

Regression Analysis

Linear model

Recall a linear model :

$$\mathbf{Y} = \mathbf{X}\boldsymbol{\beta} + \boldsymbol{\epsilon},$$

where $\mathbb{E}[\boldsymbol{\epsilon}] = 0$, $\text{Var}(\boldsymbol{\epsilon}) = \sigma^2 \mathbf{I}_n$, and $\mathbf{X} = [\mathbf{X}_1, \dots, \mathbf{X}_p]$ is a design matrix which we assume to be of full rank.

Model Diagnostics

Having fitted a classical linear model, we assess the validity of the model using diagnostic tools.

- Examination of the model assumptions
 - 1. linearity of the predictors, 2. constant variance, 3. uncorrelated, and 4. normally distributed errors
- Outliers
 - "unusual" points in feature or response spaces
- Collinearity
 - highly correlated predictors

Residuals

Recall, for each $i = 1, \dots, n$, the i th residual is defined as

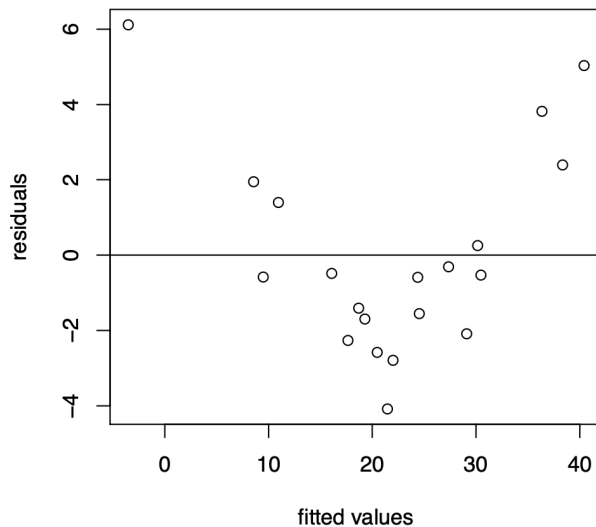
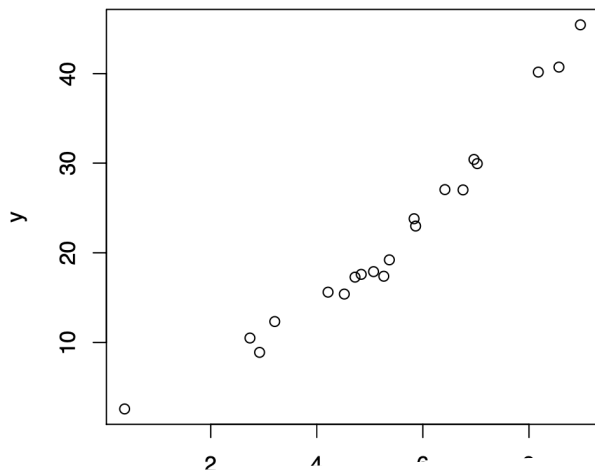
$$\hat{\epsilon}_i = \mathbf{Y}_i - \widehat{\mathbf{Y}}_i,$$

where $\widehat{\mathbf{Y}}_i = \mathbf{x}_i^\top \widehat{\boldsymbol{\beta}}$

Regression Diagnostics

1. linearity of the predictors

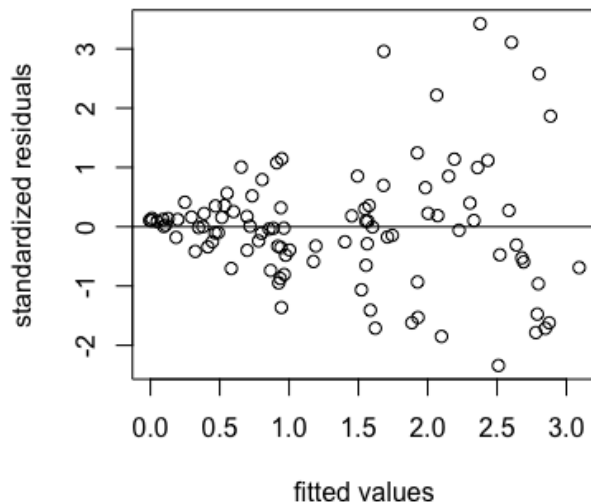
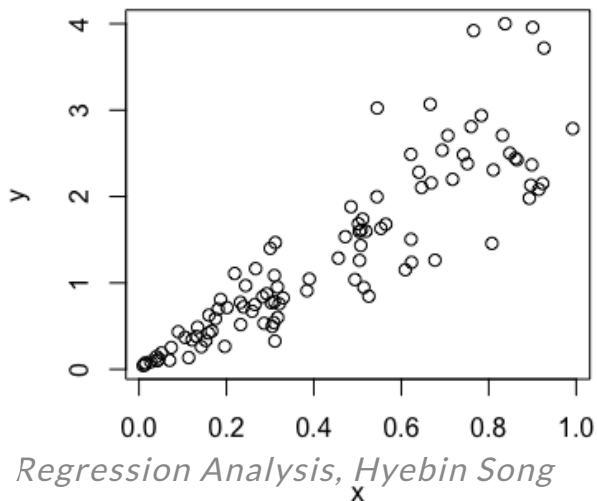
- plot $\hat{\epsilon}_i$ against the fitted values \hat{Y}_i .
- we expect a plot with **no discernible trend or pattern**



