Outliers & Robust Bayesian Regression

Readings: Hoff Chapter 9, West JRSSB 1984, Fúquene, Pérez & Pericchi 2015

STA 601 Duke University

Duke University

October 27, 2021

► Hoeting, Madigan and Raftery (in various permutations) consider the problem of simultaneous variable selection and outlier identification.

- Hoeting, Madigan and Raftery (in various permutations) consider the problem of simultaneous variable selection and outlier identification.
- ► This is implemented in the package BMA in the function MC3.REG. This has the advantage that more than 2 points may be considered as outliers at the same time.

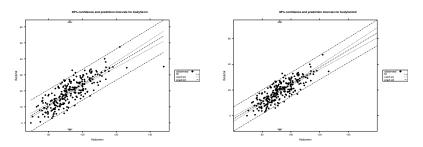
- Hoeting, Madigan and Raftery (in various permutations) consider the problem of simultaneous variable selection and outlier identification.
- ► This is implemented in the package BMA in the function MC3.REG. This has the advantage that more than 2 points may be considered as outliers at the same time.
- The function uses a Markov chain to identify both important variables and potential outliers, but is coded in Fortran so should run reasonably quickly.

- ► Hoeting, Madigan and Raftery (in various permutations) consider the problem of simultaneous variable selection and outlier identification.
- ➤ This is implemented in the package BMA in the function MC3.REG. This has the advantage that more than 2 points may be considered as outliers at the same time.
- ► The function uses a Markov chain to identify both important variables and potential outliers, but is coded in Fortran so should run reasonably quickly.
- ► Can also use BAS or other variable selection programs

- ► Hoeting, Madigan and Raftery (in various permutations) consider the problem of simultaneous variable selection and outlier identification.
- ➤ This is implemented in the package BMA in the function MC3.REG. This has the advantage that more than 2 points may be considered as outliers at the same time.
- ► The function uses a Markov chain to identify both important variables and potential outliers, but is coded in Fortran so should run reasonably quickly.
- ► Can also use BAS or other variable selection programs

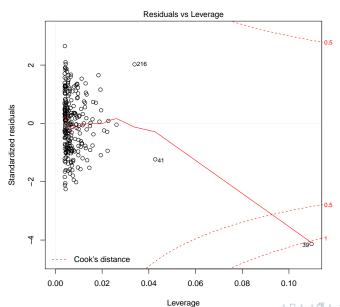
Body Fat Data: Intervals w/ All Data

Response % Body Fat and Predictor Waist Circumference



Which analysis do we use? with Case 39 or not – or something different?

Cook's Distance



Im/Podufot I/Abdomon 2 E4 * 24\\

▶ Are there scientific grounds for eliminating the case?

- ▶ Are there scientific grounds for eliminating the case?
- ► Test if the case has a different mean than population

- Are there scientific grounds for eliminating the case?
- Test if the case has a different mean than population
- Report results with and without the case

- Are there scientific grounds for eliminating the case?
- Test if the case has a different mean than population
- Report results with and without the case
- Model Averaging to Account for Model Uncertainty?

- Are there scientific grounds for eliminating the case?
- Test if the case has a different mean than population
- Report results with and without the case
- Model Averaging to Account for Model Uncertainty?
- Full model $Y = X\beta + I_n\delta + \epsilon$

- Are there scientific grounds for eliminating the case?
- ► Test if the case has a different mean than population
- Report results with and without the case
- Model Averaging to Account for Model Uncertainty?
- Full model $Y = X\beta + I_n\delta + \epsilon$
- $ightharpoonup 2^n$ submodels $\gamma_i = 0 \Leftrightarrow \delta_i = 0$
- ▶ If $\gamma_i = 1$ then case *i* has a different mean "mean shift" outliers.

