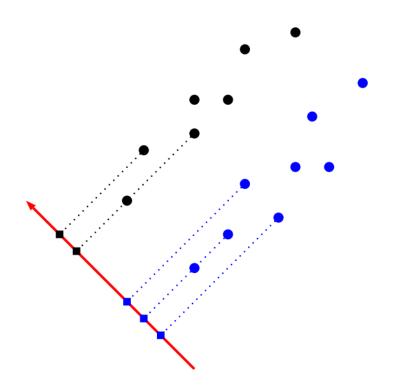
Lecture 5 Linear Discriminant Analysis

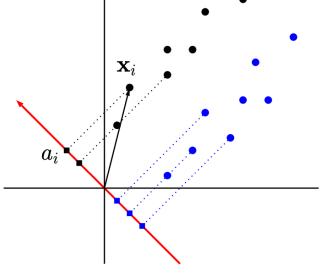
Given a training data set $x_1, ..., x_n \in \mathbb{R}^d$ consisting of two classes, find a (unit-vector) direction that "best" discriminates between the two classes.



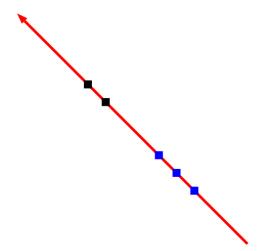
LDA is a supervised dimensionality reduction method that tries to preserve the discriminatory information between classes.

- First, observe that projections of the two classes onto parallel lines always have "the same amount of separation".
- This time we are going to focus on lines that pass through the origin.
- Consider any unit vector $\mathbf{w} \in \mathbb{R}^d$, the 1D projections of the points are

$$a_i = \mathbf{w}^T \mathbf{x}_i, i = 1, ..., n$$



Now the data look like this:



How do we quantify the separation between the two classes (in order to compare different directions v and select the best one)?

One (naive) idea is to measure the distance between the two class means in the 1D projection space: $|\mu_1 - \mu_2|$, where

$$egin{aligned} \mu_1 &= rac{1}{n_1} \sum_{\mathbf{x}_i \in X_1} a_i = rac{1}{n_1} \sum_{\mathbf{x}_i \in X_1} \mathbf{w}^T \mathbf{x}_i \ &= \mathbf{w}^T \cdot rac{1}{n_1} \sum_{\mathbf{x}_i \in X_1} \mathbf{x}_i = \mathbf{w}^T \mathbf{m}_1 \end{aligned}$$

and similarly,

$$\mu_2 = \mathbf{w}^T \mathbf{m}_2, \quad \mathbf{m}_2 = rac{1}{n_2} \sum_{\mathbf{x}_i \in X_2} \mathbf{x}_i$$

That is, we solve the following problem

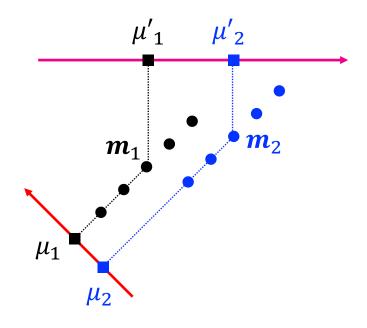
$$\max_{\mathbf{w}:\|\mathbf{w}\|=1} \lvert \mu_1 - \mu_2
vert$$

where

$$\mu_j = \mathbf{w}^T \mathbf{m}_j, \quad j=1,2.$$

However, this criterion does not always work (as shown in the right plot).

What else do we need to control?



It turns out that we should also pay attention to the variances of the projected classes:

$$s_1^2 = \sum_{\mathbf{x}_i \in X_1} (a_i - \mu_1)^2, \quad s_2^2 = \sum_{\mathbf{x}_i \in X_2} (a_i - \mu_2)^2$$

Ideally, the projected classes have both faraway means and small variances.

This can be achieved through the following modified formulation:

$$\max_{\mathbf{w}:\|\mathbf{w}\|=1}rac{(\mu_1-\mu_2)^2}{s_1^2+s_2^2}$$

The optimal **w** should be such that

- $(u_1 u_2)^2$: large
- s_1^2, s_2^2 : both small



First, we derive a formula for the distance between the two projected centroids:

$$egin{aligned} (\mu_1 - \mu_2)^2 &= \left(\mathbf{w}^T \mathbf{m}_1 - \mathbf{w}^T \mathbf{m}_2
ight)^2 = \left(\mathbf{w}^T (\mathbf{m}_1 - \mathbf{m}_2) + \mathbf{w}^T \mathbf{m}_1 - \mathbf{m}_2
ight)^2 \\ &= \mathbf{w}^T (\mathbf{m}_1 - \mathbf{m}_2) \cdot (\mathbf{m}_1 - \mathbf{m}_2)^T \mathbf{w} \\ &= \mathbf{w}^T \mathbf{S}_b \mathbf{w} \end{aligned}$$

where

$$\mathbf{S}_b = (\mathbf{m}_1 - \mathbf{m}_2)(\mathbf{m}_1 - \mathbf{m}_2)^T \in \mathbb{R}^{d imes d}$$

is called the **between-class scatter matrix**.

