STA 601: Lecture 3

The Normal Model & Prior/Posterior Predictive Distributions

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Outline

- Normal Model
- Predictive Distributions
 - Prior Predictive; useful for prior elicitation
- Posterior Predictive
 - Predicting/forecasting future events
- Comparing Estimators



Normal Model Setup

Suppose we have independent observations

$$\mathbf{y} = (y_1, y_2, \dots, y_n)^T$$

where each $y_i \sim \mathsf{N}(\theta, \sigma^2)$ (iid)

- We will see that it is more convenient to work with $\tau = 1/\sigma^2$ (precision)
- reparameterizing the model for the data we have

$$y_i \sim \mathcal{N}(heta, au^{-1})$$

• for simplicity we will treat τ as known initially.



Marginal Distribution

Recall that the marginal distribution is

$$p(y) = p(y_1, \dots, y_n) = \int_{\Theta} p(y_1, \dots, y_n \mid heta) \pi(heta) \, d heta$$

- this is also called the **prior predictive** distribution and is independent of any unknown parameters
- We may care about making predictions before we even see any data.
- This is often useful as a way to see if the sampling distribution or prior we have chosen is appropriate, after integrating out all unknown parameters.
- Need to specify a prior for θ on \mathbb{R}



Prior for a Normal Mean

Natural choice is a Normal/Gaussian distribution (Conjugate prior)

$$heta \sim \mathsf{N}(heta_0, 1/ au_0)$$

- θ_0 is the prior mean best guess for θ using information other than y
- Prior variance $\sigma_0^2 = 1/\tau_0$
- τ_0 is the prior precision and expresses our certainty about this guess
- one notion of non-informative is to let $\tau_0 \to 0$
- better justification is as Jeffreys' prior (uniform measure) **Derive**

$$\pi(\theta) \propto 1$$

 parameterization invariant and invariant to shift changes in the data (group invariance)



Prior Predictive for a Single Case

$$egin{split} p(y) & \propto \int_{\mathbb{R}} p(y \mid heta) \pi(heta) \, d heta \ & \propto \int_{\mathbb{R}} \expigg\{ -rac{1}{2} au(y- heta)^2 igg\} \expigg\{ -rac{1}{2} au_0 (heta- heta_0)^2 igg\} \, d heta \end{split}$$

Quadratic $au_0(heta- heta_0)^2= au_0 heta^2-2 au_0 heta_0 heta+ au_0 heta_0^2$

- 1) **Expand** quadratics
- 2) **Group** terms with θ^2 and θ
- 3) Read off posterior precision and posterior mean
- 4) Complete the square
- 5) **Integrate** out θ to obtain marginal!



Try it!

$$p(y) \propto \int_{\mathbb{R}} \exp iggl\{ -rac{1}{2} [au(y- heta)^2 + au_0 (heta - heta_0)^2 iggr\} \, d heta$$



Results

Posterior for θ based on a single observation (Conjugate family)

$$heta \mid y \sim \mathsf{N}\left(\hat{ heta}, rac{1}{ au_0 + au}
ight)$$

- posterior mean $\hat{\theta} = \frac{\tau_0}{\tau_0 + \tau} \theta_0 + \frac{\tau}{\tau_0 + \tau} y$
- precision weighted average of prior mean and MLE (based on 1 observation)
- posterior precision is the sum of prior precision and data precision
- marginal distribution for Y (prior predictive)

$$Y \sim \mathsf{N}\left(heta_0, rac{1}{ au_0} + rac{1}{ au}
ight) ext{ or } \mathsf{N}(heta_0, \sigma^2 + \sigma_0^2)$$

 two sources of variability: variability from the model for the data and prior variability



Prior Predictive

- useful to think about observable quantities when choosing the prior
- sample directly from the prior predictive and assess whether the samples are consistent with our prior knowledge
- if not, go back and modify the prior & repeat
- sequential substitution sampling (repeat T times)
 - 1) draw $\theta^{(t)} \sim \pi(\theta)$
 - 2) draw $y^{(t)} \sim p(y \mid \theta^{(t)})$
- takes into account uncertain about *θ* and variability in *Y*!



