

# Statistics 305/605: Introduction to Biostatistical Methods for Health Sciences

## Chapter 16: Multiple $2 \times 2$ Contingency Tables

Jinko Graham

# Associations and extraneous variables

- ▶ Want to study the association between two dichotomous variables: an exposure variable and a disease outcome.
- ▶ However, a third, extraneous variable with  $g$  different values (levels) is also present
  - ▶ Does ignoring it lead to incorrect conclusions?
- ▶ When associating the disease outcome and the exposure variable, need to account for the possibility that the extraneous variable is either
  1. an **effect modifier** or
  2. a **confounder**.

## Example: Aortic Stenosis (AS) Study

- ▶ What is the association between AS and smoking? How does sex affect this association?
  - ▶ AS: disease outcome
  - ▶ smoking: exposure variable
  - ▶ sex: extraneous variable with  $g = 2$  levels.
- ▶ First few lines of the data:

##	smoke	AS	sex
## 1	1	1	1
## 2	1	1	1
## 3	1	1	1
## 4	1	1	1
## 5	1	1	1
## 6	1	1	1

- ▶ When associating AS and smoking, we should account for the possibility that sex is either
  1. an **effect modifier** or
  2. a **confounder**.

## Effect Modification

# Effect Modification

- ▶ When the ORs for different strata of the extraneous variable are different, the extraneous variable is an **effect modifier**.
- ▶ Inference (estimates, CIs, tests) of these different ORs must then be based on the stratified tables.
- ▶ That is, we estimate the ORs, perform tests, and construct CIs with each stratum-specific table, separately.

# Testing for Effect Modification

- ▶ When the ORs for different strata of the extraneous variable are the same, they are said to be **homogeneous**.
- ▶ Can test for this with.
  - ▶  $H_0$ : homogeneity of ORs, versus
  - ▶  $H_a$ : effect modification.
- ▶ If we reject  $H_0$ , declare the extraneous variable to be an effect modifier.

## Test of Homogeneity (Text, Section 16.2.1)

- ▶ Suppose the extraneous variable has  $g$  levels or **strata**
- ▶ Let  $\widehat{OR}_i$  be the estimated OR between the exposure and the disease outcome in the  $i$ th stratum of the extraneous variable.
- ▶ Let  $y_i = \log_e \widehat{OR}_i$ , with variance estimated by

$$se^2(\log_e \widehat{OR}_i) = 1/a + 1/b + 1/c + 1/d.$$

- ▶ Let  $w_i = 1/se^2(\log_e \widehat{OR}_i)$ .
  - ▶ The larger the  $se(\log_e \widehat{OR}_i)$ , the smaller the weight  $w_i$  accorded to stratum  $i$  of the extraneous variable.
- ▶ A weighted average of the  $y_i$ 's that diminishes the contribution of values with high standard errors is:

$$Y = \sum_{i=1}^g \left( \frac{w_i}{\sum_{j=1}^g w_j} \right) y_i = \frac{\sum_{i=1}^g w_i y_i}{\sum_{j=1}^g w_j}.$$

- ▶ The test statistic tallies up the deviations of the estimated log-ORs from this average:

$$X^2 = \sum_{i=1}^g w_i (y_i - Y)^2$$

Large deviations suggest heterogeneous ORs.

- ▶ Under  $H_0$ : homogeneous ORs, the statistic  $X^2$  has a chi-square distribution with  $g - 1$  df.



# Test of Homogeneity in the AS Example

- In the AS example, the gender-specific tables are:

```
## , , sex = 0
##
##      smoke
## AS    0  1
##    0 47 19
##    1 29 14
##
## , , sex = 1
##
##      smoke
## AS    0  1
##    0 20 24
##    1 25 37
```

- ▶ We test for homogeneity of the ORs for AS across the strata defined by sex at level 5%.
- ▶ We get (see R demo):

```
## $X2  
## [1] 0.003072521  
##  
## $pval  
## [1] 0.9557956
```

- ▶ Retain the null hypothesis of homogeneous ORs at 5% level.  
Insufficient evidence that sex modifies the effect of smoking on AS.
  - ▶ See text, section 16.2.1 for another example.
- ▶ Next question (and topic): Sex doesn't appear to be an effect modifier but is it a confounder?

