# Lecture 13: Bayesian Multiple Testing

**Merlise Clyde** 

October 14



#### Recall normal model with

$$Y_i \mid \mu_i, \sigma^2 \stackrel{ind}{\sim} \mathsf{N}(\mu_i, \sigma^2)$$

 $lacksquare H_{0i}: \mu_i = 0 ext{ Versus } H_{1i}: \mu_i 
eq 0$ 



#### Recall normal model with

$$Y_i \mid \mu_i, \sigma^2 \stackrel{ind}{\sim} \mathsf{N}(\mu_i, \sigma^2)$$

- $lacksquare H_{0i}: \mu_i = 0 ext{ Versus } H_{1i}: \mu_i 
  eq 0$
- Spike & Slab Prior:

$$\mu_i \stackrel{iid}{\sim} \pi_0 \delta_0 + (1-\pi_0) g(\mu_i \mid 0, au, H_{i1})$$



#### Recall normal model with

$$Y_i \mid \mu_i, \sigma^2 \stackrel{ind}{\sim} \mathsf{N}(\mu_i, \sigma^2)$$

- $lacksquare H_{0i}: \mu_i = 0 ext{ Versus } H_{1i}: \mu_i 
  eq 0$
- Spike & Slab Prior:

$$\mu_i \stackrel{iid}{\sim} \pi_0 \delta_0 + (1-\pi_0) g(\mu_i \mid 0, au, H_{i1})$$

- need to specify
  - $\blacksquare$   $\pi_0$
  - **■** g

#### Recall normal model with

$$Y_i \mid \mu_i, \sigma^2 \stackrel{ind}{\sim} \mathsf{N}(\mu_i, \sigma^2)$$

- $lacksquare H_{0i}: \mu_i = 0 ext{ Versus } H_{1i}: \mu_i 
  eq 0$
- Spike & Slab Prior:

$$\mu_i \stackrel{iid}{\sim} \pi_0 \delta_0 + (1-\pi_0) g(\mu_i \mid 0, au, H_{i1})$$

- need to specify
  - $\blacksquare$   $\pi_0$
  - **■** g
- concern: is that # errors blows up with n (\$n\$ = # tests = dimension of  $\{\mu_i\}$ )



 $\pi_0=\Pr(H_{0i})$ 



$$\pi_0=\Pr(H_{0i})$$

seemingly non-informative choice?



$$\pi_0=\Pr(H_{0i})$$

seemingly non-informative choice?

$$\pi_0 = 0.5$$



$$\pi_0=\Pr(H_{0i})$$

seemingly non-informative choice?

$$\pi_0 = 0.5$$

• What does imply about the "model size" prior? The number of times  $H_{1i}$  is true?



$$\pi_0=\Pr(H_{0i})$$

seemingly non-informative choice?

$$\pi_0 = 0.5$$

- What does imply about the "model size" prior? The number of times  $H_{1i}$  is true?
- Let

$$\gamma_i = \left\{ egin{array}{l} 1 ext{ if } H_{1i} ext{ is true} \ 0 ext{ if } H_{0i} ext{ is true} \end{array} 
ight.$$



$$\pi_0=\Pr(H_{0i})$$

seemingly non-informative choice?

$$\pi_0 = 0.5$$

- What does imply about the "model size" prior? The number of times  $H_{1i}$  is true?
- Let

$$\gamma_i = \left\{ egin{aligned} 1 & ext{if } H_{1i} & ext{is true} \ 0 & ext{if } H_{0i} & ext{is true} \end{aligned} 
ight.$$

$$\gamma^{(n)} = (\gamma_1, \gamma_2, \dots, \gamma_n)^2$$

• e.g.  $\gamma^{(n)} = (0, 1, 0, 0, \dots, 1)^T$ 



$$\pi_0=\Pr(H_{0i})$$

seemingly non-informative choice?

$$\pi_0 = 0.5$$

- What does imply about the "model size" prior? The number of times  $H_{1i}$  is true?
- Let

$$\gamma_i = \left\{ egin{aligned} 1 & ext{if } H_{1i} & ext{is true} \ 0 & ext{if } H_{0i} & ext{is true} \end{aligned} 
ight.$$

$$\gamma^{(n)}=(\gamma_1,\gamma_2,\ldots,\gamma_n)^2$$

- e.g.  $\gamma^{(n)} = (0, 1, 0, 0, \dots, 1)^T$
- model size  $p_{\gamma} = \sum_{i=1}^{n} \gamma$  is the number of non-zero values



$$\pi_0=\Pr(H_{0i})$$

seemingly non-informative choice?

