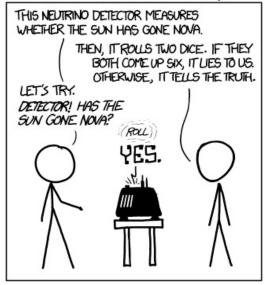
# Classical statistical inference

Part 2

Associated notebook:

04-Basic statistical inference frequentists 2/Frequentist inference 01.ipynb

## DID THE SUN JUST EXPLODE? (IT'S NIGHT, SO WE'RE NOT SURE.)



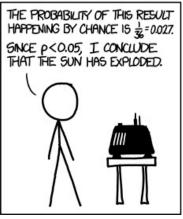
## What is inference?

Derive **INFORMATION** based on **DATA** Examples:

- Exoplanet transit  $\Rightarrow$  M, d  $\Rightarrow$  P(M | d)
- Supernovae distances ⇒ Expansion rate H<sub>0</sub>

Inference generally implies an underlying *statistical* model: PDF or regression laws with parameters  $\theta$ 

#### FREQUENTIST STATISTICIAN:



BAYESIAN STATISTICIAN:

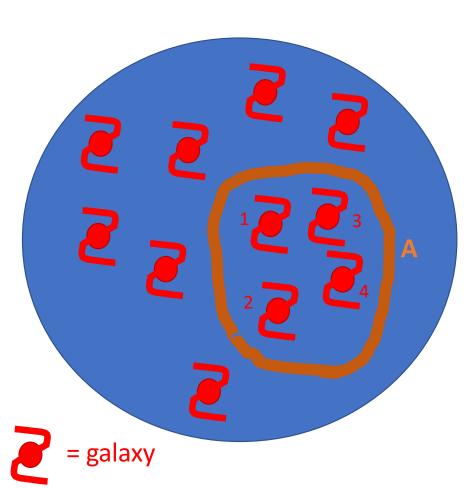


Three types of inference:

- Point estimation: "best"  $\theta$
- Confidence interval: Confidence around  $\theta$
- Hypothesis testing: data OK w. model?

https://xkcd.com/1132/

## Point estimate $\hat{\theta}$



### **Example:**

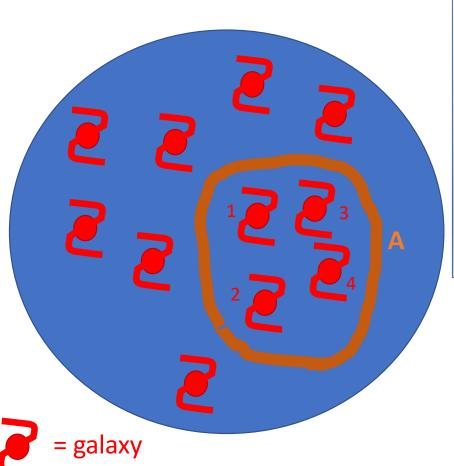
 $\theta$  = mean mag. of a population of galaxies

A: Your data set = subsample of measurements:  $A = \{X_1, X_2, X_3, X_4\}$  where X = mag. (this is a RV)

$$\hat{\theta} = \frac{X_1 + X_2 + X_3 + X_4}{4} \equiv \text{Point estimate of } \theta$$

If you do the experiment with another sample  $(\Rightarrow \text{ different } \frac{\text{realisation}}{\text{on }})$  you will get another  $\hat{\theta}$ 

