Bayesian Statistics I

Lecture 2 - Normal and Poisson data. Prior elicitation.

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Lecture overview

- The Normal model with known variance
- The Poisson model
- Conjugate priors
- Prior elicitation
- Jeffreys' prior

Normal data, known variance - uniform prior

Model

$$x_1, ..., x_n | \theta, \sigma^2 \stackrel{iid}{\sim} N(\theta, \sigma^2).$$

Prior

$$p(\theta) \propto c$$
 (a constant)

Likelihood

$$p(x_1, ..., x_n | \theta, \sigma^2) = \Pi_{i=1}^n (2\pi\sigma^2)^{-1/2} \exp\left[-\frac{1}{2\sigma^2} (x_i - \theta)^2\right]$$

$$\propto \exp\left[-\frac{1}{2(\sigma^2/n)} (\theta - \bar{x})^2\right].$$

Posterior

$$\theta | x_1, ..., x_n \sim N(\bar{x}, \sigma^2/n)$$

Normal data, known variance - normal prior

Prior

$$\theta \sim N(\mu_0, \tau_0^2)$$

Posterior

$$p(\theta|x_1,...,x_n) \propto p(x_1,...,x_n|\theta,\sigma^2)p(\theta)$$

$$\propto N(\theta|\mu_n,\tau_n^2),$$

where

$$\frac{1}{\tau_n^2} = \frac{n}{\sigma^2} + \frac{1}{\tau_0^2},$$

$$u_n = w\bar{x} + (1 - w)u_0,$$

and

$$W = \frac{\frac{n}{\sigma^2}}{\frac{n}{\sigma^2} + \frac{1}{\tau_0^2}}.$$

Normal data, known variance - normal prior

$$\theta \sim N(\mu_0, \tau_0^2) \stackrel{x_1, \dots, x_n}{\Longrightarrow} \theta | x \sim N(\mu_n, \tau_n^2).$$

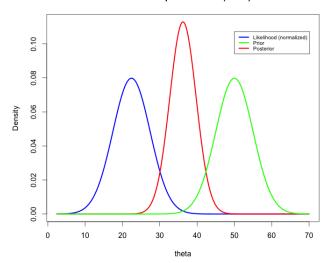
Posterior precision = Data precision + Prior precision

Posterior mean =

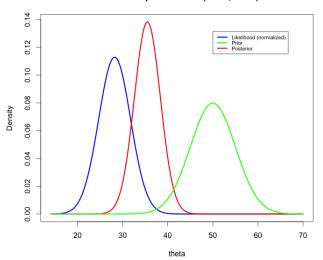
$$\frac{\text{Data precision}}{\text{Posterior precision}} \left(\text{Data mean} \right) + \frac{\text{Prior precision}}{\text{Posterior precision}} \left(\text{Prior mean} \right)$$

- **Data**: x = (22.42, 34.01, 35.04, 38.74, 25.15) Mbit/sec.
- Model: $X_1, ..., X_5 \sim N(\theta, \sigma^2)$.
- Assume $\sigma = 5$ (measurements can vary ± 10 MBit with 95% probability)
- My prior: $\theta \sim N(50, 5^2)$.

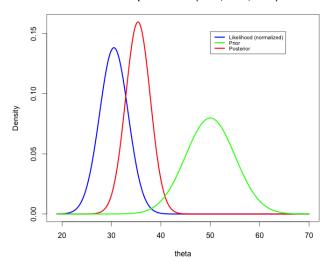
Download speed data: x=(22.42)



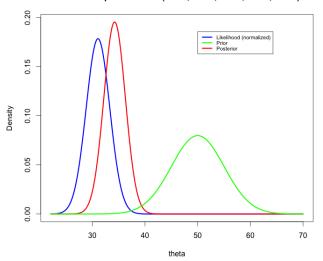
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Download speed data: x=(22.42, 34.01, 35.04)

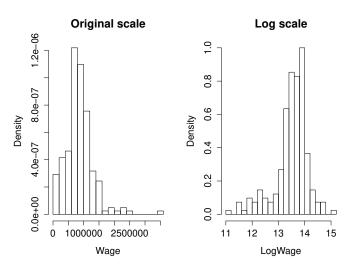


Download speed data: x=(22.42, 34.01, 35.04, 38.74, 25.15)



Canadian wages data

Data on wages for 205 Canadian workers.



