Lecture 17: Bayesian Variable Selection and Model Averaging

STA702

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Diabetes Example

```
1 set.seed(8675309)
2 source("yX.diabetes.train.txt")
3 diabetes.train = as.data.frame(diabetes.train)
4 source("yX.diabetes.test.txt")
5 diabetes.test = as.data.frame(diabetes.test)
6 colnames(diabetes.test)[1] = "y"
7
8 str(diabetes.train)
```

```
'data.frame':
                               65 variables:
                342 obs. of
                  -0.0147 -1.0005 -0.1444 0.6987 -0.2222 ...
$ у
          : num
$ age
                  0.7996 -0.0395 1.7913 -1.8703 0.113 ...
          : num
                  1.064 -0.937 1.064 -0.937 -0.937 ...
$ sex
          : num
$ bmi
                  1.296 - 1.081 0.933 - 0.243 - 0.764 \dots
          : num
$ map
                  0.459 - 0.553 - 0.119 - 0.77 0.459 \dots
          : num
$ tc
                  -0.9287 -0.1774 -0.9576 0.256 0.0826
          : num
$ 1d1
                 -0.731 - 0.402 - 0.718 0.525 0.328 \dots
          : num
                  -0.911 1.563 -0.679 -0.757 0.171 ...
$ hdl
          : num
$ tch
                  -0.0544 -0.8294 -0.0544 0.7205 -0.0544 ...
          : num
$ ltg
                  0.4181 - 1.4349 \ 0.0601 \ 0.4765 - 0.6718 \dots
          : num
$ glu
                  -0.371 - 1.936 - 0.545 - 0.197 - 0.979 \dots
          : num
$ age^2
                  _0.312 _0.867 1.925 2.176 _0.857 ...
          • niim
```

MCMC with BAS

```
1 library(BAS)
 diabetes.bas = bas.lm(y ~ ., data=diabetes.train,
2
                         prior = "JZS",
3
                         method="MCMC",
4
5
                         n.models = 10000,
                         MCMC.iterations=500000,
7
                         thin = 10,
                         initprobs="eplogp",
8
9
                         force.heredity=FALSE)
```

```
user system elapsed 19.523 0.951 20.530
```

- [1] "number of unique models 5008"
- increase MCMC.iterations?
- · check diagnostics

Estimates of Posterior Probabilities

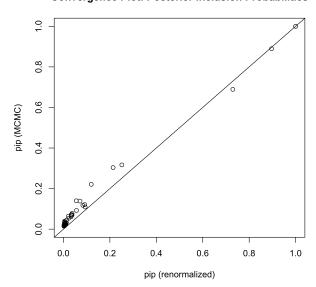
- ullet relative frequencies $\hat{P}_{RF}(oldsymbol{\gamma}\mid \mathbf{Y})=rac{\# ext{ times }oldsymbol{\gamma}\in S}{S}$
 - lacksquare Ergodic average converges to $p(oldsymbol{\gamma} \mid \mathbf{Y})$ as $S o \infty$
 - asymptoptically unbaised
- renormalized posterior probabilities $\hat{P}_{RN}(m{\gamma}\mid \mathbf{Y}) = rac{p(\mathbf{Y}|m{\gamma})p(m{\gamma})}{\sum_{m{\gamma}\in S}p(\mathbf{Y}|m{\gamma})p(m{\gamma})}$
 - also asymptoptically unbaised
 - Fisher consistent (e.g if we happen to enumerate all models in S iterations we recover the truth)
- if we run long enough the two should agree
- also look at other summaries i.e posterior inclusion probabilities

$$\hat{p}(\gamma_j = 1 \mid \mathbf{Y}) = \sum_S \gamma_j \hat{P}(oldsymbol{\gamma} \mid \mathbf{Y})$$

Diagnostic Plot

diagnostics(diabetes.bas, type="pip")

Convergence Plot: Posterior Inclusion Probabilities



• model probabilities converge much slower!

