# An Introduction To Bayesian **Statistics**

Simon Thornewill von Essen

Data Analyst, Goodgame Studios

@sthornewillve 🐍



"[Bayesian Stats.] is one of those ideas that seems hard when you first encounter it. Then, at some point there is a breakthrough and then it seems obvious.

Once you've got it, it's such a beautiful idea that it changes how you see everything."

Prof. Allan Downey, Data Framed, 2018



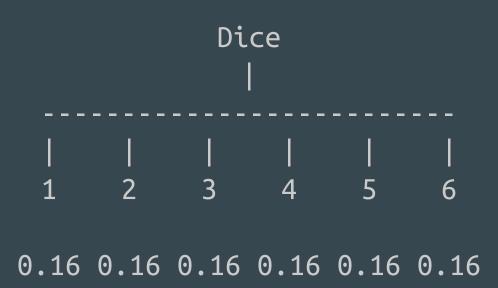
**Quick Revision** 

- Classical: By considering equal outcomes
- Frequentist: Relative Frequency over time
- Bayesian: By quantifying our uncertainties

#### **Coin Toss: Classical Est.**



#### Dice: Classical Est.



#### Classical Stats

#### Requirements

- All Outcomes are known
- Outcomes are assumed to be equally likely

#### Advantages

- Fast Estimation
- Easy to understand

#### Disadvantages

- Outcomes must be known
- Even if outcomes are finite, doing the combinatorics can still be relatively difficult
- Often created overly simplified models when applied to complex phenomena

## How do we estimate probability?

- Classical
- Frequentist
- Bayesian

#### Frequentist Est.

Check to see if thermometer is properly calibrated

#### Frequentist Approach:

Take many readings and use the expectation value (mean) and std for sample

Calculate the probability of your data given your data following some parameter.



#### **Frequentist Stats**

#### Requirements

- Possibility to perform experiments indefinitely
- Parameters are assumed to be fixed
- Able to estimate params given enough experiments

#### Advantages

- Works well with simulations
- "Objective"

#### Disadvantages

- Requires large sample size to be meaningful
- Does not allow for integration of domain knowledge
- P-values and confidence intervals are unintuitive
- Difficult to communicate to non-statisticians

#### Thermometer Calibration: Test

? 1. If P-value = 0.001 (highly significant), is the probability of getting this result or a more extreme one given our data 0.001?

**2** 2. For a given confidence interval, does the parameter lie within it 95% of the time?



#### Thermometer Calibration: Test

1. If P-value = 0.001 (highly significant), is the probability of getting this result or a more extreme one given our data 0.001?

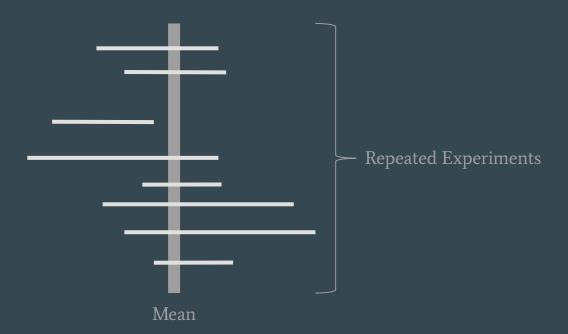
Prob. of this result or a more extreme one given a certain mean/std (i.e. Means of two groups are the same,  $H_0$ )

**2** 2. For a given confidence interval, does the parameter lie within it 95% of the time?

<u>Intervals of repeated experiments</u> will contain the parameter 95% of the time

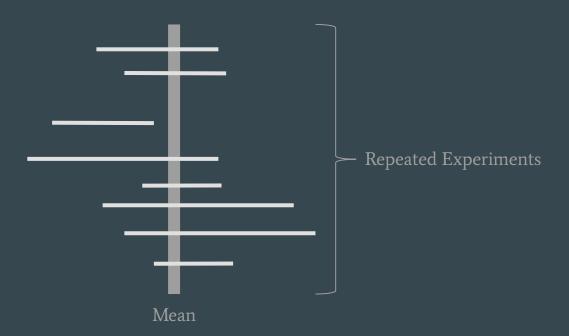
# Thermometer Calibration: Test Learnings

Frequentism expects that parameters exist and are fixed, the probabilities are the likelihood of our data given these expectations

