

# STA 601: Random Effects

# STA 601 Fall 2021

# Merlise Clyde

# Nov 1, 2021



# Hierarchical Models Continued

- Models for Gaussian Data with no Covariates

$$y_{ij} \sim ? \quad i = 1, \dots, n; j = 1, \dots, n_i$$



# Hierarchical Models Continued

- Models for Gaussian Data with no Covariates

$$y_{ij} \sim ? \quad i = 1, \dots, n; j = 1, \dots, n_i$$

- $i$  "block" - schools, counties, etc



# Hierarchical Models Continued

- Models for Gaussian Data with no Covariates

$$y_{ij} \sim ? \quad i = 1, \dots, n; j = 1, \dots, n_i$$

- $i$  "block" - schools, counties, etc
- $j$  observations within a block - students within schools, households within counties, etc



# Hierarchical Models Continued

- Models for Gaussian Data with no Covariates

$$y_{ij} \sim ? \quad i = 1, \dots, n; j = 1, \dots, n_i$$

- $i$  "block" - schools, counties, etc
- $j$  observations within a block - students within schools, households within counties, etc
- potentially there may be within block dependence in the observations due to unmeasured covariates



# Hierarchical Models Continued

- Models for Gaussian Data with no Covariates

$$y_{ij} \sim ? \quad i = 1, \dots, n; j = 1, \dots, n_i$$

- $i$  "block" - schools, counties, etc
- $j$  observations within a block - students within schools, households within counties, etc
- potentially there may be within block dependence in the observations due to unmeasured covariates
- structure?



# Models

- Naive model (baseline)

$$y_{ij} \stackrel{iid}{\sim} N(\mu, \sigma^2)$$



# Models

- Naive model (baseline)

$$y_{ij} \stackrel{iid}{\sim} N(\mu, \sigma^2)$$

- issue: no systematic variation across blocks





# Models

- Naive model (baseline)

$$y_{ij} \stackrel{iid}{\sim} N(\mu, \sigma^2)$$

- issue: no systematic variation across blocks
- Fixed Effects model:

$$y_{ij} \stackrel{ind}{\sim} N(\mu_i, \sigma^2)$$



# Models

- Naive model (baseline)

$$y_{ij} \stackrel{iid}{\sim} N(\mu, \sigma^2)$$

- issue: no systematic variation across blocks
- Fixed Effects model:

$$y_{ij} \stackrel{ind}{\sim} N(\mu_i, \sigma^2)$$

- Common reparameterization

$$y_{ij} \stackrel{ind}{\sim} N(\alpha + \beta_i, \sigma^2)$$

- $\mu$  intercept
- $\beta_i$  shift for school



# Models

- Naive model (baseline)

$$y_{ij} \stackrel{iid}{\sim} N(\mu, \sigma^2)$$

- issue: no systematic variation across blocks
- Fixed Effects model:

$$y_{ij} \stackrel{ind}{\sim} N(\mu_i, \sigma^2)$$

- Common reparameterization

$$y_{ij} \stackrel{ind}{\sim} N(\alpha + \beta_i, \sigma^2)$$

- $\mu$  intercept
- $\beta_i$  shift for school
- Identifiability ?



# Non-Identifiability

- Example:  $y_i \sim N(\alpha + \beta, \sigma^2)$  overparameterized



# Non-Identifiability

- Example:  $y_i \sim N(\alpha + \beta, \sigma^2)$  overparameterized
- $\mu = \alpha + \beta$  and  $\sigma^2$  are uniquely estimated, but not  $\alpha$  or  $\beta$