$$\pi_0 = 0.5$$

- What does imply about the "model size" prior? The number of times  $H_{1i}$  is true?
- Let

$$\gamma_i = \left\{ egin{array}{l} 1 ext{ if } H_{1i} ext{ is true} \ 0 ext{ if } H_{0i} ext{ is true} \end{array} 
ight.$$

$$\gamma^{(n)} = (\gamma_1, \gamma_2, \dots, \gamma_n)^2$$

- e.g.  $\gamma^{(n)} = (0, 1, 0, 0, \dots, 1)^T$
- model size  $p_{\gamma} = \sum_{i=1}^{n} \gamma$  is the number of non-zero values



• Distribution of  $p_{\gamma}$ ?

#### **Induced Distribution**

 $p_{\gamma} \sim \mathsf{Binomial}(\mathsf{n,}\ 1/2)$ 



#### **Induced Distribution**

 $p_{\gamma} \sim \mathsf{Binomial}(\mathsf{n,}\ 1/2)$ 

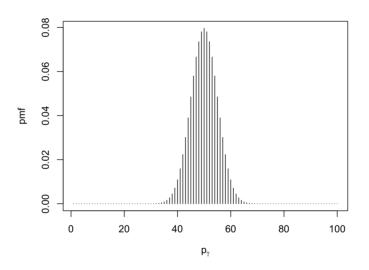
■ Expected 1/2 of the hypotheses to be true a priori



#### **Induced Distribution**

$$p_{\gamma} \sim \mathsf{Binomial}(\mathsf{n,}\ 1/2)$$

■ Expected 1/2 of the hypotheses to be true a priori





## Probabities of no or at least 1 signal

$$p_{\gamma} \sim \mathsf{Binomial}(\mathsf{n,}\ 1/2)$$

lacktriangledown probability of observing no signals  $\gamma^{(n)}=(0,0,0,\dots,0)^T$  or  $p_\gamma=0$ 

$$\Pr(p_{\gamma}=0)=\pi_0^n=0.5^n$$



## Probabities of no or at least 1 signal

$$p_{\gamma} \sim \mathsf{Binomial}(\mathsf{n,}\ 1/2)$$

lacktriangledown probability of observing no signals  $\gamma^{(n)}=(0,0,0,\dots,0)^T$  or  $p_\gamma=0$ 

$$\Pr(p_{\gamma}=0)=\pi_0^n=0.5^n$$

■ approximately 0 for large n



## Probabities of no or at least 1 signal

$$p_{\gamma} \sim \mathsf{Binomial}(\mathsf{n,}\ 1/2)$$

lacktriangledown probability of observing no signals  $\gamma^{(n)}=(0,0,0,\dots,0)^T$  or  $p_\gamma=0$ 

$$\Pr(p_{\gamma}=0)=\pi_0^n=0.5^n$$

- approximately 0 for large n
- probability of at least one signal is  $1 0.5^n \approx 1$



■ Suppose we want to fix  $\pi_0$  that protexts against Type I errors blowing up as n increases



■ Suppose we want to fix  $\pi_0$  that protexts against Type I errors blowing up as n increases

$$\Pr(p_{\gamma}=\mathbf{0}_n)=rac{1}{2}=\pi_0^n$$



• Suppose we want to fix  $\pi_0$  that protexts against Type I errors blowing up as n increases

$$\Pr(p_{\gamma}=\mathbf{0}_n)=rac{1}{2}=\pi_0^n$$

■ "Bayesian Bonferroni Prior"



■ Suppose we want to fix  $\pi_0$  that protexts against Type I errors blowing up as n increases

$$\Pr(p_{\gamma}=\mathbf{0}_n)=rac{1}{2}=\pi_0^n$$

- "Bayesian Bonferroni Prior"
- so  $\pi_0 = 0.5^{1/n}$  very close to 1! Need overwhelming evidence in the data for  $\Pr(H_{1i} \mid y^{(n)})$  to not be  $\approx 0$ !



