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Rate-of-change analysis in palaeoecology revisited: a new approach

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ABSTRACT

1. Dynamics in the rate of compositional change beyond the time of human observation are uniquely preserved in palaeoecological sequences from peat or lake sediments. Changes in sedimentation rates and sampling strategies result in an uneven distribution of time intervals within stratigraphical data, which makes assessing rates of compositional change and the detection of periods with a high rate-of-change (RoC) or ‘peak-points’ challenging. Despite these known issues and their importance, and the frequent use of RoC in palaeoecology, there has been relatively little exploration of differing approaches to quantifying RoC.

2. Here, we introduce R-Ratepol (an easy to use R package) that provides a robust numerical technique for detecting and summarising RoC patterns in complex multivariate time-ordered stratigraphical sequences. We compare the performance of common methods of estimating RoC using simulated pollen-stratigraphical data with known patterns of compositional change and temporal resolution. In addition, we apply our new methodology to four representative European pollen sequences.

3. Simulated data show large differences in the successful detection of known patterns in RoC peak-point detection depending on the smoothing methods and dissimilarity coefficients used, and the level density and their taxonomic richness. Building on these results, we propose a new method of binning with a moving window in combination with a generalised additive model for peak-point detection. The method shows a 22% increase in the correct detection of peak-points and 4% lower occurrence of false positives compared to the more traditional way of peak selection by individual levels, as well as achieving a reasonable compromise between type I and type II errors. The four representative pollen sequences from Europe show that our methodological combination also performs well in detecting periods of significant compositional change including the onset of human activity, early land-use transformation, and changes in fire frequency.

4. Expanding the approach using R-Ratepol to the increasingly available stratigraphical data on pollen, chironomids, or diatoms will allow future palaeoecological and macroecological studies to quantify, and then attribute, major changes in biotic composition across broad spatial areas through time.

KEYWORDS: Compositional change; peak-point detection; pollen analysis; rate-of-change; simulated data

INTRODUCTION

Quantifying spatio-temporal changes in biological composition or diversity is essential for monitoring the current biodiversity crisis and for disentangling underlying drivers like climate, land-use change, pollution, or introduction of invasive species, as well as intrinsic population dynamics driven by biotic interactions (e.g. Wolfe, Champ, Flemer, & Mearns, 1987; Richardson et al., 2006; Silvertown et al., 2006; Gotelli, Dorazio, Ellison, & Grossman, 2010; Magurran et al., 2010; Monchamp et al., 2018; Steinbauer et al., 2018). The importance of such ecological long-term observational studies (50-100 years) is increasingly recognised (Dornelas et al., 2013, 2018; Hillebrand et al., 2018; Magurran, Dornelas, Moyes, & Henderson, 2019). Substantial compositional turnover has been observed during recent decades (e.g. Gibson-Reinemer, Sheldon, & Rahel, 2015; Steinbauer et al., 2018; Feeley, Bravo-Avila, Fadrique, Perez, & Zuleta, 2020) and “The Great Acceleration” following human influences (Steffen, Broadgate, Deutsch, Gaffney, & Ludwig, 2015), either as a result of human impact earlier within the Holocene (Stephens et al., 2019), or in response to regional climate change (Shuman, Bartlein, & Webb, 2005; Seddon, Macias-Fauria, & Willis, 2015). However, similar rates of environmental change may also have occurred in former time periods (Kemp, Eichenseer, & Kiessling, 2015). To understand the impacts of humans on ecosystems it is essential to compare temporal changes in composition and diversity through human history and also to investigate whether such

changes are unique to the epoch of human impact ('Anthropocene') or whether they precede human-dominated systems (Birks, Felde, & Seddon, 2016).

Palaeoecological sequences of pollen, diatoms, chironomids, cladocerans, etc. preserved in peat or lake sediments are a unique resource for quantifying spatio-temporal changes in biological composition and diversity beyond the time period of human observations. Rate-of-change (RoC) analysis was introduced into palaeoecology by Jacobson and Grimm (1986) to quantify the rate of compositional change and the magnitude of change within a Holocene pollen sequence. It was extended by Jacobson, Webb and Grimm (1987) to quantify and compare rates of change within and between sequences (see also Grimm & Jacobson, 1992) in an attempt to identify regional-scale and local-scale patterns in rates of change. RoC analysis estimates compositional change or temporal beta-diversity between adjacent stratigraphical levels and hence between times. Unlike other estimates of beta-diversity, RoC analysis specifically estimates the magnitude of compositional change per *unit time*. Therefore, an essential requirement in RoC analysis is a robust age-depth model for the stratigraphical sequence to precisely estimate the age of levels within the sequence.

Despite the wide use of RoC analysis in palaeoecology (e.g. Laird, Fritz, & Cumming, 1998; H. H. Birks & Ammann, 2000; H. J. B. Birks & Birks, 2008; Solovieva, Jones, Birks, Appleby, & Nazarova, 2008; Urrego et al., 2009; Correa-Metrio et al., 2012; Grindean, Tanțău, & Feurdean, 2019; Szabó et al., 2020), there has been relatively little exploration of the underlying methodology of RoC analysis and how the various methodologies and sequence properties influence RoC estimates (but see Lotter, Ammann, & Sturm, 1992; Bennett & Humphry, 1995; Birks, 2012). Different approaches to estimate RoC (Bennett & Humphry, 1995; Birks, 2012) involve choices in terms of approach and technique, such as transforming the stratigraphical levels and associated fossil data (using binning or interpolation to constant time intervals), smoothing the data, and the metric used to estimate the amount of compositional dissimilarity between adjacent levels (Birks, 2012).

The variety of approaches available to RoC analysis can, however, create problems. First, Anderson et al. (2020) stress the sensitivity of RoC to temporal sampling variation within and between records, consequently hampering conclusive comparisons. Second, there has been a lack of standardisation between studies (i.e. incorporating different dissimilarity metrics, different temporal procedures) which means that it is not possible to compare RoC between sequences (Birks, 2012). Third, intrinsic properties of an individual sequence, such as its taxonomic richness and the density and distribution of levels within the sequence, can influence the estimated RoC. The end result is that, since the expected patterns of RoC in a sequence are unknown, the most appropriate choices or methods to undertake RoC analysis are also unknown (Bennett & Humphry, 1995; Birks, 2012).

Here, we present a new algorithm for RoC analysis as a standardised and robust methodology, and hence current best practice, for estimating and comparing RoC estimates within and between stratigraphical sequences. Applying our approach, we compare the performance of various methods of estimating RoC using simulated pollen-stratigraphical data (Blaauw, Bennett, & Christen, 2010) with known patterns of compositional change and resolution. We compare results from different smoothing methods and dissimilarity coefficients. In addition, we introduce a method for detecting ‘peak-points’, defined as a rapid change in taxonomic composition or relative abundances of pollen taxa within the sequence, which provides a means of comparing RoC between sequences and interpreting the potential drivers of assemblage change. We illustrate our approach by estimating RoC values for four palynological sequences with different densities of stratigraphical levels and pollen richness, incorporating age uncertainties from Bayesian age-depth modelling and standardisation of pollen richness. Our approach can provide important information for future palaeoecological and macroecological studies attempting to quantify, and then attribute, major changes in biotic composition across broad spatial areas, and to compare the observed changes in recent decades (‘Anthropocene’) with changes that occurred within the Holocene (e.g. Gibson-Reinemer et al., 2015).

