## Classical statistical inference

Regression and Model fitting

Associated notebook:

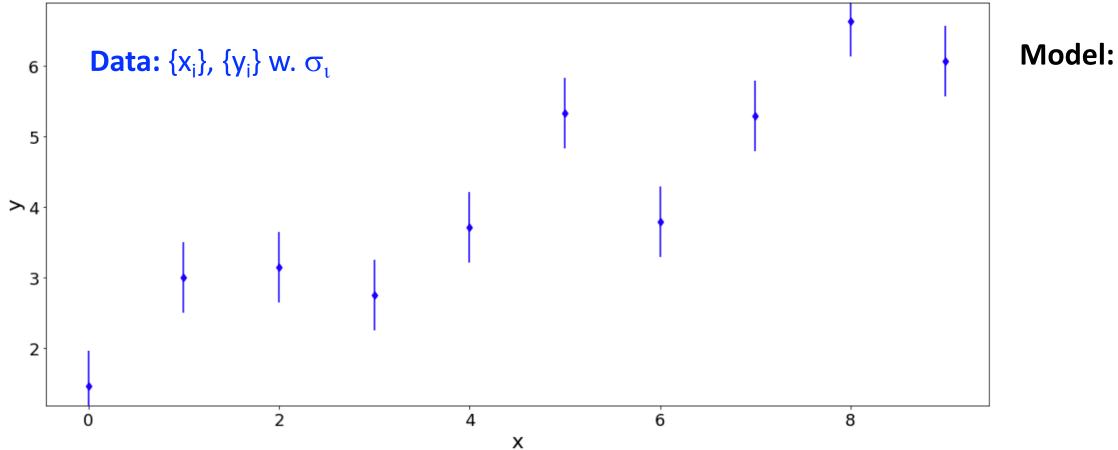
05-MLE\_and\_regression/Regression\_short.ipynb

**Problem:** the quantities of interest are parameters of a model, not the RV that you measure

#### Examples

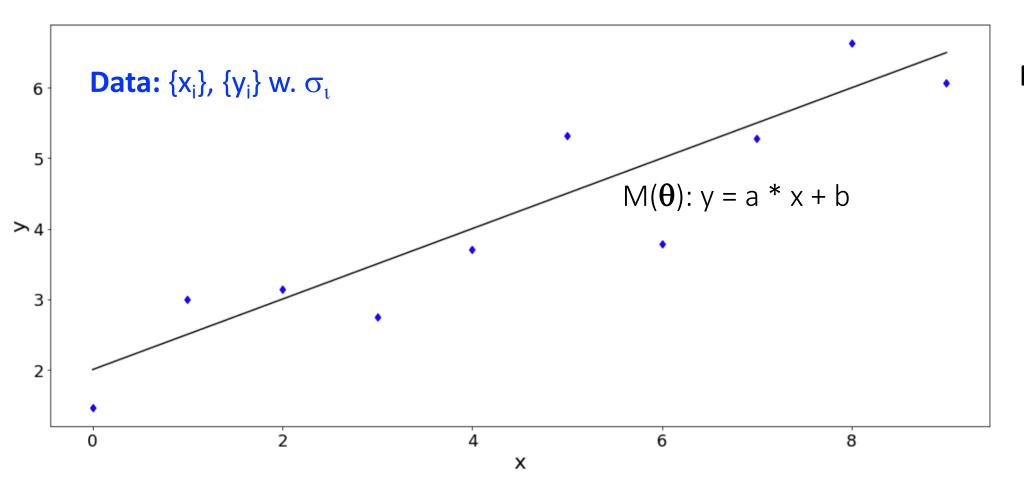
Observation	Quantity of interest	Model
Position of a star: x(t)	Proper motion (velocity) of the star	V = f(x, t,)
Photometry of an asteroid: mag(t)	P (period of rotation)	mag = f(t, P,)
Transit of a planet: mag(t)	P (period), e (eccentricity), D (dist to star)	$\Delta$ m = f (t, P, e, D,)
Spectrum of a QSO: F ( $\lambda$ )	M <sub>BH</sub> (Black hole mass of QSO)	$FWHM = f(M_{BH}, L,)$

