Classical statistical inference

Uncertainties on arbitrary RV

Part 4

Associated notebook:

04-Basic_statistical_inference_frequentists_2/Frequentist_Monte_Carlo.ipynb

How to calculate stde (or simply std) on a RV if it is a function of one or multiple RV?

Case 1:
$$y = \phi(x)$$

$$p(x)$$
 known

Case 2:
$$z = \phi(x, y)$$

$$p(x)$$
 and $p(y)$ known

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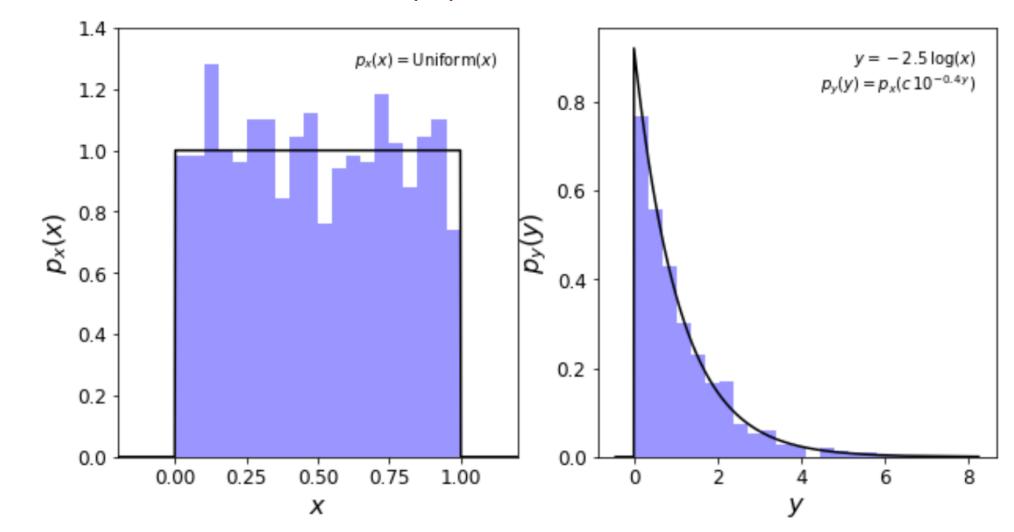
$$x = \phi^{-1}(y) = 10^{-0.4*y} \qquad p(x) \sim U(0,1) < \int_{0 \text{ if } x \in [0,1]}^{1 \text{ if } x \in [0,1]}$$

$$p(y) = 1 \times 0.4 \ln(10) 10^{-0.4 y}$$

$$0 < y < \infty$$

Go to: Sect. II.4.1 of the notebook

Example: y = -2.5 * log(x)



Case 2:
$$z = \phi(x, y)$$

$$p(x)$$
 and $p(y)$ known

Error propagation formula

$$\sigma_z^2 = \left(\frac{\partial \phi}{\partial x}\right)_{\bar{x}}^2 \sigma_x^2 + \left(\frac{\partial \phi}{\partial y}\right)_{\bar{y}}^2 \sigma_y^2 + 2\left(\frac{\partial \phi}{\partial x}\right)_{\bar{x}} \left(\frac{\partial \phi}{\partial y}\right)_{\bar{y}} \sigma_{xy}$$

Results from Taylor expanding z around $ar{x}$ and $ar{y}$ => neglects some high order terms

Go to: Sect. II.4.2 of the notebook for the demo

The expectation ("average") of a function f(x) of a RV x can be approximated by drawing a virtually infinite sample from x

$$E(f(x)) = \int_{-\infty}^{+\infty} f(x)h(x)dx \to \frac{1}{N} \sum_{i=0}^{N} f(x_i),$$

You can specialise f() to the calculation of the mean, or the variance of a RV.