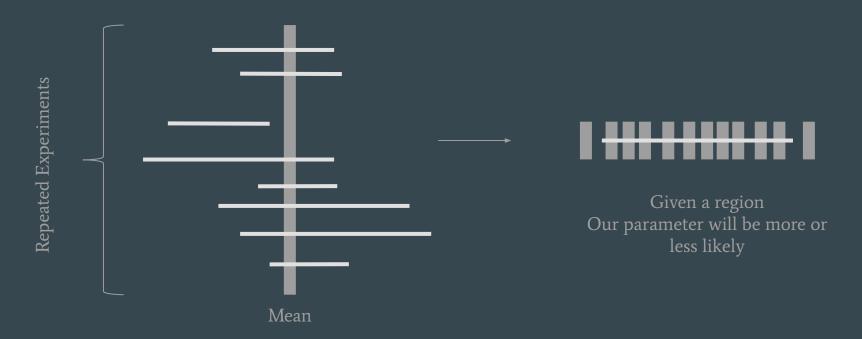


# Thermometer Calibration: Test Learnings

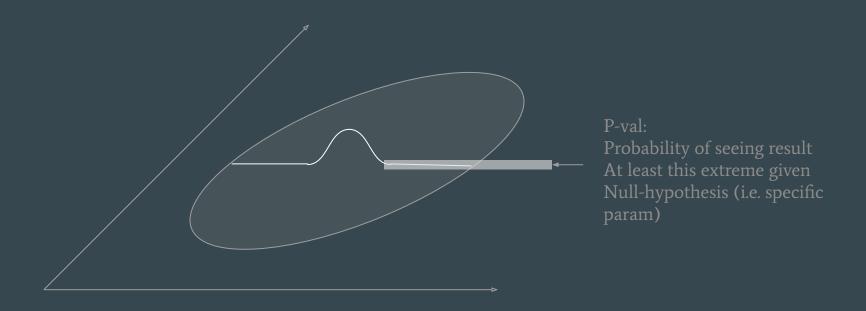
Frequentism expects that parameters exist and are fixed, the probabilities are the likelihood of our data given these expectations



#### Wait, wasn't this was we were doing with frequentism?"



#### Wait, wasn't this was we were doing with frequentism?"



#### Thermometer Calibration: Test Learnings



Child doesn't move, your repeated photos contain them 95% of the time

#### Frequentist Stats Disav. Cont.

#### What if?

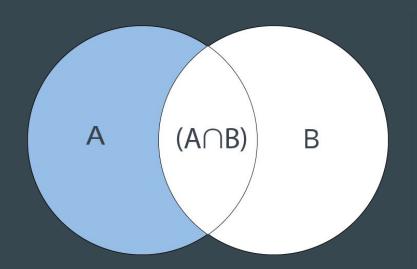
- Amount of data you have is limited?
- You have relevant and applicable prior information
- "Infinite" experiments are not possible? (Cost, feasibility)
- Stakeholders have a hard time understanding frequentist logic? 🗸
- Children never stay still and assuming they do is blasphemy

## How do we estimate probability?

- Classical
- Frequentist
- Bayesian

### Bayes Theorem: Derivation

$$P(A|B) = \frac{P(A \cap B)}{P(B)}$$



#### **Bayes Theorem: Derivation**

The Same

$$P(A|B) = \frac{P(A \cap B)}{P(B)} \qquad P(B|A) = \frac{P(B \cap A)}{P(A)}$$
$$\therefore P(A|B) = \frac{P(B|A)P(A)}{P(B)}$$

#### **Bayes Theorem: Alternate View**

 $\theta$  = Parameter,

X = Data

posterior  $p(\theta|X) = \frac{p(X|\theta)p(\theta)}{p(X)}$ 

#### **Bayes Theorem: Alternate View**

$$p(\theta|X) = \frac{p(X|\theta)p(\theta)}{p(X)^{?}}$$

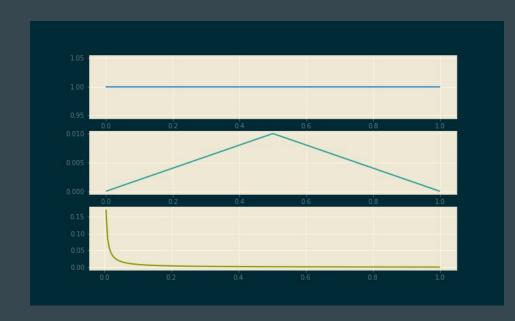
- ? Problem:
  - How to calculate p(X)
  - How to calculate  $p(\theta)$

1. Create Your Own

2. Take Previous posterior

$$p(\theta|X) = \frac{p(X|\theta)p(\theta)}{p(X)}$$

- 1. Create Your Own
- 2. Take Previous posterior



Problem: Are you Baking your biases into your model?