Posterior Updating

- Sequential updating using the previous result as our prior!
- New prior after seeing 1 observation is

$${\sf N}(\theta_1,1/\tau_1)$$

prior mean weighted average

$$heta_1 \equiv rac{ au_0 heta_0 + au y_1}{ au_0 + au_1}$$

prior precision after 1 observation

$$au_1 \equiv au_0 + au$$

• prior variance is now $\sigma_1^2 = 1/\tau_1$



Posterior Predictive for y_2 given y_1

- Conditional $p(y_2 | y_1) = p(y_2, y_1)/p(y_1)$ (Hard way!)
- Use latent variable representation

$$p(y_2 \mid y_1) = \int_{\Theta} rac{p(y_2, \mid heta) p(y_1 \mid heta) \pi(heta) \, d heta}{p(y_1)}$$

simplify to previous problem and use results

$$p(y_2 \mid y_1) = \int_{\Theta} p(y_2 \mid heta) \pi(heta \mid y_1) \, d heta$$

■ (Posterior) Predictive

$$y_2 \mid y_1 \sim \mathsf{N}(heta_1, \sigma^2 + \sigma_1^2)$$



Iterated Expectations

Based on expressions we have an exponential of a quadratic in y_2 so know that distribution is Gaussian

- Find the mean and variance using iterated expectations:
- mean

$$\mathsf{E}[Y_2 \mid y_1] = \mathsf{E}_{\theta \mid y_1}[\mathsf{E}_{Y_2 \mid y_1, \theta}[Y_2 \mid y_1, \theta] \mid y_1]$$



Variance via Iterated Expectations

 $\mathsf{Var}[Y_2 \mid y_1] =$

 $\mathsf{E}_{\theta \mid y_1}[\mathsf{Var}_{Y_2 \mid y_1, \theta}[Y_2 \mid y_1, \theta] \mid y_1] + \mathsf{Var}_{\theta \mid y_1}[\mathsf{E}_{Y_2 \mid y_1, \theta}[Y_2 \mid y_1, \theta] \mid y_1]$



Updated Posterior for θ

$$p(\theta \mid y_1, y_2) \propto p(y_2 \mid \theta) p(y_1 \mid \theta) \pi(\theta)$$

$$p(\theta \mid y_1, y_2) \propto p(y_2 \mid \theta) p(\theta \mid y_1)$$

Apply previous updating rules

new posterior mean

$$heta_2=rac{ au_1 heta_1+ au y_2}{ au_1+ au}=rac{ au_0 heta_0+2 auar y}{ au_0+2 au}$$

new precision

$$au_2= au_1+ au= au_0+2 au$$



After n observations

Posterior for θ

$$heta \mid y_1, \dots, y_n \sim \mathsf{N}\left(rac{ au_0 heta_0 + n auar{y}}{ au_0 + n au}, rac{1}{ au_0 + n au}
ight)$$

Posterior Predictive Distribution for Y_{n+1}

$$Y_{n+1} \mid y_1, \dots, y_n \sim \mathsf{N}\left(rac{ au_0 heta_0 + n auar{y}}{ au_0 + n au}, rac{1}{ au} + rac{1}{ au_0 + n au}
ight)$$

- Shrinkage of the MLE to the prior mean
- More accurate estimation of θ as $n \to \infty$ (reducible error)
- Cannot reduce the error for prediction Y_{n+1} due to σ^2
- predictive distribution for a next observation given everything we know - prior and likelihood



Results with Jeffreys' Prior

- What if $au_0 o 0$? (or $\sigma_0^2 o \infty$)
- Prior predictive $N(\theta_0, \sigma_0^2 + \sigma^2)$ (not proper in the limit)
- Posterior for θ (formal posterior)

$$egin{aligned} heta \mid y_1, \dots, y_n &\sim \mathsf{N}\left(rac{ au_0 heta_0 + n au ar{y}}{ au_0 + n au}, rac{1}{ au_0 + n au}
ight) \ &
ightarrow & heta \mid y_1, \dots, y_n &\sim \mathsf{N}\left(ar{y}, rac{1}{n au}
ight) \end{aligned}$$

- Recovers the MLE as the posterior mode!
- Posterior variance of $\theta = \sigma^2/n$ (same as variance of the MLE)

Posterior Predictive Distribution

Posterior predictive distribution for Y_{n+1}

$$Y_{n+1} \mid y_1, \dots, y_n \sim \mathsf{N}\left(rac{ au_0 heta_0 + n auar{y}}{ au_0 + n au}, rac{1}{ au} + rac{1}{ au_0 + n au}
ight).$$

Under Jeffreys' prior

$$Y_{n+1}\mid y_1,\ldots,y_n \sim \mathsf{N}\left(ar{y},\sigma^2(1+rac{1}{n})
ight)$$

Captures extra uncertainty due to not knowing θ (compared to plug-in approach where we plug in MLE in sampling model!