## Point estimate $\hat{\theta}$

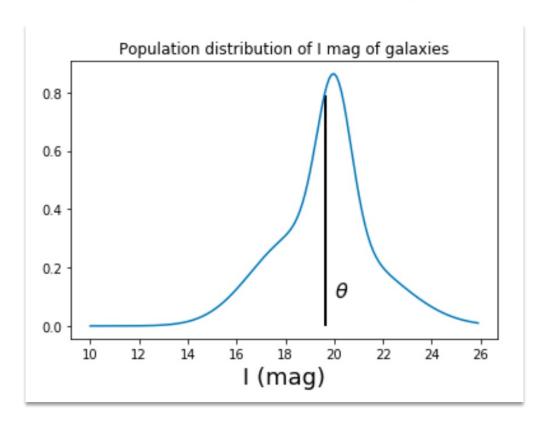


### **Generalisation:**

$$\hat{\theta} = g(X_1, X_2, X_3, \dots X_n)$$

- Point estimate of a param. is a *function* of RV X<sub>1</sub>, ...
- It is as well a Random Variable (RV)
- It can be biased, is characterized by a variance but should ideally be consistent (converges towards  $\theta$ )
- Distribution of  $\hat{\theta}$  is called sampling distribution

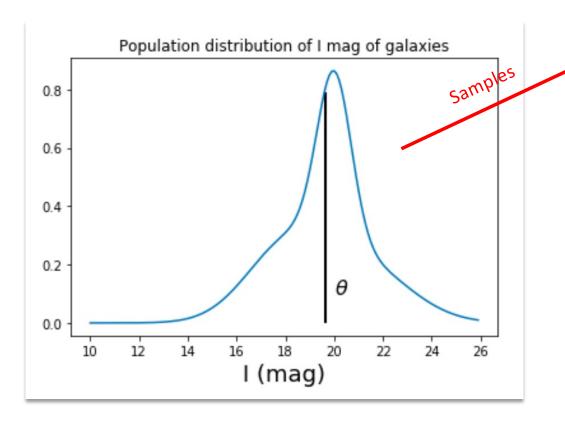
# Point estimate $\hat{\theta}$

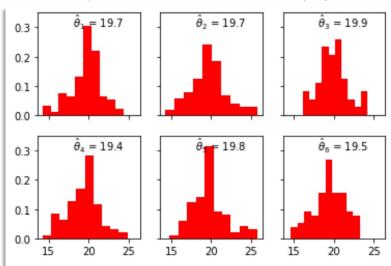


Population mean:  $\theta = 19.66$ 

### Six different *samples* drawn from the true population

# Point estimate $\hat{\theta}$

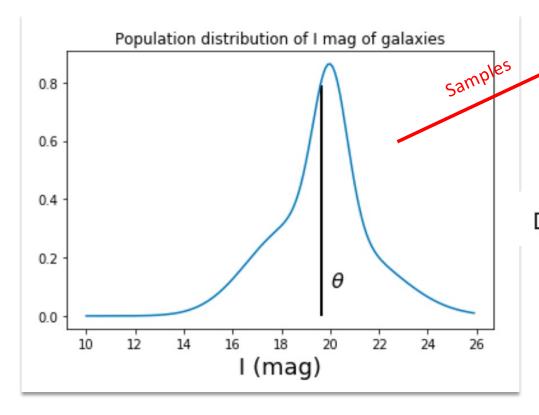




Population mean:  $\theta = 19.66$ 

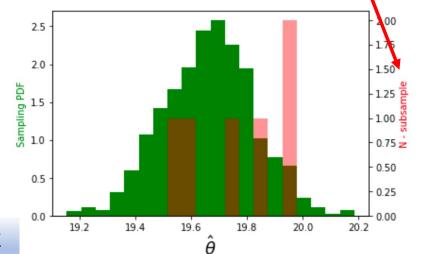
Six different samples drawn from the true population

# Point estimate $\hat{\theta}$



 $\hat{\theta}_{1} = 19.7$   $\hat{\theta}_{2} = 19.7$   $\hat{\theta}_{3} = 19.9$   $\hat{\theta}_{3} = 19.9$   $\hat{\theta}_{3} = 19.9$   $\hat{\theta}_{3} = 19.9$   $\hat{\theta}_{4} = 19.4$   $\hat{\theta}_{5} = 19.8$   $\hat{\theta}_{6} = 19.5$   $\hat{\theta}_{6} = 19.5$   $\hat{\theta}_{1} = 19.8$   $\hat{\theta}_{2} = 19.8$   $\hat{\theta}_{3} = 19.9$   $\hat{\theta}_{4} = 19.4$   $\hat{\theta}_{5} = 19.8$   $\hat{\theta}_{6} = 19.5$   $\hat{\theta}_{7} = 19.8$   $\hat{\theta}_{8} = 19.9$ 