Mean Shift = Variance Inflation

- $ightharpoonup Model Y = X\beta + I_n\delta + \epsilon$
- Prior

$$\delta_i \mid \gamma_i \sim N(0, V\sigma^2\gamma_i)$$

 $\gamma_i \sim \text{Ber}(\pi)$

Then ϵ_i given σ^2 is independent of δ_i and

$$\epsilon_i^* \equiv \epsilon_i + \delta_i \mid \sigma^2 \left\{ egin{array}{ll} N(0, \sigma^2) & \textit{wp} & (1 - \pi) \\ N(0, \sigma^2(1 + V)) & \textit{wp} & \pi \end{array}
ight.$$

Model $Y = X\beta + \epsilon^*$ "variance inflation" V+1=K=7 in the paper by Hoeting et al. package BMA





Simultaneous Outlier and Variable Selection

```
MC3.REG(all.y = bodyfat$Bodyfat, all.x = as.matrix(bodyfat$Abdom
       num.its = 10000, outliers = TRUE)
Model parameters: PI=0.02 K=7 nu=2.58 lambda=0.28 phi=2.85
     models were selected
Best 5 models (cumulative posterior probability = 0.9939):
          prob
                model 1 model 2 model 3 model 4 model 5
variables
 all.x
                   х
                           X
                                    х
                                                    х
                                            Х
outliers
 39
      0.94932
                           х
                                            х
 204
       0.04117
                                            Х
         0.10427
 207
                           х
                                                     Х
              0.815
                        0.095
                                 0.044
                                         0.035
                                                  0.004
post prob
```

$$Y_i \stackrel{\text{ind}}{\sim} t(\nu, \alpha + \beta x_i, 1/\phi)$$

$$Y_i \stackrel{\text{ind}}{\sim} t(\nu, \alpha + \beta x_i, 1/\phi)$$

$$L(\alpha, \beta, \phi) \propto \prod_{i=1}^n \phi^{1/2} \left(1 + \frac{\phi(y_i - \alpha - \beta x_i)^2}{\nu} \right)^{-\frac{(\nu+1)}{2}}$$

$$Y_i \stackrel{\text{ind}}{\sim} t(\nu, \alpha + \beta x_i, 1/\phi)$$

$$L(\alpha, \beta, \phi) \propto \prod_{i=1}^n \phi^{1/2} \left(1 + \frac{\phi(y_i - \alpha - \beta x_i)^2}{\nu} \right)^{-\frac{(\nu+1)}{2}}$$

Use Prior $p(\alpha, \beta, \phi) \propto 1/\phi$

$$Y_i \stackrel{\text{ind}}{\sim} t(\nu, \alpha + \beta x_i, 1/\phi)$$

$$L(\alpha, \beta, \phi) \propto \prod_{i=1}^n \phi^{1/2} \left(1 + \frac{\phi(y_i - \alpha - \beta x_i)^2}{\nu} \right)^{-\frac{(\nu+1)}{2}}$$

Use Prior $p(\alpha, \beta, \phi) \propto 1/\phi$

Posterior distribution

$$p(\alpha, \beta, \phi \mid Y) \propto \phi^{n/2-1} \prod_{i=1}^{n} \left(1 + \frac{\phi(y_i - \alpha - \beta x_i)^2}{\nu} \right)^{-\frac{(\beta+1)}{2}}$$

$$Y_i \stackrel{\text{ind}}{\sim} t(\nu, \alpha + \beta x_i, 1/\phi)$$

$$L(\alpha, \beta, \phi) \propto \prod_{i=1}^n \phi^{1/2} \left(1 + \frac{\phi(y_i - \alpha - \beta x_i)^2}{\nu} \right)^{-\frac{(\nu+1)}{2}}$$

Use Prior $p(\alpha, \beta, \phi) \propto 1/\phi$

Posterior distribution

$$p(\alpha, \beta, \phi \mid Y) \propto \phi^{n/2-1} \prod_{i=1}^{n} \left(1 + \frac{\phi(y_i - \alpha - \beta x_i)^2}{\nu} \right)^{-\frac{(\beta+1)}{2}}$$

Treat σ^2 as given, then *influence* of individual observations on the posterior distribution of β in the model where $E[Y_i] = x_i^T \beta$ is investigated through the score function:

