sigma\mathbf{a})$$

Define, $Y_j = \mathbf{a}'\mathbf{X}_j \forall j = 1(1)n$. Then, $Y_j \sim (0, \sigma_j^2)$ where $\sigma_j^2 = \mathbf{a}'\Sigma_j\mathbf{a}$, $\forall j = 1(1)n$. Say,

$$\tilde{S}_n = \sum_{j=1}^n Y_j, \quad s_n^2 = \sum_{j=1}^n \sigma_j^2 = \mathbf{a}' \left(\sum_{j=1}^n \Sigma_j \right) \mathbf{a}$$

For any $\epsilon > 0$

$$\begin{aligned} & \frac{1}{s_n^2} \sum_{j=1}^n \mathbb{E} (Y_j^2 \mathbf{1}_{\{|Y_j| > \epsilon s_n\}}) \\ & \frac{1}{s_n^2} \sum_{j=1}^n \mathbb{E} ((\mathbf{a}'\mathbf{X}_j\mathbf{X}_j'\mathbf{a}) \mathbf{1}_{\{(\mathbf{a}'\mathbf{X}_j\mathbf{X}_j'\mathbf{a}) > \epsilon^2 s_n^2\}}) \\ & \leq \frac{\|\mathbf{a}\|^2}{s_n^2} \sum_{j=1}^n \mathbb{E} \left(\|\mathbf{X}_j\|^2 \mathbf{1}_{\{\|\mathbf{X}_j\|^2 \geq \epsilon'^2 s_n^2\}} \right) \text{ where } \epsilon'^2 = \frac{\epsilon^2}{\|\mathbf{a}\|^2} \\ & = \frac{1}{n} \sum_{j=1}^n \mathbb{E} \left(\|\mathbf{X}_j\|^2 \mathbf{1}_{\{\|\mathbf{X}_j\|^2 \geq \epsilon'^2 s_n^2\}} \right) \frac{\|\mathbf{a}\|^2}{s_n^2/n} \end{aligned}$$

Now,

$$\frac{s_n}{\sqrt{n}} \rightarrow \sqrt{\mathbf{a}'\Sigma\mathbf{a}} = c \text{ (say)}$$

Then, $\exists K \in \mathbb{N} \ni \forall n \geq K, \frac{s_n}{\sqrt{n}} > \frac{c}{2}$

$$\implies \forall n \geq K, \frac{1}{n} \sum_{j=1}^n \mathbb{E} \left(\|\mathbf{X}_j\|^2 \mathbf{1}_{\{\|\mathbf{X}_j\|^2 \geq \epsilon'^2 s_n^2\}} \right) \leq \frac{1}{n} \sum_{j=1}^n \mathbb{E} \left(\|\mathbf{X}_j\|^2 \mathbf{1}_{\{\|\mathbf{X}_j\|^2 \geq \epsilon''^2 n\}} \right) \text{ where } \epsilon'' = \epsilon' c/2$$

Taking lim sup as $n \rightarrow \infty$ on both side, we get,

$$\frac{1}{n} \sum_{j=1}^n \mathbb{E} \left(\|\mathbf{X}_j\|^2 \mathbf{1}_{\{\|\mathbf{X}_j\|^2 \geq \epsilon'^2 s_n^2\}} \right) \longrightarrow 0 \text{ as } n \rightarrow \infty$$

from the given conditions.

$$\therefore \sum_{j=1}^n \frac{Y_j}{s_n} \xrightarrow{\mathcal{D}} \mathcal{N}(0, 1)$$

also,

$$\frac{s_n}{\sqrt{n}} \longrightarrow \sqrt{\mathbf{a}' \Sigma \mathbf{a}}$$

Combining, by *Slutsky's Theorem*, we get,

$$\sum_{j=1}^n \frac{Y_j}{\sqrt{n}} \xrightarrow{\mathcal{D}} \mathcal{N}(0, \mathbf{a}' \Sigma \mathbf{a})$$

i.e.

$$\frac{\mathbf{a}' \mathbf{S}_n}{\sqrt{n}} \xrightarrow{\mathcal{D}} \mathcal{N}(0, \mathbf{a}' \Sigma \mathbf{a})$$