Exponential Random Graph Models for Social Networks

ERGM Introduction

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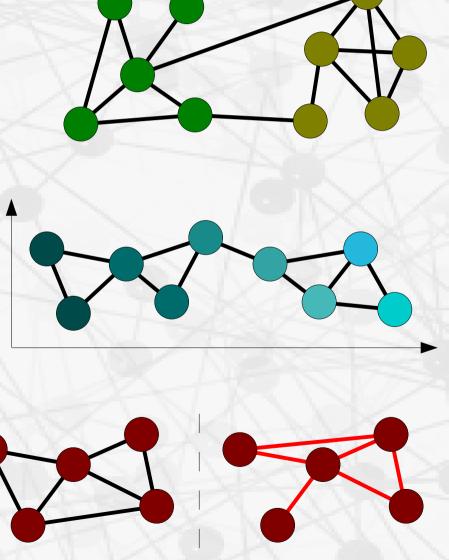
2011 Political Networks Conference June 15, 2011 Ann Arbor, MI http://polnet2011.statnet.org/

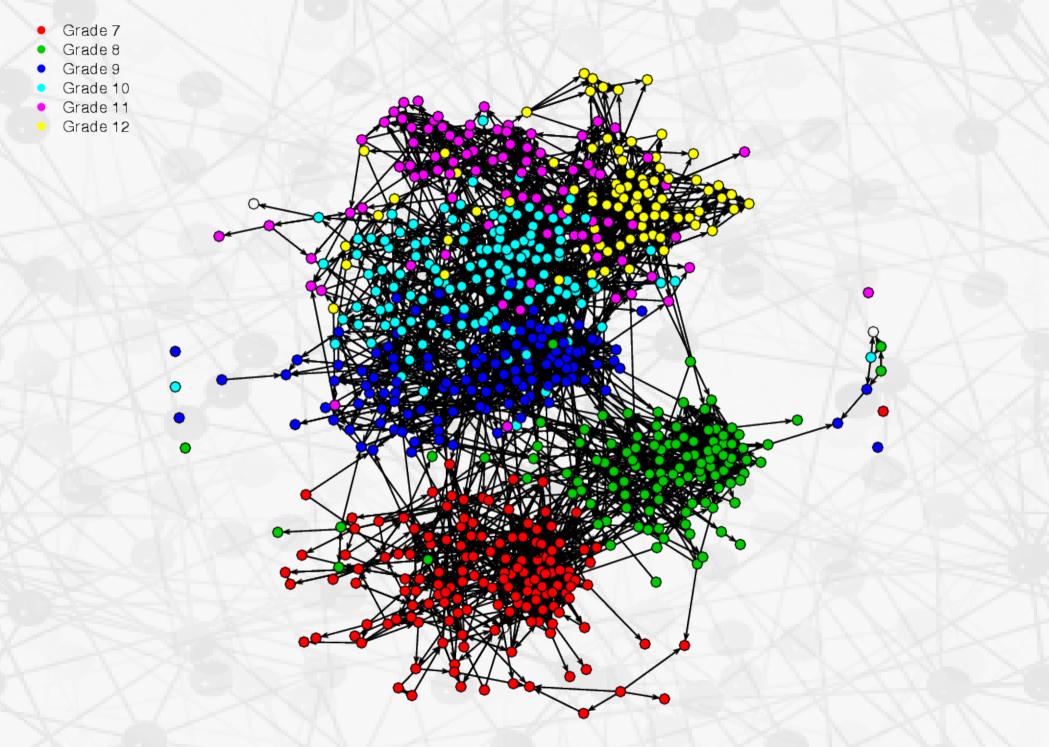
From Description to Modeling

- Ultimately, want to do more than describe networks
- Network modeling: predict the formation and structure of social networks
- Many examples
 - Conditional uniform graphs, Bernoulli graphs
 - Holland and Leinhardt's p_1
 - Degree distribution models, growth models, etc.
- ERGM: a general representation for such models
 - Draws on theory of statistical exponential families
 - Not really a "type" of model (in a scientific sense), but a way of representing and working with new and existing models!