Treat σ^2 as given, then *influence* of individual observations on the posterior distribution of β in the model where $E[Y_i] = x_i^T \beta$ is investigated through the score function:

$$\frac{d}{d\beta}\log p(\beta \mid Y) = \frac{d}{d\beta}\log p(\beta) + \sum_{i=1}^{n} x_{i}g(y_{i} - x_{i}^{T}\beta)$$

Treat σ^2 as given, then *influence* of individual observations on the posterior distribution of β in the model where $E[Y_i] = x_i^T \beta$ is investigated through the score function:

$$\frac{d}{d\beta}\log p(\beta \mid \mathsf{Y}) = \frac{d}{d\beta}\log p(\beta) + \sum_{i=1}^{n} \mathsf{x}_{i}g(y_{i} - \mathsf{x}_{i}^{\mathsf{T}}\beta)$$

where

$$g(\epsilon) = -\frac{d}{d\epsilon} \log p(\epsilon)$$

is the influence function of the error distribution (unimodal, continuous, differentiable, symmetric)

Treat σ^2 as given, then *influence* of individual observations on the posterior distribution of β in the model where $E[Y_i] = x_i^T \beta$ is investigated through the score function:

$$\frac{d}{d\beta}\log p(\beta \mid Y) = \frac{d}{d\beta}\log p(\beta) + \sum_{i=1}^{n} x_{i}g(y_{i} - x_{i}^{T}\beta)$$

where

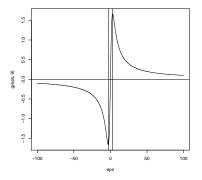
$$g(\epsilon) = -\frac{d}{d\epsilon} \log p(\epsilon)$$

is the influence function of the error distribution (unimodal, continuous, differentiable, symmetric)

An outlying observation y_j is accommodated if the posterior distribution for $p(\beta \mid Y_{(i)})$ converges to $p(\beta \mid Y)$ for all β as $|Y_i| \to \infty$. Requires error models with influence functions that go to zero such as the Student t (O'Hagan, 1979)

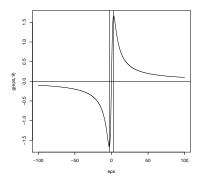
 \blacktriangleright Score function for t with α degrees of freedom has turning points at $\pm\sqrt{\alpha}$

 \blacktriangleright Score function for t with α degrees of freedom has turning points at $\pm\sqrt{\alpha}$



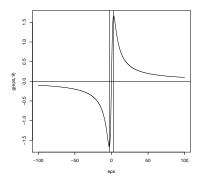
• $g'(\epsilon)$ is negative when $\epsilon^2 > \alpha$ (standardized errors)

Score function for t with α degrees of freedom has turning points at $\pm \sqrt{\alpha}$



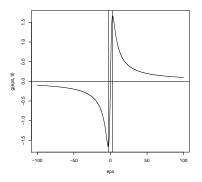
- $g'(\epsilon)$ is negative when $\epsilon^2 > \alpha$ (standardized errors)
- Contribution of observation to information matrix is negative and the observation is doubtful

 \triangleright Score function for t with α degrees of freedom has turning points at $\pm \sqrt{\alpha}$



- $ightharpoonup g'(\epsilon)$ is negative when $\epsilon^2 > \alpha$ (standardized errors)
- Contribution of observation to information matrix is negative and the observation is doubtful
- Suggest taking $\alpha = 8$ or $\alpha = 9$ to reject errors larger than $\sqrt{8}$ or 3 sd.

 \triangleright Score function for t with α degrees of freedom has turning points at $\pm \sqrt{\alpha}$



- $ightharpoonup g'(\epsilon)$ is negative when $\epsilon^2 > \alpha$ (standardized errors)
- Contribution of observation to information matrix is negative and the observation is doubtful
- Suggest taking $\alpha = 8$ or $\alpha = 9$ to reject errors larger than $\sqrt{8}$ or 3 sd.