Canadian wages

Model

$$X_1, ..., X_n | \theta \sim N(\theta, \sigma^2), \ \sigma^2 = 0.4$$

Prior

$$\theta \sim \textit{N}(\mu_0, \tau_0^2),~\mu_0 = 12~\text{and}~\tau_0 = 10$$

Posterior

$$\theta|x_1,...,x_n \sim N(\mu_n,\tau_n^2)$$
,

where $\mu_n = w\bar{x} + (1 - w)\mu_0$.

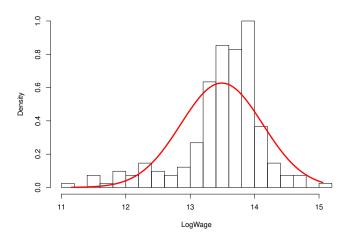
For the Canadian wage data:

$$w = \frac{\sigma^{-2}n}{\sigma^{-2}n + \tau_0^{-2}} = \frac{2.5 \cdot 205}{2.5 \cdot 205 + 1/100} = 0.999.$$

$$\mu_n = w\bar{x} + (1 - w)\mu_0 = 0.999 \cdot 13.489 + (1 - 0.999) \cdot 12 \approx 13.489$$

$$\tau_n^2 = (2.5 \cdot 205 + 1/100)^{-1} = 0.00195$$

Canadian wages data - model fit



Poisson model

Model

$$y_1, ..., y_n | \theta \stackrel{iid}{\sim} Pois(\theta)$$

Poisson distribution

$$p(y) = \frac{\theta^y e^{-\theta}}{y!}$$

Likelihood from iid Poisson sample $y = (y_1, ..., y_n)$

$$p(y|\theta) = \left[\prod_{i=1}^{n} p(y_i|\theta)\right] \propto \theta^{\left(\sum_{i=1}^{n} y_i\right)} \exp(-\theta n),$$

Prior

$$p(\theta) \propto \theta^{\alpha - 1} \exp(-\theta \beta) \propto Gamma(\alpha, \beta)$$

which contains the info: $\alpha - 1$ counts in β observations.



Poisson model, cont.

Posterior

$$p(\theta|y_1, ..., y_n) \propto \left[\prod_{i=1}^n p(y_i|\theta)\right] p(\theta)$$

$$\propto \theta^{\sum_{i=1}^n y_i} \exp(-\theta n) \theta^{\alpha-1} \exp(-\theta \beta)$$

$$= \theta^{\alpha + \sum_{i=1}^n y_i - 1} \exp[-\theta (\beta + n)],$$

which is proportional to the $Gamma(\alpha + \sum_{i=1}^{n} y_i, \beta + n)$ distribution.

■ Prior-to-Posterior mapping

Model:
$$y_1, ..., y_n | \theta \stackrel{iid}{\sim} Pois(\theta)$$

Prior: $\theta \sim Gamma(\alpha, \beta)$

Posterior:
$$\theta|y_1,...,y_n \sim Gamma(\alpha + \sum_{i=1}^n y_i, \beta + n)$$
.

Example - Number of bids in eBay auctions

Data:

- Number of placed bids in n = 1000 eBay coin auctions.
- ▶ Sum of counts: $\sum_{i=1}^{n} y_i = 3635$.
- Average number bids per auction: $\bar{y} = 3635/1000 = 3.635$.
- **Prior**: $\alpha = 2$, $\beta = 1/2$.

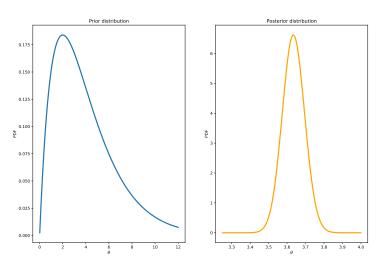
$$E(\theta) = \frac{\alpha}{\beta} = 4$$
 $SD(\theta|\mathbf{y}) = \frac{\alpha}{\beta^2} = 2.823$