Remark. Clearly, S_b is square, symmetric and positive semidefinite. Moreover, rank(S_b) = 1, which implies that it only has 1 positive eigenvalue!

Next, for each class j = 1, 2, the variance of the projection (onto w) is

$$egin{aligned} s_j^2 &= \sum_{\mathbf{x}_i \in X_j} (a_i - \mu_j)^2 = \sum_{\mathbf{x}_i \in X_j} \left(\mathbf{w}^T \mathbf{x}_i - \mathbf{w}^T \mathbf{m}_j
ight)^2 \ &= \sum_{\mathbf{x}_i \in X_j} \mathbf{w}^T (\mathbf{x}_i - \mathbf{m}_j) (\mathbf{x}_i - \mathbf{m}_j)^T \mathbf{w} \ &= \mathbf{w}^T \Bigg[\sum_{\mathbf{x}_i \in X_j} (\mathbf{x}_i - \mathbf{m}_j) (\mathbf{x}_i - \mathbf{m}_j)^T \Bigg] \mathbf{w} \ &= \mathbf{w}^T \mathbf{S}_j \mathbf{w} \end{aligned}$$

where

$$\mathbf{S}_j = \sum_{\mathbf{x}_i \in X_j} (\mathbf{x}_i - \mathbf{m}_j) (\mathbf{x}_i - \mathbf{m}_j)^T \in \mathbb{R}^{d imes d}$$

is called the **within-class scatter matrix** for class *j*.



The total within-class scatter of the two classes in the projection space is

$$s_1^2 + s_2^2 = \mathbf{w}^T \mathbf{S}_1 \mathbf{w} + \mathbf{w}^T \mathbf{S}_2 \mathbf{w} = \mathbf{w}^T (\mathbf{S}_1 + \mathbf{S}_2) \mathbf{w} = \mathbf{w}^T \mathbf{S}_w \mathbf{w}$$

where

$$\mathbf{S}_w = \mathbf{S}_1 + \mathbf{S}_2 = \sum_{\mathbf{x}_i \in X_1} (\mathbf{x}_i - \mathbf{m}_1) (\mathbf{x}_i - \mathbf{m}_1)^T + \sum_{\mathbf{x}_i \in X_2} (\mathbf{x}_i - \mathbf{m}_2) (\mathbf{x}_i - \mathbf{m}_2)^T$$

is called the **total within-class scatter matrix** of the original data.

Remark. $S_w \in \mathbb{R}^{d \times d}$ is also square, symmetric, and positive semidefinite.

Putting everything together, we have derived the following optimization problem:

$$\max_{\mathbf{w}: \|\mathbf{w}\| = 1} \frac{\mathbf{w}^T \mathbf{S}_b \mathbf{w}}{\mathbf{w}^T \mathbf{S}_w \mathbf{w}}$$
 Generalized Rayleigh quotient (广义瑞利商)

Theorem. Suppose S_w is nonsingular. The maximizer of the problem is given by the largest eigenvector w_1 of $S_w^{-1}S_b$, i.e.,

$$\mathbf{S}_w^{-1}\mathbf{S}_b\mathbf{w}_1 = \lambda_1\mathbf{w}_1 \quad \longleftarrow \quad \mathbf{S}_b\mathbf{w} = \lambda\mathbf{S}_w\mathbf{w}$$

Remark. rank($S_w^{-1}S_b$) = rank(S_b) = 1, so λ_1 is the only nonzero (positive) eigenvalue that can be found. It represents the largest amount of separation between the two classes along any single direction.

■ Generalized Rayleigh quotient (广义瑞利商)

$$J = \frac{\boldsymbol{w}^{\mathrm{T}} \mathbf{S}_{b} \boldsymbol{w}}{\boldsymbol{w}^{\mathrm{T}} \mathbf{S}_{w} \boldsymbol{w}}$$

■ Let $\mathbf{w}^T \mathbf{S}_w \mathbf{w} = 1$, maximizing generalized Rayleigh quotient is equivalent to

$$\min_{\boldsymbol{w}} - \boldsymbol{w}^{\mathrm{T}} \mathbf{S}_b \boldsymbol{w}$$

s.t. $\boldsymbol{w}^{\mathrm{T}} \mathbf{S}_w \boldsymbol{w} = 1$

Using the method of Lagrangian multipliers

$$\mathbf{S}_b \boldsymbol{w} = \lambda \mathbf{S}_w \boldsymbol{w}$$



Instead of solving a generalized eigenvalue problem, a smartest way is to rewrite as

$$egin{aligned} \lambda \mathbf{w} &= \mathbf{S}_w^{-1} \underbrace{(\mathbf{m}_1 - \mathbf{m}_2)(\mathbf{m}_1 - \mathbf{m}_2)^T}_{\mathbf{S}_b} \mathbf{w} \ &= \mathbf{S}_w^{-1} (\mathbf{m}_1 - \mathbf{m}_2) \cdot \underbrace{(\mathbf{m}_1 - \mathbf{m}_2)^T}_{\mathrm{scalar}} \mathbf{w} \end{aligned}$$

This implies that

$$\mathbf{w} \propto \mathbf{S}_w^{-1} (\mathbf{m}_1 - \mathbf{m}_2)$$

and it can be computed from $\mathbf{w} \propto \mathbf{S}_w^{-1}(\mathbf{m}_1 - \mathbf{m}_2)$ through rescaling!