Regression Diagnostics

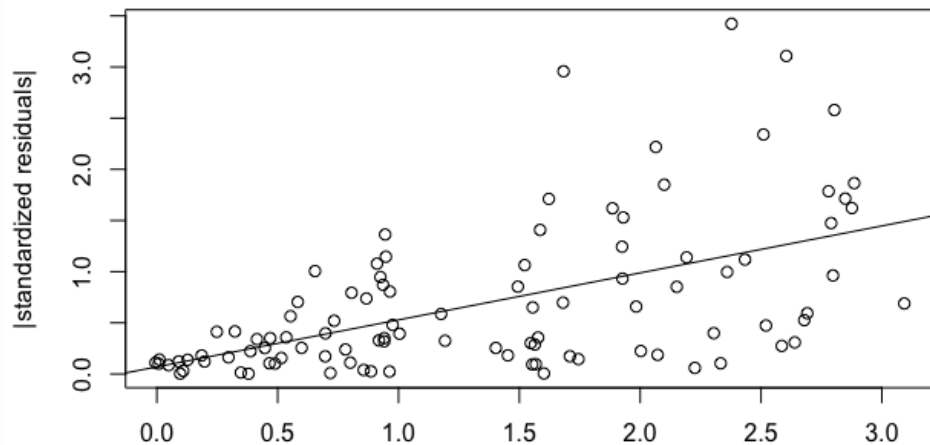
2. heteroscedasticity

- plot the standardized residuals r_i against the fitted values (or predictors)
- we expect a **"horizontal band" centered at zero.**



2. heteroscedasticity

- we may also plot $|r_i|$ against the fitted values \widehat{Y}_i or r_i^2 against the fitted values \widehat{Y}_i .
 - If we add a least-squares line to this plot, we see whether there is any tendency for $|r_i|$ to increase or decrease with \widehat{Y}_i .



Regression Diagnostics

3. Nonindependence of the error terms

- Possible forms of nonindependence:
 - Observations collected over time and/or across space.
 - Study done on sets of related subjects
- Correlated errors may indicate misspecification (especially through omitted explanatory variables and unnoticed nonlinearity). We should first examine whether model specification can be improved.

3. Nonindependence of the error terms

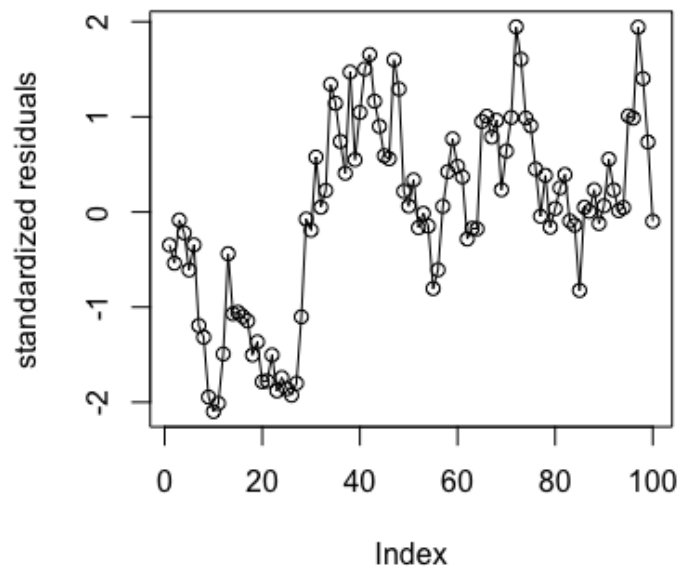
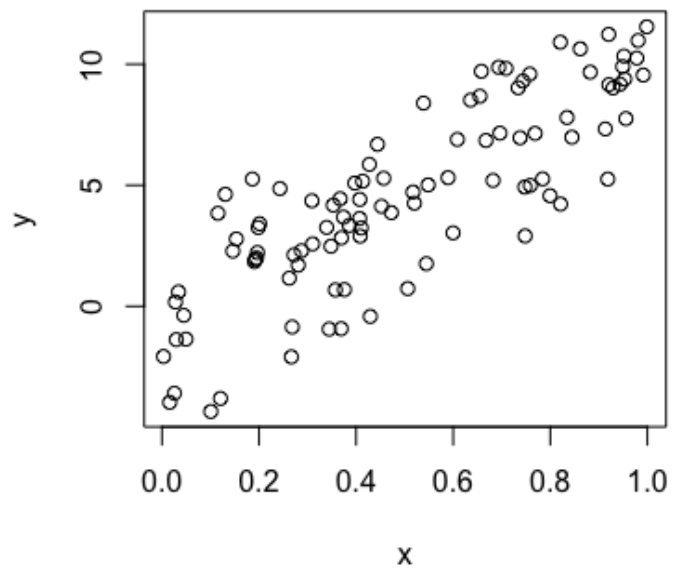
Example: Corn Yield