Distribution of  $\{\hat{\theta}_1, \hat{\theta}_2, ..., \hat{\theta}_k\} = \text{sample distribution}$ 

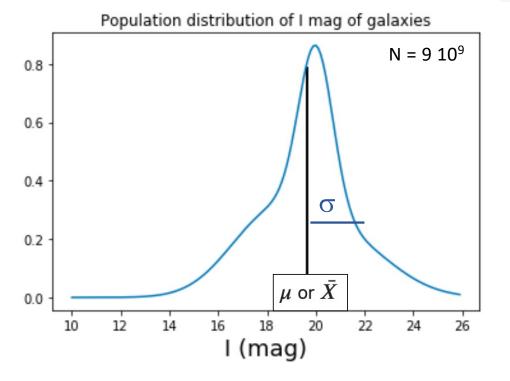


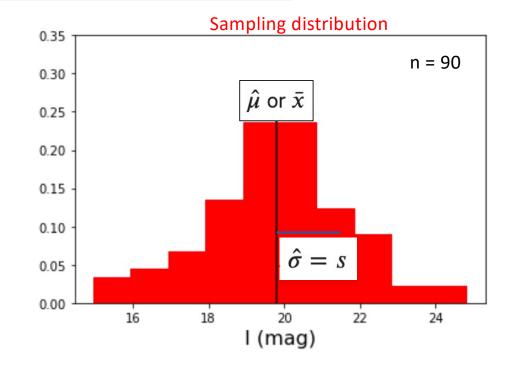
Population mean:  $\theta = 19.66$ 

Go to: Sect. II.1 of the notebook

# Summary statistics

Name	Population Statistics	Sample Statistics
size	N	n
mean	$\mu = \bar{X} = \frac{\sum_{i} X_{i}}{N}$	$\hat{\mu} = \bar{x} = \frac{\sum_{i} x_{i}}{n}$
Variance	$\sigma^2 = \frac{\sum_i (X_i - \bar{X})^2}{N}$	$s^2 = \frac{\sum_i (x_i - \bar{x})^2}{n - 1}$
Standard deviation	$\sigma = \sqrt{\sigma^2}$	$\hat{\sigma} = s = \sqrt{s^2}$





# Summary (sample) statistics: standard error

Standard error (stde) ≠ Standard deviation (std)

Name	Formula
Standard error on the mean	$stde(\bar{x}) = \frac{s}{\sqrt{n}}$
Standard error on the sample stdev	$stde(s) = s/\sqrt{2(n-1)}$
Standard error on proportions	$stde(\hat{p}) = \frac{\sqrt{p(1-p)}}{\sqrt{n}}$
Standard error on $p^{th}$ percentile ( $h_p = PDF$ at $p^{th}$ percentile )	$stde(q^p) = \frac{1}{h_p} \frac{\sqrt{p(1-p)}}{\sqrt{n}}$

### Central limit theorem

When independent random variables are added, their sum tends towards a normal distribution (if n >>)

- This is true even if the original RV are not normally distributed
- Sampling dist. of mean tends (for large *n* ) towards a Normal distribution
- Sampling dist. of variance (for large *n*) does NOT tend towards a Normal distribution