Remark. Here, inverting S_w can be done through singular value decomposition $S_w = \mathbf{U} \mathbf{\Sigma} \mathbf{V}^T$



Singular Value Decomposition

Singular Value Decomposition (奇异值分解,SVD)

Any real matrix $\mathbf{A} \in \mathbb{R}^{m \times n}$ can be decomposed as

$$\mathbf{A} = \mathbf{U} \mathbf{\Sigma} \mathbf{V}^{\mathrm{T}}$$

 $\mathbf{U} \in \mathbb{R}^{m \times m}$ is a unitary matrix of order m satisfying $\mathbf{U}^{\mathrm{T}}\mathbf{U} = \mathbf{I}$,

 $\mathbf{V} \in \mathbb{R}^{n \times n}$ is a unitary matrix of order n satisfying $\mathbf{V}^T \mathbf{V} = \mathbf{I}$,

 $\Sigma \in \mathbb{R}^{m \times n}$ is a $m \times n$ matrix with all of its elements take the value 0 except $(\Sigma)_{ii} = \sigma_i$, where σ_i are non-negative real numbers and $\sigma_1 \geqslant \sigma_2 \geqslant \ldots \geqslant 0$.

Singular Value Decomposition

Applications of SVD, e.g., low-rank matrix approximation

to approximate a r-rank matrix **A** with a k-rank matrix **A**, where $k \leq r$,

$$\min_{\widetilde{\mathbf{A}} \in \mathbb{R}^{m \times n}} \|\mathbf{A} - \widetilde{\mathbf{A}}\|_{F}$$
s.t. $\operatorname{rank}(\widetilde{\mathbf{A}}) = k$.

SVD can provide an analytical solution:

- Performing SVD on A;
- obtain a matrix Σ_k by setting the r-k smallest singular values in Σ to zero, (i.e., keep only the k largest singular values)
- the optimal solution: $\mathbf{A}_k = \mathbf{U}_k \mathbf{\Sigma}_k \mathbf{V}_k^{\mathrm{T}}$, \mathbf{U}_k and \mathbf{V}_k are, respectively, the first k columns of \mathbf{U} and \mathbf{V}

Linear Discriminant Analysis

- **The idea of LDA:** project the same class samples onto a line, while samples of different classes are far away from each other.
 - To make the projection of similar samples as close as possible, we can make the covariance of the projection points of similar samples as small as possible
 - To make the projection of samples from different classes as far away as possible, we can make the distance between the class centers as large as possible

The global scatter matrix:

$$\mathbf{S}_w = \sum_{j=1}^c \sum_{\mathbf{x} \in X_j} (\mathbf{x} - \mathbf{m}_j) (\mathbf{x} - \mathbf{m}_j)^T, \quad \mathbf{S}_b = \sum_{j=1}^c n_j (\mathbf{m}_j - \mathbf{m}) (\mathbf{m}_j - \mathbf{m})^T.$$

We have arrived at the same kind of problem:

$$\max_{\mathbf{w}:\|\mathbf{w}\|=1} rac{\mathbf{w}^T \mathbf{S}_b \mathbf{w}}{\mathbf{v}^T \mathbf{S}_w \mathbf{w}}$$

The solution is given by the largest eigenvector of $S_w^{-1}S_b$ (S_b is nonsingular):

$$\mathbf{S}_w^{-1}\mathbf{S}_b\mathbf{w}_1=\lambda_1\mathbf{w}_1$$

However, the formula $\mathbf{w} \propto \mathbf{S}_w^{-1}(\mathbf{m}_1 - \mathbf{m}_2)$ is no longer valid:

$$\lambda_1 \mathbf{w}_1 = \mathbf{S}_w^{-1} \mathbf{S}_b \mathbf{w}_1 = \mathbf{S}_w^{-1} \sum_j n_j (\mathbf{m}_j - \mathbf{m}) \underbrace{(\mathbf{m}_j - \mathbf{m})^T \mathbf{w}_1}_{ ext{scalar}}$$