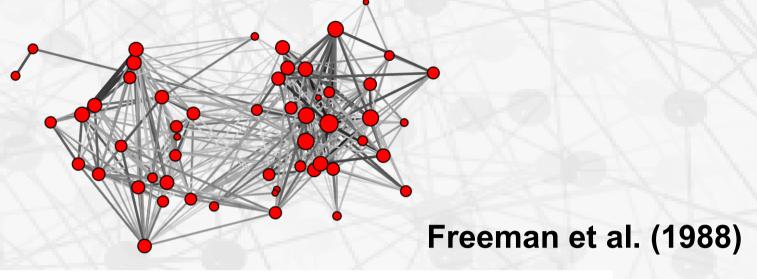
Initial Intuition: Factors in Tie Formation

- All ties are not equally probable
 - Chance of an (i,j) edge may depend on properties of i and j
 - Can also depend on other (i,j) relationships
- Some examples:
 - Homophily
 - Propinquity
 - Multiplexity





AddHealth Friendship Network, by Grade



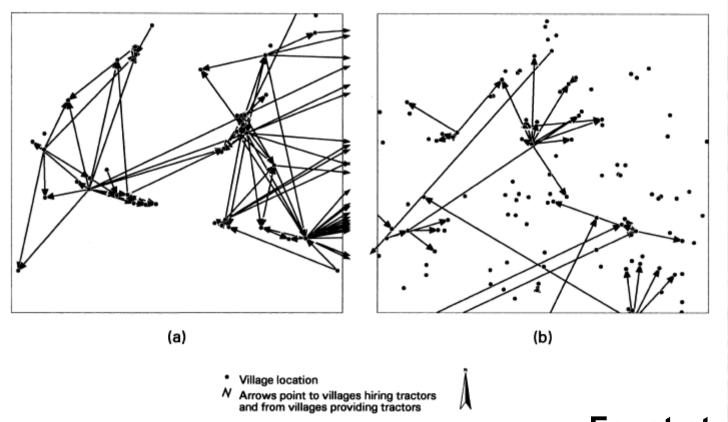
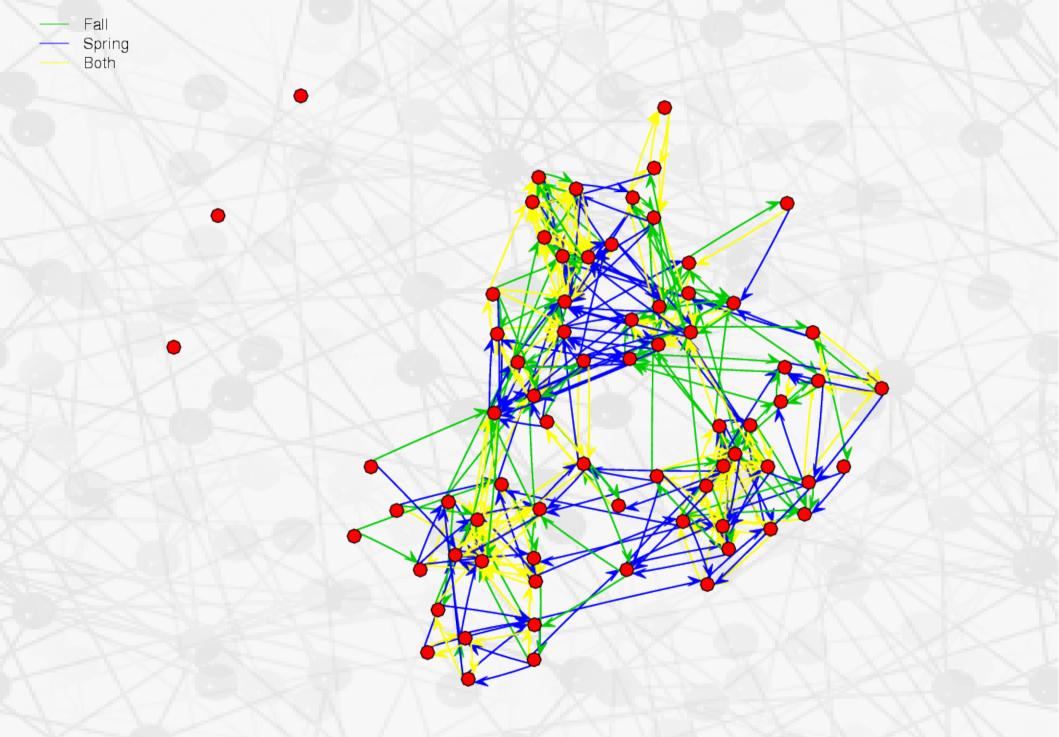


Fig. 3. Tractor hiring network in two regions of Nang Rong.