# Non-Identifiability

- Example:  $y_i \sim N(\alpha + \beta, \sigma^2)$  overparameterized
- $\mu = \alpha + \beta$  and  $\sigma^2$  are uniquely estimated, but not  $\alpha$  or  $\beta$
- $x_i \in \{1, \dots, d\}$  factor levels

$$y_i \sim N(\mu + \sum_j \beta_j 1(x_i = j), \sigma^2)$$



# Non-Identifiability

- Example:  $y_i \sim N(\alpha + \beta, \sigma^2)$  overparameterized
- $\mu = \alpha + \beta$  and  $\sigma^2$  are uniquely estimated, but not  $\alpha$  or  $\beta$
- $x_i \in \{1, \dots, d\}$  factor levels

$$y_i \sim N(\mu + \sum_j \beta_j 1(x_i = j), \sigma^2)$$

$\theta_j = \mu + \beta_j$  identifiable -  $d$  equations but  $d + 1$  unknowns



# Non-Identifiability

- Example:  $y_i \sim N(\alpha + \beta, \sigma^2)$  overparameterized
- $\mu = \alpha + \beta$  and  $\sigma^2$  are uniquely estimated, but not  $\alpha$  or  $\beta$
- $x_i \in \{1, \dots, d\}$  factor levels

$$y_i \sim N(\mu + \sum_j \beta_j 1(x_i = j), \sigma^2)$$

$\theta_j = \mu + \beta_j$  identifiable -  $d$  equations but  $d + 1$  unknowns

- Put constraints on parameters
  - $\alpha = 0$
  - $\beta_d = 0$
  - $\sum \beta_j = 0$





# Bayesian Notion of Identifiability

- Bayesian Learning



# Bayesian Notion of Identifiability

- Bayesian Learning
- the posterior distribution differs from the prior



# Bayesian Notion of Identifiability

- Bayesian Learning
- the posterior distribution differs from the prior
- **Note:** In general, it's good to avoid working with non-identifiable models;

**Caveats:**



# Bayesian Notion of Identifiability

- Bayesian Learning
- the posterior distribution differs from the prior
- **Note:** In general, it's good to avoid working with non-identifiable models;

## Caveats:

- Forcing identifiability may involve (complex) constraints that bias parameter estimates and make MCMC less efficient



# Bayesian Notion of Identifiability

- Bayesian Learning
- the posterior distribution differs from the prior
- **Note:** In general, it's good to avoid working with non-identifiable models;

## Caveats:

- Forcing identifiability may involve (complex) constraints that bias parameter estimates and make MCMC less efficient
- sometimes purposely introduce non-identifiability to improve computation (parameter expansion PX)



# Bayesian Notion of Identifiability

- Bayesian Learning
- the posterior distribution differs from the prior
- **Note:** In general, it's good to avoid working with non-identifiable models;

## Caveats:

- Forcing identifiability may involve (complex) constraints that bias parameter estimates and make MCMC less efficient
- sometimes purposely introduce non-identifiability to improve computation (parameter expansion PX)
- run non-identifiable model and focus on identifiable parameters or functions of them



# Bayesian Notion of Identifiability

- Bayesian Learning
- the posterior distribution differs from the prior
- **Note:** In general, it's good to avoid working with non-identifiable models;

## Caveats:

- Forcing identifiability may involve (complex) constraints that bias parameter estimates and make MCMC less efficient
- sometimes purposely introduce non-identifiability to improve computation (parameter expansion PX)
- run non-identifiable model and focus on identifiable parameters or functions of them
- post-processing of output



# Issue with Fixed Effect Approach

- What if  $n_i$ , number of observations per block, are small?





# Issue with Fixed Effect Approach

- What if  $n_i$ , number of observations per block, are small?
- Estimated uncertainty/variances are large based on MLE using group specific means



# Issue with Fixed Effect Approach

- What if  $n_i$ , number of observations per block, are small?
- Estimated uncertainty/variances are large based on MLE using group specific means
- What if blocks might be viewed as a sample from some larger population? Sample of schools?



# Issue with Fixed Effect Approach

- What if  $n_i$ , number of observations per block, are small?
- Estimated uncertainty/variances are large based on MLE using group specific means
- What if blocks might be viewed as a sample from some larger population? Sample of schools?
- May want inference about the larger population and say things about future blocks!