MATERIALS AND METHODS

We develop a novel statistical approach to evaluate RoC in a single stratigraphical sequence using pollen-count data and age uncertainties for each level. R-Ratepol is written as an R package (R Core Team 2018). The process of computation in R-Ratepol is summarised in Fig. 1 (a detailed description of each step and all the formulae for the calculations are given in the Supplementary Methods)

Selection of working units

RoC is calculated between consecutive levels or working units (WUs). Traditionally, these WUs represent individual stratigraphical levels. However, changes in sedimentation rates and sampling strategies can result in an uneven distribution of levels within a time sequence, which in turn makes the comparison of RoC between sequences problematic. There are various methods that attempt to minimise such problems. The first is **interpolation of levels** to evenly spaced time intervals, and the use of the interpolated data as WUs. This can lead to a loss of information when the density of levels is high. Second is **binning of levels**: pollen data are pooled into age brackets of various size (bins) and these serve as WUs. Here, the issue is a lower resolution of WUs and their uneven size in terms of total pollen count (bins with more levels have higher pollen counts). Third is **selective binning**. Like classical binning, bins of selected size are created, but instead of pooling pollen data together, only one level is selected as representative of each bin. This results in an even number of WUs in bins with a similar number of pollen grains. However, the issue of low resolution remains.

We propose a **new method of binning with a moving window**, which is a compromise between using individual levels and selective binning. This method follows a simple sequence: bins are created, levels are selected as in selective binning, and RoC between bins is calculated. However, the brackets of the time bin (window) are then moved forward by a selected amount of time (Z), levels are selected again, and RoC calculated for the new set of WUs. This is repeated X times (where X is the bin size divided by Z) while retaining all the results (See Fig. 1 and the Supplementary Methods).

Simulated data

We simulated pollen datasets following Blaauw *et al.* (2010) based on generating pollen-assemblage data in response to known changes in environmental conditions and compositional properties (Fig. S1). The settings in the simulation algorithm were changed according to: i) richness – *low richness* (LR) datasets contain 5 pollen taxa, *high richness* (HR) datasets contain 50 pollen taxa; and ii) position of change in external environmental properties – *recent* (R) has a sudden increase of environmental properties at 2000 calibrated years before present (1950 CE: cal yr BP) and a decrease at 3000 cal yr BP, *late* (L) has a sudden increase at 5500 and a decrease at 6500 cal yr BP (Fig. S1). This results in two abrupt changes in pollen composition in each dataset. Note that the different timings of the change in environmental properties (hereafter called ‘focal area’) were selected to illustrate the effects of level density (datasets with the *later* manual edit have a low density of levels within the focal area and datasets with a *recent* manual edit have a high density). The combination of these two settings results in four types of simulated datasets: *low richness – recent change* (LR-R), *low richness – late change* (LR-L), *high richness – recent change* (HR-R), and *high richness – late change* (HR-L). To increase the reliability of the simulated data, we used the density of levels from a European pollen sequence, Glendalough (Sequence A) as a template for the number of levels and their corresponding ages to each level. We then adjusted the total values of each simulated taxon so that the pollen assemblage resembles a log-scale rank distribution and added jitter to each of the pollen taxa (function *jitter*, factor = 1.5).

Using the simulated datasets to test analytical performance with different settings

Using R-Ratepol, we implemented three types of WU selection (individual stratigraphical levels, selective binning, and binning with moving window), five types of smoothing of pollen data (none, moving average, Grimm’s smoothing (Grimm & Jacobson, 1992), age-weighted average, Shepard’s 5-term filter (Davis, 1986; Wilkinson, 2005)), four types of dissimilarity coefficient (Euclidean distance, standardised Euclidean distance, Chord distance, Chi-squared coefficient (Prentice, 1980)), and three types of peak-point detection (Threshold, generalised additive model (GAM), Signal-to-Noise Index (SNI; Kelly, Higuera, Barrett, & Feng Sheng, 2011)). In order to investigate the differences between the various settings, we used the four types of simulated datasets (LR-R, LR-L, HR-R, HR-L; see above),

simulated 100 datasets for each dataset type, and compared the success rate of detecting peak-points in the expected time period for all combinations of WU selection, smoothing type, dissimilarity coefficient, and peak-point detection.

To detect significant peak-points of RoC scores in each sequence, each point is tested to see if it represents a significant increase in RoC values (the criteria vary between the peak-point detection methods, see Supplementary Methods for details). For all simulated datasets with all WU selections, smoothing, and dissimilarity coefficient combinations, we extracted a) the number of WUs where R-Ratepol detects a significant peak-point during the focal area (± 500 yr), and b) the number of samples, where R-Ratepol detects a peak-point during a time of no expected change ('false positives'). First, we test if the various WU selection methods affect the successful detection of peak-points within the focal area. We pooled all randomisations from all dataset types and created generalised linear mixed models using the Template Model Builder (Brooks et al., 2017) `glmmTMBsuccess-method` with the ratio of levels marked as peak-points to all levels in the focal area ($R_{success-method}$) as the dependent variable with a beta-binomial error distribution. Independent variables were: WU (3 level factor), peak-point detection method (3 level factor), position of environmental change (i.e. density of levels; 2 level factor), richness of dataset (2 level factor), and all their interactions. Individual dataset ID was selected as the random factor. We then used the *dredge* function from the *MuMIn* package (Barton, 2020) to reduce unnecessary predictors using a test of parsimony (AICc). We used the *emmeans* package (Russell, 2020) to obtain estimated marginal means and the 95% quantile of the independent variables from the final model. Similarly, we built `glmmTMBFalsePositive-method` using the ratio of WU marked as peak-points to all WU outside the focal area ($R_{FalsePositives-method}$).

Next, for more a detailed exploration of different settings in R-Ratepol, we divided the data to only include the best performing WU selection and peak-point detection combination (i.e. binning with moving window and GAM peak-point detection, see Results) and created `glmmTMBsuccess-settings` and `glmmTMBFalsePositive-settings` with $R_{success-settings}$ and $R_{FalsePositives-settings}$ as the dependent variables, respectively. Both models were fitted with a beta-binomial error distribution and the independent variables were data-smoothing type (5 level factor), dissimilarity coefficient (4 level factor), position of environmental change (i.e. density of levels; 2 level factor), richness of dataset (2 level factor), and all their interactions. Individual dataset ID was selected as the random factor. The *dredge* function was used to reduce unnecessary predictors.

Examples of rate-of-change results using real palynological data

After assessing the differences between the individual settings of R-Ratepol based on the results from the simulated datasets, we examined the RoC score for real palynological data representing different

pollen richness and composition, and different density of levels. We obtained pollen data from the Neotoma database (Williams, Grimm, et al., 2018) using the *Neotoma R package* (Goring et al., 2015). We chose four European sequences (A–D; Fig 2). Taxa were standardised to the taxonomically highest pollen morphotype (Level = MHVar2) using the pollen harmonisation table in Giesecke et al. (2019).

To develop age-depth models, we used the pre-selected radiometric control points provided in Giesecke et al. (2014) and calibrated the radiocarbon dates using the IntCal13 Northern Hemisphere calibration curve (Reimer et al., 2013). For each sequence, we constructed an age-depth model using the *Bchron R package* (Haslett & Parnell, 2008) to generate 1000 possible age predictions (i.e. age uncertainties) for all levels. We calculated the median of all the uncertainties for each level to give the most probable age (default age) in cal yr BP.

In each sequence, we excluded all levels that contained less than 150 pollen counts of terrestrial taxa, and all levels beyond a 3000-years extrapolation of the oldest chronological control point. In addition, we excluded all levels with an age older than 8500 cal yr BP to ensure focus on the period with substantial human impact.