For $i = 1, \dots, n$,

- i = the index of a $2m^2$ patch planted to corn and the patches are arranged in a long line at the edge of a field.
- x_i = the amount of fertilizer applied to the i th patch.
- Y_i = the corn yield in the i th patch.

3. Nonindependence of the error terms

Example: Corn Yield



Regression Diagnostics

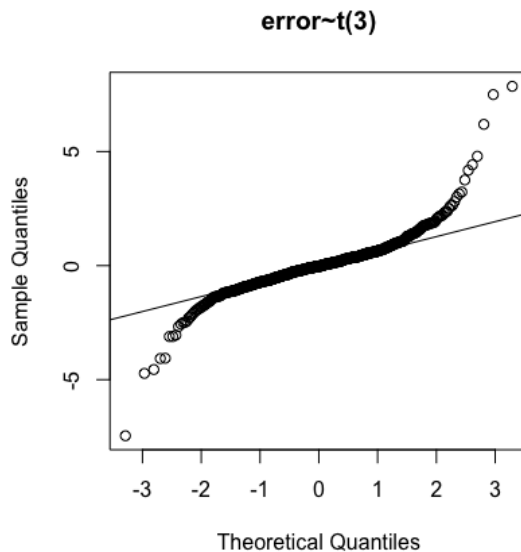
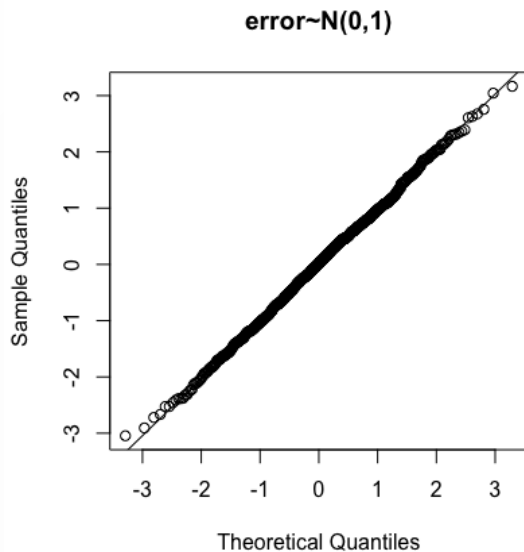
4. Nonnormality of error terms

- Normal QQ (quantile-quantile) plot of the standardized residuals.
 - **Idea:** if two distributions are similar, their quantiles should be similar.
 - We compare empirical quantiles of the standardized residuals with theoretical quantiles from $N(0,1)$
 - If the residuals are approximately normal, the normal QQ plot should be **approximately linear**

4. Nonnormality of error terms

Example: two Normal QQ plots for the standardized residuals

- Model1: $Y_i = \mathbf{x}_i^\top \boldsymbol{\beta} + \epsilon_i, \epsilon_i \sim N(0, 3), \text{ iid}$
- Model2: $Y_i = \mathbf{x}_i^\top \boldsymbol{\beta} + \epsilon_i, \epsilon_i \sim t(3), \text{ iid}$



Types of outliers

- An **outlier** is a data point which comes from a different distribution than the rest of the data
 - a big difference between the explanatory vector \mathbf{x}_i for the i th case and the rest of the \mathbf{x} -data : "outlier in the x-direction" or "high-leverage point"
 - a large difference between the response Y_i and the mean $\mathbf{x}_i^\top \boldsymbol{\beta}$: "outlier in the y-direction", "error outlier", or "outlier"
- We say a point is a **high-influence point** if the regression result changes markedly after refitting without the observation of interest