$$Z_i \stackrel{\mathrm{iid}}{\sim} t(\nu, 0, \sigma^2) \Leftrightarrow$$

$$Z_i \stackrel{\mathrm{iid}}{\sim} t(\nu, 0, \sigma^2) \Leftrightarrow$$

$$Z_i \mid \lambda_i \stackrel{\mathrm{ind}}{\sim} N(0, \sigma^2/\lambda_i)$$

$$Z_i \stackrel{\mathrm{iid}}{\sim} t(\nu, 0, \sigma^2) \Leftrightarrow$$

$$Z_i \mid \lambda_i \stackrel{\mathrm{ind}}{\sim} N(0, \sigma^2/\lambda_i)$$

$$\lambda_i \stackrel{\mathrm{iid}}{\sim} G(\nu/2, \nu/2)$$

$$Z_i \stackrel{\mathrm{iid}}{\sim} t(\nu, 0, \sigma^2) \Leftrightarrow$$

$$Z_i \mid \lambda_i \stackrel{\mathrm{ind}}{\sim} N(0, \sigma^2/\lambda_i)$$

$$\lambda_i \stackrel{\mathrm{iid}}{\sim} G(\nu/2, \nu/2)$$

Integrate out "latent" λ 's to obtain marginal distribution.

Latent Variable Model

$$Y_i \mid \alpha, \beta, \phi, \lambda \stackrel{\text{ind}}{\sim} N(\alpha + \beta x_i, \frac{1}{\phi \lambda_i})$$

Latent Variable Model

$$Y_i \mid \alpha, \beta, \phi, \lambda \stackrel{\text{ind}}{\sim} N(\alpha + \beta x_i, \frac{1}{\phi \lambda_i})$$

$$\lambda_i \stackrel{\text{iid}}{\sim} G(\nu/2, \nu/2)$$

$$Y_i \mid \alpha, \beta, \phi, \lambda \stackrel{\text{ind}}{\sim} N(\alpha + \beta x_i, \frac{1}{\phi \lambda_i})$$

$$\lambda_i \stackrel{\text{iid}}{\sim} G(\nu/2, \nu/2)$$

$$p(\alpha, \beta, \phi) \propto 1/\phi$$

$$Y_i \mid \alpha, \beta, \phi, \lambda \stackrel{\text{ind}}{\sim} N(\alpha + \beta x_i, \frac{1}{\phi \lambda_i})$$

$$\lambda_i \stackrel{\text{iid}}{\sim} G(\nu/2, \nu/2)$$

$$p(\alpha, \beta, \phi) \propto 1/\phi$$

$$Y_i \mid \alpha, \beta, \phi, \lambda \stackrel{\text{ind}}{\sim} N(\alpha + \beta x_i, \frac{1}{\phi \lambda_i})$$

$$\lambda_i \stackrel{\text{iid}}{\sim} G(\nu/2, \nu/2)$$

$$p(\alpha, \beta, \phi) \propto 1/\phi$$

$$p((\alpha, \beta, \phi, \lambda_1, \dots, \lambda_n \mid Y) \propto \phi^{n/2} \exp \left\{-\frac{\phi}{2} \sum_i \lambda_i (y_i - \alpha - \beta x_i)^2\right\} \times \phi^{n/2}$$

$$Y_i \mid \alpha, \beta, \phi, \lambda \stackrel{\text{ind}}{\sim} N(\alpha + \beta x_i, \frac{1}{\phi \lambda_i})$$

$$\lambda_i \stackrel{\text{iid}}{\sim} G(\nu/2, \nu/2)$$

$$p(\alpha, \beta, \phi) \propto 1/\phi$$

$$p((\alpha, \beta, \phi, \lambda_1, \dots, \lambda_n \mid Y) \propto \phi^{n/2} \exp \left\{ -\frac{\phi}{2} \sum_i \lambda_i (y_i - \alpha - \beta x_i)^2 \right\} \times \phi^{-1}$$

$$Y_i \mid \alpha, \beta, \phi, \lambda \stackrel{\text{ind}}{\sim} N(\alpha + \beta x_i, \frac{1}{\phi \lambda_i})$$