# Issue with Fixed Effect Approach

- What if  $n_i$ , number of observations per block, are small?
- Estimated uncertainty/variances are large based on MLE using group specific means
- What if blocks might be viewed as a sample from some larger population? Sample of schools?
- May want inference about the larger population and say things about future blocks!
- fixed effects do not allow us to say anything about blocks not in our sample!



# Issue with Fixed Effect Approach

- What if  $n_i$ , number of observations per block, are small?
- Estimated uncertainty/variances are large based on MLE using group specific means
- What if blocks might be viewed as a sample from some larger population? Sample of schools?
- May want inference about the larger population and say things about future blocks!
- fixed effects do not allow us to say anything about blocks not in our sample!
- how to address this?



# Random Effects

$$y_{ij} = \alpha + \beta_i + \epsilon_{ij}, \quad \epsilon_{ij} \stackrel{iid}{\sim} N(0, \sigma^2)$$
$$\beta_i \stackrel{iid}{\sim} N(0, \tau^2)$$

- random effects  $\beta_j$



# Random Effects

$$y_{ij} = \alpha + \beta_i + \epsilon_{ij}, \quad \epsilon_{ij} \stackrel{iid}{\sim} N(0, \sigma^2)$$
$$\beta_i \stackrel{iid}{\sim} N(0, \tau^2)$$

- random effects  $\beta_j$



# Random Effects

$$y_{ij} = \alpha + \beta_i + \epsilon_{ij}, \quad \epsilon_{ij} \stackrel{iid}{\sim} N(0, \sigma^2)$$
$$\beta_i \stackrel{iid}{\sim} N(0, \tau^2)$$

- random effects  $\beta_j$
- Random effect distributions should be viewed as part of the model specification (likelihood)





# Random Effects

$$y_{ij} = \alpha + \beta_i + \epsilon_{ij}, \quad \epsilon_{ij} \stackrel{iid}{\sim} N(0, \sigma^2)$$
$$\beta_i \stackrel{iid}{\sim} N(0, \tau^2)$$

- random effects  $\beta_j$
- Random effect distributions should be viewed as part of the model specification (likelihood)
- We've specified the likelihood in a hierarchical manner to induce desirable structure



# Random Effects

$$y_{ij} = \alpha + \beta_i + \epsilon_{ij}, \quad \epsilon_{ij} \stackrel{iid}{\sim} N(0, \sigma^2)$$
$$\beta_i \stackrel{iid}{\sim} N(0, \tau^2)$$

- random effects  $\beta_j$
- Random effect distributions should be viewed as part of the model specification (likelihood)
- We've specified the likelihood in a hierarchical manner to induce desirable structure
- unknown parameters are population parameters  $\alpha$ ,  $\tau$  and  $\sigma^2$



# Random Effects

$$y_{ij} = \alpha + \beta_i + \epsilon_{ij}, \quad \epsilon_{ij} \stackrel{iid}{\sim} N(0, \sigma^2)$$
$$\beta_i \stackrel{iid}{\sim} N(0, \tau^2)$$

- random effects  $\beta_j$
- Random effect distributions should be viewed as part of the model specification (likelihood)
- We've specified the likelihood in a hierarchical manner to induce desirable structure
- unknown parameters are population parameters  $\alpha$ ,  $\tau$  and  $\sigma^2$
- Bayesians put prior distributions on  $\alpha$ ,  $\tau$  and  $\sigma^2$ ; frequentists don't!



