#### STA 601: Lecture 4

# Comparing Estimators & Prior/Posterior Checks

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#### **Normal Model Setup from Last Class**

- independent observations  $\mathbf{y} = (y_1, y_2, \dots, y_n)^T$  where each  $y_i \sim \mathsf{N}(\theta, 1/\tau)$  (iid)
- The likelihood for θ is proportional to the sampling model

$$\mathcal{L}(\theta) \propto \tau^{\frac{n}{2}} \exp\left\{-\frac{1}{2}\tau \sum_{i=1}^{n} (y_{i} - \theta)^{2}\right\}$$

$$\propto \tau^{\frac{n}{2}} \exp\left\{-\frac{1}{2}\tau \sum_{i=1}^{n} [(y_{i} - \bar{y}) - (\theta - \bar{y})]^{2}\right\}$$

$$\propto \tau^{\frac{n}{2}} \exp\left\{-\frac{1}{2}\tau \left[\sum_{i=1}^{n} (y_{i} - \bar{y})^{2} + \sum_{i=1}^{n} (\theta - \bar{y})^{2}\right]\right\}$$

$$\propto \tau^{\frac{n}{2}} \exp\left\{-\frac{1}{2}\tau \left[\sum_{i=1}^{n} (y_{i} - \bar{y})^{2} + n(\theta - \bar{y})^{2}\right]\right\}$$

$$\propto \exp\left\{-\frac{1}{2}\tau n(\theta - \bar{y})^{2}\right\}$$



#### **Exercises for Practice**

#### Try this

- 1) Use  $\mathcal{L}(\theta)$  based on n observations to find  $\pi(\theta \mid y_1, \dots, y_n)$  based on the sufficient statistics and prior  $\theta \sim \mathsf{N}(\theta_0, 1/\tau_0)$
- 2) Use  $\pi(\theta \mid y_1, \dots, y_n)$  to find the posterior predictive distribution for  $Y_{n+1}$



#### After n observations

Posterior for  $\theta$ 

$$heta \mid y_1, \dots, y_n \sim \mathsf{N}\left(rac{ au_0 heta_0 + n au ar{y}}{ au_0 + n au}, rac{1}{ au_0 + n au}
ight)$$

Posterior Predictive Distribution for  $Y_{n+1}$ 

$$Y_{n+1} \mid y_1, \dots, y_n \sim \mathsf{N}\left(rac{ au_0 heta_0 + n auar{y}}{ au_0 + n au}, rac{1}{ au} + rac{1}{ au_0 + n au}
ight)$$

Shrinkage of the MLE to the prior mean



# **Results with Jeffreys' Prior**

- What if  $au_0 o 0$ ? (or  $\sigma_0^2 o \infty$ )
- Prior predictive  $N(\theta_0, \sigma_0^2 + \sigma^2)$  (not proper in the limit)
- Posterior for  $\theta$  (formal posterior)

$$heta \mid y_1, \dots, y_n \sim \mathsf{N}\left(rac{ au_0 heta_0 + n auar{y}}{ au_0 + n au}, rac{1}{ au_0 + n au}
ight)$$

$$egin{aligned} 
ightarrow & heta \mid y_1, \ldots, y_n \sim \mathsf{N}\left(ar{y}, rac{1}{n au}
ight). \end{aligned}$$

Posterior Predictive  $Y_{n+1} \mid y_1, \dots, y_n \sim \mathsf{N}\left(ar{y}, \sigma^2(1+rac{1}{n})
ight)$ 



# **Comparing Estimators**

Expected loss (from frequentist perspective) of using Bayes Estimator

Posterior mean is optimal under squared error loss (min Bayes Risk)
 [also absolute error loss]

Compute Mean Square Error (or Expected Average Loss)

$$\mathsf{E}_{ar{y}| heta}\left[\left(\hat{ heta}- heta
ight)^2\mid heta
ight]$$

$$=\mathsf{Bias}(\hat{ heta})^2+\mathsf{Var}(\hat{ heta})$$

■ For the MLE  $\bar{Y}$  this is just the variance of  $\bar{Y}$  or  $\sigma^2/n$ 

#### **MSE for Bayes**

$$\mathsf{E}_{ar{y}| heta}\left[\left(\hat{ heta}- heta
ight)^2\mid heta
ight]=\mathsf{MSE}=\mathsf{Bias}(\hat{ heta})^2+\mathsf{Var}(\hat{ heta})$$

Bias of Bayes Estimate

$$\mathsf{E}_{ar{Y}| heta}\left[rac{ au_0 heta_0+ au nar{Y}}{ au_0+ au n}
ight]=rac{ au_0( heta_0- heta)}{ au_0+ au n}$$

Variance

$$\mathsf{Var}\left(rac{ au_0 heta_0+ au nar{Y}}{ au_0+ au n}- heta\mid heta
ight)=rac{ au n}{( au_0+ au n)^2}$$

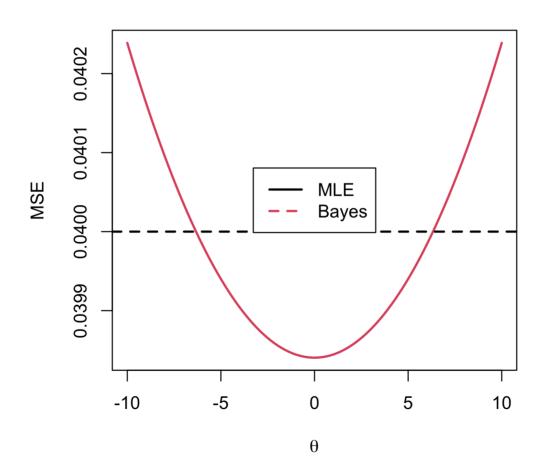
(Frequentist) expected Loss when truth is  $\theta$ 

$$\mathsf{MSE} = rac{ au_0^2( heta - heta_0)^2 + au n}{( au_0 + au n)^2}$$

Behavior?