$$\lambda_i \stackrel{\text{iid}}{\sim} G(\nu/2, \nu/2)$$

$$p(\alpha, \beta, \phi) \propto 1/\phi$$

$$p((\alpha, \beta, \phi, \lambda_1, \dots, \lambda_n \mid Y) \propto \phi^{n/2} \exp\left\{-\frac{\phi}{2} \sum_i \lambda_i (y_i - \alpha - \beta x_i)^2\right\} \times \phi^{-1}$$

$$\prod_{i=1}^n \lambda_i^{\nu/2-1} \exp(-\lambda_i \nu/2)$$

Model Specification via R2jags

```
rr.model = function() {
  for (i in 1:n) {
    mu[i] <- alpha0 + alpha1*(X[i] - Xbar)</pre>
    lambda[i] ~ dgamma(9/2, 9/2)
    prec[i] <- phi*lambda[i]</pre>
    Y[i] ~ dnorm(mu[i], prec[i])
  }
  phi ~ dgamma(1.0E-6, 1.0E-6)
  alpha0 ~ dnorm(0, 1.0E-6)
  alpha1 \sim dnorm(0,1.0E-6)
```

The parameters to be monitored and returned to R are specified with the variable parameters

```
parameters = c("beta0", "beta1", "sigma",
               "mu34", "y34", "lambda[39]")
```

All of the above (except lambda) are calculated from the other parameters. (See R-code for definitions of these parameters.)

The parameters to be monitored and returned to R are specified with the variable parameters

```
parameters = c("beta0", "beta1", "sigma",
               "mu34", "y34", "lambda[39]")
```

- All of the above (except lambda) are calculated from the other parameters. (See R-code for definitions of these parameters.)
- ▶ mu34 and y34 are the mean functions and predictions for a man with a 34 in waist.
- ▶ lambda[39] saves only the 39th case of λ

The parameters to be monitored and returned to R are specified with the variable parameters

- ➤ All of the above (except lambda) are calculated from the other parameters. (See R-code for definitions of these parameters.)
- mu34 and y34 are the mean functions and predictions for a man with a 34 in waist.
- ▶ lambda[39] saves only the 39th case of λ
- ► To save a whole vector (for example all lambdas, just give the vector name)

The parameters to be monitored and returned to R are specified with the variable parameters

- ➤ All of the above (except lambda) are calculated from the other parameters. (See R-code for definitions of these parameters.)
- mu34 and y34 are the mean functions and predictions for a man with a 34 in waist.
- ▶ lambda[39] saves only the 39th case of λ
- ► To save a whole vector (for example all lambdas, just give the vector name)

Output

	mean	sd	2.5%	50%	97.5%
beta0	-41.70	2.75	-46.91	-41.67	-36.40
beta1	0.66	0.03	0.60	0.66	0.71
sigma	4.48	0.23	4.05	4.46	4.96
mu34	15.10	0.35	14.43	15.10	15.82
y34	14.94	5.15	4.37	15.21	24.65
lambda[39]	0.33	0.16	0.11	0.30	0.72

95% HPD interval for expected bodyfat (14.5, 15.8) 95% HPD interval for bodyfat (5.1, 25.3)

▶ 95% Probability Interval for β is (0.60, 0.71) with t_9 errors

- ▶ 95% Probability Interval for β is (0.60, 0.71) with t_9 errors
- ▶ 95% Confidence Interval for β is (0.58, 0.69) (all data normal model)

- ▶ 95% Probability Interval for β is (0.60, 0.71) with t_0 errors
- ▶ 95% Confidence Interval for β is (0.58, 0.69) (all data normal model)
- ▶ 95% Confidence Interval for β is (0.61, 0.73) (normal model without case 39)

- ▶ 95% Probability Interval for β is (0.60, 0.71) with t_9 errors
- ▶ 95% Confidence Interval for β is (0.58, 0.69) (all data normal model)
- ▶ 95% Confidence Interval for β is (0.61, 0.73) (normal model without case 39)

Results intermediate without having to remove any observations

- ▶ 95% Probability Interval for β is (0.60, 0.71) with t_9 errors
- ▶ 95% Confidence Interval for β is (0.58, 0.69) (all data normal model)
- ▶ 95% Confidence Interval for β is (0.61, 0.73) (normal model without case 39)