Sequence A is from Glendalough, Ireland (53°00'10.0"N 6°22'09.0"W; Haslett et al., 2006). It has a very uneven distribution of levels (N = 101) with the highest density around 2000 cal yr BP (Fig. 2A). The sequence contains 80 pollen taxa with abundant *Corylus*, *Quercus*, *Alnus*, *Betula*, and Poaceae. Cyperaceae, Poaceae, and Ericales all increase in the last 1000 cal yr BP, preceded by a period with high values of *Quercus*, *Betula*, and *Alnus* until ~ 5000 cal yr BP. Before 5000 cal yr BP the sequence has large variations in *Pinus*, an increase of *Alnus*, and a large decrease in *Corylus*. **Sequence B** is from Dallican Water, Scotland (60°23'14.5"N 1°05'47.3"W; Bennett, Boreham, Sharp, & Switsur, 1992). It has a relatively even distribution of levels (N = 60; Fig. 2B) and contains 50 pollen taxa with Ericales, *Betula*, and Poaceae being the most abundant. The pollen record shows sudden increases of Ericales after 4000 cal yr BP. **Sequence C** is from Steerenmoos, Germany (47°48'20.0"N 8°12'01.6"E; Rösch, 2000). It is an example of a very detailed sequence, but with an uneven distribution of levels (N = 256; Fig. 2C). The sequence contains 103 pollen taxa with *Abies*, *Fagus*, *Corylus*, *Betula*, and *Quercus* being the most abundant. As with Sequence A, the pollen stratigraphy can be separated into three major parts: i) a recent period until 1000 cal yr BP typified by high values of *Pinus*; ii) 2000–6000 cal yr BP typified by *Abies* and *Fagus*; and iii) 6000 cal yr BP to the base with abundant *Betula*, *Corylus*, and *Quercus*. **Sequence D** is from Alanen Laanijärvi, a boreal-forest lake in Swedish Lapland (67°58'00.9"N 20°27'37.9"W; Heinrichs, Peglar, Bigler, & Birks, 2005). It is an example of a sequence with 54 levels across the time of interest, containing 44 pollen taxa with high values of *Pinus* and *Betula* throughout the whole sequence (Fig. 2D).

We calculated RoC scores for the four selected European sequences (A-D) using all three methods of WU selection with age-weighted average smoothing and Chi-squared coefficient selected as the dissimilarity coefficient. A GAM approach was used for detecting peak-points in each sequence, the number of randomisations was set to 10,000, and the size of time bin was selected as 500 yr (see Fig. S2 for example of changes of RoC values with size of time bin). In addition, to provide examples of the differences between different settings of R-Ratepol, we explored data from sequence A using binning with the moving window method of WU selection and calculated RoC scores for all data smoothing and dissimilarity coefficient combinations.

RESULTS

Comparison of success rates in peak-point detection for simulated datasets

Successful detection of peak-points ($R_{\text{success-method}}$) and the number of incorrectly (false positives) detected peak-points ($R_{\text{FalsePositives-method}}$) are both significantly affected by WU selection, method of peak-point detection, density of levels in the sequence, and taxonomic richness of the dataset (Fig. 3A; see Table S1 and Table S2 for the best selected models). For detecting peak-points successfully within the focal area, the GAM method scores relatively highly (estimated marginal means are 60% higher than the SNI method; Fig. 3A), and with a relatively low false positive rate compared to the other peak-point detection methods (estimated marginal means are 40% lower than the threshold method; Fig. 3A). Simulations show that using our approach of binning with a moving window results in an overall better performance in successful peak-point detection than selective binning or the use of individual levels (Fig. 3). The GAM method has 20 and 22% higher marginal means of success detection than selective binning and the use of individual levels, respectively.

While using only the binning with the moving-window approach and the GAM method for peak-point detection, the successful detection of peak-points ($R_{\text{success-settings}}$) and the number of false positives ($R_{\text{FalsePositives-settings}}$) are both influenced by the density of levels, dataset richness, type of dissimilarity coefficient, and data-smoothing type (Fig. 3B; see Tables S3 and S4 for the best selected models). The datasets with recent environmental change (i.e. change in the time of high density of levels) show estimated marginal means with a 28% higher success of peak-point detection and 3% lower false positive detection than datasets with environmental change later in the sequence (i.e. change at the time of low density of levels; Fig. S3). Datasets with a high richness of pollen taxa have estimated marginal means with a 5% higher success in the detection of peak-points than datasets with low richness (Fig. S3). Datasets using the Chi-squared coefficient as the dissimilarity coefficient have the highest number of correctly detected peak-points (3% higher estimated marginal means than the second best, Chord distance; Fig. S3). All smoothing types show similar results, except Grimm's

smoothing, which has the lowest success in the correct detection of peak-points and the highest number of false positives compared to other smoothers (Fig. S3).

Examples of rate-of-change results using real palynological data

The selection of WUs affected not only the overall shape of the RoC curve in all sequences but also the timing of the period of significant increase in the RoC (Fig. 2). Binning with a moving window detected significant periods of increased RoC scores at sequence A (Glendalough) between 4200–5000 cal yr BP; sequence B (Dallican Water) between 3200–3500 cal yr BP; sequence C (Steerenmoos) between 5800–6400 cal yr BP; and sequence D (Alanen Laanijärvi) between 350–580 cal yr BP (Fig. 2). These patterns remain relatively similar with selective binning but not when individual levels are used as WUs.

DISCUSSION

Based on the results of our simulations, the choice of methods and parameters is critical in the analysis of rate-of-change in pollen sequences. Comparison of methods using simulated datasets shows that selection of WUs, peak-point detection method, dissimilarity coefficient, smoothing of the data, density of the levels in the period of environmental change, and pollen richness significantly influence the detection of sudden changes in stratigraphical sequences. Our proposed method of binning with a moving window in combination with GAM peak-point detection improves the detection of increased RoC. The method shows a 22% increase in the correct detection of peak-points and a 4% lower occurrence of false positives compared to the more traditional approach of selection by individual levels. Peak-point detection using GAM provides a reasonable compromise between statistical error type I (*Threshold* method) and type II (*SNi* method).

Using the Chi-squared coefficient as the dissimilarity coefficient shows overall the best results, and we therefore recommend using this measure for palaeoecological data sequences that are expressed as percentages or proportions. Recommendation of a specific smoothing technique of data is less straightforward. There is a trade-off between the degree of smoothing of the data and the density of samples in the expected period of higher RoC. For example, Shepard's 5-term filter (relatively soft smoothing of data) and data without smoothing do not perform well in the specific case where environmental change occurred during the period with a low density of samples. Detailed exploration of the RoC pattern in sequence A (Fig. 4) shows that all smoothing methods detected the RoC increase in the period between 4000–6000 cal yr BP but also others in the recent period (0–500 yr BP). The latter are likely to be false positives as this period is characterised by a low density of levels and used Shepard's 5-term filter data smoothing. Smoothing using age-weighted average is able to perform better in such a scenario but at the cost of an increased number of false positives. Therefore, there is

a need for critical evaluation in future studies to assess the risks of using either Shepard's 5-term filter or age-weighted average data for smoothing, depending on the research questions of interest.

A multitude of appropriate data-smoothing approaches (Wilkinson, 2005), dissimilarity coefficients (e.g. Prentice, 1980; Legendre & Legendre, 2012), techniques for creating working units, and methods of peak-point detection (e.g. Simpson, 2018) are potentially available. Of the four selected dissimilarity methods and four smoothing algorithms, all commonly used in palaeoecology, we conclude that datasets with a high density of levels tend to have a higher chance of successfully detecting peak-points. Anderson et al. (2020) reached similar conclusions in a regional-scale synthesis of fossil pollen data from California. Here the authors suggest maintaining a consistent temporal spacing within records, and the use of probabilistic models explicitly incorporating age-model uncertainties to increase the precision of age estimates of each level. We show that our method of binning with a moving window yields better results than the traditional selective binning method. Using R-Ratepol is an advantage in this case as such age-model uncertainties can be incorporated.