**Problem:** You measure  $D \equiv (\{y_i\}, \{x_i\})$ 



Model:  $M(\theta)$ 

**Problem:** You measure  $D \equiv (\{y_i\}, \{x_i\})$ 



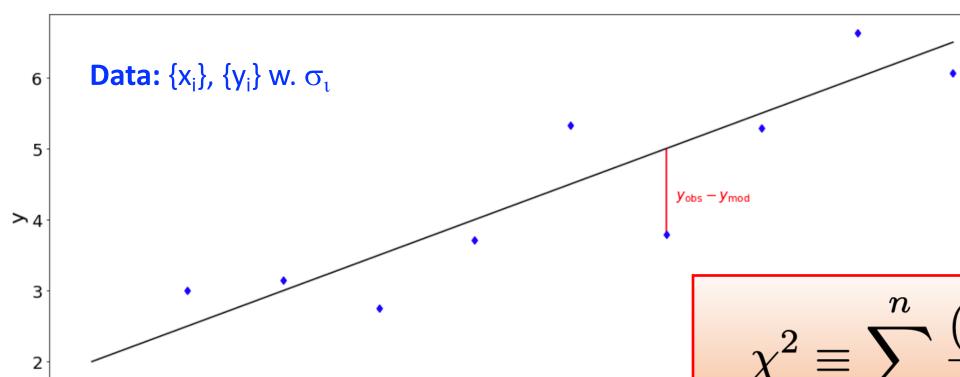
Model:  $M(\theta)$ 

$$y = a * x + b$$

$$\theta$$
 = a, b

#### How to find a good model?

2



Х

Model:  $M(\theta)$ 

$$y = a * x + b$$
  
=  $f(x | \theta)$ 

$$\theta$$
 = a, b

Minimize

$$\chi^2 \equiv \sum_{i=1}^n \frac{(y_i - y_{i,mod})^2}{\sigma_i^2}$$

$$\chi^2 \equiv \sum_{i=1}^n \frac{(y_i - y_{i,mod})^2}{\sigma_i^2}$$

If  $\sigma_i = 1$ : Least square regression

If  $\sigma_i \neq 1$ : chi-square regression

The  $\chi^2$  is called a **merit** function

When uncertainties between variables are correlated, the  $\chi^2$  is expressed:

$$\chi^{2} = \sum_{i=1}^{n} \sum_{l=1}^{n} (y_{i} - y_{i,mod}) F_{i,l} (y_{l} - y_{l,mod})$$

Where *F* is the Fisher matrix

$$[F]^{-1} = [C] = \begin{bmatrix} \sigma_k^2 & \sigma_{kl} \\ \sigma_{kl} & \sigma_l^2 \end{bmatrix}$$

# Link between $\chi^2$ and likelihood $L = p(D \,|\, M(oldsymbol{ heta}))$

$$L = p(D \mid M(\boldsymbol{\theta}))$$

Case of a straight line:  $y_i = heta_0 + heta_1 \, x_i + \epsilon_i$  with  $\epsilon_i \sim N(0, \sigma_i)$ 

For each y<sub>k</sub> we have: 
$$p(y_k \mid \mu, \sigma) = \frac{1}{\sigma \sqrt{2 \, \pi}} \, \exp \left[ -0.5 \left( \frac{y_k - \mu}{\sigma} \right)^2 \right]$$

Hence, we have for our data set D:

$$L \equiv p(\{y_i\} | \{x_i\}, \boldsymbol{\theta}) = \prod_{i=1}^{N} \frac{1}{\sigma_i \sqrt{2\pi}} \exp \left[ \left( \frac{-(y_i - (\theta_0 + \theta_1 x_i))^2}{2\sigma_i^2} \right) \right]$$