# Plot





#### **Exercise**

Repeat this for estimating a future Y under squared error loss using a proper prior and Jeffreys' prior

$$\mathsf{E}_{Y_{n+1}\mid heta}\left[(Y_{n+1} - \mathsf{E}[Y_{n+1}\mid y_1,\ldots,n])^2]
ight]$$



#### **Uses of Posterior Predictive**

- Plot the entire density or summarize
- Available analytically for conjugate families
- Monte Carlo Approximation

$$p(y_{n+1} \mid y_1, \dots y_n) pprox rac{1}{T} \sum_{t=1}^T p(y_{n+t} \mid heta^{(t)})$$

where  $heta^{(t)} \sim \pi( heta \mid y_1, \dots y_n)$  for  $t = 1, \dots, T$ 

- T samples from the posterior distribution
- Empirical Estimates & Quantiles from Monte Carlo Samples



#### **Model Diagnostics**

- Need an accurate specification of likelihood function (and reasonable prior)
- George Box: *All models are wrong but some are useful*
- "Useful" → model provides a good approximation; there aren't clear aspects of the data that are ignored or misspecified



#### **Example**

$$Y_i \sim \mathsf{Poisson}(\theta)$$
  $i = 1, \dots, n$ 

How might our model be misspecified?

- Poisson assumes that  $E(Y_i) = Var(Y_i) = \theta$
- it's *very* common for data to be **over-dispersed**  $E(Y_i) < Var(Y_i)$
- zero-inflation many more zero values than consistent with the poisson model
- Can we use the Posterior Predictive to diagnose whether these are issues with our observed data?



#### **Posterior Predictive (PP) Checks**

- $y^{(n)}$  is observed & fixed training data
- $p(y_{n+1} \mid y^{(n)})$  is PP distributoin
- lacksquare  $ilde{y}_t^{(n)}$  is  $t^{ ext{th}}$  new dataset sampled from the PP of size n (same as training)
- $p(\tilde{y}_t^{(n)} \mid y^{(n)})$  is PP of new data sets
- compare some feature of the observed data to the datasets simulated from the PP



#### **Formally**

- choose a "test statistic"  $t(\cdot)$  that captures some summary of the data, e.g.  $Var(y^{(n)})$  for over-dispersion
- $t(y^{(n)}) \equiv t_{\text{obs}}$  value of test statistic in observed data
- $t(\tilde{y}^{(n)}) \equiv t_{\mathrm{pred}}$  value of test statistic for a random dataset drawn from the posterior predictive
- plot posterior predictive distribution of  $t(\tilde{y}^{(n)})$
- add  $t_{\rm obs}$  to plot
- How *extreme* is  $t_{obs}$  compared to the distribution of  $t(\tilde{y}^{(n)})$



# **Example Over Dispersion**



#### **Posterior Predictive p-values (PPPs)**

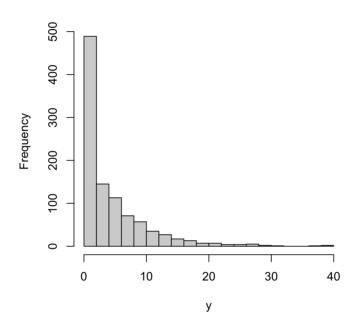
- p-value is probability of seeing something as extreme or more so under a hypothetical "null" model & are uniformally distributed under the "null" model
- PPPs advocated by Gelman & Rubin in papers and BDA are not valid p-values. They are do not have a uniform distribution under the hypothesis that the model is correctly specified
- the PPPs tend to be concentrated around 0.5, tends not to reject (conservative)
- theoretical reason for the incorrect distribution is due to double use of the data

**DO NOT USE as a formal test!** use as a diagnostic plot to see how model might fall flat



Better approach is to split the data use one piece to learn  $\theta$  and the other to calculate  $t_{\rm obs}$ 

#### **Zero Inflated Distribution**



• Let the t() be the proportion of zeros

$$t(y) = \frac{\sum_{i=1}^n \mathbb{1}(y_i = 0)}{n}$$



#### **Posterior Predictive Distribution**



# **Modeling Over-Dispersion**

- Original Model  $Y_i \mid \theta \sim \mathsf{Poisson}(\theta)$
- cause of overdispersion is variation in the rate

$$Y_i \mid heta \sim \mathsf{Poisson}( heta_i)$$
  $heta_i \sim \pi_{ heta}()$ 

- $\pi_{\theta}()$  characterizes variation in the rate parameter across inviduals
- Simple Two Stage Hierarchical Model



# **Example**

$$heta_i \sim \mathsf{Gamma}(\phi\mu,\phi)$$

- Find pmf for  $Y_i \mid \mu, \phi$
- Find  $E[Y_i \mid \mu, \phi]$  and  $Var[Y_i \mid \mu, \phi]$
- Homework:

$$heta_i \sim \mathsf{Gamma}(\phi, \phi/\mu)$$

- Can either of these model zero-inflation?
- See Bayarri & Berger (2000) for more discussion about why PPP should not be used as a test