# Equivalent Model

$$y_i = (y_{i1}, y_{i2}, \dots, y_{in_i})$$

$$y_i \stackrel{ind}{\sim} N_{n_i} \left( \alpha 1_{n_i}, \begin{pmatrix} \sigma^2 + \tau & \tau & \dots & \tau \\ \tau & \ddots & & \tau \\ \vdots & & \ddots & \vdots \\ \tau & \dots & \tau & \sigma^2 + \tau \end{pmatrix} \right)$$

within-block correlation



# Equivalent Model

$$y_i = (y_{i1}, y_{i2}, \dots, y_{in_i})$$

$$y_i \stackrel{ind}{\sim} N_{n_i} \left( \alpha 1_{n_i}, \begin{pmatrix} \sigma^2 + \tau & \tau & \dots & \tau \\ \tau & \ddots & & \tau \\ \vdots & & \ddots & \vdots \\ \tau & \dots & \tau & \sigma^2 + \tau \end{pmatrix} \right)$$

within-block correlation

- algorithmically we can use either the latent variable model or the collapsed (marginal) model for inferences;



# Equivalent Model

$$y_i = (y_{i1}, y_{i2}, \dots, y_{in_i})$$

$$y_i \stackrel{ind}{\sim} N_{n_i} \left( \alpha 1_{n_i}, \begin{pmatrix} \sigma^2 + \tau & \tau & \dots & \tau \\ \tau & \ddots & & \tau \\ \vdots & & \ddots & \vdots \\ \tau & \dots & \tau & \sigma^2 + \tau \end{pmatrix} \right)$$

within-block correlation

- algorithmically we can use either the latent variable model or the collapsed (marginal) model for inferences;
- often latent variable is easier to work with!



# Simple Gibbs Sampler

$$\theta = (\alpha, \tau, \sigma^2, \beta_1, \dots, \beta_n)$$

$$\alpha \sim N(\alpha_0, V_0)$$

$$\tau^{-1} \sim \text{Gamma}(a_\tau/2, b_\tau/2)$$

$$\sigma^{-2} \sim \text{Gamma}(a_\sigma/2, b_\sigma/2)$$



# Simple Gibbs Sampler

$$\theta = (\alpha, \tau, \sigma^2, \beta_1, \dots, \beta_n)$$

$$\alpha \sim N(\alpha_0, V_0)$$

$$\tau^{-1} \sim \text{Gamma}(a_\tau/2, b_\tau/2)$$

$$\sigma^{-2} \sim \text{Gamma}(a_\sigma/2, b_\sigma/2)$$

**Full Conditionals:**

$$\alpha \mid \tau, \sigma^2, \beta_1, \dots, \beta_n \sim N(\hat{\alpha}, \hat{V}_n)$$

$$\hat{V}_n = \left( \frac{1}{V_0} + \sum_i \frac{n_i}{\sigma^2} \right)^{-1}$$

$$\hat{\alpha} = \frac{\frac{\alpha_0}{V_0} + \frac{\sum_i n_i \bar{y}_i^*}{\sigma^2}}{\hat{V}_n^{-1}}$$

$$y_{ij}^* \equiv y_{ij} - \beta_i \quad \bar{y}_i^* \equiv \frac{\sum_j (y_{ij} - \beta_i)}{n_i}$$





# Full Conditional Continued

$$\sigma^{-2} \mid \alpha, \tau, \beta_1, \dots, \beta_n \sim \text{Gamma} \left( \frac{a_\sigma + \sum_i n_i}{2}, \frac{b_\sigma + \sum_i \sum_j (y_{ij} - \alpha - \beta_i)^2}{2} \right)$$



# Full Conditional Continued

$$\sigma^{-2} \mid \alpha, \tau, \beta_1, \dots, \beta_n \sim \text{Gamma} \left( \frac{a_\sigma + \sum_i n_i}{2}, \frac{b_\sigma + \sum_i \sum_j (y_{ij} - \alpha - \beta_i)^2}{2} \right)$$

$$\tau^{-1} \mid \alpha, \sigma^2, \beta_1, \dots, \beta_n \sim \text{Gamma} \left( \frac{a_\tau + n}{2}, \frac{b_\tau + \sum_i \beta_i^2}{2} \right)$$



