

$$f_X(x; \theta) = \frac{1}{\theta} \mathbb{1}(0 \leq x \leq \theta)$$

(الف - 1)

$$\begin{aligned} \Rightarrow f_{\text{Data}}(\text{Data}; \theta) &= \prod_{i=1}^n \frac{1}{\theta} \mathbb{1}(0 \leq x_i \leq \theta) \\ &= \frac{1}{\theta^n} \prod_{i=1}^n \mathbb{1}(0 \leq x_i \leq \theta) \end{aligned}$$

$$\underset{\theta \in \mathbb{R}_{>0}}{\operatorname{argmax}} \frac{1}{\theta^n} \prod_{i=1}^n \mathbb{1}(0 \leq x_i \leq \theta) \rightarrow \begin{array}{ll} \min \theta & \text{s.t.} \\ \theta > 0 & \theta \geq x_i \\ & i=1, \dots, n \end{array}$$

$$\Rightarrow \boxed{\hat{\theta}_{ML} = \max_{i=1, \dots, n} x_i} \quad (2)$$

$$a(1 - \frac{x}{\theta}) \quad 0 \leq x \leq \theta \rightarrow \text{PDF}$$

(ب)

$$\Rightarrow \int_0^{\theta} a(1 - \frac{x}{\theta}) dx = a\theta - a \frac{\theta^2}{2} = 1$$

$$\Rightarrow a\theta - a \frac{\theta}{2} = \frac{a\theta}{2} = 1$$

$$\Rightarrow \boxed{a = \frac{2}{\theta}} \quad (0.5)$$

$$\hat{\theta}_{ML} = \underset{\theta \geq 0}{\operatorname{argmax}} \prod_{i=1}^n \frac{2}{\theta} \left(1 - \frac{x_i}{\theta}\right)$$

(ج)

$$\left(\frac{2}{\theta}\right)^n \prod_{i=1}^n \left(1 - \frac{x_i}{\theta}\right)$$

$$LLH(\theta) = \log 2^n - n \log \theta + \sum_{i=1}^n \log \left(1 - \frac{x_i}{\theta}\right)$$

$$\frac{\partial}{\partial \theta} LLH(\theta) = 0 \Rightarrow \frac{-n}{\theta} + \sum_{i=1}^n \frac{-\frac{x_i}{\theta}}{1 - \frac{x_i}{\theta}} \frac{1}{\theta^2} = 0$$

$$\left[\sum_{i=1}^n \frac{x_i}{\theta - x_i} = n \right] \quad (2.5)$$

$$g(\theta) \triangleq \sum_{i=1}^n \frac{x_i}{\theta - x_i} \quad M \triangleq \max_i x_i \quad (3)$$

$g(\cdot) \rightarrow$ Decreasing function

$$\hat{\theta}_m \triangleq M(1 + \frac{1}{n}) \Rightarrow$$

$$g(\hat{\theta}_m) = \frac{M}{M(1 + \frac{1}{n}) - M} + \underbrace{\sum_{\substack{i=1 \\ i \neq i_{\max}}}^n \frac{x_i}{\hat{\theta}_m - x_i}}_{\geq 0} \geq n$$

$$\Rightarrow \hat{\theta}_{ML} \geq M(1 + \frac{1}{n})$$

$$\hat{\theta}_M = 2M \Rightarrow$$

$$g(\hat{\theta}_M) = \sum_{i=1}^n \frac{x_i}{\hat{\theta}_M - x_i} \leq \sum_{i=1}^n \frac{M}{2M - x_i} \leq \sum_{i=1}^n \frac{M}{M} = n$$

$$\Rightarrow \hat{\theta}_{ML} \leq \hat{\theta}_M = 2M$$

$$\Rightarrow \boxed{M(1 + \frac{1}{n}) \leq \hat{\theta}_{ML} \leq 2M} \quad *$$

$$\mathbb{P}(X > a) = \mathbb{P}(e^{tX} > e^{ta}) \leq \frac{\mathbb{E}(e^{tX})}{e^{ta}} \quad t > 0 \quad \text{۲- الف)}$$

$$\Rightarrow \mathbb{P}(X > a) \leq e^{-ta} M_X(t) \quad \text{if } M_X(t) \text{ exists for } \exists t > 0 \quad \text{2)}$$

ب) توزیع کوئی $M_X(t)$ ندارد. به ازای هیچ مداری از $t > 0$. 1)

if $Z \sim \text{Unif}(0,1)$ and $X \sim f_X(x)$ ۳-
 $\hookrightarrow \text{CDF} \rightarrow F_X(x)$