■ Suppose we want to fix  $\pi_0$  that protexts against Type I errors blowing up as n increases

$$\Pr(p_{\gamma}=\mathbf{0}_n)=rac{1}{2}=\pi_0^n$$

- "Bayesian Bonferroni Prior"
- so  $\pi_0 = 0.5^{1/n}$  very close to 1! Need overwhelming evidence in the data for  $\Pr(H_{1i} \mid y^{(n)})$  to not be  $\approx 0$ !
- not a great idea to prespecify  $\pi_0$ !





$$Y_i \mid \mu_i, \sigma^2 \stackrel{ind}{\sim} \mathsf{N}(\mu_i, \sigma^2)$$

$$\mu_i \stackrel{iid}{\sim} \pi_0 \delta_0 + (1-\pi_0) \mathsf{N}(\mu_i;0, au)$$



$$egin{aligned} Y_i \mid \mu_i, \sigma^2 \stackrel{ind}{\sim} \mathsf{N}(\mu_i, \sigma^2) \ & \mu_i \stackrel{iid}{\sim} \pi_0 \delta_0 + (1-\pi_0) \mathsf{N}(\mu_i; 0, au) \ & \mathcal{L}(\pi_0, au) = \int_{\mathbb{R}^n} \prod_{i=1}^n \mathsf{N}(y_i; \mu_i, \sigma^2) \left\{ \pi_0 \delta_0(\mu_i) + (1-\pi_0) \mathsf{N}(\mu_i; 0, au) 
ight\} d\mu_1 \ldots d\mu_n \ & = \prod_{i=1}^n \int_{\mathbb{R}} \mathsf{N}(y_i; \mu_i, \sigma^2) \left\{ \pi_0 \delta_0(\mu_i) + (1-\pi_0) \mathsf{N}(\mu_i; 0, au) 
ight\} d\mu_i \end{aligned}$$



We could try to maximize the marginal likelihood

$$egin{aligned} Y_i \mid \mu_i, \sigma^2 \stackrel{ind}{\sim} \mathsf{N}(\mu_i, \sigma^2) \ & \mu_i \stackrel{iid}{\sim} \pi_0 \delta_0 + (1-\pi_0) \mathsf{N}(\mu_i; 0, au) \ & \mathcal{L}(\pi_0, au) = \int_{\mathbb{R}^n} \prod_{i=1}^n \mathsf{N}(y_i; \mu_i, \sigma^2) \left\{ \pi_0 \delta_0(\mu_i) + (1-\pi_0) \mathsf{N}(\mu_i; 0, au) 
ight\} d\mu_1 \ldots d\mu_n \ & = \prod_{i=1}^n \int_{\mathbb{R}} \mathsf{N}(y_i; \mu_i, \sigma^2) \left\{ \pi_0 \delta_0(\mu_i) + (1-\pi_0) \mathsf{N}(\mu_i; 0, au) 
ight\} d\mu_i \end{aligned}$$

■ Conjugate or nice setups we can integrate out  $\mu_i$  and then maximize marginal likelihood for  $\pi_0$  and  $\tau$ 



$$egin{aligned} Y_i \mid \mu_i, \sigma^2 \stackrel{ind}{\sim} \mathsf{N}(\mu_i, \sigma^2) \ & \mu_i \stackrel{iid}{\sim} \pi_0 \delta_0 + (1-\pi_0) \mathsf{N}(\mu_i; 0, au) \ & \mathcal{L}(\pi_0, au) = \int_{\mathbb{R}^n} \prod_{i=1}^n \mathsf{N}(y_i; \mu_i, \sigma^2) \left\{ \pi_0 \delta_0(\mu_i) + (1-\pi_0) \mathsf{N}(\mu_i; 0, au) 
ight\} d\mu_1 \ldots d\mu_n \ & = \prod_{i=1}^n \int_{\mathbb{R}} \mathsf{N}(y_i; \mu_i, \sigma^2) \left\{ \pi_0 \delta_0(\mu_i) + (1-\pi_0) \mathsf{N}(\mu_i; 0, au) 
ight\} d\mu_i \end{aligned}$$