RoC analysis has several applications in palaeoecological research. It is a useful numerical tool to detect patterns in stratigraphical data that cannot readily be seen by visual inspection (see Jacobson et al., 1987 for an example at Gould Pond, Maine). It is also useful to compare the rates of change in different proxies (e.g. pollen, diatoms, chironomids, etc.) studied in the same stratigraphical sequence (e.g. Birks & Ammann, 2000; Szabó et al., 2020). Moreover, it is most useful when RoC results are compared from several sequences (e.g. Jacobson et al., 1987; Grimm & Jacobson, 1992). Consistent patterns in RoC peaks in several sequences may potentially indicate response to extrinsic drivers such as regional climate change, pathogenic attacks, or widespread human activity. Ecologists are recognising the importance of quantifying RoC of extrinsic drivers such as temperature, toxins, and salinity on ecosystems (Pineck et al., 2020) and are developing new and powerful numerical tools for space-time analysis of community compositional data over time intervals of decades (e.g. Legendre & Gauthier, 2014). RoC peaks unique to individual sequences may, in contrast, reflect local factors such as sequence hiatuses, sedimentary reworking, or spatially restricted human impact or disturbance.

A connection between RoC and local factors can be seen in sequences used in this study (Fig. 2). Sequence A (Glendalough) shows significant peak-points at the time of the onset of human activity and the expansion of grassland and heath (Haslett et al., 2006). Sequence B (Dallican Water) has significant peak-points at times when major changes in land-use occurred as the island was abandoned by humans and then reinhabited 800 years later (Bennett et al., 1992). Sequence C (Steerenmoos) has several significant peak-points at the time of the expansion of *Abies* pollen in the late Neolithic period

and with frequent fires (Rösch, 2000). Sequence D (Alanen Laanijärvi) has significant peak-points in the very recent period, possibly associated with changes in organic deposition, fertilisation, and timber harvesting (Heinrichs et al., 2005). Here we can assume that our approach of binning with a moving window is closest to the expected pattern, further supported by the relatively high detection of peak-points in the simulated datasets (Fig. 3). In contrast, the traditional use of individual levels fails to detect the period of significant change (5800–6400 cal yr BP) in sequence C, or falsely assigns peak-points to other, non-relevant periods (sequence A: 2000 cal yr BP; sequence C; 0–500 cal yr BP; sequence D: 0–4000 cal yr BP), or a combination of both (sequence B). Most of these errors occur, as expected, with low density sample levels. These sequences and the RoC outcomes exemplify the usefulness of quantifying rates of compositional change estimated through time.

In summary, we have developed a framework for the robust estimation of rate-of-change in stratigraphical time-ordered palaeoecological sequences of, for example, pollen, chironomids, or diatoms. Our overall framework consists of four major parts:

1. Establishing a robust age-depth model for the stratigraphical sequence of interest with age uncertainties for all individual levels.
2. Selecting consecutive working units prior to RoC estimation to allow for uneven temporal distribution of the analysed levels within the sequence using selective binning with a moving window.
3. Estimating compositional dissimilarity between working units.
4. Detecting statistically significant peak-points in the RoC estimates within the sequence.

RoC analysis as implemented in R-Ratepol is a significant advance over existing methods for RoC analysis that do not incorporate methods to detect statistically significant peak-points (Birks, 2012). R-Ratepol is, we believe, a useful addition to the toolkit of robust numerical techniques available to palaeoecologists and palaeolimnologists for detecting and summarising patterns in complex multivariate time-ordered stratigraphical sequences (Birks, 2010).

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AUTHORS' CONTRIBUTION

OM and HJBB conceived the ideas and designed the study. OM led the manuscript writing, analyses and interpretation of the data. OM, AWRS, KPB, VAF, and SGAF designed the methodology for extracting and processing palaeoecological sequences from Neotoma. OM, JAG, AWRS, MJS, and HJBB contributed to methodology development. HJBB acquired the funding. All authors contributed to the writing of the text and interpretation of the results and gave final approval for publication.

DATA AVAILABILITY

R-Ratepol is available as an R package at GITHUB (<https://github.com/HOPE-UIB-BIO/R-Ratepol-package>), including all the data used in this study. Script for all analyses used in this study can be found at GITHUB repository (<https://github.com/HOPE-UIB-BIO/RateOfChange>).

REFERENCES

- Anderson, L., Wahl, D. B., & Bhattacharya, T. (2020). Understanding rates of change: A case study using fossil pollen records from California to assess the potential for and challenges to a regional data synthesis. *Quaternary International*. doi:10.1016/j.quaint.2020.04.044
- Barton, K. (2020). MuMIn: Multi-Model Inference. R package version 1.43.17.
- Bennett, K. D., Boreham, S., Sharp, M. J., & Switsur, V. R. (1992). Holocene history of environment, vegetation and human settlement on Catta Ness, Lunnasting, Shetland. *Journal of Ecology*, 80(2), 241. doi:10.2307/2261010
- Bennett, K. D., & Humphry, R. W. (1995). Analysis of late-glacial and Holocene rates of vegetational change at two sites in the British Isles. *Review of Palaeobotany and Palynology*, 85(3–4), 263–287. doi:10.1016/0034-6667(94)00132-4
- Birks, H. H., & Ammann, B. (2000). Two terrestrial records of rapid climatic change during the Glacial-Holocene Transition (14,000–9,000 calendar years B.P.) from Europe. *Proceedings of the National Academy of Sciences of the United States of America*, 97(4), 1390–1394.
- Birks, H. J. B. (2010). Numerical methods for the analysis of diatom assemblage data. In J. P. Smol & E. F. Stoermer (Eds.), *The diatoms: applications for the environmental and earth sciences* (2nd edn., pp. 23–54). Cambridge University Press, Cambridge.
- Birks, H. J. B. (2012). Analysis of stratigraphical data. In H. J. B. Birks, A. F. Lotter, S. Juggins, & J. P. Smol (Eds.), *Tracking Environmental Change Using Lake Sediments Volume 5: Data Handling and Numerical Techniques* (pp. 355–378). Dordrecht: Springer.
- Birks, H. J. B., & Birks, H. H. (2008). Biological responses to rapid climate change at the Younger Dryas-Holocene transition at Kråkenes, western Norway. *The Holocene*, 18(1), 19–30. doi:10.1177/0959683607085572
- Birks, H. J. B., Felde, V. A., & Seddon, A. W. R. (2016). Biodiversity trends within the Holocene. *The Holocene*, 26(6), 994–1001. doi:10.1177/0959683615622568
- Blaauw, M., Bennett, K. D., & Christen, J. A. (2010). Random walk simulations of fossil proxy data. *The Holocene*, 20(4), 645–649. doi:10.1177/0959683609355180
- Brooks, M. E., Kristensen, K., van Benthem, K. J., Magnusson, A., Berg, C. W., Nielsen, A., & Skaug, H. J. (2017). glmmTMB Balances Speed and Flexibility Among Packages for Zero-inflated Generalized Linear Mixed Modeling. *The R Journal*, 9(2), 378–400. doi:10.3929/ethz-b-000240890

