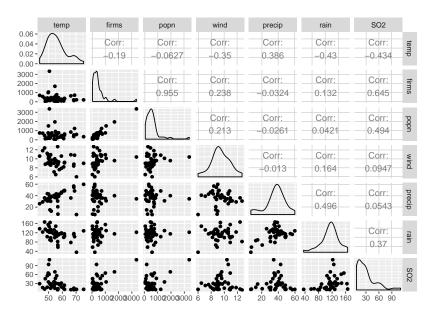
Bayesian Variable Selection

Hoff Chapter 9, Mixtures of g-Priors Liang et al JASA

February 28, 2017

US Air Example



Im summary

```
lm(formula = log(SO2) \sim temp + log(firms) + log(popn) + win
   precip + rain, data = usair)
Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept) 6.7142760 1.6475086 4.075 0.000261 ***
     temp
log(firms) 0.3698588 0.1934076 1.912 0.064289 .
log(popn) -0.1771293 0.2335520 -0.758 0.453428
wind -0.1738606 0.0656713 -2.647 0.012204 *
precip 0.0156032 0.0132718 1.176 0.247893
rain 0.0009153 0.0057335 0.160 0.874104
Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 '
Residual standard error: 0.5108 on 34 degrees of freedom
```

Multiple R-squared: 0.5503, ^IAdjusted R-squared: 0.471 F-statistic: 6.936 on 6 and 34 DF, p-value: 7.12e-05

Zellner's g-prior

Centered model:

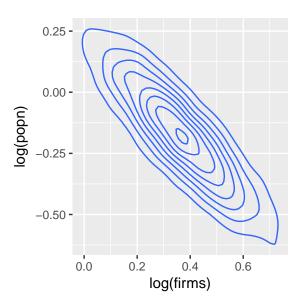
$$\mathbf{Y} = \mathbf{1}_{n}\alpha + \mathbf{X}^{c}\boldsymbol{\beta} + \epsilon$$

where \mathbf{X}^c is the centered design matrix where all variables have had their mean subtracted

- $ightharpoonup p(\phi) \propto 1/\phi$
- $ightharpoonup p(\alpha \mid \phi) \propto 1$

$$eta \mid \mathbf{Y}, lpha, \phi \sim N\left(rac{g}{1+g}\hat{oldsymbol{eta}}, \phi^{-1}rac{g}{1+g}(\mathbf{X}^T\mathbf{X})^{-1}
ight)$$
 $\phi \mid \mathbf{Y} \sim \mathsf{Gamma}\left(rac{n-1}{2}, rac{\mathsf{SSE} - rac{1}{1+g}\hat{oldsymbol{eta}}^T(\mathbf{X}^T\mathbf{X})\hat{oldsymbol{eta}}}{2}
ight)$

joint posterior draws of beta's



Bayesian Variable Selection

- Avoid the use of redundant variables (problems with interpretations)
- Inclusion of un-necessary terms yields less precise estimates, particularly if explanatory variables are highly correlated with each other
- reduced MSE: reduced variance but possibly higher bias
- it is too "expensive" to use all variables

Bayesian Model Choice

- Models for the variable selection problem are based on a subset of the X₁,...X_p variables
- ▶ Encode models with a vector $\boldsymbol{\gamma} = (\gamma_1, \dots \gamma_p)$ where $\gamma_j \in \{0,1\}$ is an indicator for whether variable \mathbf{X}_j should be included in the model \mathcal{M}_{γ} . $\gamma_j = 0 \Leftrightarrow \beta_j = 0$
- **Each** value of γ represents one of the 2^p models.
- ▶ Under model \mathcal{M}_{γ} :

$$\mathbf{Y} \mid \alpha, \boldsymbol{\beta}, \sigma^2, \boldsymbol{\gamma} \sim \mathsf{N}(\mathbf{1}\alpha + \mathbf{X}_{\boldsymbol{\gamma}}\boldsymbol{\beta}_{\boldsymbol{\gamma}}, \sigma^2 \mathbf{I})$$

Where \mathbf{X}_{γ} is design matrix using the columns in \mathbf{X} where $\gamma_j=1$ and $\boldsymbol{\beta}_{\gamma}$ is the subset of $\boldsymbol{\beta}$ that are non-zero.

Posterior Probabilities of Models

Posterior model probabilities

$$p(\mathcal{M}_j \mid \mathbf{Y}) = \frac{p(\mathbf{Y} \mid \mathcal{M}_j)p(\mathcal{M}_j)}{\sum_j p(\mathbf{Y} \mid \mathcal{M}_j)p(\mathcal{M}_j)}$$

Marginal likelihod of a model is proportional to

$$p(\mathbf{Y} \mid \mathcal{M}_{\gamma}) = \iint p(\mathbf{Y} \mid \boldsymbol{\beta}_{\gamma}, \sigma^{2}) p(\boldsymbol{\beta}_{\gamma} \mid \boldsymbol{\gamma}, \sigma^{2}) p(\sigma^{2} \mid \boldsymbol{\gamma}) d\boldsymbol{\beta} d\sigma^{2}$$

▶ Probability $\beta_j \neq 0$: $\sum_{\mathcal{M}_j:\beta_j\neq 0} p(\mathcal{M}_j \mid \mathbf{Y})$ (marginal posterior inclusion probability)

Prior Distributions

- Bayesian Model choice requires proper prior distributions on regression coefficients (exception parameters that are included in all models)
- Vague but proper priors may lead to paradoxes!
- Conjugate Normal-Gammas lead to closed form expressions for marginal likelihoods, Zellner's g-prior is one of the most popular.