- Conjugate or nice setups we can integrate out  $\mu_i$  and then maximize marginal likelihood for  $\pi_0$  and  $\tau$
- Numerical integration or EM algorithms to get  $\hat{\pi}_0^{\mathtt{EB}}$  and  $\hat{ au}^{\mathtt{EB}}$
- Clyde & George (2000) Silverman & Johnstone (2004) for wavelet regression



• introduce latent variables so that "complete" data likelihood is nice! e.g.  $\gamma$ :

$$y_i \mid \gamma_i, au \stackrel{ind}{\sim} \mathsf{N}(0,1)^{1-\gamma_i} \mathsf{N}(0,1+ au)^{\gamma_i}$$
  $\gamma_i \stackrel{iid}{\sim} \mathsf{Ber}(1-\pi_0)$ 



• introduce latent variables so that "complete" data likelihood is nice! e.g.  $\gamma$ :

$$y_i \mid \gamma_i, au \stackrel{ind}{\sim} \mathsf{N}(0,1)^{1-\gamma_i} \mathsf{N}(0,1+ au)^{\gamma_i}$$
  $\gamma_i \stackrel{iid}{\sim} \mathsf{Ber}(1-\pi_0)$ 

• Iterate: For  $t = 1, \dots$ 



• introduce latent variables so that "complete" data likelihood is nice! e.g.  $\gamma$ :

$$y_i \mid \gamma_i, au \stackrel{ind}{\sim} \mathsf{N}(0,1)^{1-\gamma_i} \mathsf{N}(0,1+ au)^{\gamma_i}$$
  $\gamma_i \stackrel{iid}{\sim} \mathsf{Ber}(1-\pi_0)$ 

- Iterate: For  $t = 1, \dots$ 
  - M-step: Solve for  $(\hat{\pi}_0^{(t)}, \hat{\tau}^{(t)}) = \arg \max \mathcal{L}(\pi_o, \tau \mid \hat{\gamma}^{(t-1)})$



• introduce latent variables so that "complete" data likelihood is nice! e.g.  $\gamma$ :

$$y_i \mid \gamma_i, au \stackrel{ind}{\sim} \mathsf{N}(0,1)^{1-\gamma_i} \mathsf{N}(0,1+ au)^{\gamma_i}$$
  $\gamma_i \stackrel{iid}{\sim} \mathsf{Ber}(1-\pi_0)$ 

- Iterate: For  $t = 1, \ldots$ 
  - M-step: Solve for  $(\hat{\pi}_0^{(t)}, \hat{\tau}^{(t)}) = \arg \max \mathcal{L}(\pi_o, \tau \mid \hat{\gamma}^{(t-1)})$
  - E-step: find the expected values of the latent sufficient statistics given the data,  $\hat{\pi}_0^{(t)}$ ,  $\hat{\tau}^{(t)}$

$$\hat{\gamma}^{(t)} = \mathsf{E}[\gamma_i \mid y, \hat{\pi}_0^{(t)}, \hat{ au}^{(t)}]$$



# **Adding Noise**

What happens to  $\hat{\pi}_0^{\text{EB}}$  as we add more and more noise to a fixed number of signals?



## **Adding Noise**

What happens to  $\hat{\pi}_0^{\text{EB}}$  as we add more and more noise to a fixed number of signals?

•  $\hat{\pi}_0^{\mathsf{EB}}$  becomes closer to one as  $n \to \infty$  for a fixed number of true signals



#### **Adding Noise**

What happens to  $\hat{\pi}_0^{\text{EB}}$  as we add more and more noise to a fixed number of signals?

- $\hat{\pi}_0^{\mathsf{EB}}$  becomes closer to one as  $n \to \infty$  for a fixed number of true signals
- This is good as it protects against Type I errors blowing up as *n* increases!



## **Adding Noise**

What happens to  $\hat{\pi}_0^{\text{EB}}$  as we add more and more noise to a fixed number of signals?

- $\hat{\pi}_0^{\mathsf{EB}}$  becomes closer to one as  $n \to \infty$  for a fixed number of true signals
- This is good as it protects against Type I errors blowing up as *n* increases!
- However it becomes more and more difficult to find the few needles in a haystack!