- Correa-Metrio, A., Bush, M. B., Cabrera, K. R., Sully, S., Brenner, M., Hodell, D. A., ... Guilderson, T. (2012). Rapid climate change and no-analog vegetation in lowland Central America during the last 86,000 years. *Quaternary Science Reviews*, 38, 63–75. doi:10.1016/j.quascirev.2012.01.025
- Davis, J. C. (1986). *Statistics and Data Analysis in Geology* (2nd edn.). J. Wiley & Sons, New York.
- Dornelas, M., Magurran, A. E., Buckland, S. T., Chao, A., Chazdon, R. L., Colwell, R. K., ... Vellend, M. (2013). Quantifying temporal change in biodiversity: Challenges and opportunities. *Proceedings of the Royal Society B: Biological Sciences*, 280(1750), 20121931. doi:10.1098/rspb.2012.1931
- Dornelas, M., Antão, L. H., Moyes, F., Bates, A. E., Magurran, A. E., Adam, D., ... Zettler, M. L. (2018). BioTIME: A database of biodiversity time series for the Anthropocene. *Global Ecology and Biogeography*, 27(7), 760-786. doi:10.1111/geb.12729
- Feeley, K. J., Bravo-Avila, C., Fadrique, B., Perez, T. M., & Zuleta, D. (2020). Climate-driven changes in the composition of New World plant communities. *Nature Climate Change*. doi:10.1038/s41558-020-0873-2
- Gibson-Reinemer, D. K., Sheldon, K. S., & Rahel, F. J. (2015) Climate change creates rapid species turnover in montane communities. *Ecology and Evolution*, 5(12), 2340-2347. doi:10.1002/ece3.1518
- Giesecke, T., Davis, B., Brewer, S., Finsinger, W., Wolters, S., Blaauw, M., ... Bradshaw, R. H. W. (2013). Towards mapping the late Quaternary vegetation change of Europe. *Vegetation History and Archaeobotany*, 23(1), 75–86. doi:10.1007/s00334-012-0390-y
- Giesecke, T., Wolters, S., van Leeuwen, J. F. N., van der Knaap, P. W. O., Leydet, M., & Brewer, S. (2019). Postglacial change of the floristic diversity gradient in Europe. *Nature Communications*, 10(1). doi:10.1038/s41467-019-13233-y
- Goring, S., Dawson, A., Simpson, G. L., Ram, K., Graham, R. W., Grimm, E. C., & Williams, J. W. (2015). Neotoma: A programmatic interface to the Neotoma paleoecological database. *Open Quaternary*, 1, 1–17. doi:10.5334/oq.ab
- Gotelli, N. J., Dorazio, R. M., Ellison, A. M., & Grossman, G. D. (2010). Detecting temporal trends in species assemblages with bootstrapping procedures and hierarchical models. *Philosophical Transactions of the Royal Society B: Biological Sciences*, 365(1558), 3621–3631. doi:10.1098/rstb.2010.0262
- Grimm, E. C., & Jacobson, G. L. (1992). Fossil-pollen evidence for abrupt climate changes during the past 18000 years in eastern North America. *Climate Dynamics*, 6, 179–184.

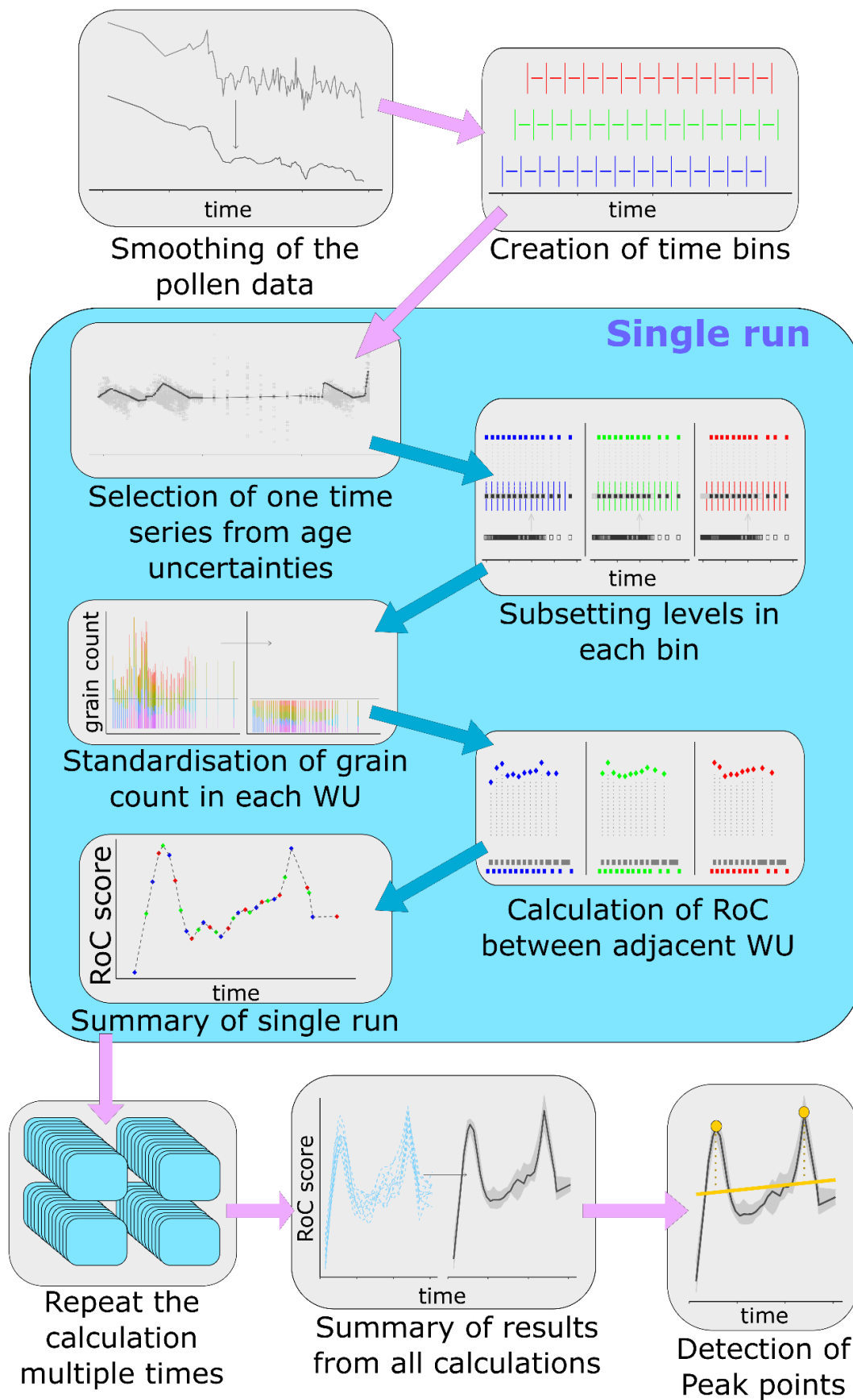
- Grindean, R., Tanțău, I., & Feurdean, A. (2019). Linking vegetation dynamics and stability in the old-growth forests of Central Eastern Europe: Implications for forest conservation and management. *Biological Conservation*, 229, 160–169. doi:10.1016/j.biocon.2018.11.019
- Haslett, J., & Parnell, A. (2008). A simple monotone process with application to radiocarbon-dated depth chronologies. *Journal of the Royal Statistical Society. Series C: Applied Statistics*, 57(4), 399–418. doi:10.1111/j.1467-9876.2008.00623.x
- Haslett, J., Whitley, M., Bhattacharya, S., Salter-Townshend, M., Wilson, S. P., Allen, J. R. M., ... Mitchell, F. J. G. (2006). Bayesian palaeoclimate reconstruction. *Journal of the Royal Statistical Society. Series A: Statistics in Society*, 169(3), 395–438. doi:10.1111/j.1467-985X.2006.00429.x
- Heinrichs, M., Peglar, S. M., Bigler, C., & Birks, H. J. B. (2005). A multi-proxy palaeoecological study of Alanen Laanijärvi, a boreal-forest lake in Swedish Lapland. *Boreas*, 34(2), 192–206. doi:10.1080/03009480510012881
- Hillebrand, H., Blasius, B., Borer, E. T., Chase, J. M., Downing, J. A., Eriksson, B. K., ... Ryabov, A.B. (2018). Biodiversity change is uncoupled from species richness trends: Consequences for conservation and monitoring. *Journal of Applied Ecology*, 55, 169–184. 10.1111/1365-2664.12959
- Jacobson, G. L., & Grimm, E. C. (1986). A numerical analysis of Holocene forest and prairie vegetation in Central Minnesota. *Ecology*, 67(4), 958–966.
- Jacobson, G. L., Webb, T., & Grimm, E. C. (1987). Patterns and rates of vegetation change during the deglaciation of eastern North America. In W. F. Ruddiman & H. E. Wright (Eds.), *North America and Adjacent Oceans During the Last Deglaciation* (pp. 277–288). Geological Society of America, Boulder, Colorado. doi:10.1130/dnag-gna-k3.277
- Kelly, R. F., Higuera, P. E., Barrett, C. M., & Feng Sheng, H. (2011). A signal-to-noise index to quantify the potential for peak detection in sediment-charcoal records. *Quaternary Research*, 75(1), 11–17. doi:10.1016/j.yqres.2010.07.011
- Kemp, D. B., Eichenseer, K., & Kiessling, W. (2015). Maximum rates of climate change are systematically underestimated in the geological record. *Nature Communications*, 6, 1–6. doi:10.1038/ncomms9890
- Laird, K. R., Fritz, S. C., & Cumming, B. F. (1998). A diatom-based reconstruction of drought intensity, duration, and frequency from Moon Lake, North Dakota: A sub-decadal record of the last 2300 years. *Journal of Paleolimnology*, 19(2), 161–179. doi:10.1023/A:1007929006001

