

STA 702: Random Effects

Merlise Clyde

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Building Hierarchical Models

- Models for Gaussian Data with no Covariates

$$y_{ij} \sim ? \quad i = 1, \dots, n; j = 1, \dots, n_i$$



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- structure?



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- Identifiability ?



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- Put constraints on parameters

- $\alpha = 0$
- $\beta_d = 0$
- $\sum \beta_j = 0$



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- post-processing of output



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- May want inference about the larger population and say things about future blocks!
- fixed effects do not allow us to say anything about blocks not in our sample!
- how to address this?



Random Effects

$$y_{ij} = \alpha + \beta_i + \epsilon_{ij}, \quad \epsilon_{ij} \stackrel{iid}{\sim} N(0, \sigma^2)$$
$$\beta_i \stackrel{iid}{\sim} N(0, \tau^2)$$

- random effects β_j



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- random effects β_j
 - Random effect distributions should be viewed as part of the model specification
 - We've specified the likelihood in a hierarchical manner to induce desirable structure
 - unknown parameters are population parameters α , τ and σ^2
 - Bayesians put prior distributions on α , τ and σ^2 ; frequentists don't!



Equivalent Model

$$y_i = (y_{i1}, y_{i2}, \dots, y_{in_i})$$

$$y_i \stackrel{ind}{\sim} N_{n_i} \left(\alpha \mathbf{1}_{n_i}, \begin{pmatrix} \sigma^2 + \tau & \tau & \dots & \tau \\ \tau & \ddots & & \tau \\ \vdots & & \ddots & \vdots \\ \tau & \dots & \tau & \sigma^2 + \tau \end{pmatrix} \right)$$

within-block correlation

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within-block correlation

- algorithmically we can use either the latent variable model or the collapsed (marginal) model for inferences;
- often latent variable is easier to work with!



Simple Gibbs Sampler

$$\theta = (\alpha, \tau, \sigma^2, \beta_1, \dots, \beta_n)$$

$$\alpha \sim N(\alpha_0, V_0)$$

$$\tau^{-1} \sim \text{Gamma}(a_\tau/2, b_\tau/2)$$

$$\sigma^{-2} \sim \text{Gamma}(a_\sigma/2, b_\sigma/2)$$



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Full Conditionals:

$$\begin{aligned}\alpha \mid \tau, \sigma^2, \beta_1, \dots, \beta_n &\sim N(\hat{\alpha}, \hat{V}_n) \\ \hat{V}_n &= \left(\frac{1}{V_0} + \sum_i \frac{n_i}{\sigma^2} \right)^{-1} \\ \hat{\alpha} &= \frac{\frac{\alpha_0}{V_0} + \frac{\sum_i n_i \bar{y}_i^*}{\sigma^2}}{\hat{V}_n^{-1}} \\ y_{ij}^* &\equiv y_{ij} - \beta_i \quad \bar{y}_i^* \equiv \frac{\sum_j (y_{ij} - \beta_i)}{n_i}\end{aligned}$$



Full Conditional Continued

$$\sigma^{-2} \mid \alpha, \tau, \beta_1, \dots, \beta_n \sim \text{Gamma} \left(\frac{a_\sigma + \sum_i n_i}{2}, \frac{b_\sigma + \sum_i \sum_j (y_{ij} - \alpha - \beta_i)^2}{2} \right)$$



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$$\beta_j \mid \alpha, \tau, \sigma^2 \stackrel{ind}{\sim} N(\hat{b}_i, \hat{V}_{\beta_i})$$

$$\hat{V}_{\beta_i} = \left(\frac{1}{\tau} + \frac{n_i}{\sigma^2} \right)^{-1}$$

$$\hat{b}_i = \frac{\frac{0}{\tau} + \frac{n_i \bar{y}_i^*}{\sigma^2}}{\hat{V}_{\beta_i}^{-1}}$$

$$y_{ij}^{**} \equiv y_{ij} - \alpha \quad \bar{y}_i^{**} \equiv \frac{\sum_j (y_{ij} - \alpha)}{n_i}$$



Complications Relative to Usual Regression

1. Prior Choice



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2. Mixing and its dependence on parameterization



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 2. Mixing and its dependence on parameterization
- Early recommendation after Gibbs Sampler introduced non-informative priors

$$\pi(\alpha) \propto 1$$

$$\pi(\sigma^{-2}) \sim \text{Gamma}(\epsilon/2, \epsilon/2) \quad \pi(\sigma^{-2}) \propto 1/\sigma^{-2} \text{ as } \epsilon \rightarrow 0$$

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- Are full conditionals proper ?
- Is joint posterior proper ?