Posterior

$$E(\theta|\mathbf{y}) = \frac{\alpha + \sum_{i=1}^{n} y_i}{\beta + n} = \frac{2 + 3635}{1/2 + 1000} \approx 3.635.$$

$$SD(\theta|\mathbf{y}) = \left(\frac{\alpha + \sum_{i=1}^{n} y_i}{(\beta + n)^2}\right)^{1/2} \approx 0.060.$$

eBay data - Posterior of θ



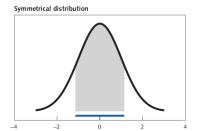
Posterior intervals

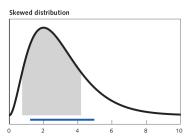
- Bayesian 95% credible interval: the probability that the unknown parameter θ lies in the interval is 0.95.
- Approximate 95% credible interval for θ

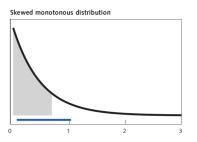
$$E(\theta|y) \pm 1.96 \cdot SD(\theta|y) = [3.517; 3.753]$$

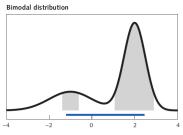
- An exact 95% equal-tail interval is [3.518; 3.754]
- Highest Posterior Density (HPD) interval contains the θ values with highest pdf.

Illustration of different interval types

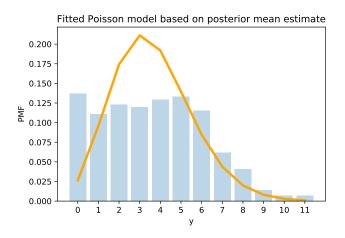








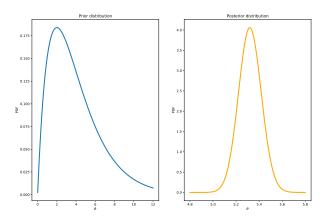
eBay dat - Fit



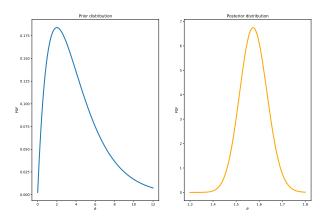
eBay - low/high seller's reservation price

- The data is very heterogenous. Some auctions start with very high reservations prices (lowest price accepted by the seller).
- Split the data into auctions with low/high reservation prices.
- Low reservation price auctions:
 - ightharpoonup n = 550 eBay coin auctions.
 - Posterior mean: 5.321 bids.
- High reservation price auctions:
 - ightharpoonup n = 450 eBay coin auctions.
 - Posterior mean: 1.576 bids.

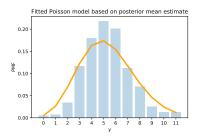
eBay - low seller's reservation price

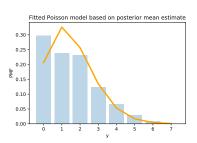


eBay - high seller's reservation price



eBay - fit low/high reservation prices





- Better fits, but still not good enough.
- Lab: Fit Poisson regression with reservation price as continuous covariate.

Conjugate priors

- Normal likelihood: Normal prior→Normal posterior.
- Bernoulli likelihood: Beta prior→Beta posterior.
- Poisson likelihood: Gamma prior→Gamma posterior.
- Conjugate priors: A prior is conjugate to a model if the prior and posterior belong to the same distributional family.
- Formal definition: Let $\mathcal{F} = \{p(y|\theta), \theta \in \Theta\}$ be a class of sampling distributions. A family of distributions \mathcal{P} is conjugate for \mathcal{F} if

$$p(\theta) \in \mathcal{P} \Rightarrow p(\theta|x) \in \mathcal{P}$$

holds for all $p(y|\theta) \in \mathcal{F}$.

Prior elicitation

- The prior should be determined (elicited) by an expert. Typically, expert≠statistician.
- Elicit the prior on a quantity that the expert knows well. Convert afterwards.
- Ask probabilistic questions to the expert:
 - \triangleright $E(\theta) = ?$
 - \triangleright $SD(\theta) = ?$
 - \triangleright $Pr(\theta < c) = ?$
 - ▶ Pr(y > c) = ?
- Show some consequences of the elicitated prior to the expert.
- Beware of psychological effects, such as anchoring.