- 465 Legendre, P., & Gauthier, O. (2014). Statistical methods for temporal and space–time analysis of
466 community composition data. *Proceedings of the Royal Society B: Biological Sciences*,
467 281(1778), 20132728. doi:10.1098/rspb.2013.2728
- 468 Legendre, P., & Legendre, L. (2012). *Numerical ecology* (3rd English edn.). Elsevier, Amsterdam.
469 doi:10.1016/B978-0-12-409548-9.10595-0
- 470 Lotter, A. F., Ammann, B., & Sturm, M. (1992). Rates of change and chronological problems during the
471 late-glacial period. *Climate Dynamics*, 6(3–4), 233–239. doi:10.1007/BF00193536
- 472 Magurran, A. E., Baillie, S. R., Buckland, S. T., Dick, J. M. P., Elston, D. A., Scott, E. M., ... Watt, A. D.
473 (2010). Long-term datasets in biodiversity research and monitoring: Assessing change in
474 ecological communities through time. *Trends in Ecology and Evolution*, 25(10), 574–582.
475 doi:10.1016/j.tree.2010.06.016
- 476 Magurran, A. E., Dornelas, M., Moyes, F., & Henderson, P. A. (2019). Temporal β diversity—A
477 macroecological perspective. *Global Ecology and Biogeography*, 28(12), 1949–1960.
478 doi:10.1111/geb.13026
- 479 Monchamp, M. E., Spaak, P., Domaizon, I., Dubois, N., Bouffard, D., & Pomati, F. (2018).
480 Homogenization of lake cyanobacterial communities over a century of climate change and
481 eutrophication. *Nature Ecology and Evolution*, 2(2), 317–324. doi:10.1038/s41559-017-0407-0
- 482 Pinek, L., Mansour, I., Lakovic, M., Ryo, M., & Rillig, M. C. (2020). Rate of environmental change across
483 scales in ecology. *Biological Reviews*. doi:10.1111/brv.12639
- 484 Prentice, I. C. (1980). Multidimensional scaling as a research tool in Quaternary palynology: A review
485 of theory and methods. *Review of Palaeobotany and Palynology*, 31(C), 71–104.
486 doi:10.1016/0034-6667(80)90023-8
- 487 R Core Team. (2018). *R: A language and environment for statistical computing*. Vienna, Austria: R
488 Foundation for Statistical Computing.
- 489 Reimer, P. J., Bard, E., Bayliss, A., Beck, J. W., Blackwell, P. G., Ramsey, C. B., ... van der Plicht, J. (2013).
490 IntCal13 and Marine13 radiocarbon age calibration curves 0–50,000 years cal BP. *Radiocarbon*,
491 55(4), 1869–1887. doi:10.2458/azu_js_rc.55.16947
- 492 Richardson, A. J., Walne, A. W., John, A. W. G., Jonas, T. D., Lindley, J. A., Sims, D. W., ... Witt, M. (2006).
493 Using continuous plankton recorder data. *Progress in Oceanography*, 68(1), 27–74.
494 doi:10.1016/j.pocean.2005.09.011

- Rösch, M. (2000). Long-term human impact as registered in an upland pollen profile from the southern Black Forest, south-western Germany. *Vegetation History and Archaeobotany*, 9(4), 205–218. doi:10.1007/BF01294635
- Russell, L. (2020). emmeans: Estimated Marginal Means, aka Least-Squares Means. R package version 1.4.8.
- Seddon, A. W. R., Macias-Fauria, M., & Willis, K. J. (2015). Climate and abrupt vegetation change in Northern Europe since the last deglaciation. *The Holocene*, 25(1), 25–36. doi:10.1177/0959683614556383
- Shuman, B., Bartlein, P. J., & Webb, T. (2005). The magnitudes of millennial- and orbital-scale climatic change in eastern North America during the Late Quaternary. *Quaternary Science Reviews*, 24(20–21), 2194–2206. doi:10.1016/j.quascirev.2005.03.018
- Silvertown, J., Poulton, P., Johnston, E., Edwards, G., Heard, M., & Biss, P. M. (2006). The Park Grass Experiment 1856–2006: Its contribution to ecology. *Journal of Ecology*, 94(4), 801–814. doi:10.1111/j.1365-2745.2006.01145.x
- Simpson, G. L. (2018). Modelling palaeoecological time series using generalised additive models. *Frontiers in Ecology and Evolution*, 6, 1–21. doi:10.3389/fevo.2018.00149
- Solovieva, N., Jones, V., Birks, H. J. B., Appleby, P. G., & Nazarova, L. (2008). Diatom responses to 20th century climate warming in lakes from the northern Urals, Russia. *Palaeogeography, Palaeoclimatology, Palaeoecology*, 259(2–3), 96–106. doi:10.1016/j.palaeo.2007.10.001
- Steffen, W., Broadgate, W., Deutsch, L., Gaffney, O., & Ludwig, C. (2015). The trajectory of the Anthropocene: The Great Acceleration. *Anthropocene Review*, 2(1), 81–98. doi:10.1177/2053019614564785
- Steinbauer, M. J., Grytnes, J.-A., Jurasinski, G., Kulonen, A., Lenoir, J., Pauli, H., ... Wipf, S. (2018). Accelerated increase in plant species richness on mountain summits is linked to warming. *Nature*, 556(7700), 231–234. doi:10.1038/s41586-018-0005-6
- Stephens, L., Fuller, D., Boivin, N., Rick, T., Gauthier, N., Kay, A., ... Ellis, E. (2019). Archaeological assessment reveals Earth’s early transformation through land use. *Science*, 365(6456), 897–902. doi:10.1126/science.aax1192
- Szabó, Z., Buczkó, K., Haliuc, A., Pál, I., L. Korponai, J., Begy, R.-C., ... Magyari, E. K. (2020). Ecosystem shift of a mountain lake under climate and human pressure: A move out from the safe operating space. *Science of the Total Environment*, 743, 140584. doi:10.1016/j.scitotenv.2020.140584