Zellner's g-prior within Models

Centered model:

$$\mathbf{Y} = \mathbf{1}_{n}\alpha + \mathbf{X}_{\gamma}^{c}\beta_{\gamma} + \epsilon$$

Common parameters

$$p(\alpha,\phi)\propto\phi^{-1}$$

Model Specific parameters

$$oldsymbol{eta}_{\gamma} \mid lpha, \phi, oldsymbol{\gamma} \sim \mathsf{N}(0, g\phi^{-1}(\mathbf{X}^{c}_{\gamma}{}'\mathbf{X}^{c}_{\gamma})^{-1})$$

▶ Marginal likelihood of \mathcal{M}_{γ} is proportional to

$$ho(\mathbf{Y}\mid \mathcal{M}_{\gamma}) = C(1+g)^{rac{n-p-1}{2}}(1+g(1-R_{\gamma}^2))^{-rac{(n-1)}{2}}$$

where R_{γ}^2 is the usual R^2 for model \mathcal{M}_{γ} and C is a constant that is $p(\mathbf{Y} \mid \mathcal{M}_0)$ (model with intercept alone)

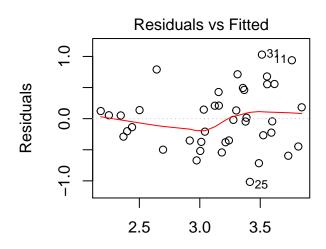
lacktriangle uniform distribution over space of models $p(\mathcal{M}_{\gamma})=1/(2^p)$

USair Data: Enumeration of All Models

```
library(devtools)
  suppressMessages(install_github("merliseclyde/BAS"))
library(BAS)
 poll.bma = bas.lm(log(SO2) \sim temp + log(firms) + temp + log(firm
                                                                                                                                                                                                                                        log(popn) + wind +
                                                                                                                                                                                                                                        precip+ rain,
                                                                                                                                                data=usair,
                                                                                                                                                prior="g-prior",
                                                                                                                                                 alpha=41, \# q = n
                                                                                                                                                n.models=2^7, # enumerate (can omit)
                                                                                                                                                method="deterministic") # fast enumera
```

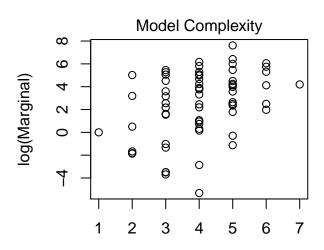
residual plot)

```
plot(poll.bma, which=1)
```



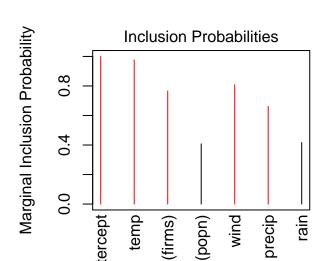
Model Complexity)

```
plot(poll.bma, which=3)
```



Inclusion Probabilities)

plot(poll.bma, which=4)



Model Space

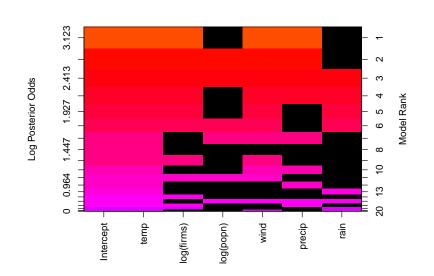
logmarg

```
summary(poll.bma)
             P(B != 0 | Y) \mod 1 \mod 2 \mod 3
##
                 1.0000000 1.000000 1.0000000 1.00000000
## Intercept
          0.9762282 1.000000 1.0000000 1.00000000
## temp
## log(firms) 0.7659857 1.000000 1.0000000 1.00000000
## log(popn) 0.4075393 0.000000 1.0000000 1.00000000
                0.8080832 1.000000 1.0000000 1.00000000
## wind
                 0.6615960 1.000000 1.0000000 1.00000000
## precip
                 0.4166394 0.000000 0.0000000 1.00000000
## rain
## BF
                        NA 1.000000 0.2093987 0.03277353 (
                       NA 0.209600 0.1097000 0.10300000
## PostProbs
## R.2
                       NA 0.542700 0.5500000 0.55030000 (
## dim
                       NA 5.000000 6.0000000 7.00000000
```

NA 7.616228 6.0527128 4.19809382 !

Summary

image(poll.bma)

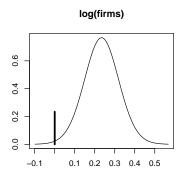


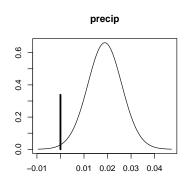
Coefficients

```
beta = coef(poll.bma, n.models=1)
beta
##
   Marginal Posterior Summaries of Coefficients:
##
##
##
  Based on the top 1 models
            post mean post SD post p(B != 0)
##
## Intercept 3.15300 0.07818 1.00000
       -0.07130 0.01268 0.97623
## temp
## log(firms) 0.23428 0.08573 0.76599
## log(popn) 0.00000 0.00000 0.40754
## wind -0.17998 0.06128 0.80808
## precip 0.01884 0.00729 0.66160
        0.00000 0.00000 0.41664
## rain
```

Coefficients

```
par(mfrow=c(2,2)); plot(beta, subset=c(3, 6))
```





Bayesian Confidence Intervals

```
confint(beta)
##
                  2.5 % 97.5 %
                                         beta
## Intercept 2.994993257 3.31101398 3.15300362
## temp -0.096926645 -0.04567203 -0.07129934
## log(firms) 0.061014518 0.40753936 0.23427694
## log(popn) 0.000000000 0.00000000 0.00000000
## wind -0.303835463 -0.05612195 -0.17997871
## precip 0.004105874 0.03357242 0.01883915
## rain 0.000000000 0.00000000 0.00000000
## attr(,"Probability")
## [1] 0.95
## attr(,"class")
## [1] "confint.bas"
```