Might as well have your explicit and tangible biases.

As the sample size increases, priors get washed out. (As long as you are "reasonable")

- Low Sample Size: Frequentist Stats is börked anyway, so why not?
- High Sample Size: Prior Doesn't matter
- "Reasonable" = Cromwell's Rule

$$p(\theta|X) = \frac{p(X|\theta)p(\theta)}{p(X)^{?}}$$

- 1. Sum over all possible hypotheses
- 2. This is the hard part
- 3. Can be ignored if comparing inside of the same distribution

$$p(X) = \sum_{i=0}^{n} p(X|\theta_i)p(\theta_i)$$
$$p(X) = \int p(X|\theta)p(\theta)d\theta$$

#### **How!?: Doctor's Diagnosis**

- Doctor's Diagnosis:
  - You are suspicious you have a rare disease
  - O Disease affects 0.1% of the population
  - Test is 99% accurate
  - What is the probability you have this disease given that you tested positive?

## **How!?: Doctor's Diagnosis**

$$p(d \mid pos.) = \frac{p(pos. \mid d)p(d)}{p(pos.)}$$

$$p(d \mid pos.) = \frac{p(pos. \mid d)p(d)}{p(pos. \mid d)p(d) + p(pos. \mid \neg d)p(\neg d))}$$

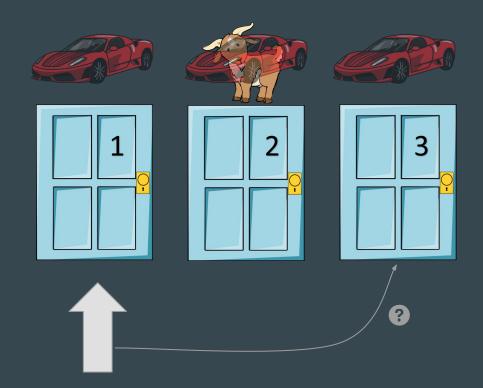
$$p(d \mid pos.) = \frac{0.99 \times 0.001}{0.99 \times 0.001 + 0.01 \times 0.999}$$

#### **How!?: Doctor's Diagnosis**

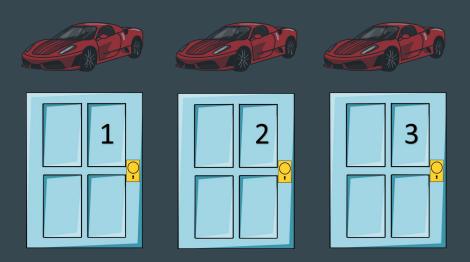
#### Doctor's Diagnosis:

- Works out to be roughly 9% chance of having disease
- Rate is lower than expected because disease is so rare
- 99% is a high accuracy, but not *that* high
- Application: Think about this example next time you build a classifier

- The Monty Hall Problem:
  - You Pick Door 1
  - o Monty opens door 2 to reveal a goat
  - Should you switch to door 3?



Hypothesis i	Prior P( $\theta_{\rm i}$ )
Car Behind 1	1/3
Car Behind 2	1/3
Car Behind 3	1/3

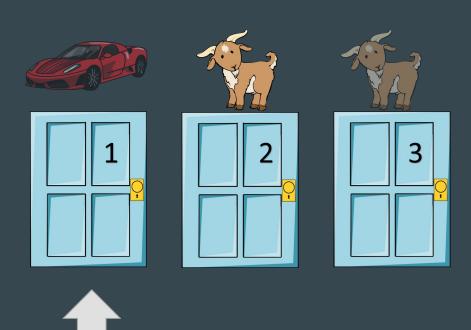


Hypothesis i Likelihood p(X |  $\theta_i$ )

Car Behind 1

Car Behind 2

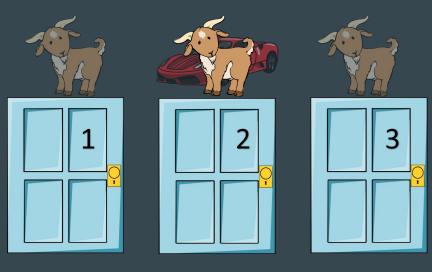
Car Behind 3



Hypothesis i Likelihood  $p(X \mid \theta_i)$ Car Behind 1 1/2

Car Behind 2

Car Behind 3





Hypothesis i	Likelihood p(X   $\theta_i$ )
Car Behind 1	1/2
Car Behind 2	0.0
Car Behind 3	