- 526 Urrego, D. H., Bush, M. B., Silman, M. R., Correa-Metrio, A. Y., Ledru, M.-P., Mayle, F. E., ... Valencia,
527 B. G. (2009). Millennial-scale ecological changes in tropical South America since the Last Glacial
528 Maximum. In F. Vimeaux, F. Sylvestre, & M. Khodri (Eds.), *Past Climate Variability in South*
529 *America and Surrounding Regions* (pp. 283–300). Springer, Dordrecht. doi:10.1007/978-90-481-
530 2672-9_12
- 531 Wilkinson, L. (2005). *The Grammar of Graphics*. Springer-Verlag, New York, USA 37.
532 doi:10.2307/2669493
- 533 Williams, J. W., Grimm, E. C., Blois, J. L., Charles, D. F., Davis, E. B., Goring, S. J., ... Takahara, H. (2018).
534 The Neotoma Paleoecology Database, a multiproxy, international, community-curated data
535 resource. *Quaternary Research*, 89(1), 156-177. doi:10.1017/qua.2017.105
- 536 Wolfe, D. A., Champ, M. A., Flemer, D. A., & Mearns, A. J. (1987). Long-term biological data sets: Their
537 role in research, monitoring, and management of estuarine and coastal marine systems.
538 *Estuaries*, 10(3), 181–193. doi:10.2307/1351847

539 FIGURES AND TABLES



540

Figure 1. Schematic visualisation of R-Ratepol calculation with binning with a moving window.

- 1) Smooth pollen data (smooth each taxon using one of five selected smoothing methods: *none*, *moving average*, *Grimm's smoothing*, *age-weighted average*, or *Shepard's 5-term filter*).
- 2) Create template for all bin brackets in all window movements.
- 3) Steps in each single run (an individual loop):
 - a. Randomly select a single age sequence from age uncertainties for all levels
 - b. Create working units (WUs) (i.e. subsetting levels in each bin)
 - c. (optional) Data standardisation (subsampling pollen data in each WU to a total count of 150 pollen grains).
 - d. Calculate the RoC as the dissimilarity coefficient (DC) between adjacent WUs (four different dissimilarity coefficients: *Euclidean distance*, *standardised Euclidean distance*, *Chord distance*, *Chi-squared coefficient*) standardised by age differences between WUs.
 - e. Summary of RoC results from all moving windows
- 4) Repeat the run multiple times (default = 10,000)
- 5) Validate and summarise results from all runs; exclude data from beyond the selected time period (i.e. 8000 cal yr BP in this study).
- 6) Detect and validate significant peak-points (three ways of testing for the significance of individual RoC points: *Threshold*, *Generalised Additive Model (GAM)*, and *Signal-to-Noise Index (SNI)*).

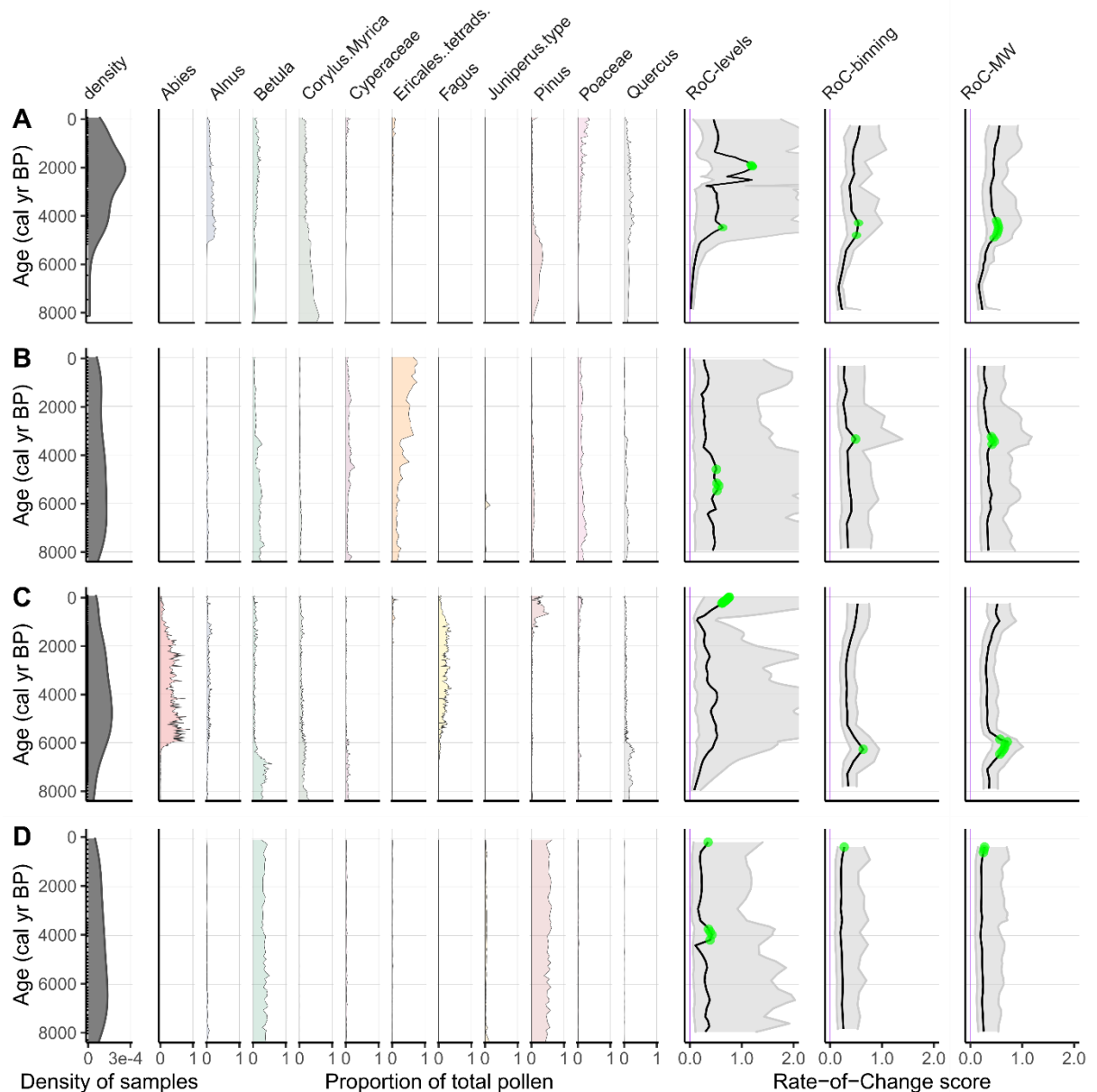


Figure 2. Summary of four sequences used in this study: Glendalough (A), Dallican Water (B), Steerenmoos (C), and Alanen Laanijärvi (D). For each sequence, three different plots are shown. First plot (left) shows the density of levels within each sequence for the last 8000 cal yr. Second plot (middle) shows pollen stratigraphies for the most common taxa in the sequences. Values are presented as proportions of pollen grains in each level. Last plot (right) shows the rate-of-change score calculated with R-Ratepol for three different methods for WU selection: RoC-levels = use of individual levels, RoC-binning = selective binning, RoC-MW = binning with moving window. Age-weighted average and Chi-squared coefficient were selected for smoothing and dissimilarity

570 measure, respectively. Green points indicate significant peak-points detected using the GAM
571 method. For a detailed explanation of the methods, see Supplementary Methods.