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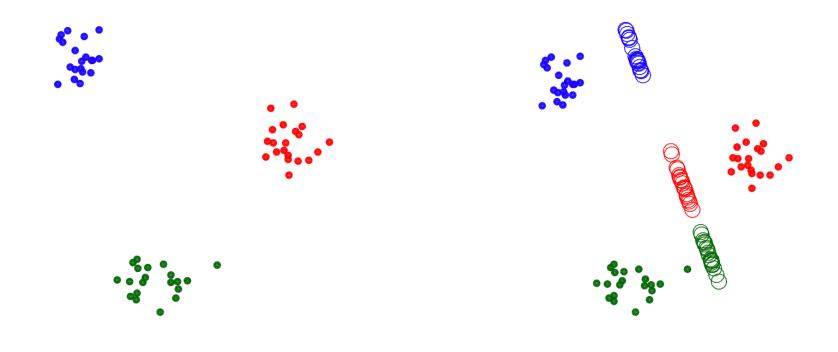
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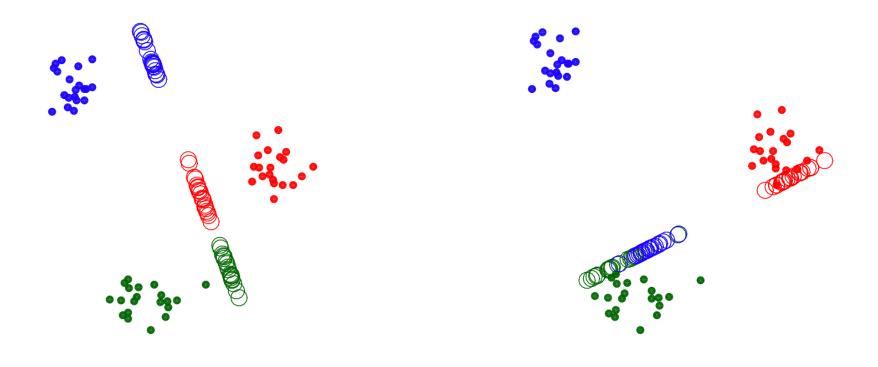
$$\lambda_1 \mathbf{w}_1 = \mathbf{S}_w^{-1} \mathbf{S}_b \mathbf{w}_1 = \mathbf{S}_w^{-1} \sum_j n_j (\mathbf{m}_j - \mathbf{m}) \underbrace{(\mathbf{m}_j - \mathbf{m})^T \mathbf{w}_1}_{ ext{scalar}}$$

So we have to find w_1 by solving a generalized eigenvalue problem:

$$\mathbf{S}_b \boldsymbol{w} = \lambda \mathbf{S}_w \boldsymbol{w}$$



What about the second eigenvector w_2 ?



How many discriminatory directions can we find?

To answer this question, we just need to count the number of nonzero eigenvalues

$$\mathbf{S}_w^{-1}\mathbf{S}_b\mathbf{w} = \lambda\mathbf{w}$$

since only the nonzero eigenvectors will be used as the discriminatory directions.

In the above equation, the within-class scatter matrix S_w is assumed to be nonsingular. However, the between-class scatter matrix S_b is of low rank:

$$egin{aligned} \mathbf{S}_b &= \sum n_i (\mathbf{m}_i - \mathbf{m}) (\mathbf{m}_i - \mathbf{m})^T \ &= \left[\sqrt{n_1} (\mathbf{m}_1 - \mathbf{m}) \cdots \sqrt{n_c} (\mathbf{m}_c - \mathbf{m})
ight] \cdot egin{bmatrix} \sqrt{n_i} (\mathbf{m}_1 - \mathbf{m})^T \ dots \ \sqrt{n_c} (\mathbf{m}_c - \mathbf{m})^T \end{bmatrix} \end{aligned}$$

Observe that the columns of the matrix

$$[\sqrt{n_1}(\mathbf{m}_1 - \mathbf{m}) \cdots \sqrt{n_c}(\mathbf{m}_c - \mathbf{m})]$$

are linearly dependent:

$$\sqrt{n_1} \cdot \sqrt{n_1} (\mathbf{m}_1 - \mathbf{m}) + \dots + \sqrt{n_c} \cdot \sqrt{n_c} (\mathbf{m}_c - \mathbf{m})$$

$$= (n_1 \mathbf{m}_1 + \dots + n_c \mathbf{m}_c) - (n_1 + \dots + n_c) \mathbf{m}$$

$$= n\mathbf{m} - n\mathbf{m}$$

$$= \mathbf{0}.$$

The shows that $rank(S_b) \le c - 1$ (where c is the number of training classes).

Therefore, one can only find at most c-1 discriminatory directions.

Objective

$$\max_{\mathbf{W}} \frac{\operatorname{tr}\left(\mathbf{W}^{\mathrm{T}}\mathbf{S}_{b}\mathbf{W}\right)}{\operatorname{tr}\left(\mathbf{W}^{\mathrm{T}}\mathbf{S}_{w}\mathbf{W}\right)}$$

where $\mathbf{W} \in \mathbb{R}^{d \times (N-1)}$

$$\mathbf{S}_b\mathbf{W} = \lambda\mathbf{S}_w\mathbf{W}$$

$$egin{aligned} \operatorname{tr}(\mathbf{A}) &= \sum_{i=1}^n A_{ii} \ & rac{\partial}{\partial \mathbf{X}} \operatorname{tr}\left(\mathbf{X}^{\mathrm{T}} \mathbf{B} \mathbf{X}
ight) = (\mathbf{B} + \mathbf{B}^{\mathrm{T}}) \mathbf{X} \end{aligned}$$

Concatenating the eigenvectors corresponding to the d' largest non-zero eigenvalues of $\mathbf{S}_w^{-1}\mathbf{S}_b$ leads to the closed-form solution of \mathbf{W} , where $d' \leq c - 1$