Choose a prior for  $\pi_0$ :  $\pi_0 \sim \mathsf{Beta}(a,b)$ 



Choose a prior for  $\pi_0$ :  $\pi_0 \sim \text{Beta}(a, b)$ 

Consider the thought experiment ...

$$\gamma^{(n)}=(?,0,,\ldots,0)^T$$

where we don't know the first hypothesis but we know that the others are all null  $\gamma_j = 0$  for j = 2, ..., n



Choose a prior for  $\pi_0$ :  $\pi_0 \sim \mathsf{Beta}(a,b)$ 

Consider the thought experiment ...

$$\gamma^{(n)}=(?,0,,\dots,0)^T$$

where we don't know the first hypothesis but we know that the others are all null  $\gamma_j = 0$  for j = 2, ..., n

lacksquare  $\gamma_i \sim \mathsf{Bernoulli}(1-\pi_0)$ 



Choose a prior for  $\pi_0$ :  $\pi_0 \sim \text{Beta}(a,b)$ 

Consider the thought experiment ...

$$\gamma^{(n)}=(?,0,,\ldots,0)^T$$

where we don't know the first hypothesis but we know that the others are all null  $\gamma_j = 0$  for j = 2, ..., n

- lacksquare  $\gamma_i \sim \mathsf{Bernoulli}(1-\pi_0)$
- lacksquare Update the prior for  $\pi_0$  to include the info  $\gamma_j=0$  for  $j=2,\ldots,n$

$$\pi(\pi_0 \mid \gamma_2, \dots, \gamma_n) \propto \pi_0^{a-1} (1-\pi_0)^{b-1} \prod_{j=2}^n \pi_0^{1-\gamma_j} (1-\pi_0)^{\gamma_j}$$

$$\pi(\pi_0 \mid \gamma_2, \dots, \gamma_n) \propto \pi_0^{a+n-1-1} (1-\pi_0)^{b-1}$$



$$\pi_0 \mid \gamma_2, \dots, \gamma_n \sim \mathsf{Beta}(\mathsf{a} + \mathsf{n} - \mathsf{1}, \, \mathsf{b})$$

with mean

$$\mathsf{E}[\pi_0 \mid \gamma_2, \dots, \gamma_n] = rac{a+n-1}{a+n-1+b}$$



$$\pi_0 \mid \gamma_2, \dots, \gamma_n \sim \mathsf{Beta}(\mathsf{a} + \mathsf{n} - \mathsf{1}, \mathsf{b})$$

with mean

$$\mathsf{E}[\pi_0 \mid \gamma_2, \dots, \gamma_n] = rac{a+n-1}{a+n-1+b}$$

• suppose a = b = 1 (Uniform prior)

$$\pi_0 \mid \gamma_2, \dots, \gamma_n \sim \mathsf{Beta}(\mathsf{a} + \mathsf{n} - \mathsf{1}, \, \mathsf{b})$$

with mean

$$\mathsf{E}[\pi_0 \mid \gamma_2, \dots, \gamma_n] = rac{a+n-1}{a+n-1+b}$$

• suppose a = b = 1 (Uniform prior)

$$\mathsf{E}[\pi_0 \mid \gamma_2, \dots, \gamma_n] = \frac{n}{n+1}$$



$$\pi_0 \mid \gamma_2, \dots, \gamma_n \sim \mathsf{Beta}(\mathsf{a} + \mathsf{n} - \mathsf{1}, \, \mathsf{b})$$

with mean

$$\mathsf{E}[\pi_0 \mid \gamma_2, \dots, \gamma_n] = rac{a+n-1}{a+n-1+b}$$

• suppose a = b = 1 (Uniform prior)

$$\mathsf{E}[\pi_0 \mid \gamma_2, \dots, \gamma_n] = rac{n}{n+1}$$

■ implies probability of  $H_{01} \to 1$  and  $H_{11} \to 0$  as  $n \to \infty$  borrowing strength from other nulls



$$\pi_0 \mid \gamma_2, \dots, \gamma_n \sim \mathsf{Beta}(\mathsf{a} + \mathsf{n} - \mathsf{1}, \, \mathsf{b})$$

with mean

$$\mathsf{E}[\pi_0 \mid \gamma_2, \dots, \gamma_n] = rac{a+n-1}{a+n-1+b}$$

• suppose a = b = 1 (Uniform prior)