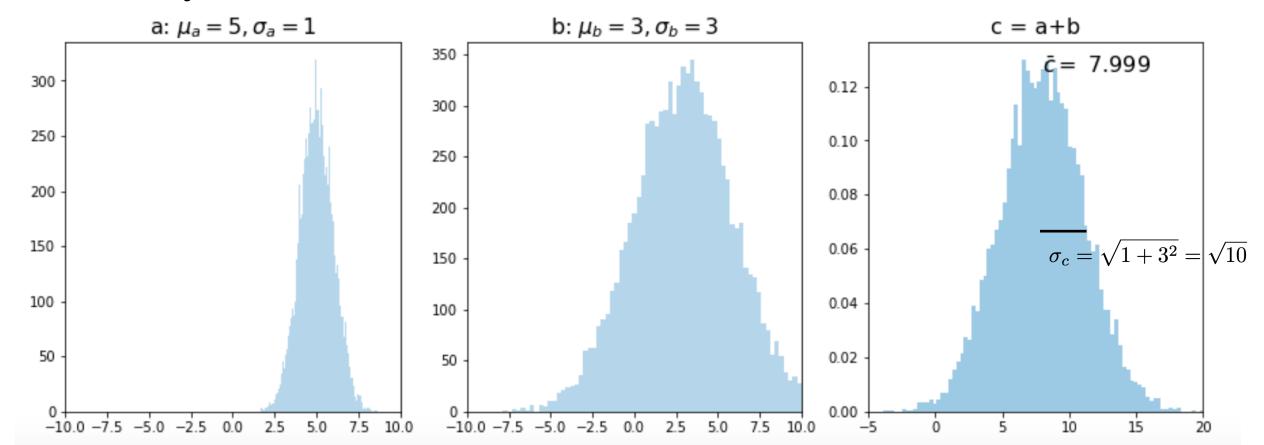
- Draw random samples from $h(a; \mu_a, \sigma_a)$, $h(b, \mu_b, \sigma_b)$
- Construct a random (monte-carlo) sample of $c_i = \Phi(a_i, b_i)$
- Derive σ_c from your Monte-Carlo sample

Monte-Carlo

$$c = a + b$$

Go to: Sect. II.4.3 of the notebook

- 1. Draw random samples from $h(a; \mu_a, \sigma_a)$, $h(b, \mu_b, \sigma_b)$
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- 3. Derive σ_c from your Monte-Carlo sample



Classical statistical inference

Bootstrap and Jacknife

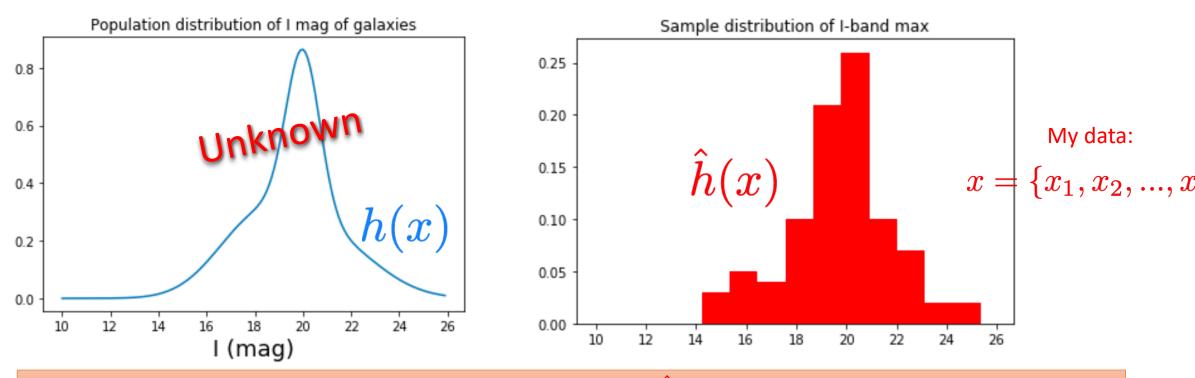
Part 5

Associated notebook:

O4-Basic statistical inference frequentists 2/ Frequentist_inference_Bootstrap.ipynb