Go to: Sect. II.1.1. of the notebook

# Distribution of sample quantities

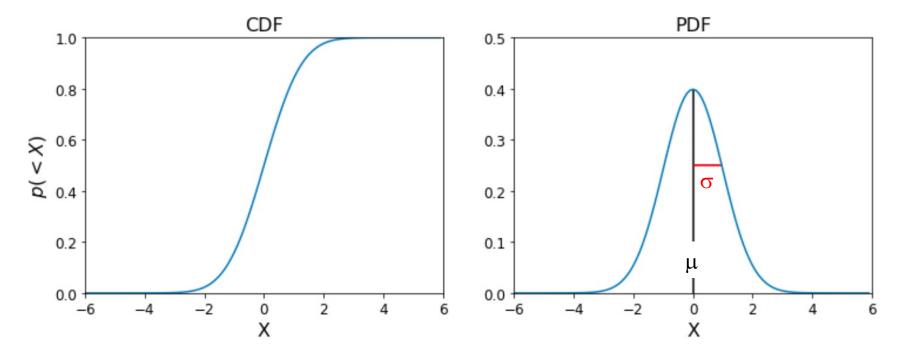
- Sample distribution of  $\hat{\mu} \sim N(\mu, \sigma/\sqrt{n}) \iff Z = \frac{\mu \hat{\mu}}{(\sigma/\sqrt{n})} \sim N(0, 1)$
- Sample distribution of  $t = \frac{\mu \hat{\mu}}{(s/\sqrt{n})} \sim t(n-1)$  s is derived from the sample dist.
- Sample distribution of  $S = \frac{(n-1)s^2}{\sigma^2} \sim \chi^2(n-1).$  Chi square distribution

# Classical statistical inference: confidence interval

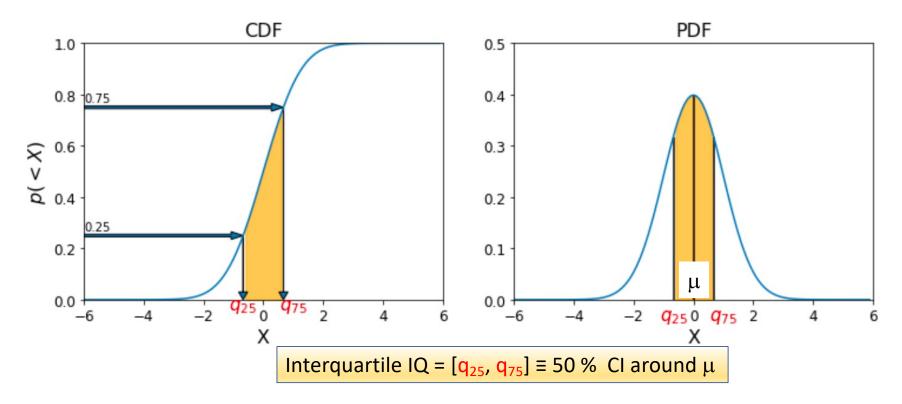
Part 3

Associated notebook:

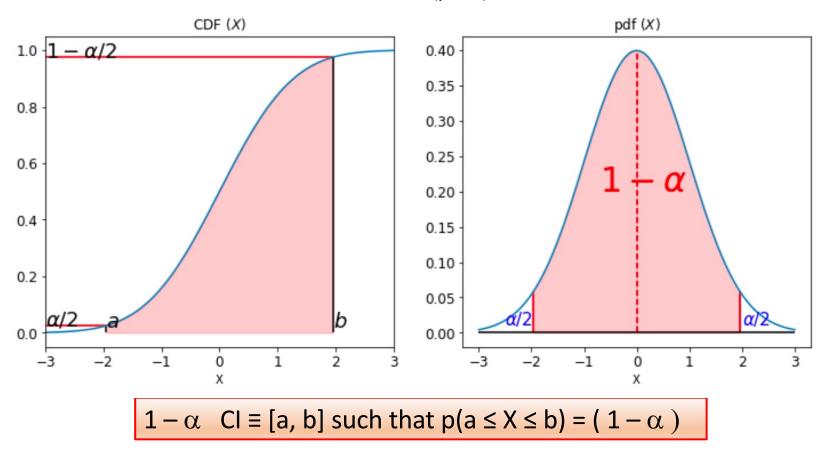
04-Basic statistical inference frequentists 2/Frequentist inference 02.ipynb