### Link between $\chi^2$ and likelihood

Hence, we have for our data set D:

$$L \equiv p(\{y_i\} | \{x_i\}, \boldsymbol{\theta}) = \prod_{i=1}^{N} \frac{1}{\sigma_i \sqrt{2\pi}} \exp \left[ \left( \frac{-(y_i - (\theta_0 + \theta_1 x_i))^2}{2\sigma_i^2} \right) \right]$$

$$\ln(L) \propto \sum_{i=1}^{N} \left( \frac{-(y_i - (\theta_0 + \theta_1 x_i))^2}{2 \sigma_i^2} \right) \qquad \chi^2 \equiv \sum_{i=1}^{n} \frac{(y_i - y_{i,mod})^2}{\sigma_i^2}$$

$$\chi^2 \equiv \sum_{i=1}^n \frac{(y_i - y_{i,mod})^2}{\sigma_i^2}$$



Minimizing  $\chi^2$  is equivalent to maximizing L

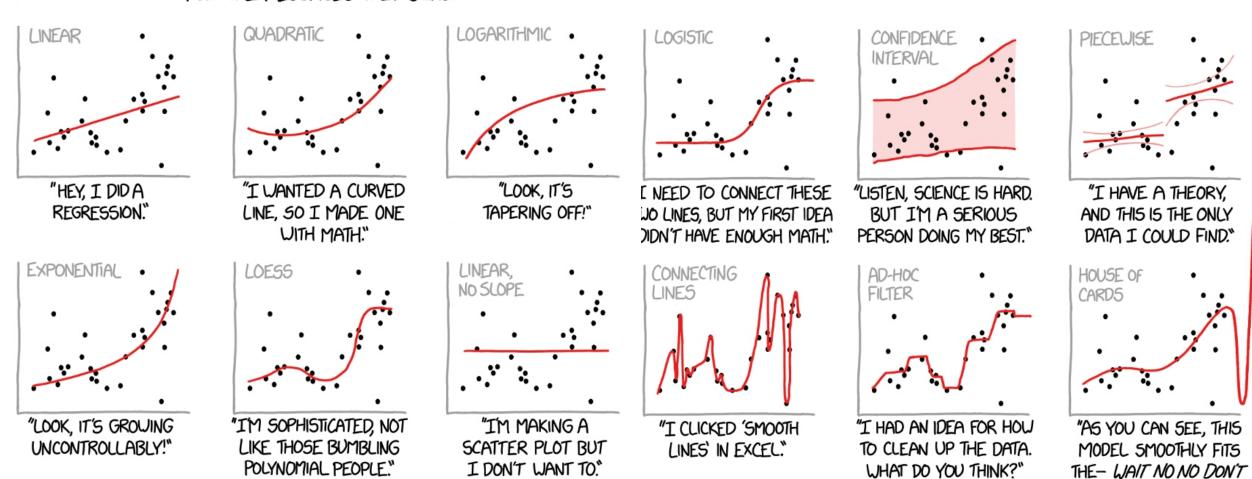
#### Regression of a straight line in python

Linear model fitting: See Sect. IV.1

Python implementation: numpy.polyfit(x, y, deg=1, w=1/sigma)

Go to Sect. IV.1.1 of the Notebook for practical example

CURVE-FITTING METHODS AND THE MESSAGES THEY SEND



Extend it aaaaaa!!"

https://xkcd.com/2048/

#### How to choose a suitable regression method?

The way you will tackle a regression problem may depend on:

- Linearity: is the model linear in its parameters?  $f(x \,|\, m{ heta}) = \sum_{p=1}^\infty \theta_p g_p(x)$
- Complexity: large number of parameters increase complexity and covariance matrix on uncertainties
- Error behaviour: uncertainties on dependent and independent variable and their correlation.

#### How to choose a suitable regression method?

Frequentist: (this lecture)	Bayes (future lecture):
Optimization with some merit function	Sampling of the likelihood
Search for <i>best</i> (fit) model <i>parameters</i>	PDF on parameters
Often when simple error behaviour	More <i>complex</i> error behaviour

#### Linear vs non linear regression

A model is **linear** if: 
$$f(x \mid \boldsymbol{\theta}) = \sum_{p=1}^k \theta_p g_p(x)$$

 $g_p(x)$  can be a non linear function of x BUT does not depend on any free parameter

In this case, the values of the parameters that yield  $\ \frac{\partial \ln(L)}{\partial \theta_i}=0$  (max. likelihood ) can be found "analytically"

When the model is **not linear**, the minimization of the  $\chi^2$  has to be performed *numerically* 