# Full Conditional Continued

$$\sigma^{-2} \mid \alpha, \tau, \beta_1, \dots, \beta_n \sim \text{Gamma} \left( \frac{a_\sigma + \sum_i n_i}{2}, \frac{b_\sigma + \sum_i \sum_j (y_{ij} - \alpha - \beta_i)^2}{2} \right)$$

$$\tau^{-1} \mid \alpha, \sigma^2, \beta_1, \dots, \beta_n \sim \text{Gamma} \left( \frac{a_\tau + n}{2}, \frac{b_\tau + \sum_i \beta_i^2}{2} \right)$$

$$\beta_j \mid \alpha, \tau, \sigma^2 \stackrel{\text{ind}}{\sim} N(\hat{b}_i, \hat{V}_{\beta_i})$$

$$\hat{V}_{\beta_i} = \left( \frac{1}{\tau} + \frac{n_i}{\sigma^2} \right)^{-1}$$

$$\hat{b}_i = \frac{\frac{0}{\tau} + \frac{n_i \bar{y}_i^*}{\sigma^2}}{\hat{V}_{\beta_i}^{-1}}$$

$$y_{ij}^{**} \equiv y_{ij} - \alpha \qquad \bar{y}_i^{**} \equiv \frac{\sum_j (y_{ij} - \alpha)}{n_i}$$



# Complications Relative to Usual Regression

## 1. Prior Choice



# Complications Relative to Usual Regression

1. Prior Choice
2. Mixing and its dependence on parameterization



# Complications Relative to Usual Regression

1. Prior Choice
  2. Mixing and its dependence on parameterization
- Early recommendation after Gibbs Sampler introduced non-informative priors

$$\pi(\alpha) \propto 1$$

$$\pi(\sigma^{-2}) \sim \text{Gamma}(\epsilon/2, \epsilon/2) \quad \pi(\sigma^{-2}) \propto 1/\sigma^{-2} \text{ as } \epsilon \rightarrow 0$$

$$\pi(\tau^{-1}) \sim \text{Gamma}(\epsilon/2, \epsilon/2) \quad \pi(\tau^{-1}) \propto 1/\tau^{-1} \text{ as } \epsilon \rightarrow 0$$



# Complications Relative to Usual Regression

1. Prior Choice
  2. Mixing and its dependence on parameterization
- Early recommendation after Gibbs Sampler introduced non-informative priors

$$\pi(\alpha) \propto 1$$

$$\pi(\sigma^{-2}) \sim \text{Gamma}(\epsilon/2, \epsilon/2) \quad \pi(\sigma^{-2}) \propto 1/\sigma^{-2} \text{ as } \epsilon \rightarrow 0$$

$$\pi(\tau^{-1}) \sim \text{Gamma}(\epsilon/2, \epsilon/2) \quad \pi(\tau^{-1}) \propto 1/\tau^{-1} \text{ as } \epsilon \rightarrow 0$$

- Are full conditionals proper ?



# Complications Relative to Usual Regression

1. Prior Choice
  2. Mixing and its dependence on parameterization
- Early recommendation after Gibbs Sampler introduced non-informative priors

$$\pi(\alpha) \propto 1$$

$$\pi(\sigma^{-2}) \sim \text{Gamma}(\epsilon/2, \epsilon/2) \quad \pi(\sigma^{-2}) \propto 1/\sigma^{-2} \text{ as } \epsilon \rightarrow 0$$

$$\pi(\tau^{-1}) \sim \text{Gamma}(\epsilon/2, \epsilon/2) \quad \pi(\tau^{-1}) \propto 1/\tau^{-1} \text{ as } \epsilon \rightarrow 0$$

- Are full conditionals proper ?
- Is joint posterior proper ?