Types of outliers

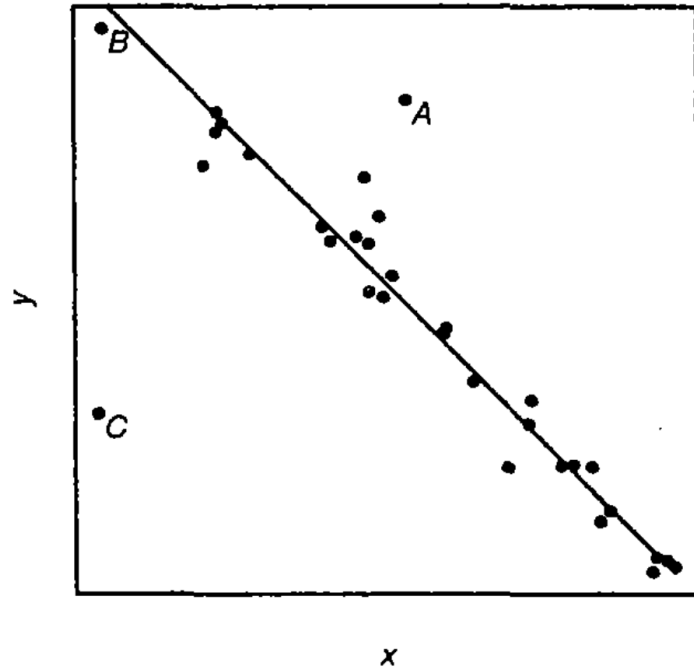


Figure credit: SL 9.1

Outliers (in the y-direction)

Say Y_n is an outlier. That is, we assume, $Y_i \sim N(\mathbf{x}_i^\top \boldsymbol{\beta}, \sigma^2)$, $i = 1, 2, \dots, n - 1$ but $Y_n \sim N(\Delta + \mathbf{x}_n^\top \boldsymbol{\beta}, \sigma^2)$ for some $\Delta \neq 0$.

1. diagnose by looking at absolute values of standardized/studentized residuals
 - Recall both standardized and studentized residuals approximately follow $N(0, 1)$ distribution.
 - Any $|r_i| \geq 3$ or $|r_i^*| \geq 3$ are potential outliers

Outliers (in the y-direction)

Remarks

- If there are outliers detected, do not just throw them away, but search for an explanation (an effect not modeled, error in data entry, etc.)
- Not all outliers have a very strong influence on the fitted regression mean. However, an observation which is both outlier and high-leverage point may have a strong effect on the response mean.
- In the presence of an outlier, estimate model parameters **with** and **without** the outliers. Report both results.

Cook's Distance

Recall a $100(1 - \alpha)\%$ confidence ellipsoid for β is

$$\left\{ \mathbf{b} : (\mathbf{b} - \widehat{\beta})^\top \mathbf{X}^\top \mathbf{X} (\mathbf{b} - \widehat{\beta}) \leq pS^2 F_{p, n-p, \alpha} \right\}$$

Cook [1977] suggested measuring the distance of $\widehat{\beta}_{(i)}$ from $\widehat{\beta}$ by using the measure

$$D_i = \frac{(\widehat{\beta}_{(i)} - \widehat{\beta})^\top \mathbf{X}^\top \mathbf{X} (\widehat{\beta}_{(i)} - \widehat{\beta})}{pS^2}$$

Cook's Distance

Note,

$$D_i = \frac{(\widehat{\boldsymbol{\beta}}_{(i)} - \widehat{\boldsymbol{\beta}})^\top \mathbf{X}^\top \mathbf{X} (\widehat{\boldsymbol{\beta}}_{(i)} - \widehat{\boldsymbol{\beta}})}{pS^2} = \frac{\|\widehat{\mathbf{Y}}_{(i)} - \widehat{\mathbf{Y}}\|^2}{pS^2}$$

- Some recommend considering points for which $D_i > 1$ to be influential. Others suggest points for which $D_i > F_{p, n-p, .1}$