Faust et al. (1999)

June 1. Julio, 1 JINet2011, 06/15/11



Boy's School Friendship Network (Coleman, 1964)

Carter T. Butts, PolNet2011, 06/15/11

Logistic Network Regression

- A classic starting point: why not treat edges as independent, w/log-odds as a linear function of covariates?
 - Special case of standard logistic regression
 - Dependent variable is a network adjacency matrix
- Model form:

$$\log \left(\frac{\Pr(\mathbf{Y_{ij}} = 1)}{\Pr(\mathbf{Y_{ij}} = 0)} \right) = \theta_1 \mathbf{X_{ij1}} + \theta_2 \mathbf{X_{ij2}} + \dots + \theta_m \mathbf{X_{ijm}} = \theta^T \mathbf{X_{ij}}$$

- where \mathbf{Y}_{ij} is the value of the edge from i to j on the dependent relation, \mathbf{X}_{ijk} is the value of the kth predictor for the (i,j) ordered pair, and $\theta_{1},...,\theta_{m}$ are coefficients
 - $\log(p/(1-p)) = \log(p)$, maps (0,1) to $(-\infty,\infty)$

Moving Beyond the Logistic Case

- The logistic model can be quite powerful, but still very limiting
 - No way to model conditional dependence among edges
 - E.g., true triad closure bias, reciprocity
 - Cannot handle exotic support constraints
 - What if your network must be transitive (e.g., sports contests, entailments), an interval graph (e.g., life history graphs), etc?
- A more general framework: discrete exponential families
 - Very general way of representing discrete distributions
 - Turns up frequently in statistics, physics, etc.

Exponential Families for Random Graphs (w/Covariates)

• For random graph G w/countable support G and covariate set X, pmf can be written in ERG form:

$$\Pr[G = g | \mathbf{t}, \theta, \mathcal{G}, X] = \frac{\exp[\theta^T \mathbf{t}(g, X)]}{\sum_{g' \in \mathcal{G}} \exp[\theta^T \mathbf{t}(g', X)]} I_{\mathcal{G}}(g)$$

- θ^{T} t: linear predictor
 - $t: \mathcal{G} \rightarrow \mathbb{R}^m$: vector of sufficient statistics
 - $\theta \in \mathbb{R}^m$: vector of parameters
 - $\sum_{g' \in \mathcal{G}} \exp(\theta^{T} \mathbf{t}(g', X))$: normalizing factor
- Intuition: ERG placed more/less weight on structures with certain features, as determined by t, θ
 - Framework is complete for pmfs on G, few constraints on t

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Source of great difficulty!

- Intuition: ERG placed more/less weight on structures with certain features, as determined by t, θ
 - Framework is complete for pmfs on G, few constraints on t

Equivalent Expression: Modeling the Adjacency Matrix

• For adjacency matrix Y w/countable support \mathcal{Y} and covariate set X, any pmf can be written in ERG form:

$$\Pr\left(\mathbf{Y} = \mathbf{y} \middle| \mathbf{t}, \theta, \mathcal{Y}, X\right) = \frac{\exp\left(\theta^{T} \mathbf{t}(\mathbf{y}, X)\right)}{\sum_{\mathbf{y}' \in \mathcal{Y}} \exp\left(\theta^{T} \mathbf{t}(\mathbf{y}', X)\right)} I_{\mathcal{Y}}(\mathbf{y})$$

- θ^{T} t: linear predictor
 - $t: y \rightarrow \mathbb{R}^m$: vector of sufficient statistics
 - $\theta \in \mathbb{R}^m$: vector of parameters
 - $\sum_{\mathbf{y'} \in y} \exp(\theta^{\mathrm{T}} \mathbf{t}(\mathbf{y'}, X))$: normalizing factor
- Additional notation:
 - $\mathbf{y}_{ij}^+, \mathbf{y}_{ij}^-$: y w/ijth cell set to 1 or 0 (respectively)
 - y_{ij}^c : all elements of y other than the ijth