# MCMC and Priors

- proper full conditionals



# MCMC and Priors

- proper full conditionals
- joint is improper



# MCMC and Priors

- proper full conditionals
- joint is improper
- MCMC won't converge to the stationary distribution (doesn't exist)



# MCMC and Priors

- proper full conditionals
- joint is improper
- MCMC won't converge to the stationary distribution (doesn't exist)
- may not notice it!



# Diffuse But Proper

$$\alpha \sim N(0, 10^{-6})$$

$$\pi(\sigma^{-2}) \sim \text{Gamma}(10^{-6}, 10^{-6})$$

$$\pi(\tau^{-1}) \sim \text{Gamma}(10^{-6}, 10^{-6})$$



# Diffuse But Proper

$$\alpha \sim N(0, 10^{-6})$$

$$\pi(\sigma^{-2}) \sim \text{Gamma}(10^{-6}, 10^{-6})$$

$$\pi(\tau^{-1}) \sim \text{Gamma}(10^{-6}, 10^{-6})$$

- Nearly improper priors lead to terrible performance! highly sensitive to just how vague the prior is!



# Diffuse But Proper

$$\alpha \sim N(0, 10^{-6})$$

$$\pi(\sigma^{-2}) \sim \text{Gamma}(10^{-6}, 10^{-6})$$

$$\pi(\tau^{-1}) \sim \text{Gamma}(10^{-6}, 10^{-6})$$

- Nearly improper priors lead to terrible performance! highly sensitive to just how vague the prior is!



# Alternative Priors

- Choose a flat or heavy tailed prior for random effect standard deviation  $\tau^{1/2}$

$$y_{ij} = \alpha + \beta_i + \epsilon_{ij}$$

$$y_{ij} = \alpha + \lambda\eta_i + \epsilon_{ij}$$

$\Leftrightarrow$

$$\beta_i \stackrel{iid}{\sim} N(0, \tau)$$

$$\eta_i \stackrel{iid}{\sim} N(0, 1)$$





# Alternative Priors

- Choose a flat or heavy tailed prior for random effect standard deviation  $\tau^{1/2}$

$$y_{ij} = \alpha + \beta_i + \epsilon_{ij}$$

$$y_{ij} = \alpha + \lambda\eta_i + \epsilon_{ij}$$

$$\Leftrightarrow$$

$$\beta_i \stackrel{iid}{\sim} N(0, \tau)$$

$$\eta_i \stackrel{iid}{\sim} N(0, 1)$$

- Reparameterization

$$\eta_i = \frac{\beta_i}{\tau^{1/2}} \Rightarrow \frac{\beta_i}{\lambda} \sim N(0, 1)$$



# Alternative Priors

- Choose a flat or heavy tailed prior for random effect standard deviation  $\tau^{1/2}$

$$y_{ij} = \alpha + \beta_i + \epsilon_{ij}$$

$$y_{ij} = \alpha + \lambda\eta_i + \epsilon_{ij}$$

$$\Leftrightarrow$$

$$\beta_i \stackrel{iid}{\sim} N(0, \tau)$$

$$\eta_i \stackrel{iid}{\sim} N(0, 1)$$

- Reparameterization

$$\eta_i = \frac{\beta_i}{\tau^{1/2}} \Rightarrow \frac{\beta_i}{\lambda} \sim N(0, 1)$$

- $\pi(\lambda) \propto 1(\lambda > 0)$  (improper prior)



# Alternative Priors

- Choose a flat or heavy tailed prior for random effect standard deviation  $\tau^{1/2}$

$$y_{ij} = \alpha + \beta_i + \epsilon_{ij}$$

$$y_{ij} = \alpha + \lambda\eta_i + \epsilon_{ij}$$

$$\Leftrightarrow$$

$$\beta_i \stackrel{iid}{\sim} N(0, \tau)$$

$$\eta_i \stackrel{iid}{\sim} N(0, 1)$$

- Reparameterization

$$\eta_i = \frac{\beta_i}{\tau^{1/2}} \Rightarrow \frac{\beta_i}{\lambda} \sim N(0, 1)$$