#### Linear vs non linear regression

Linear model fitting: See Sect. IV.1

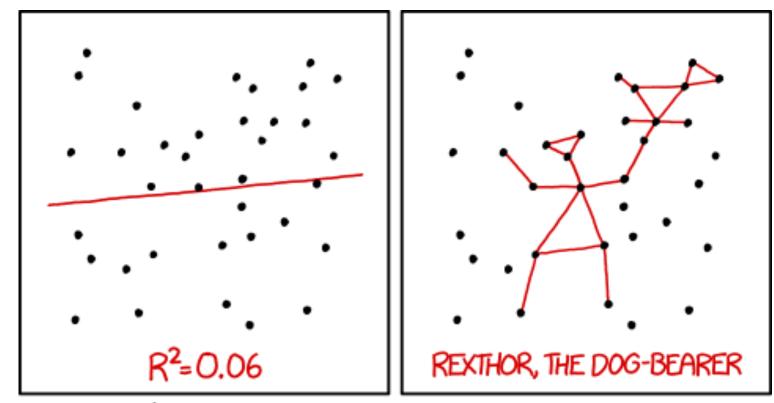
Python implementation: numpy.polyfit(x, y, deg=1, w=1/sigma)

NON Linear model fitting: See Sect. IV.3

Python implementation: scipy.optimize.curvefit()

Go to Sect. IV.1.1 of the Notebook for practical example

#### Quality of the regression



I DON'T TRUST LINEAR REGRESSIONS WHEN IT'S HARDER TO GUESS THE DIRECTION OF THE CORRELATION FROM THE SCATTER PLOT THAN TO FIND NEW CONSTELLATIONS ON IT.

#### Quality of the regression

Your  $\chi^2$  is a random variable!

$$Q = \sum_{i=1}^{k} z_i^2 \to p(Q|k) = \frac{1}{(2\Gamma(k/2))} (Q/2)^{k/2-1} \exp(-Q/2)$$

k = degree of freedom = N points - n parameters

If you fit a model with 2 parameters on a set of 100 points => 98 d.o.f.

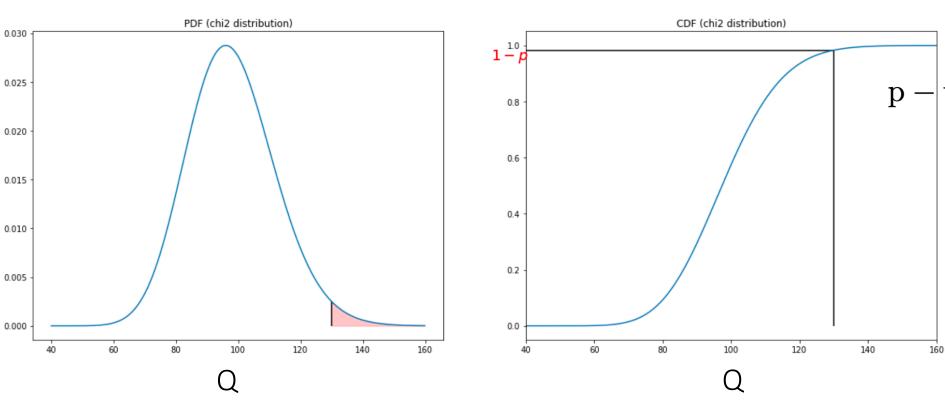
Expectation 
$$E(\chi^2) = 100 - 2 = 98$$

Reduced 
$$\chi^2: \chi^2_{\text{red}} = \chi^2_{\text{dof}} / \text{d.o.f.}$$
  $\Rightarrow$  Reduced  $\chi^2 \equiv 1$ . if good fit

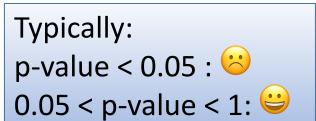
See also the Notebook 03-Basic\_statistics\_and\_proba\_concepts/Descriptive\_statistics\_02.ipynb

#### Quality of the regression

$$p(Q|k) = \frac{1}{(2\Gamma(k/2))} (Q/2)^{k/2-1} \exp(-Q/2)$$



p – value = p(Q 
$$\geq \chi^2_{\rm obs}$$
)  
= 1 –  $p(Q \leq \chi^2_{\rm obs})$ 