$$\mathsf{E}[\pi_0 \mid \gamma_2, \dots, \gamma_n] = rac{n}{n+1}$$

- implies probability of  $H_{01} \to 1$  and  $H_{11} \to 0$  as  $n \to \infty$  borrowing strength from other nulls
- Multiplicity adjustment as in the EB case

$$\pi_0 \mid \gamma_2, \dots, \gamma_n \sim \mathsf{Beta}(\mathsf{a} + \mathsf{n} - \mathsf{1}, \, \mathsf{b})$$

with mean

$$\mathsf{E}[\pi_0 \mid \gamma_2, \dots, \gamma_n] = rac{a+n-1}{a+n-1+b}$$

• suppose a = b = 1 (Uniform prior)

$$\mathsf{E}[\pi_0 \mid \gamma_2, \dots, \gamma_n] = \frac{n}{n+1}$$

- implies probability of  $H_{01} \to 1$  and  $H_{11} \to 0$  as  $n \to \infty$  borrowing strength from other nulls
- Multiplicity adjustment as in the EB case
- Scott & Berger (2006 JSPI, 2010 AoS) show that above framework protects against increasing Type I errors with n; We also get FDR control automatically



Exercise: If  $p_\gamma \mid \pi_0 \sim \mathsf{Binomial}(n, 1 - \pi_0)$  and  $\pi_0 \sim \mathsf{Beta}(1, 1)$ 

--

• What is the probability that  $p_{\gamma}=0$ 



Exercise: If  $p_\gamma \mid \pi_0 \sim \mathsf{Binomial}(n, 1 - \pi_0)$  and  $\pi_0 \sim \mathsf{Beta}(1, 1)$ 

--

- What is the probability that  $p_{\gamma}=0$
- What is the probability that  $p_{\gamma} = n$



Exercise: If  $p_{\gamma} \mid \pi_0 \sim \mathsf{Binomial}(n, 1 - \pi_0)$  and  $\pi_0 \sim \mathsf{Beta}(1, 1)$ 

\_\_

- What is the probability that  $p_{\gamma} = 0$
- What is the probability that  $p_{\gamma} = n$
- What is the distribution of  $p_{\gamma}$ ?

This is a Beta-Binomial and in the special case a = b = 1 this is a discrete uniform on model size!

Exercise: If  $p_{\gamma} \mid \pi_0 \sim \mathsf{Binomial}(n, 1 - \pi_0)$  and  $\pi_0 \sim \mathsf{Beta}(1, 1)$ 

--

- What is the probability that  $p_{\gamma} = 0$
- What is the probability that  $p_{\gamma} = n$
- What is the distribution of  $p_{\gamma}$ ?

This is a Beta-Binomial and in the special case a = b = 1 this is a discrete uniform on model size!

Bottomline: We need to "learn" key parameters in our hierarchical prior or the magic doesn't work! Borrowing comes through using all the data to inform about "global" parameters in the prior, in this case  $\pi_0$ 



■ Joint posterir distribution of  $\mu_1, \ldots, \mu_n$  averaged over hypotheses "Model averaging"



- Joint posterir distribution of  $\mu_1, \ldots, \mu_n$  averaged over hypotheses "Model averaging"
  - distribution is a spike at 0 and continous distribution



- Joint posterir distribution of  $\mu_1, \ldots, \mu_n$  averaged over hypotheses "Model averaging"
  - distribution is a spike at 0 and continous distribution
- select a hypothesis



- Joint posterir distribution of  $\mu_1, \ldots, \mu_n$  averaged over hypotheses "Model averaging"
  - distribution is a spike at 0 and continous distribution
- select a hypothesis
- report posterior (summaries) conditional on a hypothesis



- Joint posterir distribution of  $\mu_1, \ldots, \mu_n$  averaged over hypotheses "Model averaging"
  - distribution is a spike at 0 and continous distribution
- select a hypothesis
- report posterior (summaries) conditional on a hypothesis
- issue winner's curse!
- need to have coherent conditional inference given that you selected a hypothesis.