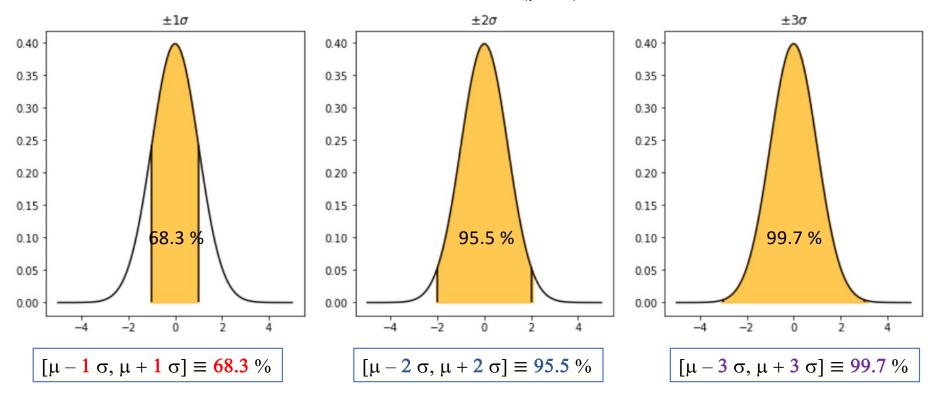


If RV  $\{X_i\}$  whose population is distributed as  $N(\mu, \sigma)$ 

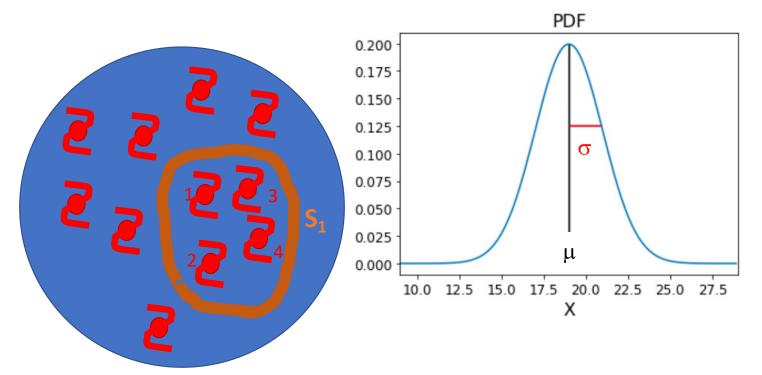


 $1-\alpha$  CI = [a, b] such that p(a  $\leq$  X  $\leq$  b) = (1- $\alpha$ )



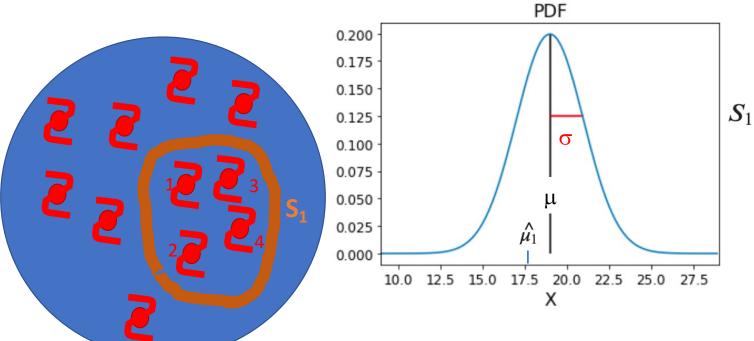


## Confidence interval from an estimator



$$S_1 = \{x_1, x_2, x_3, x_4\}$$

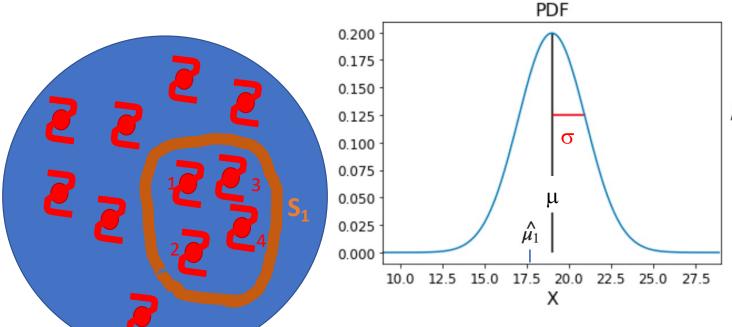
## Confidence interval from an estimator