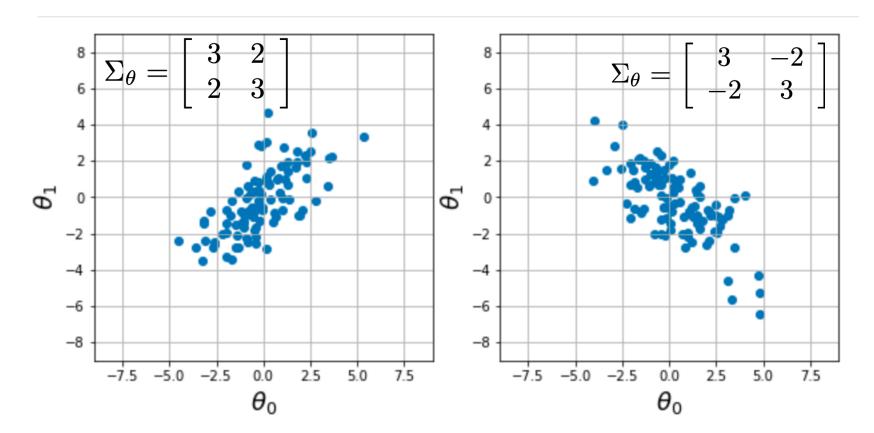
p-value close to 1:

1-scipy.stats.chi2.cdf(chi2\_data, df= len(data)-nparam)

#### Uncertainty on the fitted parameters

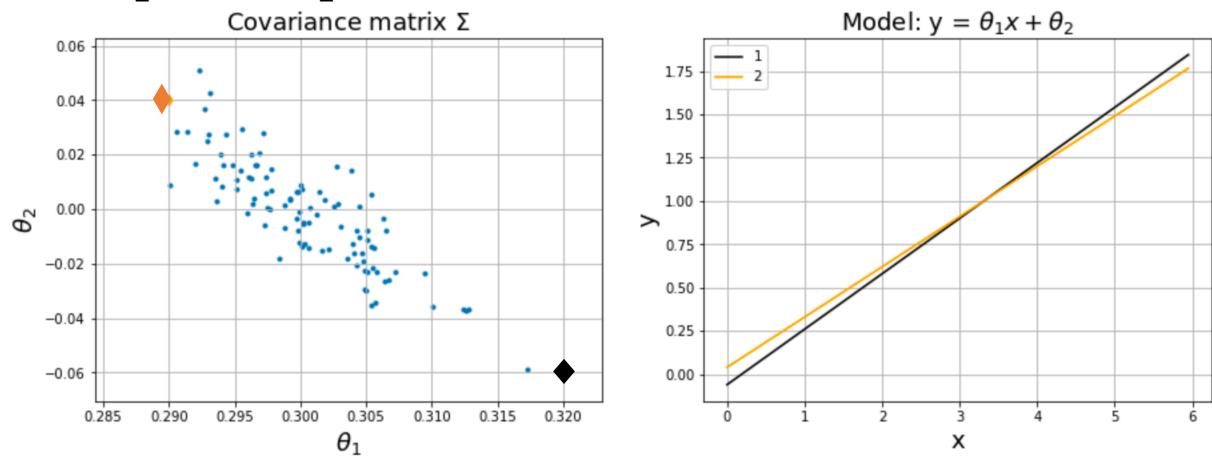
The python functions return a covariance matrix (Warning: use arg. cov=True)

The diagonal elements of the matrix give the variance on the parameters (uncertainty<sup>2</sup>)



#### Uncertainty on the fitted parameters

$$\Sigma_{ heta} = \left[ egin{array}{ccc} + & - \ - & + \end{array} 
ight]$$



#### Uncertainty on the fitted parameters

$$\Sigma_{ heta} = \left[ egin{array}{ccc} \sigma_{ heta_0}^2 & \sigma_{ heta_0 heta_1} \ \sigma_{ heta_0 heta_1} & \sigma_{ heta_1}^2 \end{array} 
ight]$$