LDA for classification

- Explain from the aspect of Bayesian decision theory
 - It can be proved that we have the optimal solution of LDA when both classes follow Gaussian distribution with the same prior and covariance

Bayes Theorem for Classification

Thomas Bayes was a famous mathematician whose name represents a big subfield of statistical and probabilistic modeling. Here we focus on a simple result, known as Bayes theorem:

$$\Pr(Y = k | X = x) = \frac{\Pr(X = x | Y = k) \cdot \Pr(Y = k)}{\Pr(X = x)}$$

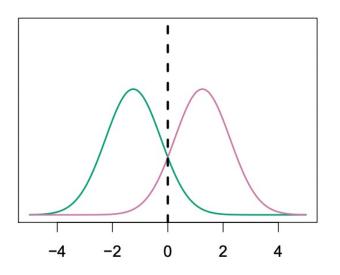
One writes this slightly differently for discriminant analysis:

$$\Pr(Y = k | X = x) = \frac{\pi_k f_k(x)}{\sum_{l=1}^{K} \pi_l f_l(x)},$$
 where

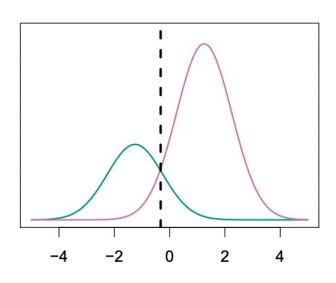
- $f_k(x) = \Pr(X = x \mid Y = k)$ is the density for X in class k. Here we will use normal densities for these, separately in each class.
- $\pi_k = \Pr(Y = k)$ is the marginal or prior probability for class k.

Classify to the highest density

$$\pi_1$$
=.5, π_2 =.5



$$\pi_1$$
=.3, π_2 =.7



- We classify a new point according to which density is highest
- When the priors are different, we take them into account as well, and compare $\pi_k f_k(x)$. On the right, we favor the pink class the decision boundary has shifted to the left.

Linear Discriminant Analysis (p = 1)

The Gaussian density has the form

$$f_k(x) = \frac{1}{\sqrt{2\pi}\sigma_k} e^{-\frac{1}{2}\left(\frac{x-\mu_k}{\sigma_k}\right)^2}$$

Here μ_k is the mean, and σ_k^2 the variance (in class k). We will assume that all the $\sigma_k = \sigma$ are the same.

Plugging this into Bayes formula, we get a rather complex expression for $p_k(x) = Pr(Y = k \mid X = x)$:

$$p_k(x) = \frac{\pi_k \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{1}{2} \left(\frac{x-\mu_k}{\sigma}\right)^2}}{\sum_{l=1}^K \pi_l \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{1}{2} \left(\frac{x-\mu_l}{\sigma}\right)^2}}$$

Happily, there are simplifications and cancellations.



Discriminant functions

To classify at the value X = x, we need to see which of the $p_k(x)$ is largest. Taking logs, and discarding terms that do not depend on k, we see that this is equivalent to assigning x to the class with the largest *discriminant score*:

$$\delta_k(x) = x \cdot \frac{\mu_k}{\sigma^2} - \frac{\mu_k^2}{2\sigma^2} + \log(\pi_k)$$

Note that $\delta_k(x)$ is a *linear* function of x.

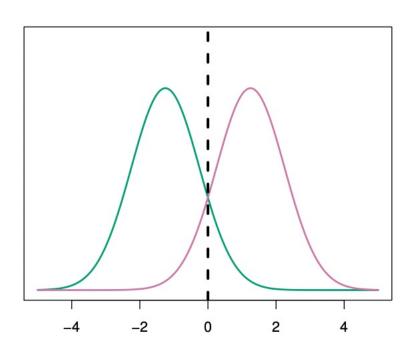
If there are K=2 classes and $\pi_1=\pi_2=0.5$, then one can see that the decision boundary is at

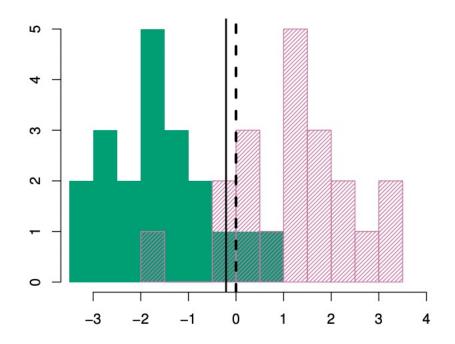
$$x = \frac{\mu_1 + \mu_2}{2}.$$

(See if you can show this)