Hypothesis i	Prior $p(\theta_i)$	Likelihood p(X   $\theta_i$ )
Car Behind 1	1/3	1/2
Car Behind 2	1/3	0.0
Car Behind 3	1/3	1.0

# **How!?: Monty Hall Problem**

Hypothesis i	Prior p( $\theta_{\rm i}$ )	Likelihood p(X   $\theta_i$ )	Prior * Likelihood
Car Behind 1	1/3	1/2	1/6
Car Behind 2	1/3	0.0	0.0
Car Behind 3	1/3	1.0	2/6

$$P(X) = \sum_{i=0}^{n} p(X|\theta_i)p(\theta_i) = 1/6 + 0 + 2/6$$

$$= 3/6$$

$$= 1/2$$

<sup>\*</sup> This is the Dot Product of  $p(\theta_i)$  and  $p(X \mid \theta_i)$ 

# **How!?: Monty Hall Problem**

Hypothesis i	Prior $p(\theta_i)$	Likelihood p(X   $\theta_i$ )	Prior * Likelihood	Posterior
Car Behind 1	1/3	1/2	1/6	1/3
Car Behind 2	1/3	0.0	0.0	0.0
Car Behind 3	1/3	1.0	2/6	2/3

#### Key to problem:

Monty does not choose doors at random and so opening a door provides you with information

You see a train labeled 60 (labels in ascending order of creation)

• How many trains does this company own given that you've seen #60?

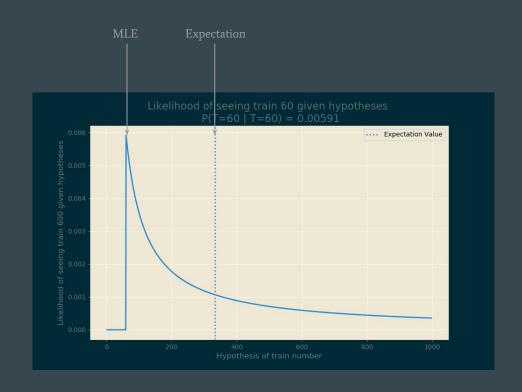


Hypothesis i	Prior $p(\theta_i)$	Likelihood p(X   $\theta_i$ )	Prior * Likelihood	Posterior
1 Train	1/1000	0.0	0.0	post <sub>1</sub>
2 Trains	1/1000	0.0	0.0	post <sub>2</sub>
60 Trains	1/1000	1/60	1/(6*10 <sup>4</sup> )	post <sub>60</sub>
1000 Trains	1/1000	1/1000	1/106	post <sub>1000</sub>

$$\Sigma = P(Train)$$

- What if we change priors?
  - Posterior changes

- What if we increase the max number of trains
  - Posterior changes



# How!?: Part 1 - Discrete Case Recap

- Remember the table calculation!
- Steps:
  - Pick Prior (Often Uniform)
  - Multiply by Frequentist Likelihood
  - O Divide by Normalisation constant
- Similar things are true for the cont. case

$$p(\theta|X) = \frac{p(X|\theta)p(\theta)}{\sum_{i=0}^{n} p(X|\theta_i)p(\theta_i)}$$

#### AB Testing:

- a. Two variants
- b. What is the probability of A being better than B?



#### Challenges w/ Freq. AB Tests:

- Test needs to reach pre-defined sample size
- Need to adjust α for multiple tests (Bonferroni Corrections)
- People start accepting null hypotheses (Logically illegal)
- Can only reject/fail to reject null hypothesis, (leads to p-hacking)
- People peek at tests before tests are over, (more p-hacking)

#### Bayesian AB Tests:

- Tests can be started and stopped as necessary
- No need for adjustments, everything is included in calculation
- Probabilistic statements give better indication of risk
- No p-hacking

$$P(\theta|X) = Beta(\alpha, \beta)$$

$$\alpha = 1 + \sum y_i,$$

$$\beta = n - 1 + \sum y_i$$

# Conjugate Priors

#### Derivation:

- Choose Prior: Uniform
- Choose Likelihood: Bernoulli
- Do some math magic
- Get Beta distribution out

$$P(\theta|X) = Beta(\alpha, \beta)$$

# **Conjugate Priors**

Likelihood	Model parameters	Conjugate prior distribution	Prior hyperparameters	Posterior hyperparameters	Interpretation of hyperparameters <sup>[note 1]</sup>	Posterior predictive[note 2]
Bernoulli	p (probability)	Beta	$\alpha, \beta$	$\alpha + \sum_{i=1}^n x_i, \ \beta + n - \sum_{i=1}^n x_i$	lpha-1 successes, $eta-1$ failures [note 1]	$p( ilde{x}=1)=rac{lpha'}{lpha'+eta'}$