# Confounding

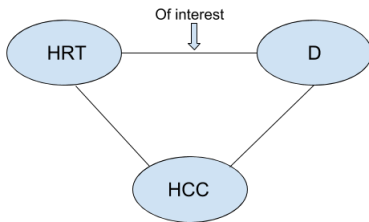
# What is a Confounder?

- ▶ A confounding variable is an extraneous variable that is associated with both the disease outcome and the exposure of interest.
- ▶ Why is this important?
- ▶ Ignoring a confounder leads to biased inference of the effect of the exposure on the disease outcome (i.e., incorrect estimate, test and CI) and potentially **wrong conclusions**.
- ▶ Confounding is thus a **big concern** in any observational study.

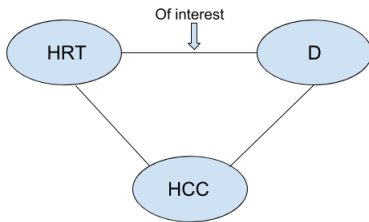
## WHI Example

- ▶ Observational studies leading up to the WHI (a clinical trial) suggested that postmenopausal women on hormone-replacement therapy (HRT) had better disease outcomes.
- ▶ Drug companies were actively promoting synthetic hormones in their marketing targetted to the large demographic group of postmenopausal women.
  - ▶ Doctors were prescribing HRT for their postmenopausal patients.
- ▶ What everyone missed was the potential confounding by health-care coverage (in the US).
- ▶ Women with better health-care coverage tend to have better disease outcomes and also to have better access to HRT. Hence HRT seemed to be associated with better disease outcomes.

- ▶ In the diagram below, let  $D$  be disease outcome,  $HRT$  be hormone replace therapy (the exposure of interest) and  $HCC$  be health-care coverage.
  - ▶ Associations between variables are indicated by lines
  - ▶  $HCC$  is associated with better disease outcomes (i.e. with  $D$ ).
  - ▶  $HCC$  is also associated with access to  $HRT$ .



Confounding of association between HRT and D by HCC



Confounding of association between *HRT* and *D* by *HCC*

- ▶ To properly understand the effect of *HRT* on *D*, it is important to control/adjust for the potential confounding effect of *HCC*.
- ▶ The WHI was a randomized, controlled, clinical trial of *HRT*.
  - ▶ The WHI investigators used treatment randomization to control/adjust for the potential confounding effects of *HCC*.
  - ▶ The randomization of *HRT* to women regardless of their *HCC* breaks the association between *HRT* and *HCC*.
- ▶ The trial established the harmful effects of *HRT*. Doctor's stopped prescribing and companies stopped marketing it.

# Simpson's Paradox

- ▶ Confounding explains “Simpson's paradox”:
  - ▶ When the effect of an exposure on the odds of the disease outcome differs in the stratified and pooled analyses.
- ▶ In the AS example, the estimated ORs summarizing the association between AS and smoking, stratified by sex, are about the same: 1.19 for females and 1.23 for males.
  - ▶ These results suggest that, for both females and males, smoking *slightly* increases the odds of AS.
- ▶ However, the estimated OR from the pooled table across genders is 1.47.
  - ▶ This result suggests a *stronger* effect of smoking than we obtain from the stratified tables!



# Effect Modifier vs. Confounder

- ▶ When considering the OR summarizing the association between the disease outcome and the exposure variable, the extraneous variable is:
  1. an *effect modifier* if the ORs for different strata of the extraneous variable are **different**.
  2. a *confounder* if the ORs for different strata of the extraneous variable are the **same** but different from the OR that we get when we ignore the extraneous variable altogether.

# The Mantel-Haenszel (MH) estimator

- ▶ To understand the effect of the exposure on disease outcome, we need to adjust the OR for the effects of a confounding variable.
- ▶ We can adjust the OR by using the Mantel-Haenszel (MH) estimator.
- ▶ The MH estimator of the OR adjusts for a confounding variable when
  - ▶ the disease outcome and exposure variable are dichotomous (i.e. have 2 categories) and
  - ▶ the confounding variable is categorical with 2 or more categories.