Bootstrap

Go to: Sect. II.5 of the notebook



Bootstrap \equiv Draw samples from the sample PDF $\hat{h}(x)$, allowing from replacements

$$B = \{x_1^*, x_2^*, ..., x_n^*\}$$

$$\mathbf{w}. x^* \text{ from } \{x_1, x_2, ..., x_n\}$$

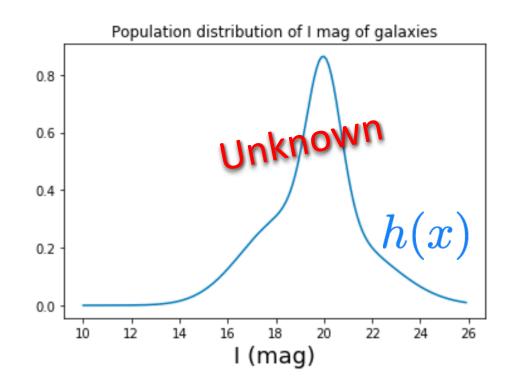
e.g.
$$B1 = \{x_1, x_1, x_7, ..., x_{28}\}$$

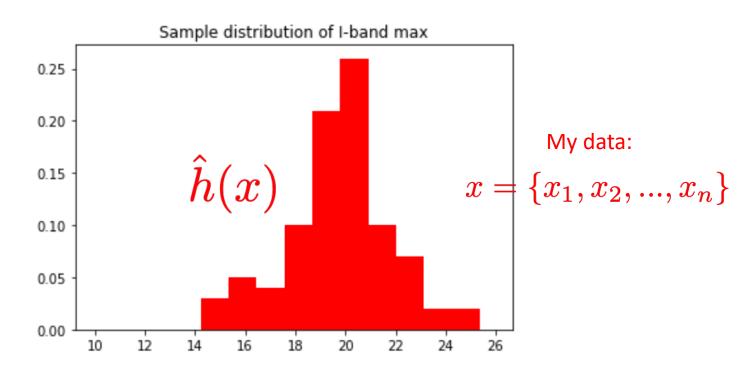
$$B2 = \{x_2, x_{36}, x_9, ..., x_8\}$$

$$Bk = \{x_{16}, x_{12}, x_3, ..., x_{10}\}$$

Bootstrap Confidence Interval

Go to: Sect. II.5.1 of the notebook





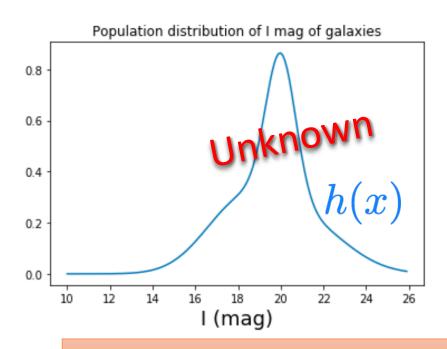
Calculate the estimate of your statistics q from all the bootstrapped samples + its associated $stde^{B}(q)$

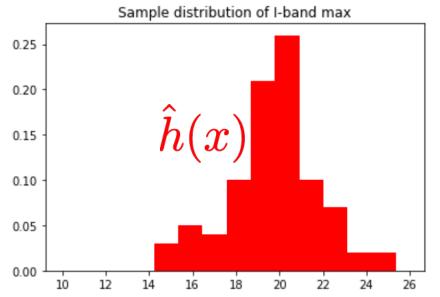
Normal CI:

$$[\hat{q} - z_{\alpha/2} \operatorname{std}e^B(q), \hat{q} + z_{\alpha 2} \operatorname{std}e^B(q)]$$

Percentile CI:

$$[q_{\alpha/2}^*, q_{1-\alpha/2}^*]$$





My data:

$$x = \{x_1, x_2, ..., x_n\}$$

$$J1 = \{x_1, ..., x_{n-2}, x_{n-1}\}$$
$$J2 = \{x_1, ..., x_{n-2}, x_n\}$$

 $Jn = \{x_2, ..., x_{n-1}, x_n\}$

With n-1 point per sample

Jacknife ≡ Remove 1 data point from your sample, n times

WARNING: statistics q_n from Jacknife is biased

$$\left[q^{J} = n \, q_{n} \, - \, (n-1) \, ar{q}_{n}
ight]_{ar{q}_{n} \, = \, n^{-1}} \sum_{i=1}^{n} \, q_{i}^{*} \left[\sigma_{q}^{2} \, = \, rac{n-1}{n} \sum_{i=1}^{n} (q_{i}^{*} \, - \, ar{q}_{n})^{2}
ight]_{i=1}^{n}$$