# **Conjugate Priors**

	111			$k, \theta$	$k+\sum_{i=1}^n x_i,\;rac{ heta}{n heta+1}$		$k$ total occurrences in $\frac{1}{ heta}$ intervals		$ ext{NB}( ilde{x} \mid k',  heta')$ (negative binomial)
Poisson	oisson A (rate)	Gamma	$lpha,eta^{ ext{[note 3]}}$	$\alpha + \sum_{i=1}^n x_i, \ \beta + n$		lpha total occurrences in $eta$ intervals		$\mathrm{NB}\Big( ilde{x} \mid lpha', rac{1}{1+eta'}\Big)$ (negative binomial)	
Exponential	Λ (rate)	Gamma	$\alpha,eta$ [note 3]	$\alpha+n,\beta+\sum_{i=1}^n x_i$		lpha observations that sum to ,	) [v]	$\operatorname{Lomax}(\tilde{x}\mid eta', lpha')$ (Lomax distribution)	

# Demo: Conjugate Priors

### Posteriors - What Now?

#### Do we calculate P-values now?

- No need, just calculate  $P(\theta_2 > \theta_1)$
- Takes some calculation, but the result is nicer

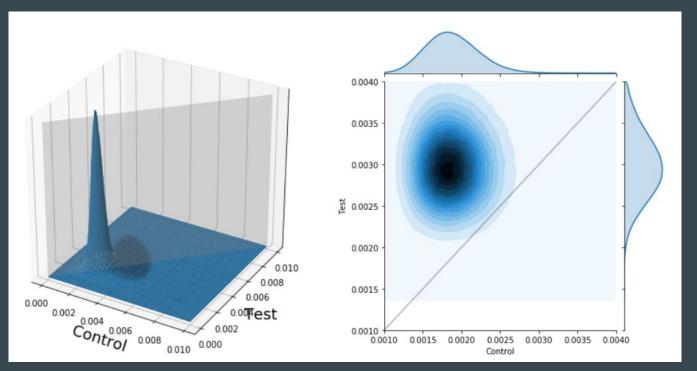
#### Can we test multiple versions

Absolutely (Still, try and be reasonable)

#### Can we peek at results or stop at any time?

• Yes!

### Posteriors - What Now?



# Demo: Bayesian AB Testing

### Bayesian Stats

#### Advantages

- a. Estimation in the case of little data (specific circumstances)
- b. Parameters are distributions
- c. Probabilistic answers answer questions people tend to have

#### Disadvantages

- a. Steeper learning curve
- b. Integrals are hard (MCMC computationally expensive)
- c. Criticisms of being less "objective" due to use of priors
- d. Point estimates become the same as frequentist estimations with high sample sizes

# How do we estimate probability?

- Classical: By considering equal outcomes
- Frequentist: Relative Frequency over time
- Bayesian: By quantifying our uncertainties

# How do we estimate probability?

- Classical: By considering equal outcomes
- Frequentist: Relative Frequency over time
- Bayesian: By quantifying our uncertainties
- There is only one field of statistics

### Conclusion and "Call to Action"

#### <u>Understanding Bayes vs Freq. is key to understanding a lot of the field stats</u>

- Beginner scientists: Check my sources as a jump off point
- Experienced scientists: Come and tell me how you applied this knowledge.
- Decision-makers: Be aware of these kinds of analyses and when the strengths benefit your use-case.
- General Advice: Try to think about conditional probability more often (doctor's diagnosis) in relation to drawing conclusions

# Resources for further learning

Mathematical Understanding:

```
"Bayesian Stats: From Concept to Data Analysis", U of Santa Cruz
```

Intuition between Bayesianism & Frequentism:

```
"Frequentism and Bayesianism", 
Scipy - Jake Vander Plas
```