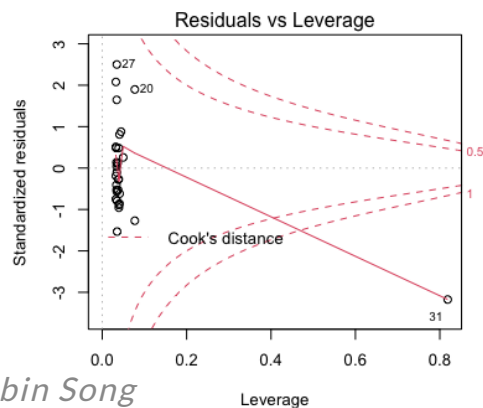
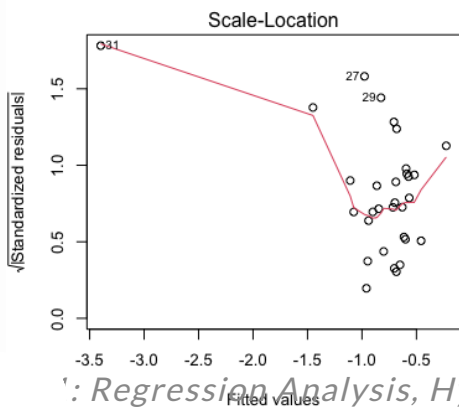
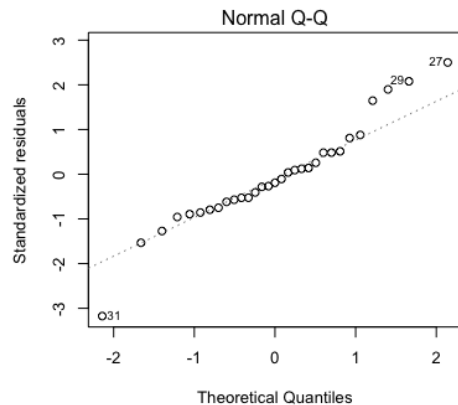
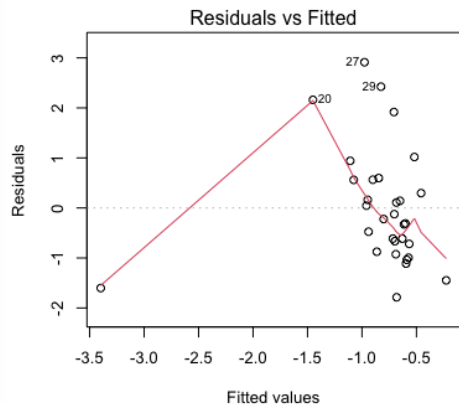
We can show that

$$D_i = r_i^2 \frac{\mathbf{H}_{ii}}{p(1 - \mathbf{H}_{ii})}$$

- captures the distance of the i th observation from the other points in both x and y directions

Example `plot(lm(fit))`

$\text{lm}(y \sim x)$



Colinearity

One important assumption in the classical linear model is that \mathbf{X} is full rank. In practice, however, the columns of \mathbf{X} could be *almost* linearly dependent or colinear.

- In such case, $\mathbf{X}^\top \mathbf{X}$ is close to singular, i.e., the smallest eigenvalue of the matrix $\mathbf{X}^\top \mathbf{X} \approx 0$
- Since $\hat{\boldsymbol{\beta}} \sim N_p(\boldsymbol{\beta}, \sigma^2(\mathbf{X}^\top \mathbf{X})^{-1})$, the precision of $\hat{\boldsymbol{\beta}}$ is determined by $(\mathbf{X}^\top \mathbf{X})^{-1}$
 - near collinearity = large variances of the estimated coefficients

Remedies

Examination of the model assumptions → [Revise a model assumption](#)

1. linearity of the predictors

- transform the response or predictors
- include additional predictors into the model to account for the non-linear relationships
- non-parametric regression (e.g., polynomial regression, splines)

2. constant variance

- variable transformation (choose h so that $\text{Var}(h(\mathbf{Y})) \approx \text{const}$)
- weighted least squares

Remedies

Examination of the model assumptions → [Revise a model assumption](#)

3. uncorrelated errors

- model error variances ("general" linear model)

4. normally distributed errors

- mild non-normality can be safely ignored (especially with light-tailed distributions)
- for heavy-tailed distributions, base the inference on the assumption of another distribution, or use resampling methods for the inference

Remedies

Outliers, leverage points, and influential points → [Examine each point](#)

- do not throw them away, but search for an explanation
- estimate model parameters **with** and **without** the observations of interest. Report both results.

Remedies

- Collinearity → Revise \mathbf{X} or use shrinkage methods
 - drop variables, ideally based on scientific knowledge.
 - obtain more data if possible.
 - create composite variables of collinear predictors, e.g., using a principle component analysis.
 - use regularization techniques.