Results intermediate without having to remove any observations Case 39 down weighted by λ_{39}

Full Conditional for λ_i

$$p(\lambda_j \mid \text{rest}, Y) \propto p(\alpha, \beta, \phi, \lambda_1, \dots, \lambda_n \mid Y)$$

Full Conditional for λ_i

$$p(\lambda_{j} \mid \mathsf{rest}, Y) \propto p(\alpha, \beta, \phi, \lambda_{1}, \dots, \lambda_{n} \mid Y)$$

$$\propto \phi^{n/2-1} \prod_{i=1}^{n} \exp\left\{-\frac{\phi}{2} \lambda_{i} (y_{i} - \alpha - \beta x_{i})^{2}\right\} \times$$

$$\prod_{i=1}^{n} \lambda_{i}^{\frac{\nu+1}{2}-1} \exp(-\lambda_{i} \frac{\nu}{2})$$

Full Conditional for λ_j

$$\begin{aligned} p(\lambda_j \mid \mathsf{rest}, Y) & \propto & p(\alpha, \beta, \phi, \lambda_1, \dots, \lambda_n \mid Y) \\ & \propto & \phi^{n/2 - 1} \prod_{i = 1}^n \exp\left\{ -\frac{\phi}{2} \lambda_i (y_i - \alpha - \beta x_i)^2 \right\} \times \\ & \prod_{i = 1}^n \lambda_i^{\frac{\nu + 1}{2} - 1} \exp(-\lambda_i \frac{\nu}{2}) \end{aligned}$$

Ignore all terms except those that involve λ_j

Full Conditional for λ_i

$$\begin{split} p(\lambda_j \mid \mathsf{rest}, Y) & \propto & p(\alpha, \beta, \phi, \lambda_1, \dots, \lambda_n \mid Y) \\ & \propto & \phi^{n/2 - 1} \prod_{i = 1}^n \exp\left\{ -\frac{\phi}{2} \lambda_i (y_i - \alpha - \beta x_i)^2 \right\} \times \\ & \prod_{i = 1}^n \lambda_i^{\frac{\nu + 1}{2} - 1} \exp(-\lambda_i \frac{\nu}{2}) \end{split}$$

Ignore all terms except those that involve λ_i

$$\lambda_j \mid \mathsf{rest}, Y \sim G\left(rac{
u+1}{2}, rac{\phi(y_j - lpha - eta x_j)^2 +
u}{2}
ight)$$



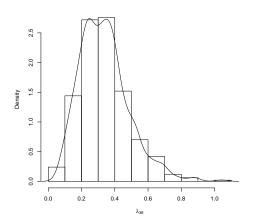
Weights

Under prior $E[\lambda_i] = 1$

Weights

Under prior $E[\lambda_i]=1$ Under posterior, large residuals are down-weighted (approximately those bigger than $\sqrt{\nu}$)

Posterior Distribution



As a general recommendation, the prior distribution should have "heavier" tails than the likelihood

ightharpoonup with t_9 errors use a t_{α} with $\alpha < 9$

- \blacktriangleright with t_0 errors use a t_α with $\alpha < 9$
- also represent via scale mixture of normals

- \blacktriangleright with t_0 errors use a t_α with $\alpha < 9$
- also represent via scale mixture of normals
- Horseshoe, Double Pareto, Cauchy all have heavier tails

- \blacktriangleright with t_0 errors use a t_α with $\alpha < 9$
- also represent via scale mixture of normals
- Horseshoe, Double Pareto, Cauchy all have heavier tails

Sumary

- Classical diagnostics useful for EDA (checking data, potential outliers/influential points) or posterior predictive checks
- ▶ BMA/BVS and Bayesian robust regression avoid interactive decision making about outliers
- Robust Regression (Bayes) can still identify outliers through distribution on weights
- continuous versus mixture distribution on scale parameters
- \triangleright Other mixtures (sub populations?) on scales and β ?