- $\pi(\lambda) \propto 1(\lambda > 0)$  (improper prior)
- $\pi(\lambda) \propto 1(\lambda > 0)N(0, 1)$  folded standard normal (half-normal)



# Alternative Priors

- Choose a flat or heavy tailed prior for random effect standard deviation  $\tau^{1/2}$

$$y_{ij} = \alpha + \beta_i + \epsilon_{ij}$$

$$y_{ij} = \alpha + \lambda\eta_i + \epsilon_{ij}$$

$\Leftrightarrow$

$$\beta_i \stackrel{iid}{\sim} N(0, \tau)$$

$$\eta_i \stackrel{iid}{\sim} N(0, 1)$$

- Reparameterization

$$\eta_i = \frac{\beta_i}{\tau^{1/2}} \Rightarrow \frac{\beta_i}{\lambda} \sim N(0, 1)$$

- $\pi(\lambda) \propto 1(\lambda > 0)$  (improper prior)
- $\pi(\lambda) \propto 1(\lambda > 0)N(0, 1)$  folded standard normal (half-normal)
- $\pi(\lambda) \propto 1(\lambda > 0)N(0, 1/\psi)$        $\psi \sim \text{Gamma}(\nu/2, \nu/2)$  folded t or half t



# Proper Posterior

Work with

$$\pi(\mu, \tau, \sigma^2 \mid y) \propto \pi(\mu, \tau, \sigma^2) \prod_{i=1}^n N \left( y_i; \alpha 1_{n_i}, \begin{pmatrix} \sigma^2 + \tau & \tau & \dots & \tau \\ \tau & \ddots & & \tau \\ \vdots & & \ddots & \vdots \\ \tau & \dots & \tau & \sigma^2 + \tau \end{pmatrix} \right)$$

- take  $\pi(\mu, \tau^{1/2}, \sigma^2) \propto \sigma^{-2} \mathbf{t}_1^+(\tau^{1/2}; 0, 1)$



# Proper Posterior

Work with

$$\pi(\mu, \tau, \sigma^2 \mid y) \propto \pi(\mu, \tau, \sigma^2) \prod_{i=1}^n N \left( y_i; \alpha 1_{n_i}, \begin{pmatrix} \sigma^2 + \tau & \tau & \dots & \tau \\ \tau & \ddots & & \tau \\ \vdots & & \ddots & \vdots \\ \tau & \dots & \tau & \sigma^2 + \tau \end{pmatrix} \right)$$

- take  $\pi(\mu, \tau^{1/2}, \sigma^2) \propto \sigma^{-2} \mathbf{t}_1^+(\tau^{1/2}; 0, 1)$
- take  $\pi(\mu, \tau^{1/2}, \sigma^2) \propto \sigma^{-2}$



# Proper Posterior

Work with

$$\pi(\mu, \tau, \sigma^2 \mid y) \propto \pi(\mu, \tau, \sigma^2) \prod_{i=1}^n N \left( y_i; \alpha 1_{n_i}, \begin{pmatrix} \sigma^2 + \tau & \tau & \dots & \tau \\ \tau & \ddots & & \tau \\ \vdots & & \ddots & \vdots \\ \tau & \dots & \tau & \sigma^2 + \tau \end{pmatrix} \right)$$

- take  $\pi(\mu, \tau^{1/2}, \sigma^2) \propto \sigma^{-2} \mathbf{t}_1^+(\tau^{1/2}; 0, 1)$
- take  $\pi(\mu, \tau^{1/2}, \sigma^2) \propto \sigma^{-2}$
- Show joint posterior is proper !