MCMC and Priors

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 - joint is improper
 - MCMC won't converge to the stationary distribution (doesn't exist)
 - may not notice it!



Diffuse But Proper

$$\begin{aligned}\alpha &\sim N(0, 10^{-6}) \\ \pi(\sigma^{-2}) &\sim \text{Gamma}(10^{-6}, 10^{-6}) \\ \pi(\tau^{-1}) &\sim \text{Gamma}(10^{-6}, 10^{-6})\end{aligned}$$



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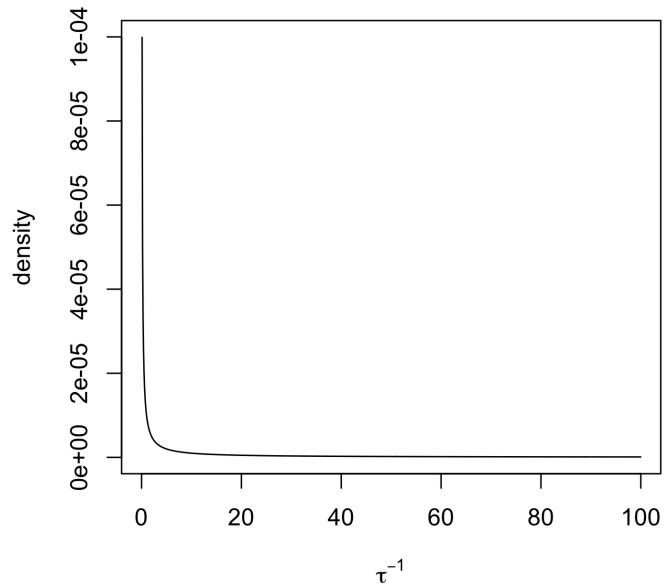
- Nearly improper priors lead to terrible performance! highly sensitive to just how vague the prior is!



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Alternative Priors

- Choose a flat or heavy tailed prior for random effect standard deviation $\tau^{1/2}$

$$\begin{array}{ll} y_{ij} = \alpha + \beta_i + \epsilon_{ij} & y_{ij} = \alpha + \lambda\eta_i + \epsilon_{ij} \\ \Leftrightarrow & \\ \beta_i \stackrel{iid}{\sim} N(0, \tau) & \eta_i \stackrel{iid}{\sim} N(0, 1) \end{array}$$



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- $\pi(\lambda) \propto 1(\lambda > 0)N(0, 1/\psi) \quad \psi \sim \text{Gamma}(\nu/2, \nu/2)$ folded t or half t



Proper Posterior

Work with

$$\pi(\mu, \tau, \sigma^2 | y) \propto \pi(\mu, \tau, \sigma^2) \prod_{i=1}^n N\left(y_i; \alpha 1_{n_i}, \begin{pmatrix} \sigma^2 + \tau & \tau & \dots & \tau \\ \tau & \ddots & & \tau \\ \vdots & & \ddots & \vdots \\ \tau & \dots & \tau & \sigma^2 + \tau \end{pmatrix}\right)$$

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- Show joint posterior is proper !



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- See Gelman 2005 discussion of Draper paper in Bayesian Analysis



Propriety

- need expression for likelihood; requires determinant and inverse of intra-class correlation matrix! Write covariance as $\sigma^2 I_{n_i} + \tau n_1 P_1$ and find spectral decomposition to provide determinant and inverse!
- integrate out α (messy)
- determine if integrals are finite (what happens at 0 and infinity ?)
- look at special case when n_i are all equal.



Bayes Estimates of Variances

- Avoids issues when estimate of variance is on the boundary of the parameter space
- Do not have to use asymptotics to construct CI!



Linear Mixed Effects

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- Fixed effects $X_{ij}^T B$
- Random effects $z_{ij}^T \beta_i$ with $\beta_i \stackrel{iid}{\sim} N(0, \Psi)$



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 - If θ is population parameters, $\theta = (B, \Psi, \sigma^2)$, find the marginal distribution for y_i given θ !