ERGs and Conditional Odds of an Edge

Can easily specify the conditional odds of an edge:

$$\frac{\Pr(\mathbf{Y} = \mathbf{y}_{ij}^{+} | \mathbf{t}, \theta, \mathcal{Y}, X)}{\Pr(\mathbf{Y} = \mathbf{y}_{ij}^{-} | \mathbf{t}, \theta, \mathcal{Y}, X)} = \frac{\exp(\theta^{T} \mathbf{t}(\mathbf{y}_{ij}^{+}, X))}{\sum_{\mathbf{y}' \in \mathcal{Y}} \exp(\theta^{T} \mathbf{t}(\mathbf{y}', X))} \frac{\sum_{\mathbf{y}' \in \mathcal{Y}} \exp(\theta^{T} \mathbf{t}(\mathbf{y}', X))}{\exp(\theta^{T} \mathbf{t}(\mathbf{y}', X))}$$

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= \frac{\exp(\theta^{T} \mathbf{t}(\mathbf{y}_{ij}^{+}, X))}{\exp(\theta^{T} \mathbf{t}(\mathbf{y}_{ij}^{+}, X))} = \exp(\theta^{T} \mathbf{t}(\mathbf{y}_{ij}^{+}, X) - \mathbf{t}(\mathbf{y}_{ij}^{-}, X))$$

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$$= \frac{\exp(\theta^{T} \mathbf{t}(\mathbf{y}_{ij}^{+}, X))}{\exp(\theta^{T} \mathbf{t}(\mathbf{y}_{ij}^{-}, X))} = \exp(\theta^{T} \mathbf{t}(\mathbf{y}_{ij}^{+}, X) - \mathbf{t}(\mathbf{y}_{ij}^{-}, X))$$

- Log-odds depend on the changescore, $\Delta_{ij} = \mathbf{t}(\mathbf{y}_{ij}^+, X) \mathbf{t}(\mathbf{y}_{ij}^-, X)$
- Useful implication: each unit change in $\mathbf{t_k}$ for (i,j) edge present (versus absent) increases the conditional log-odds of (i,j) by $\boldsymbol{\theta_k}$
 - Important: this is only conditionally true! The marginal log-odds of an (i,j) edge can depend on a complex way on other aspects of the graph

$$\Pr\left(\mathbf{Y}_{ij}=1\middle|\mathbf{Y}_{ij}^{c}=\mathbf{y}_{ij}^{c},\mathbf{t},\theta,\mathcal{Y},X\right) = \Pr\left(\mathbf{Y}=\mathbf{y}_{ij}^{+}\middle|\mathbf{Y}_{ij}^{c}=\mathbf{y}_{ij}^{c},\mathbf{t},\theta,\mathcal{Y},X\right)$$

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$$= \frac{\Pr(\mathbf{Y}=\mathbf{y}_{ij}^{+}\big|\mathbf{t},\theta,\mathcal{Y},X)}{\Pr(\mathbf{Y}=\mathbf{y}_{ij}^{-}\big|\mathbf{t},\theta,\mathcal{Y},X)} \left[1 + \frac{\Pr(\mathbf{Y}=\mathbf{y}_{ij}^{+}\big|\mathbf{t},\theta,\mathcal{Y},X)}{\Pr(\mathbf{Y}=\mathbf{y}_{ij}^{-}\big|\mathbf{t},\theta,\mathcal{Y},X)}\right]^{-1}$$

$$= \frac{\exp(\theta^{T}(\mathbf{t}(\mathbf{y}_{ij}^{+},X) - \mathbf{t}(\mathbf{y}_{ij}^{-},X)))}{1 + \exp(\theta^{T}(\mathbf{t}(\mathbf{y}_{ij}^{+},X) - \mathbf{t}(\mathbf{y}_{ij}^{-},X)))}$$

$$Pr(\mathbf{Y}_{ij} = \mathbf{Y}_{ij}^{c} = \mathbf{y}_{ij}^{c}, \mathbf{t}, \theta, \mathcal{Y}, X) = Pr(\mathbf{Y} = \mathbf{y}_{ij}^{+} | \mathbf{Y}_{ij}^{c} = \mathbf{y}_{ij}^{c}, \mathbf{t}, \theta, \mathcal{Y}, X)$$