# Proper Posterior

Work with

$$\pi(\mu, \tau, \sigma^2 \mid y) \propto \pi(\mu, \tau, \sigma^2) \prod_{i=1}^n N \left( y_i; \alpha 1_{n_i}, \begin{pmatrix} \sigma^2 + \tau & \tau & \dots & \tau \\ \tau & \ddots & & \tau \\ \vdots & & \ddots & \vdots \\ \tau & \dots & \tau & \sigma^2 + \tau \end{pmatrix} \right)$$

- take  $\pi(\mu, \tau^{1/2}, \sigma^2) \propto \sigma^{-2} \mathbf{t}_1^+(\tau^{1/2}; 0, 1)$
- take  $\pi(\mu, \tau^{1/2}, \sigma^2) \propto \sigma^{-2}$
- Show joint posterior is proper !
- See Gelman 2005 discussion of Draper paper in Bayesian Analysis





# Propriety

- need expression for likelihood; requires determinant and inverse of intra-class correlation matrix! Write covariance as  $\sigma^2 I_{n_i} + \tau n_1 P_1$  and find spectral decomposition to provide determinant and inverse!
- integrate out  $\alpha$  (messy)
- determine if integrals are finite (what happens at 0 and infinity ?)
- look at special case when  $n_i$  are all equal.



# Linear Mixed Effects

$$y_{ij} = X_{ij}^T B + z_{ij}^T \beta_i + \epsilon_{ij}$$



# Linear Mixed Effects

$$y_{ij} = X_{ij}^T B + z_{ij}^T \beta_i + \epsilon_{ij}$$

- Fixed effects  $X_{ij}^T B$



# Linear Mixed Effects

$$y_{ij} = X_{ij}^T B + z_{ij}^T \beta_i + \epsilon_{ij}$$

- Fixed effects  $X_{ij}^T B$
- Random effects  $z_{ij}^T \beta_i$  with  $\beta_i \stackrel{iid}{\sim} N(0, \Psi)$



# Linear Mixed Effects

$$y_{ij} = X_{ij}^T B + z_{ij}^T \beta_i + \epsilon_{ij}$$

- Fixed effects  $X_{ij}^T B$
- Random effects  $z_{ij}^T \beta_i$  with  $\beta_i \stackrel{iid}{\sim} N(0, \Psi)$
- Designed to accomodate correlated data due to nested/hierarchical structure/repeated measurements



# Linear Mixed Effects

$$y_{ij} = X_{ij}^T B + z_{ij}^T \beta_i + \epsilon_{ij}$$

- Fixed effects  $X_{ij}^T B$
- Random effects  $z_{ij}^T \beta_i$  with  $\beta_i \stackrel{iid}{\sim} N(0, \Psi)$
- Designed to accomodate correlated data due to nested/hierarchical structure/repeated measurements
- students w/in schools; patients w/in hospitals



# Linear Mixed Effects

$$y_{ij} = X_{ij}^T B + z_{ij}^T \beta_i + \epsilon_{ij}$$

- Fixed effects  $X_{ij}^T B$
- Random effects  $z_{ij}^T \beta_i$  with  $\beta_i \stackrel{iid}{\sim} N(0, \Psi)$
- Designed to accomodate correlated data due to nested/hierarchical structure/repeated measurements
- students w/in schools; patients w/in hospitals
- As before not inherently Bayesian! It's just a model/likelihood specification!



# Linear Mixed Effects

$$y_{ij} = X_{ij}^T B + z_{ij}^T \beta_i + \epsilon_{ij}$$

- Fixed effects  $X_{ij}^T B$
- Random effects  $z_{ij}^T \beta_i$  with  $\beta_i \stackrel{iid}{\sim} N(0, \Psi)$
- Designed to accomodate correlated data due to nested/hierarchical structure/repeated measurements
- students w/in schools; patients w/in hospitals
- As before not inherently Bayesian! It's just a model/likelihood specification!
- If  $\theta$  is population parameters,  $\theta = (B, \Psi, \sigma^2)$ , find the marginal distribution for  $y_i$  given  $\theta$ !