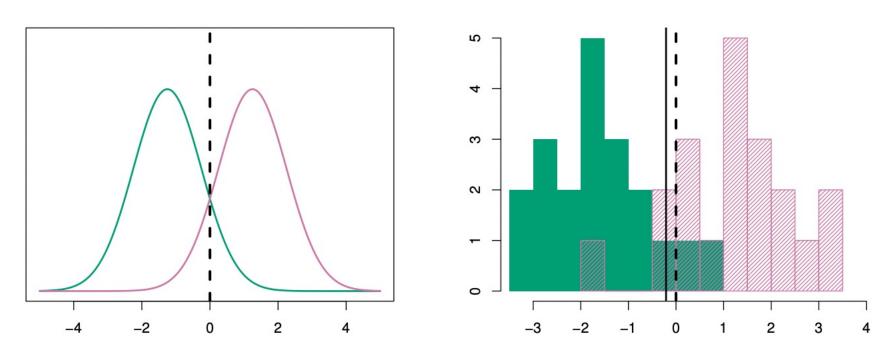
Discriminant functions





Example with $\mu_1 = -1.5$, $\mu_2 = 1.5$, $\pi_1 = \pi_2 = 0.5$, and $\sigma^2 = 1$

Discriminant functions



- Example with $\mu_1 = -1.5$, $\mu_2 = 1.5$, $\pi_1 = \pi_2 = 0.5$, and $\sigma^2 = 1$
- Typically we don't know these parameters; we just have the training data. In that case we simply estimate the parameters and plug them into the rule.

Estimating the parameters

$$\hat{\pi}_k = \frac{n_k}{n}$$

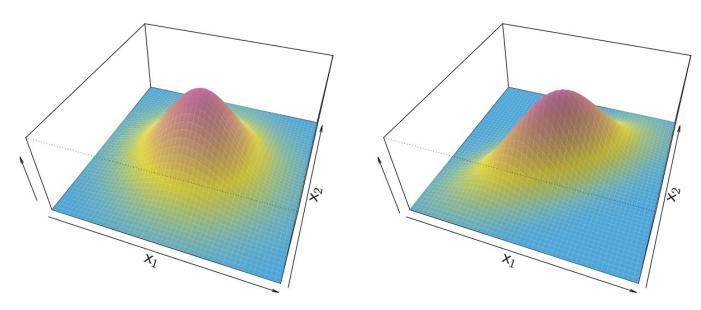
$$\hat{\mu}_k = \frac{1}{n_k} \sum_{i: y_i = k} x_i$$

$$\hat{\sigma}^2 = \frac{1}{n - K} \sum_{k=1}^K \sum_{i: y_i = k} (x_i - \hat{\mu}_k)^2$$

$$= \sum_{k=1}^K \frac{n_k - 1}{n - K} \cdot \hat{\sigma}_k^2$$

where $\hat{\sigma}_k^2 = \frac{1}{n_k - 1} \sum_{i:y_i = k} (x_i - \hat{\mu}_k)^2$ is the usual formula for the estimated variance in the k-th class.

Linear Discriminant Analysis (p > 1)



Density:
$$f(x) = \frac{1}{(2\pi)^{p/2} |\mathbf{\Sigma}|^{1/2}} e^{-\frac{1}{2}(x-\mu)^T \mathbf{\Sigma}^{-1}(x-\mu)}$$

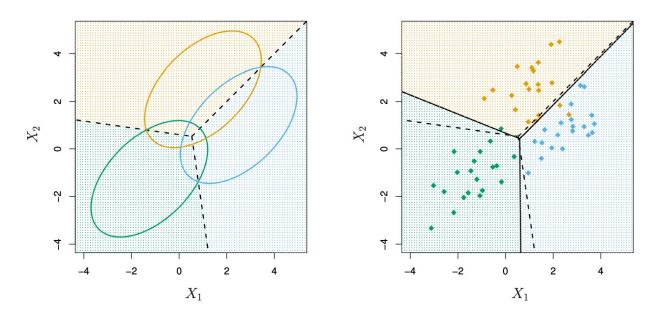
Discriminant function:
$$\delta_k(x) = x^T \mathbf{\Sigma}^{-1} \mu_k - \frac{1}{2} \mu_k^T \mathbf{\Sigma}^{-1} \mu_k + \log \pi_k$$

Despite its complex form,

$$\delta_k(x) = c_{k0} + c_{k1}x_1 + c_{k2}x_2 + \ldots + c_{kp}x_p$$
 — a linear function.



Illustration: p = 2 and K = 3 classes



Here
$$\pi_1 = \pi_2 = \pi_3 = 1/3$$
.

The dashed lines are known as the *Bayes decision boundaries*. Were they known, they would yield the fewest misclassification errors, among all possible classifiers.

Logistic Regression versus LDA

For a two-class problem, one can show that for LDA

$$\log\left(\frac{p_1(x)}{1 - p_1(x)}\right) = \log\left(\frac{p_1(x)}{p_2(x)}\right) = c_0 + c_1x_1 + \ldots + c_px_p$$

So it has the same form as logistic regression.

The difference is in how the parameters are estimated.