# MH Summary OR

- ▶ The MH summary estimator is a weighted average of the stratum-specific OR estimators
  - ▶ The MH estimator was developed over 50 years ago, is common practice in Epidemiology, and has good statistical properties.
  - ▶ Hence it is worth knowing.
- ▶ The MH estimate of the OR is

$$\widehat{OR} = \frac{\sum_{i=1}^g a_i d_i / T_i}{\sum_{i=1}^g b_i c_i / T_i}$$

where  $a_i$ ,  $b_i$ ,  $c_i$ ,  $d_i$  are from the table for the  $i$ th stratum of the extraneous variable, and  $T_i$  is the table total for this  $i$ th table.

- ▶ See the text, page 381 for a worked example of this calculation.

# The MH Test

- ▶ Once we have the MH estimate of the common OR across the strata of the confounding variable, we may test the null hypothesis that this common OR is one.
  - ▶ i.e., that there is no association between the disease outcome and the exposure variable after adjusting for the confounding variable.
- ▶ The text, section 16.2.3, goes into the details of the derivation, if you are interested.
- ▶ Instead, we'll use R to apply this test (see R demo).

## MH Test on AS Example

```
##
## Mantel-Haenszel chi-squared test without continuity correction
##
## data:  smoke and AS and sex
## Mantel-Haenszel X-squared = 0.44568, df = 1, p-value = 0.5044
## alternative hypothesis: true common odds ratio is not equal to 1
## 95 percent confidence interval:
##  0.6877312 2.1460496
## sample estimates:
## common odds ratio
##           1.214868
```

- ▶ Referring to the p-value of 0.5044 reported in the output, we see no statistical evidence that smoking affects the odds at the 5% level, once we adjust for sex.
- ▶ The MH estimate of the OR adjusted for sex is about 1.21.

*"We estimate that smoking is associated with a 1.21-fold increase in the odds of developing AS, **after adjusting for the effects of sex.**"*

# CI for the Common OR

- ▶ The MH output from R returns a 95% CI of about (0.69, 2.1).

*"With 95% confidence, we estimate that smoking is associated with between a 0.69- to 2.1-fold increase in the odds of developing AS, after adjusting for the effects of sex."*

- ▶ Notes

1. The association between smoking and AS is not significant at level 5%.
2. Therefore, the 95% CI covers 1
3. The lower limit of the CI corresponds to an estimated 0.69-fold "increase" in the odds of AS from smoking.
  - ▶ Re-express as a  $1/0.69 = 1.45$ -fold **decrease** in the odds of AS.
4. Seems strange that smoking would decrease the odds of AS.
  - ▶ But remember that the CI is an interval estimate of the population OR that is subject to sampling variability.

## Calculating CIs

- ▶ Though not discussed, can extend the MH test to arbitrary values of the common OR.
  - ▶ We only covered the special case where the value hypothesized is 1; i.e.  $H_0 : OR = 1$  (unassociated) vs.  $H_a : OR \neq 1$  (associated).
- ▶ The 95% CI is obtained by finding the set of all OR values that would be retained in a test at the 5% level.
  - ▶ E.G., With the AS data, if we were to test  $H_0 : OR = 0.69$  vs.  $H_a : OR \neq 0.69$ , we'd retain  $H_0$  at 5% level.
  - ▶ Similarly, if we were to test  $H_0 : OR = 2.1$  vs.  $H_a : OR \neq 2.1$ , we'd retain  $H_0$  at 5% level.
  - ▶ And the same for all values in between.
  - ▶ Any values smaller than 0.69 or larger than 2.1 for  $H_0$  would lead to rejection with these data.
- ▶ Thus the 95% CI obtained from the AS data is (0.69, 2.1).
  - ▶ This is called “inverting the test” and is how R gets the CI (but not how the text gets it, which we'll skip).

# Summary

- ▶ Inference of an association between an exposure variable and disease outcome should consider extraneous variables that could be effect modifiers or confounders.
- ▶ We first look for effect modification by testing for homogeneous ORs across the strata of the extraneous variable.
- ▶ If an extraneous variable is not an effect modifier, we may adjust for it as a potential confounder by applying the MH estimator of the common OR.
  - ▶ The MH estimator automatically adjusts for the extraneous variable (but is only valid if there is no effect modification).
  - ▶ In this class, we'll omit the material on the approximate CI for the common OR given in section 16.2.2 of the text and instead use the output from R which “inverts” the hypothesis test as described on the previous page.