- Joint posterir distribution of  $\mu_1, \ldots, \mu_n$  averaged over hypotheses "Model averaging"
  - distribution is a spike at 0 and continous distribution
- select a hypothesis
- report posterior (summaries) conditional on a hypothesis
- issue winner's curse!
- need to have coherent conditional inference given that you selected a hypothesis.
- Don't report selected hypotheses but report results under model averaging!



$$\mu_i \stackrel{iid}{\sim} \pi_0 \delta_0 + (1-\pi_0) g(\mu_i \mid 0, au, H_{i1})$$



$$\mu_i \stackrel{iid}{\sim} \pi_0 \delta_0 + (1-\pi_0) g(\mu_i \mid 0, au, H_{i1})$$

• growing literature on posterior contraction in high dimensional settings as  $n \to \infty$  with "sparse signals"



$$\mu_i \stackrel{iid}{\sim} \pi_0 \delta_0 + (1-\pi_0) g(\mu_i \mid 0, au, H_{i1})$$

- growing literature on posterior contraction in high dimensional settings as  $n \to \infty$  with "sparse signals"
- posterior  $\pi(\mu^{(n)}) \mid y^{(n)})$



$$\mu_i \stackrel{iid}{\sim} \pi_0 \delta_0 + (1-\pi_0) g(\mu_i \mid 0, au, H_{i1})$$

- growing literature on posterior contraction in high dimensional settings as  $n \to \infty$  with "sparse signals"
- posterior  $\pi(\mu^{(n)}) \mid y^{(n)})$

Want

$$\Pr(\mu^{(n)} \in \mathcal{N}_{\epsilon_n}(\mu^{(n)}_{ ext{o}}) \mid y^{(n)}) o ext{1}$$



$$\mu_i \stackrel{iid}{\sim} \pi_0 \delta_0 + (1-\pi_0) g(\mu_i \mid 0, au, H_{i1})$$

- growing literature on posterior contraction in high dimensional settings as  $n \to \infty$  with "sparse signals"
- posterior  $\pi(\mu^{(n)}) \mid y^{(n)})$

Want

$$\Pr(\mu^{(n)} \in \mathcal{N}_{\epsilon_n}(\mu^{(n)}_{\mathrm{o}}) \mid y^{(n)}) o \mathtt{1}$$

■ assume that there are s signals (fixed or growing slowly)



$$\mu_i \stackrel{iid}{\sim} \pi_0 \delta_0 + (1-\pi_0) g(\mu_i \mid 0, au, H_{i1})$$

- growing literature on posterior contraction in high dimensional settings as  $n \to \infty$  with "sparse signals"
- posterior  $\pi(\mu^{(n)}) \mid y^{(n)})$

#### Want

$$\Pr(\mu^{(n)} \in \mathcal{N}_{\epsilon_n}(\mu^{(n)}_{ ext{o}}) \mid y^{(n)}) 
ightarrow 1$$

- assume that there are s signals (fixed or growing slowly)
- signal values are bounded away from zero



$$\mu_i \stackrel{iid}{\sim} \pi_0 \delta_0 + (1-\pi_0) g(\mu_i \mid 0, au, H_{i1})$$

- growing literature on posterior contraction in high dimensional settings as  $n \to \infty$  with "sparse signals"
- posterior  $\pi(\mu^{(n)}) \mid y^{(n)})$

#### Want

$$\Pr(\mu^{(n)} \in \mathcal{N}_{\epsilon_n}(\mu^{(n)}_{\scriptscriptstyle{0}}) \mid y^{(n)}) 
ightarrow 1$$

- assume that there are s signals (fixed or growing slowly)
- signal values are bounded away from zero
- Want the posterior under the Spike and Slab prior to concentrate on this neighborhood (ie. probability 1)



$$\mu_i \stackrel{iid}{\sim} \pi_0 \delta_0 + (1-\pi_0) g(\mu_i \mid 0, au, H_{i1})$$

- growing literature on posterior contraction in high dimensional settings as  $n \to \infty$  with "sparse signals"
- posterior  $\pi(\mu^{(n)}) \mid y^{(n)})$

#### Want

$$\Pr(\mu^{(n)} \in \mathcal{N}_{\epsilon_n}(\mu^{(n)}_{ ext{o}}) \mid y^{(n)}) o extbf{1}$$

- assume that there are s signals (fixed or growing slowly)
- signal values are bounded away from zero
- Want the posterior under the Spike and Slab prior to concentrate on this neighborhood (ie. probability 1)



active area of research!