$$S_1 = \{x_1, x_2, x_3, x_4\}$$
  
 $S_1 = \{19.45, 16.20, 16.43, 19.10\}$   
 $\hat{\mu}_1 = 17.8$ 

### Confidence interval from an estimator

If RV  $\{X_i\}$  whose population is distributed as  $N(\mu, \sigma)$ 



$$S_1 = \{x_1, x_2, x_3, x_4\}$$

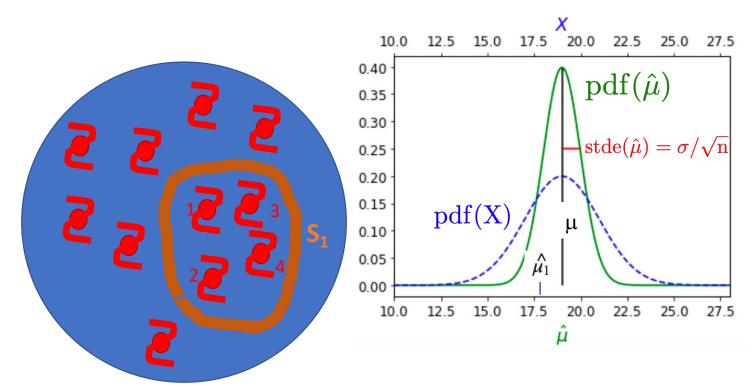
$$S_1 = \{19.45, 16.20, 16.43, 19.10\}$$
  
$$\hat{\mu}_1 = 17.8$$

95.5 % CI on sample mean?

We need to look at the distribution of the sample mean

95.5 % Cl on  $\hat{\mu}$  ?

If RV  $\{X_i\}$  whose population is distributed as  $N(\mu, \sigma)$ 



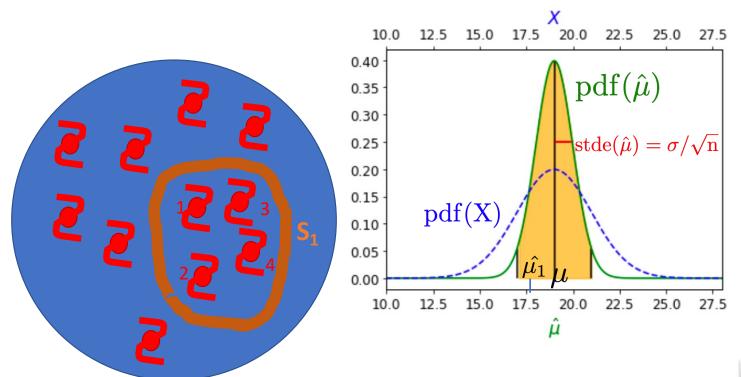
**Case 1:** σ is known

$$\hat{\mu} \sim N(\mu, \sigma/\sqrt{n})$$

$$stde(\hat{\mu}) = \sigma/\sqrt{n}$$

95.5 % Cl on  $\hat{\mu}$  ?