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$$= \frac{\exp(\theta^{T}(\mathbf{t}(\mathbf{y}_{ij}^{+}, X) - \mathbf{t}(\mathbf{y}_{ij}^{-}, X)))}{\mathbf{1} + \exp(\theta^{T}(\mathbf{t}(\mathbf{y}_{ij}^{+}, X) - \mathbf{t}(\mathbf{y}_{ij}^{-}, X)))}$$

$$= [\mathbf{1} + \exp(\theta^{T}(\mathbf{t}(\mathbf{y}_{ij}^{-}, X) - \mathbf{t}(\mathbf{y}_{ij}^{+}, X)))]^{-1}$$

$$\begin{aligned} \Pr\left(\mathbf{Y}_{ij} = 1 \middle| \mathbf{Y}_{ij}^{c} = \mathbf{y}_{ij}^{c}, \mathbf{t}, \theta, \mathcal{Y}, X\right) &= \Pr\left(\mathbf{Y} = \mathbf{y}_{ij}^{+} \middle| \mathbf{Y}_{ij}^{c} = \mathbf{y}_{ij}^{c}, \mathbf{t}, \theta, \mathcal{Y}, X\right) \\ &= \frac{\Pr\left(\mathbf{Y} = \mathbf{y}_{ij}^{+} \middle| \mathbf{t}, \theta, \mathcal{Y}, X\right)}{\Pr\left(\mathbf{Y} = \mathbf{y}_{ij}^{-} \middle| \mathbf{t}, \theta, \mathcal{Y}, X\right)} \left[1 + \frac{\Pr\left(\mathbf{Y} = \mathbf{y}_{ij}^{+} \middle| \mathbf{t}, \theta, \mathcal{Y}, X\right)}{\Pr\left(\mathbf{Y} = \mathbf{y}_{ij}^{-} \middle| \mathbf{t}, \theta, \mathcal{Y}, X\right)}\right]^{-1} \\ &= \frac{\exp\left(\theta^{T}\left(\mathbf{t}(\mathbf{y}_{ij}^{+}, X) - \mathbf{t}(\mathbf{y}_{ij}^{-}, X)\right)\right)}{1 + \exp\left(\theta^{T}\left(\mathbf{t}(\mathbf{y}_{ij}^{+}, X) - \mathbf{t}(\mathbf{y}_{ij}^{-}, X)\right)\right)} \\ &= \left[1 + \exp\left(\theta^{T}\left(\mathbf{t}(\mathbf{y}_{ij}^{+}, X) - \mathbf{t}(\mathbf{y}_{ij}^{-}, X)\right)\right)\right]^{-1} \\ &= \log i t^{-1}\left(\theta^{T}\left(\mathbf{t}(\mathbf{y}_{ij}^{+}, X) - \mathbf{t}(\mathbf{y}_{ij}^{-}, X)\right)\right) \end{aligned}$$