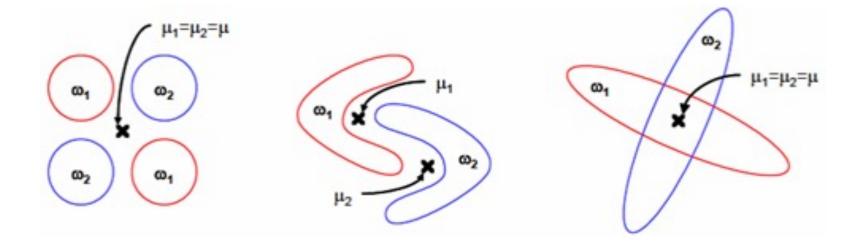
- Logistic regression uses the conditional likelihood based on Pr(Y|X) (known as *discriminative learning*).
- LDA uses the full likelihood based on Pr(*X*, *Y*) (known as *generative learning*).
- Despite these differences, in practice the results are often very similar.



LDA for classification

- When the classes are well-separated, the parameter estimates for the logistic regression model are surprisingly unstable. Linear discriminant analysis does not suffer from this problem.
- If n is small and the distribution of the predictors X is approximately normal in each of the classes, the linear discriminant model is again more stable than the logistic regression.
- Linear discriminant analysis is popular when we have more than two classes, because it also provides lowdimensional views of the data.

LDA cannot handle non-Gaussian data



Multi-class Classification

Multiclass Classification

- Multiclass Classification learning methods
 - Some binary classification methods can be directly extended to accommodate multiclass cases
 - Apply some strategies to solve multiclass classification problems with any existing binary classification methods (more general)
 - Decompose the problem and then train a binary classifier for each divided binary classification problem
 - Ensemble the outputs collected from all binary classifiers into the final multiclass predictions
- Dividing strategies
 - One vs. One (OvO)
 - One vs. Rest (OvR)
 - Many vs. Many (MvM)



Multiclass Classification - OvO

- In the decomposing phase
 - puts the N classes into pairs
 - N(N-1)/2 binary classification tasks
 - trains a classifier for each task
 - N(N-1)/2 classifiers
- In the testing phase
 - a new sample is classified by all classifiers
 - N(N-1)/2 classification outputs
 - the final prediction can be made via voting
 - the predicted class is the one received the most votes

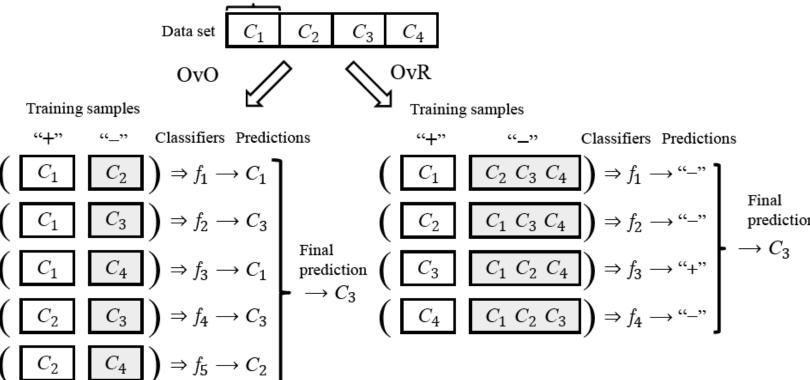
Multiclass Classification - OvR

- In the decomposing phase
 - consider each class as positive in turn, and the rest classes are considered as negative
 - N binary classification tasks
 - trains a classifier for each task
 - N classifiers
- In the testing phase
 - a new sample is classified by all classifiers
 - N classification outputs
 - the prediction confidences are usually assessed
 - the class with the highest confidence is used as the classification result



Multiclass Classification – A comparison between OvO and OvR

Samples of class C1



Multiclass Classification – A comparison between OvO and OvR

OvO

- Train N(N-1)/2 classifiers, the memory and testing time costs are often higher
- Each classifier uses only samples of two classes. Hence, the computational cost of training OvO is lower

OvR

- Train N classifiers, the memory and testing time costs are often lower
- Each classifier uses all training samples. Hence, the computational cost of training OvO is higher

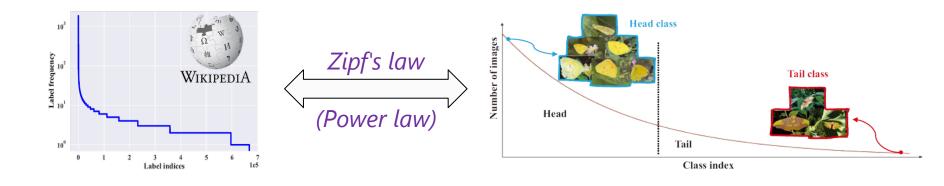
As for the prediction performance, it depends on the specific data distribution, and in most cases, the two methods have similar performance.