If RV  $\{X_i\}$  whose population is distributed as  $N(\mu, \sigma)$ 



Case 1:  $\sigma$  is known

$$\hat{\mu} \sim N(\mu, \sigma/\sqrt{n})$$

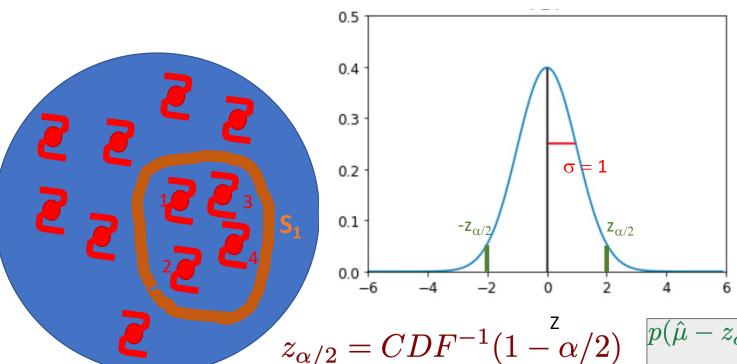
stde
$$(\hat{\mu}) = \sigma/\sqrt{n}$$
  
 $p(\hat{\mu}_1 \in [\mu \pm 2 \text{ stde}]) = 0.955$ 



$$p(\mu \in [\hat{\mu}_1 \pm 2 \text{ stde}]) = 0.955$$

95.5 % Cl on  $\mu$  ?

If RV  $\{X_i\}$  whose population is distributed as  $N(\mu, \sigma)$ 



**Case 1:** σ is known

$$stde(\hat{\mu}) = \sigma/\sqrt{n}$$

$$Z=rac{\mu-\hat{\mu}}{\sigma/\sqrt{ ext{n}}}\sim N(0,1)$$

$$p(-z_{\alpha/2} < Z < z_{\alpha/2}) = \alpha * 100 \%$$

p (-
$$\mathbf{z}_{\alpha/2}$$
 \* stde <  $\mu - \hat{\mu}$  <  $\mathbf{z}_{\alpha/2}$  \*stde)

$$= \alpha * 100 \%$$

$$p(\hat{\mu} - z_{\alpha/2} \text{ stde} < \mu < \hat{\mu} + z_{\alpha/2} \text{ stde})$$
$$= \alpha \times 100\%$$

$$CI \equiv [\hat{\mu} \pm m \times stde]$$

For 95.5 % CI, 
$$z_{\alpha/2} = m = 2$$

If RV  $\{X_i\}$  whose population is distributed as  $N(\mu, \sigma)$ 

12.5 15.0 17.5 20.0 22.5 25.0 27.5 0.40  $\mathrm{pdf}(\hat{\mu})$ 0.35 0.30 0.25 0.20 0.15 pdf(X)0.10 0.05 0.00 12.5 15.0 17.5 20.0 22.5 25.0 27.5 û

95.5 % Cl on  $\mu$  ?

Case 1:  $\sigma$  is known

$$stde(\hat{\mu}) = \sigma/\sqrt{n}$$

$$Z = rac{\mu - \hat{\mu}}{\sigma / \sqrt{\mathrm{n}}} \sim N(0, 1)$$

Case 2: σ is unknown

$$stde(\hat{\mu}) = s/\sqrt{n}$$

$$t = \frac{\mu - \hat{\mu}}{s/\sqrt{n}} \sim t(n-1)$$

 $CI \equiv [\hat{\mu} \pm m \times stde]$ 

Go to: Student\_vs\_Gauss.ipynb

### Confidence interval

### "Generic" strategy:

- Look at the PDF of a normalised/scaled estimator of known distribution.
- Define region around your (normalised) estimator (of know the distribution) that encloses (1  $\alpha$ ) X 100 % of the area under the PDF.
- For CI around the mean:

$$CI_{\alpha} = [\mu - q_{\alpha/2} \text{ stde}, \mu + q_{\alpha/2} \text{ stde}]$$

$$q_{\alpha/2} = CDF^{-1}(1 - \alpha/2)$$

For CI on other statistics (e.g. variance, difference between 2 means, proportions): see <u>Frequentist inference 03.ipynb</u> (Material in <u>Frequentist inference 03.ipynb</u> is supplementary material)

Go to: Sect. II.2. of the notebooks