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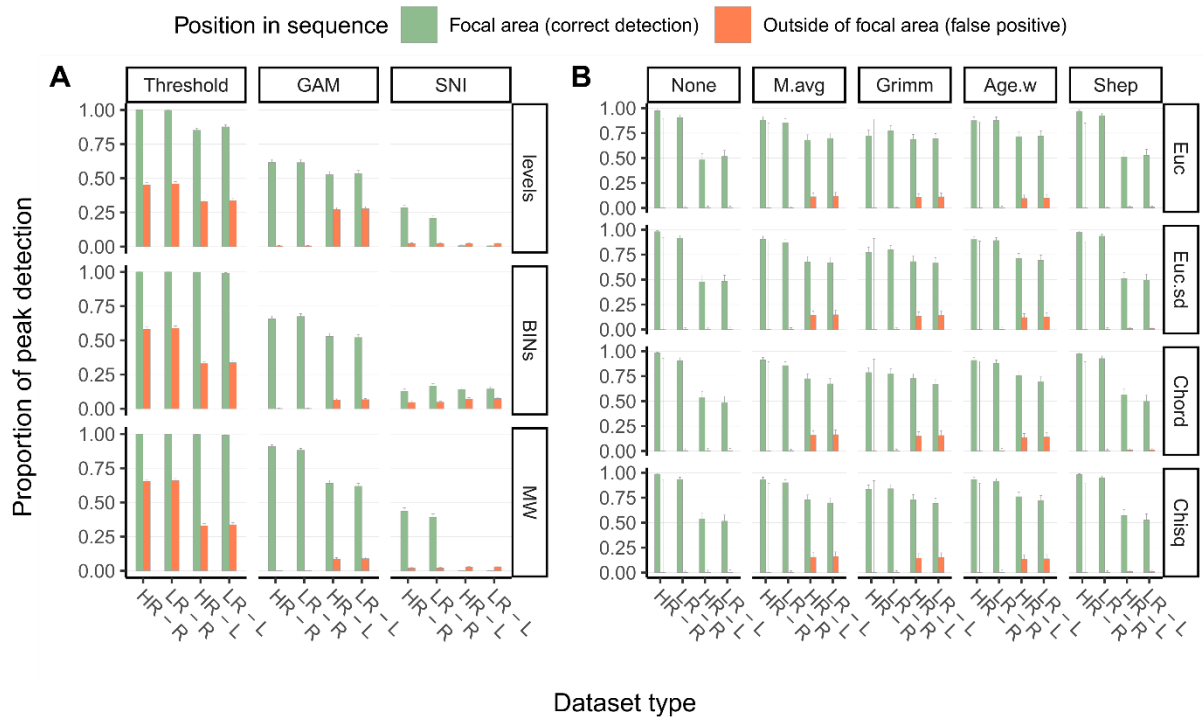


Figure 3. Comparison of the correct detection of peak-points in simulated datasets with a high or low density of levels in datasets of high or low richness. The focal area is the time of the manual edit of the environmental data (Fig. S1). Four types of simulated datasets are used: *low richness – recent manual edit of environmental properties* (LR-R), *low richness – late position of the change in environmental properties* (LR-L), *high richness – recent manual edit of environmental properties* (HR-R), and *high richness – late position of the change in environmental properties* (HR-L). A) Comparison of peak-point detection rates between different WU selections and peak-point detection methods. B) Comparison of peak-point detection rates using binning with the moving window and GAM method. Peak-point detection methods: Threshold = Threshold method, GAM = GAM method, SNI = Signal-to-Noise Index method. Working unit selection: levels = use of subsequent levels, BINs = selective binning, MW = binning with moving window. Smoothing methods: None = data without smoothing, M.avg = moving average, Grimm = Grimm’s smoothing, Age.w = age-weighted average, Shep = Shepard’s 5-term filter. Dissimilarity coefficients: Euc = Euclidian distance, Euc.sd = Standardised Euclidean distance, Chord = Chord distance, Chisq = Chi-squared coefficient. For a detailed explanation of the methods, see Supplementary Methods.

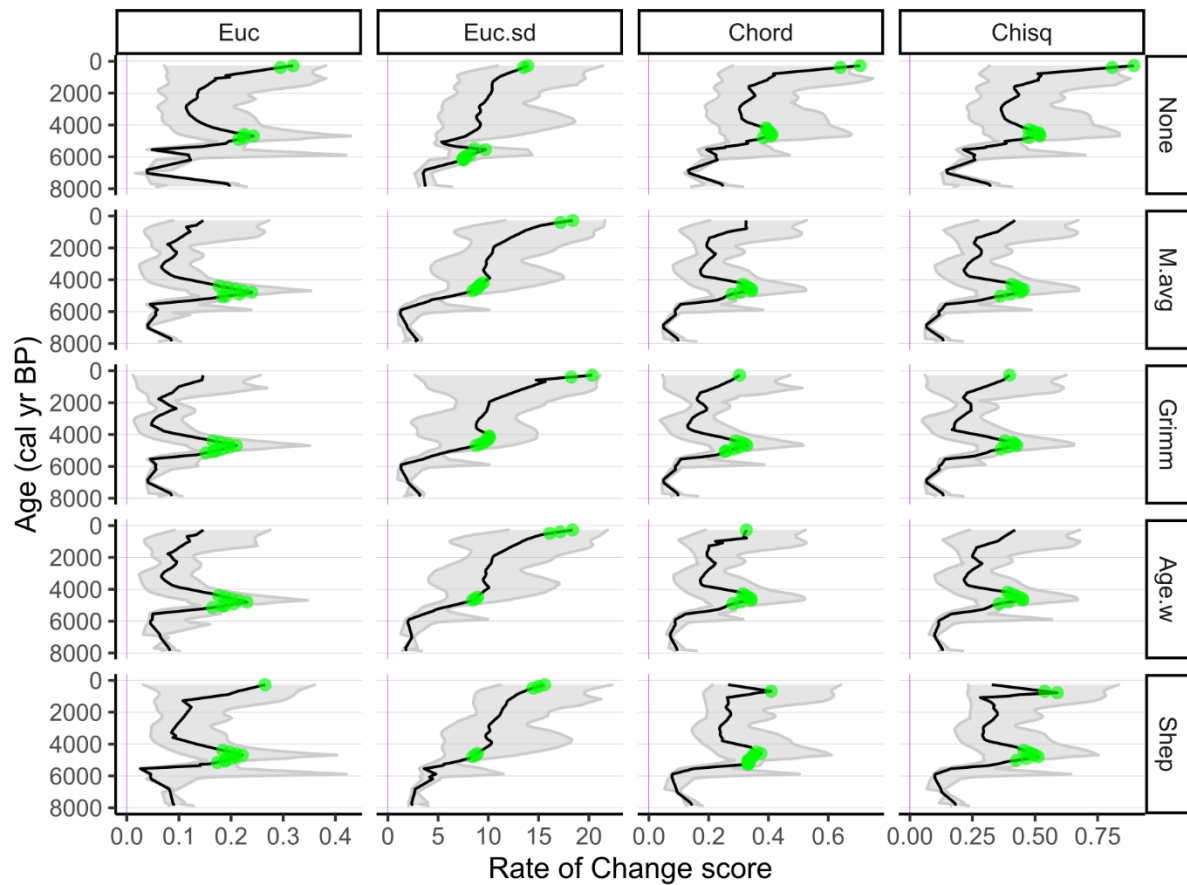


Figure 4. Rate-of-change (RoC) and peak-points detected in Sequence A (Glendalough) for four dissimilarity coefficients (columns) on five categories of data smoothing (rows). Green points indicate significant peak-points detected using the GAM approach. Dissimilarity coefficients: Euc = Euclidian distance, Euc.sd = Standardised Euclidean distance, Chord = Chord distance, Chisq = Chi-squared coefficient. Smoothing methods: None = data without smoothing, M.avg = moving average, Grimm = Grimm's smoothing, Age.w = age-weighted average, Shep = Shepard's 5-term filter. For a detailed explanation of the methods, see Supplementary Methods.