$$\begin{aligned} \Pr\left(\mathbf{Y}_{\mathbf{i}\mathbf{j}} = 1 \middle| \mathbf{Y}_{\mathbf{i}\mathbf{j}}^{\mathbf{c}} = \mathbf{y}_{\mathbf{i}\mathbf{j}}^{\mathbf{c}}, \mathbf{t}, \theta, \mathcal{Y}, X\right) &= \Pr\left(\mathbf{Y} = \mathbf{y}_{\mathbf{i}\mathbf{j}}^{+} \middle| \mathbf{Y}_{\mathbf{i}\mathbf{j}}^{\mathbf{c}} = \mathbf{y}_{\mathbf{i}\mathbf{j}}^{\mathbf{c}}, \mathbf{t}, \theta, \mathcal{Y}, X\right) \\ &= \frac{\Pr\left(\mathbf{Y} = \mathbf{y}_{\mathbf{i}\mathbf{j}}^{+} \middle| \mathbf{t}, \theta, \mathcal{Y}, X\right)}{\Pr\left(\mathbf{Y} = \mathbf{y}_{\mathbf{i}\mathbf{j}}^{-} \middle| \mathbf{t}, \theta, \mathcal{Y}, X\right)} \left[1 + \frac{\Pr\left(\mathbf{Y} = \mathbf{y}_{\mathbf{i}\mathbf{j}}^{+} \middle| \mathbf{t}, \theta, \mathcal{Y}, X\right)}{\Pr\left(\mathbf{Y} = \mathbf{y}_{\mathbf{i}\mathbf{j}}^{-} \middle| \mathbf{t}, \theta, \mathcal{Y}, X\right)} \right]^{-1} \\ &= \frac{\exp\left(\theta^{T}\left(\mathbf{t}\left(\mathbf{y}_{\mathbf{i}\mathbf{j}}^{+}, X\right) - \mathbf{t}\left(\mathbf{y}_{\mathbf{i}\mathbf{j}}^{-}, X\right)\right)\right)}{1 + \exp\left(\theta^{T}\left(\mathbf{t}\left(\mathbf{y}_{\mathbf{i}\mathbf{j}}^{+}, X\right) - \mathbf{t}\left(\mathbf{y}_{\mathbf{i}\mathbf{j}}^{+}, X\right)\right)\right)}^{-1} \\ &= \left[1 + \exp\left(\theta^{T}\left(\mathbf{t}\left(\mathbf{y}_{\mathbf{i}\mathbf{j}}^{+}, X\right) - \mathbf{t}\left(\mathbf{y}_{\mathbf{i}\mathbf{j}}^{+}, X\right)\right)\right)\right]^{-1} \\ &= \log i t^{-1} \left(\theta^{T}\left(\mathbf{t}\left(\mathbf{y}_{\mathbf{i}\mathbf{j}}^{+}, X\right) - \mathbf{t}\left(\mathbf{y}_{\mathbf{i}\mathbf{j}}^{-}, X\right)\right)\right) \end{aligned}$$

Conditional Edge Probability, Cont.

- So, the conditional probability of an (i,j) edge is simply the inverse logit of $\theta^T \Delta_{ij}$
 - Obvious idea: to find θ , why not set this up as a logistic network regression problem (regressing y on Δ)?
 - This is an "autologistic regression," and the resulting estimator is known as a pseudolikelihood estimator (Besag, 1975)
 - Problem: the probability here is only conditional can use for any one ij, but joint likelihood of y is not generally the product of $Pr(Y_{ij}=y_{ij}|Y_{ij}^c=y_{ij}^c)$
 - Another view: y appears on both sides can't regress w/out accounting for the "feedback" (i.e., dependence) among edges
 - Does work iff edges are independent i.e., the logistic case
- Still, useful aid in interpretation
 - Can consider probability of \mathbf{y}_{ij} under various scenarios to understand local model behavior

Fitting ERGs to Data

- After positing a model, generally want to estimate parameters from data
 - Benefit of the framework: standardized inferential framework
- Most common current method: maximum likelihood estimation
 - Find θ^* that maximizes $Pr(Y=y_{obs}|t,\theta^*,\mathcal{Y})$
 - Exists for regular model so long as observed data is "non-extreme" (in a specific sense to be discussed later); always unique
 - Also has first order interpretation: $E_{\theta*}t(Y)=t(y_{obs})$
 - Approximate standard errors based on Hessian of maximized likelihood
 - · Not guaranteed, but simulation studies suggest to be reasonable
- Older method: maximum pseudo-likelihood estimation
 - Based on autologistic approximation; <u>don't use</u> unless you have to can be very bad (and you can't tell)

ERG Fitting Using ergm

- Dedicated statnet package for fitting, simulating models in ERG form
- Basic call structure: ergm (y~term1 (arg))
 +term2 (arg))
 - y: network object
 - term1, term2, etc.: terms to use (see?"ergm-terms")
 - arg: where relevant, arguments to the term functions
- Output: ergm object
 - Summary, print, and other methods can be used to examine it
 - simulate command can also be used to take draws from the fitted model