Recent Progress

Deep Learning

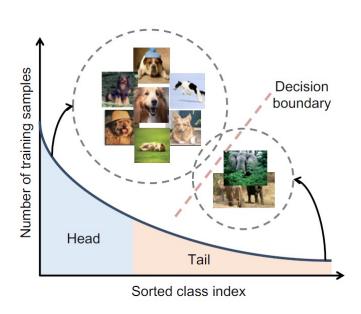
 enablers
 Algorithmic improvements
 Access to large amounts of data

Real-world class-imbalanced (long-tailed) class distribution

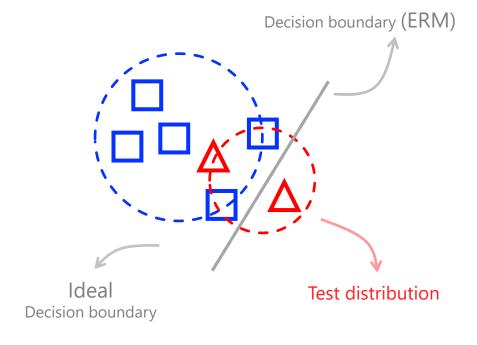


Challenge

—— model biased towards head classes



Take an example



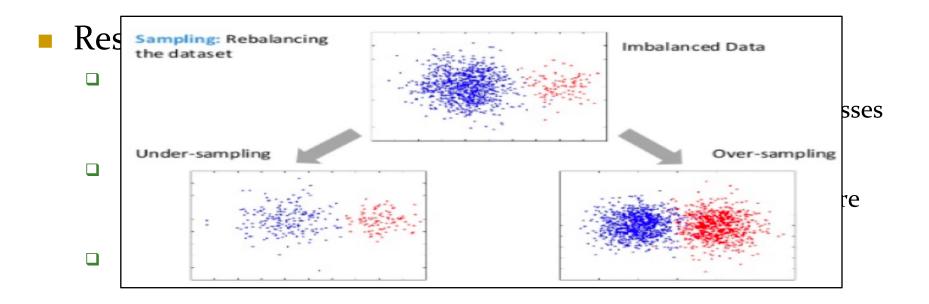
Class imbalance

a significantly different number of samples for each class. (the positive class is the minority)

$$\frac{y}{1-y} > 1$$



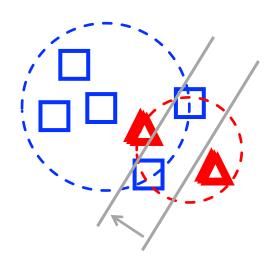
the classes are balanced
$$\frac{y}{1-y}>1$$
 $\frac{y}{1-y}>\frac{m^+}{m^-}$ the observed class ratio

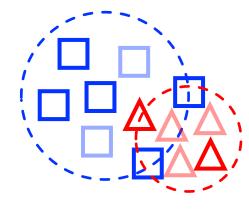


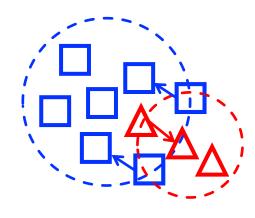
Class Re-balancing

Data Augmentation

Class-balanced Loss Function

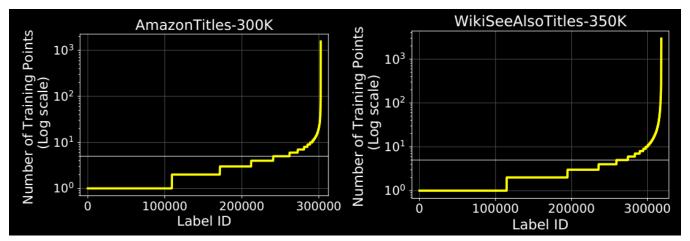




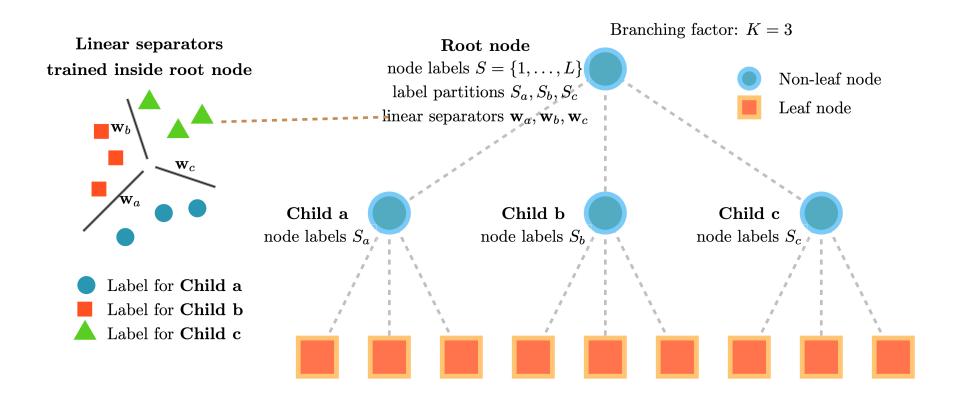


Extreme Classification

	Dataset	# of Train Points	# of Labels	# of Test Points
Benchmark*	LF-AmazonTitles-131K	294,805	131,073	134,835
	LF-WikiSeeAlsoTitles-320K	693,082	312,330	177,515
	LF-AmazonTitles-1.3M	2,248,619	1,305,265	970,237
Bing	LF-P2PTitles-300K	1,366,429	300,000	585,602
	LF-P2PTitles-2M	2,539,009	1,640,898	1,088,146



Label Tree



- recursively partition the set of classes into subsets
- train a multi-class (linear) classifier for inference