Prior elicitation - AR(p) example

Autoregressive process or order p

$$y_t = \mu + \phi_1(y_{t-1} - \mu) + ... + \phi_p(y_{t-p} - \mu) + \varepsilon_t, \ \varepsilon_t \stackrel{iid}{\sim} N(0, \sigma^2)$$

- Informative prior on the unconditional mean: $\mu \sim N(\mu_0, \tau_0^2)$.
- "Noninformative" prior on σ^2 : $p(\sigma^2) \propto 1/\sigma^2$
- Assume $\phi_i \sim N(\mu_i, \psi_i)$, i = 1, ..., p are independent a priori.
- Prior on $\phi = (\phi_1, ..., \phi_p)$ centered on persistent AR(1) process: $\mu_1 = 0.8$, $\mu_2 = ... = \mu_p = 0$
- $Var(\phi_i) = rac{c}{i^{\lambda}}$. "Longer" lags are more likely to be zero a priori.

Jeffreys' prior

Observed information

$$J_{ heta,\mathbf{x}} = -rac{\partial^2 \ln p(\mathbf{x}| heta)}{\partial heta^2}|_{ heta=\hat{ heta}}$$

Fisher information

$$I_{\theta} = E_{\mathbf{x}|\theta} \left(J_{\theta,\mathbf{x}} \right)$$

- A common non-informative prior is Jeffreys' prior $p(\theta) = |I_{\theta}|^{1/2}$.
- Invariant to 1:1 transformations of θ .
- Often non-conjugate.
- Often problematic in multiparameter settings.

Jeffreys' prior for Bernoulli sampling

$$\begin{aligned} x_1, ..., x_n | \theta \stackrel{\textit{iid}}{\sim} \textit{Bern}(\theta). \\ & \ln p(\mathbf{x}|\theta) = s \ln \theta + f \ln(1-\theta) \\ & \frac{d \ln p(\mathbf{x}|\theta)}{d\theta} = \frac{s}{\theta} - \frac{f}{(1-\theta)} \\ & \frac{d^2 \ln p(\mathbf{x}|\theta)}{d\theta^2} = -\frac{s}{\theta^2} - \frac{f}{(1-\theta)^2} \\ & I(\theta) = \frac{E_{\mathbf{x}|\theta}(s)}{\theta^2} + \frac{E_{\mathbf{x}|\theta}(f)}{(1-\theta)^2} = \frac{n\theta}{\theta^2} + \frac{n(1-\theta)}{(1-\theta)^2} = \frac{n}{\theta(1-\theta)} \end{aligned}$$

Thus, the Jeffreys' prior is

$$p(\theta) = |I(\theta)|^{1/2} \propto \theta^{-1/2} (1 - \theta)^{-1/2} \propto Beta(1/2, 1/2).$$

Jeffreys' prior for negative binomial sampling

Jeffreys' prior:

$$\begin{split} n|\theta \overset{\textit{iid}}{\sim} \textit{NegBin}(s,\theta). \\ \ln p(\mathbf{x}|\theta) &= \ln \binom{n-1}{s-1} + s \ln \theta + f \ln (1-\theta) \\ \frac{d^2 \ln p(\mathbf{x}|\theta)}{d\theta^2} &= -\frac{s}{\theta^2} - \frac{f}{(1-\theta)^2} \\ I(\theta) &= \frac{s}{\theta^2} + \frac{E_{n|\theta}(n-s)}{(1-\theta)^2} = \frac{s}{\theta^2} + \frac{s/\theta - s}{(1-\theta)^2} = \frac{s}{\theta^2(1-\theta)} \end{split}$$

Thus, the Jeffreys' prior is

$$p(\theta) = |I(\theta)|^{1/2} \propto \theta^{-1} (1 - \theta)^{-1/2} \propto Beta(\theta|0, 1/2).$$

- Jeffreys' prior is **improper**, but the posterior is proper: $\theta | n \sim \text{Beta}(s, f + 1/2)$ which is proper since $s \ge 1$.
- Jeffreys' prior violates the likelihood principle because $I(\theta)$ is sampling-based.

Different types of prior information

- Real expert information. Combo of previous studies and experience.
- Vague prior information.
- Reporting priors. Easy to understand the information they contain.
- Smoothness priors. Regularization. Shrinkage. Big thing in modern statistics/machine learning.