ERG Parameterization

- ERG form is just a way of writing models to use it, we must choose a set of terms (t)
- Some basics (dyad independence terms):
 - Edge term: $\sum_{i} \sum_{j} \mathbf{y_{ij}}$
 - Captures overall tendency of ties to form/not (density effect)
 - Row-sum term: $\sum_{i} y_{ij}$
 - Captures net tendency to send ties (sender/expansiveness effect)
 - Col-sum term: $\sum_{i} \mathbf{y_{ij}}$
 - Captures net tendency to receive ties (receiver/popularity effect)
 - Mutuality term: $\sum_{i}\sum_{j>i}y_{ij}y_{ii}$
 - Captures tendency of ties to reciprocate one another (reciprocity effect)
 - Linear covariates: $\sum_{i} \sum_{j} \mathbf{y}_{ij} X_{ij}$
 - Captures tendency of y_{ij} edges to covary with X_{ij} (covariate effect)

Some Basic Families

- Several familiar and/or famous model families can be built from the above terms...
 - *N*,*p*: edge term only
 - U|man: edge and reciprocity terms
 - netlogit: any terms other than mutuality (can rewrite as covariates using dummies)
 - p_1 : edge, row-sum, col-sum, and mutuality (or any subset thereof)
 - Can also add covariates; this was done early on for block density (selective mixing) effects
- Generally flexible, and can be fit using a multinomial logit approach
 - However, no dependence among dyads allowed...

Beyond Independence: the Star Terms

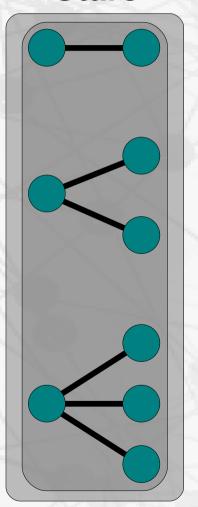
Simple subgraph census terms

- k-stars: number of subgraphs isomorphic to $K_{l,k}$
- k-in/out/mixed-stars: number of subgraphs isomorphic to orientations of $K_{I,k}$

Interpretations

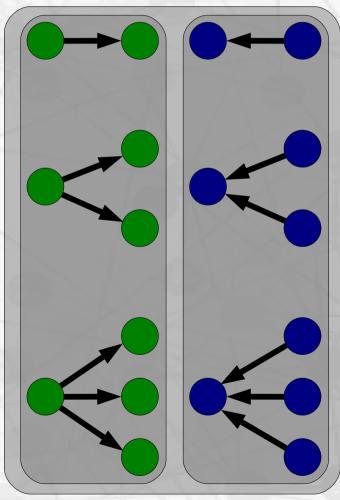
- Tendency of edges to "stick together" on endpoints ("edge clustering")
- Fixes moments of the degree distribution
 - 1-stars fix mean degree, 2stars fix variance

Stars



Undirected

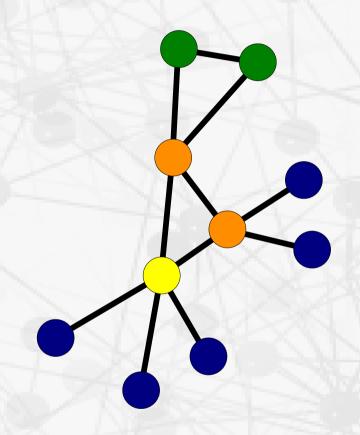
Out-stars In-stars



Directed

Another Way to See Stars: Degree Terms

- Natural reparameterization of the star terms
 - ith degree term: number of vertices of degree i
 - Likewise for indegree, outdegree terms
 - Can be derived from the full set of star terms (and vice versa)
- Interpretation
 - Non-parametric model for the degree distribution
 - Note: do not confuse with sender/receiver terms!
 - Latter refer to specific vertices, do not create dependence among edges



$$d_0=0, d_1=5, d_2=2, d_3=0,$$

 $d_4=2, d_5=1, d_6=0, d_7=0,$
 $d_8=0, d_9=0$

Triad Census Terms

Most basic terms for endogenous clustering

- Each term counts subgraphs isomorphic to triads of a given type (i.e., elements of the triad census)
- In practice, triangles, cycles, and transitives most often used

Interpretation

- Tendencies towards transitive closure, cycles, etc.
- Transitivity can be an indicator of latent hierarchy
- Cyclicity can be an indicator of extended reciprocity