Out of Sample Prediction

- ullet What is the optimal value to predict $\mathbf{Y}^{ ext{test}}$ given \mathbf{Y} under squared error?
- ullet Iterated expectations leads to BMA for ${\sf E}[{f Y}^{
 m test} \mid {f Y}]$
- Prediction under model averaging

$$\hat{Y} = \sum_{S} (\hat{lpha}_{oldsymbol{\gamma}} + \mathbf{X}^{ ext{test}}_{oldsymbol{\gamma}} \hat{oldsymbol{eta}}_{oldsymbol{\gamma}}) \hat{p}(oldsymbol{\gamma} \mid \mathbf{Y})$$

```
pred.bas = predict(diabetes.bas,
newdata=diabetes.test,
sestimator="BMA",
se=TRUE)
mean((pred.bas$fit- diabetes.test$y)^2)
```

[1] 0.4558026

Credible Intervals & Coverage

• posterior predictive distribution

$$p(\mathbf{y}^{ ext{test}} \mid \mathbf{y}) = \sum_{m{\gamma}} p(\mathbf{y}^{ ext{test}} \mid \mathbf{y}, m{\gamma}) p(m{\gamma} \mid \mathbf{y})$$

- ullet integrate out lpha and $oldsymbol{eta}_{oldsymbol{\gamma}}$ to get a normal predictive given ϕ and $oldsymbol{\gamma}$ (and $oldsymbol{ ext{y}}$)
- integrate out ϕ to get a t distribution given γ and \mathbf{y}
- credible intervals via sampling
 - sample a model from $p(\gamma \mid \mathbf{y})$
 - ullet conditional on a model sample $y \sim p(\mathbf{y}^{ ext{test}} \mid \mathbf{y}, oldsymbol{\gamma})$
 - compute quantiles from sammple y

```
1 ci.bas = confint(pred.bas);
2 coverage = mean(diabetes.test$y > ci.bas[,1] & diabetes.test$y < c
3 coverage</pre>
```

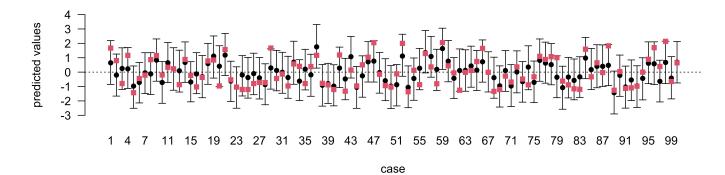
[1] 0.99

95% Prediction intervals

1 plot(ci.bas)

NULL

1 points(diabetes.test\$y, col=2, pch=15)



Selection and Prediction

• BMA - optimal for squared error loss Bayes

$$\mathsf{E}[\|\mathbf{Y}^{\text{test}} - a\|^2 \mid \mathbf{y}] = \mathsf{E}[\|\mathbf{Y}^{\text{test}} - \mathsf{E}[\mathbf{Y}^{\text{test}} \mid \mathbf{y}]\|^2 \mid \mathbf{y}] + \|\mathsf{E}[\mathbf{Y}^{\text{test}} \mid \mathbf{y}] - a\|^2$$

- What if we want to use only a single model for prediction under guared error loss?
- HPM: Highest Posterior Probability model is optimal for selection, but not prediction
- MPM: Median Probabilty model (select model where PIP > 0.5) (optimal under certain conditions; nested models)
- BPM: Best Probability Model Model closest to BMA under loss (usually includes more predictors than HPM or MPM)

Example

```
pred.bas = predict(diabetes.bas,
newdata=diabetes.test,
sestimator="BPM",
se=TRUE)
#MSE
mean((pred.bas$fit- diabetes.test$y)^2)
```

[1] 0.4740667

[1] 0.98

Theory

- consistency of mixtures of g-priors
- desire that posterior probability of model goes to 1 as $n \to \infty$
 - does not alwyas hold if the null model is true (may be highest posterior probability model)
 - ullet need prior on g to depend on n
 - asymptotically BMA collapses to the true model
- other quantities may converge i.e. posterior mean
- what if the true model is not in Γ ? What can we say?
 - M-closed; BMA converges to the model that is "closest" to the truth in Kullback-Leibler divergence
 - *M*-complete
 - **M**-open
 - latter is related to "stacking" which is a frequentist method of ensemble learning using cross-validation; see Clyde & Iversen (2013) for the curious

Summary

- Choice of prior on β_{γ}
 - orthogonally invariant priors multivariate Spike & Slab
 - products of independent Spike & Slab priors
 - non-semi-conjugate
- priors on the models (sensitivity)
- computation (MCMC, "stochastic search", variational, orthogonal data augmentation, reversible jump-MCMC)
- posterior summaries select a model or "average" over all models

Other aspects of model selection?

- ullet transformations of Y
- functions of X: interactions or nonlinear functions such as splines kernels
- choice of error distribution