Comparing Estimators

Expected loss (from frequentist perspective) of using Bayes Estimator

Posterior mean is optimal under squared error loss (min Bayes Risk)
 [also absolute error loss]

Compute Mean Square Error (or Expected Average Loss)

$$\mathsf{E}_{ar{y}| heta}\left[\left(\hat{ heta}- heta
ight)^2\mid heta
ight]$$

$$=\mathsf{Bias}(\hat{ heta})^2+\mathsf{Var}(\hat{ heta})$$

■ For the MLE \bar{Y} this is just the variance of \bar{Y} or σ^2/n

MSE for Bayes

$$\mathsf{E}_{ar{y}| heta}\left[\left(\hat{ heta}- heta
ight)^2\mid heta
ight]=\mathsf{MSE}=\mathsf{Bias}(\hat{ heta})^2+\mathsf{Var}(\hat{ heta})$$

Bias of Bayes Estimate

$$\mathsf{E}_{ar{Y}| heta}\left[rac{ au_0 heta_0+ au nar{Y}}{ au_0+ au n}
ight]=rac{ au_0(heta_0- heta)}{ au_0+ au n}$$

Variance

$$\mathsf{Var}\left(rac{ au_0 heta_0+ au nar{Y}}{ au_0+ au n}- heta\mid heta
ight)=rac{ au n}{(au_0+ au n)^2}$$

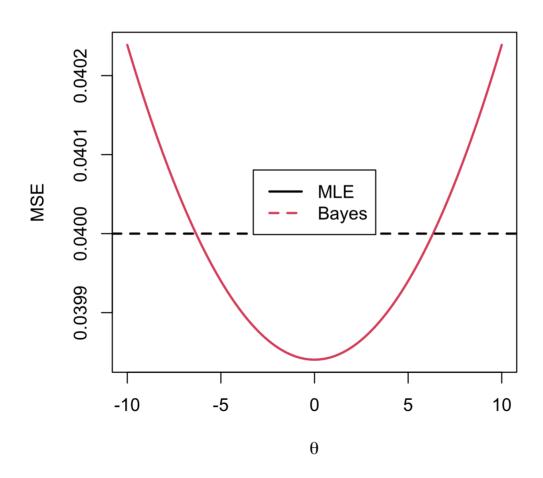
(Frequentist) expected Loss when truth is θ

$$\mathsf{MSE} = rac{ au_0^2(heta - heta_0)^2 + au n}{(au_0 + au n)^2}$$

Behavior?



Plot





Updating with _n **Observations**

• We can use the $\mathcal{L}(\theta)$ based on n observations and repeat completing the square with the original prior $\theta \sim N(\theta_0, 1/\tau_0)$



Likelihood Function

• The likelihood for θ is proportional to the sampling model

$$p(y \mid heta, au) = \prod_{i=1}^n rac{1}{\sqrt{2\pi}} au^{rac{1}{2}} \exp\left\{-rac{1}{2} au(y_i - heta)^2
ight\}.$$

Rewrite in terms of sufficient statistics!



Simplification

$$\mathcal{L}(\theta) \propto \tau^{\frac{n}{2}} \exp \left\{ -\frac{1}{2} \tau \sum_{i=1}^{n} (y_i - \theta)^2 \right\}$$

$$\propto \tau^{\frac{n}{2}} \exp \left\{ -\frac{1}{2} \tau \sum_{i=1}^{n} [(y_i - \bar{y}) - (\theta - \bar{y})]^2 \right\}$$

$$\propto \tau^{\frac{n}{2}} \exp \left\{ -\frac{1}{2} \tau \left[\sum_{i=1}^{n} (y_i - \bar{y})^2 + \sum_{i=1}^{n} (\theta - \bar{y})^2 \right] \right\}$$

$$\propto \tau^{\frac{n}{2}} \exp \left\{ -\frac{1}{2} \tau \left[\sum_{i=1}^{n} (y_i - \bar{y})^2 + n(\theta - \bar{y})^2 \right] \right\}$$

$$\propto \tau^{\frac{n}{2}} \exp \left\{ -\frac{1}{2} \tau s^2 (n-1) \right\} \exp \left\{ -\frac{1}{2} \tau n(\theta - \bar{y})^2 \right\}.$$

$$\propto \exp \left\{ -\frac{1}{2} \tau n(\theta - \bar{y})^2 \right\}$$



Exercises for Practice

Try this

- 1) Use $\mathcal{L}(\theta)$ based on n observations and $\pi(\theta)$ to find $\pi(\theta \mid y_1, \dots, y_n)$ based on the sufficient statistics
- 2) Use $\pi(\theta \mid y_1, \dots, y_n)$ to find the posterior predictive distribution for Y_{n+1}

