

STA 610L: MODULE 1.3

INTRODUCTION TO HIERARCHICAL MODELS

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INTRODUCTION TO HIERARCHICAL MODELS

The terminology **hierarchical model** is quite general and can imply anything from simple use of a prior distribution to a highly organized data hierarchy (students nested in classes nested in schools nested in school systems nested in states nested in countries).

For grouped or nested data for example, we may want to infer or estimate the relationship between a response variable and certain predictors collected across all the groups.

In that case, we should do so in a way that takes advantage of the relationship between observations in the same group, but we should also look to borrow information across groups.

INTRODUCTION TO HIERARCHICAL MODELS

Hierarchical models are often used in the following commonly-encountered settings:

- members of a "cluster" share more similarities with each other than with members of other clusters, violating the typical independence assumption of generalized linear models (like linear or logistic regression) -- examples of clusters include members of a family or students in a class
- hypotheses of interest include context-dependent associations, often across a large number of settings -- e.g., does success of a new mode of instruction depend on the individual teacher
- it is necessary to borrow information across groups in order to stabilize estimates or to obtain estimates with desirable properties -- e.g., we want to make state-specific estimates of election candidate preference by country of origin, but some states may have few immigrants from a given country

HYPOTHETICAL SCHOOL TESTING EXAMPLE

Suppose we wish to estimate the distribution of test scores for students at J different high schools.

In each school j , where $j = 1, \dots, J$, suppose we test a random sample of n_j students.

Let y_{ij} be the test score for the i th student in school j , with $i = 1, \dots, n_j$.

Option I: estimation can be done separately in each group, where we assume

$$y_{ij} | \mu_j, \sigma_j^2 \sim N(\mu_j, \sigma_j^2)$$

where for each school j , μ_j is the school-wide average test score, and σ_j^2 is the school-wide variance of individual test scores.

HYPOTHETICAL SCHOOL TESTING EXAMPLE

We can do classical inference for each school based on large sample 95% CI:

$\bar{y}_j \pm 1.96 \sqrt{s_j^2/n_j}$, where \bar{y}_j is the sample average in school j , and s_j^2 is the sample variance in school j .

Clearly, we can overfit the data within schools, for example, what if we only have 4 students from one of the schools?

Option II: alternatively, we might believe that $\mu_j = \mu$ for all j ; that is, all schools have the same mean. This is the assumption (null hypothesis) in ANOVA models for example.

Option I ignores that the μ_j 's should be reasonably similar, whereas option II ignores any differences between them.

It would be nice to find a compromise!

This is what we are able to do with **hierarchical modeling**.

HIERARCHICAL MODEL

Once again, suppose

$$y_{ij}|\mu_j, \sigma_j^2 \sim N(\mu_j, \sigma_j^2); \quad i = 1, \dots, n_j; \quad j = 1, \dots, J.$$

We can assume that the μ_j 's are drawn from a distribution based on the following: **conceive of the schools themselves as being a random sample from all possible school.**

Suppose μ_0 is the **overall mean of all school's average scores (a mean of the means)**, and τ^2 is the **variance of all school's average scores (a variance of the means)**.

Then, we can think of each μ_j as being drawn from a distribution, e.g.,

$$\mu_j|\mu_0, \tau^2 \sim N(\mu_0, \tau^2),$$

which gives us one more level, resulting in a hierarchical specification.

Usually, μ_0 and τ^2 will also be unknown so that we need to estimate them (usually MLE or Bayesian methods).

HIERARCHICAL MODEL: SCHOOL TESTING

EXAMPLE

Back to our example, it turns out that the multilevel estimate is

$$\hat{\mu}_j \approx \frac{\frac{n_j}{\sigma_j^2} \bar{y}_j + \frac{1}{\tau^2} \mu_0}{\frac{n_j}{\sigma_j^2} + \frac{1}{\tau^2}},$$

but since the unknown parameters have to be estimated, the classical estimate is

$$\hat{\mu}_j \approx \frac{\frac{n_j}{s_j^2} \bar{y}_j + \frac{1}{\hat{\tau}^2} \bar{y}_{\text{all}}}{\frac{n_j}{s_j^2} + \frac{1}{\hat{\tau}^2}},$$

where \bar{y}_{all} is the completely pooled estimate (the overall sample mean of all test scores).

HIERARCHICAL MODEL: IMPLICATIONS

Our estimate for each μ_j is a weighted average of \bar{y}_j and μ_0 , ensuring that we are borrowing information across all levels through μ_0 and τ^2 .

The weights for the weighted average is determined by relative precisions (the inverse of variance is often referred to as precision) from the data and from the second level model.

Suppose all $\sigma_j^2 \approx \sigma^2$. Then the schools with smaller n_j have estimated μ_j closer to μ_0 than schools with larger n_j .

Thus, the hierarchical model shrinks estimates with high variance towards the grand mean.

We seek to specify models like this in many different contexts, for many reasons, including the idea of "shrinkage".

We will do this over and over throughout the course.

GENERALIZED LINEAR MODELS (GLM)

The generalized linear model framework accommodates many popular statistical models, including linear regression, logistic regression, probit regression, and Poisson regression, among others.

Two popular GLM's we will use in class include the linear regression model and the logistic regression model.

LINEAR REGRESSION

Linear regression is perhaps the most widely-used statistical model.

Recall that the model is given by

$$y_i = \beta_0 + \beta_1 x_{1i} + \cdots + \beta_p x_{pi} + \varepsilon_i,$$

where

$$\varepsilon_i \sim N(0, \sigma^2).$$

If the parameter $\beta_j > 0$, then increasing levels of x_j are associated with larger expected values of y , and values of $\beta_j < 0$ are associated with smaller expected values of y .

$\beta_j = 0$ is consistent with no association between x_j and y .

LOGISTIC REGRESSION

Logistic regression is a type of generalized linear model, which generalizes the typical linear model to binary data.

Let y_i take either the value 1 or the value 0 (the labels assigned to 1 and 0 are arbitrary -- that is, we could let 1 denote voters and 0 denote non-voters, or we could exchange the labels -- we just need to remember our coding).

The logistic regression model is linear on the log of the odds:

$$\log \frac{\pi_i}{1 - \pi_i} = \beta_0 + \beta_1 x_{1i} + \cdots + \beta_p x_{pi},$$

where $\pi_i = Pr(y_i = 1)$.

If the parameter $\beta_j > 0$, then increasing levels of x_j are associated with higher probabilities that $y = 1$, and values of $\beta_j < 0$ are associated with lower probabilities that $y = 1$.

$\beta_j = 0$ is consistent with no association between x_j and y .

WHAT'S NEXT?

MOVE ON TO THE READINGS FOR THE NEXT MODULE!