Examples of Real World Applications:

```
"Think Bayes",

Allan Downey
```

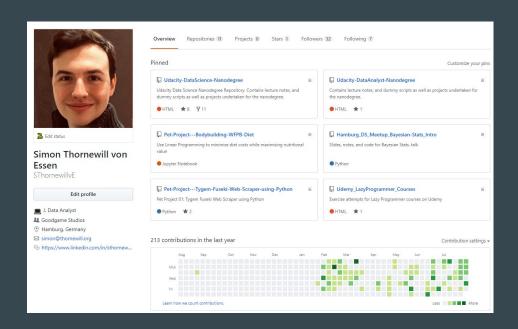
Further reading into MCMC and pyMC:

```
"Bayesian Methods for Hackers",
Cameron Davidson-Pilon
```

### Find Slides on Github

https://cutt.ly/zGqux9





# Fin!

•••

@sthornewillve



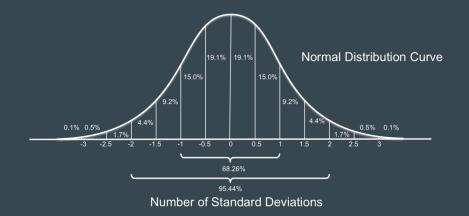




# Frequentist Est.: Confidence Intervals

#### **Confidence Interval**:

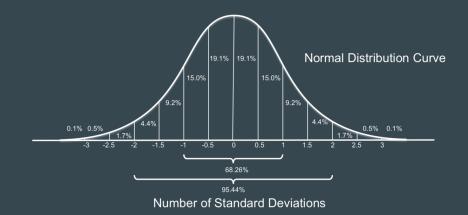
- From sample mean and standard deviation, calculate an interval
- Upon repeated experiments, intervals contain the true parameter x% of the time



# Frequentist Est.: Confidence Intervals

#### CI Intuition:

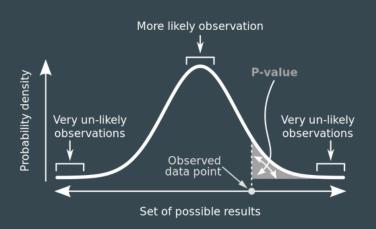
 bootstrap CI n times -> intervals would contain the mean of population 95% of the time



# Frequentist Est.: P-Values

#### P-value:

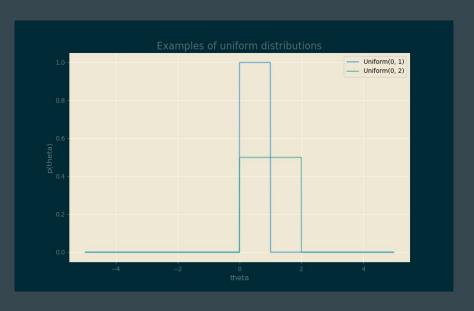
 Probability of seeing data given a parameter



#### Prior?:

- Uninformed Prior
- Uniform distribution
- Represented by Indicator Function

$$P(\theta) = I_{\{0 \le \theta \le 1\}}$$



#### AB Test:

- For people randomly placed in control/test
- Track conversions (1/0)
- What is our Likelihood?
  - o Bernoulli

$$P(X|\theta) = \theta^{\sum y_i} (1-\theta)^{n-\sum y_i}$$

$$P(\theta|X) \propto [\theta^{\sum y_i} (1-\theta)^{n-\sum y_i}][I_{\{0 \le \theta \le 1\}}]$$

$$P(\theta|X) = \frac{\theta^{\sum y_i} (1 - \theta)^{n - \sum y_i} I_{\{0 \le \theta \le 1\}}}{\int_0^1 \theta^{\sum y_i} (1 - \theta)^{n - \sum y_i} I_{\{0 \le \theta \le 1\}} d\theta}$$

$$P(\theta|X) = \frac{\theta^{\sum y_i} (1 - \theta)^{n - \sum y_i} I_{\{0 \le \theta \le 1\}}}{A^{-1} \int_0^1 A \theta^{\sum y_i} (1 - \theta)^{n - \sum y_i} I_{\{0 \le \theta \le 1\}} d\theta}$$

$$A = \frac{\Gamma(\sum n + 2)}{\Gamma(\sum y_i + 1)\Gamma(\sum n - y_i + 1)}$$

$$P(\theta|X) = A(\theta^{\sum y_i} (1 - \theta)^{n - \sum y_i} I_{\{0 \le \theta \le 1\}})$$
$$= Beta(\alpha, \beta)$$