Then $F_X^{-1}(Z) \sim f_X$

$$\begin{aligned} \rightarrow F_X(x) &= \int_{-\infty}^x \frac{u}{\sigma^2} e^{-u^2/(2\sigma^2)} du \\ &= \int_{-\infty}^x \frac{u}{\sigma^2} e^{-u^2/(2\sigma^2)} du = \int_{-\infty}^{\frac{x^2}{2\sigma^2}} e^{-u'} du' \end{aligned}$$

$$\begin{aligned} u' &= \frac{u^2}{2\sigma^2} \\ du' &= \frac{u du}{\sigma^2} \end{aligned}$$

$$= \left(1 - e^{-\frac{x^2}{2\sigma^2}}\right) \mathbb{I}(x \geq 0)$$

$$F_X^{-1}(Z) = y \Rightarrow F_X(y) = Z = 1 - e^{-\frac{y^2}{2\sigma^2}} \quad y \geq 0$$

$$\Rightarrow -\frac{y^2}{2\sigma^2} = \log(1-Z) \Rightarrow$$

$$y^2 = -2\sigma^2 \log(1-z)$$

$$\boxed{y = \sigma \sqrt{2 \log \frac{1}{1-z}} \sim f_X(y)} \quad (4)$$

$$H_0: \mu_1 = \mu_2$$

$$H_1: \mu_1 \neq \mu_2$$

(1)

4

الف

$$S = \left[\frac{1}{9} (x_1 + \dots + x_9) - \frac{1}{16} (y_1 + \dots + y_{16}) \right]$$

ب

$$\begin{aligned} E(S|H_0) &= 0 \\ \text{Var}(S|H_0) &= \frac{1}{9^2} \left(\sum_{i=1}^9 \text{Var}(X_i) \right) + \frac{1}{(16)^2} \left(\sum_{i=1}^{16} \text{Var}(Y_i) \right) \\ &= \frac{\text{Var}(X_1)}{9} + \frac{\text{Var}(Y_1)}{16} = \frac{16+9}{16 \times 9} = \frac{25}{9 \times 16} \\ &= \left(\frac{5}{12} \right)^2 \end{aligned}$$

$$\Rightarrow \bar{S} = \frac{S}{\frac{5}{12}} \rightarrow \begin{aligned} E \bar{S} &= 0 \\ \text{Var}(\bar{S}) &= 1 \end{aligned} \quad \text{قابل استبدال است}$$

$$\mathbb{P}(|\bar{S}| > \frac{13-10.5}{\frac{5}{12}} = \frac{12 \times 2.5}{5} = 6) \leq \frac{1}{6^2} = \frac{1}{36} = \frac{1}{100} \quad (1.5)$$

Chebyshev's Inq. در حد می شود

$$\bar{S} \sim \mathcal{N}(0, 1)$$

(0.5)

ج

$$\mathbb{P}(|\bar{S}| > 6) \ll \frac{1}{100} \rightarrow \text{قابل می شود}$$

$$Y_i = \mathbb{1}(X_i \geq 0)$$

5- الف)

$$Y_1, \dots, Y_n \stackrel{\text{i.i.d.}}{\sim} \text{Bern}(p) \quad p = \mathbb{P}(X \geq 0)$$

$$LH = \prod_{i=1}^n p^{Y_i} (1-p)^{1-Y_i}$$

$$\Rightarrow LLH = \sum_{i=1}^n Y_i \log p + (1-Y_i) \log(1-p)$$

$$\frac{\partial}{\partial p} LLH(p) = 0 \Rightarrow$$

$$\sum_{i=1}^n \frac{Y_i}{p^*} - \frac{1-Y_i}{1-p^*} = 0$$

$$\Rightarrow \frac{N_1}{p^*} = \frac{N_0}{1-p^*} \Rightarrow N_1 - N_1 p^* = N_0 p^*$$

$$\Rightarrow p^* = \frac{N_1}{N_0 + N_1} = \frac{N_1}{n}$$

(2)

$$\hat{p}_{ML} = \frac{N_1}{n} = \frac{\#\{Y_i = 1\}}{n}$$

$$\text{Bias}(\hat{p}_{ML}) = \mathbb{E}(\hat{p}_{ML}) - p =$$

(.)

$$\frac{\mathbb{E} N_1}{n} - p = \frac{np}{n} - p = 0$$

(1)

$$\text{Var}(\hat{p}_{ML}) = \text{Var}\left(\frac{N_1}{n}\right) = \text{Var}\left(\frac{Y_1 + \dots + Y_n}{n}\right) =$$

$$\frac{1}{n^2} \sum_{i=1}^n \text{Var}(Y_i) = \frac{np(1-p)}{n^2} = \frac{p(1-p)}{n}$$

$$\underbrace{\mathbb{P}(p|\text{Data})}_{\text{Posterior}} \propto \underbrace{\mathbb{P}(\text{Data}|p)}_{\text{Likelihood}} \underbrace{\mathbb{P}(p)}_{\text{Prior}} \quad (2)$$

$$\Rightarrow \mathbb{P}(p|\text{Data}) = \text{constant} \times \prod_{i=1}^n p^{y_i} (1-p)^{1-y_i} k p^{k-1}$$

$$\Rightarrow \hat{p}_{\text{MAP}} = \underset{p \in [0,1]}{\text{argmax}} \underbrace{\sum_{i=1}^n y_i \log p + (1-y_i) \log(1-p) + (k-1) \log p}_{\substack{\text{LP}(p) \\ \log \downarrow \text{Posterior}}}$$

$$\frac{\partial}{\partial p} \text{LP}(p) = 0$$

$$\Rightarrow \frac{1}{p^*} \sum_{i=1}^n y_i - \frac{1}{1-p^*} \sum_{i=1}^n (1-y_i) + \frac{k-1}{p^*} = 0$$

$$\Rightarrow \frac{N_1 + k - 1}{p^*} = \frac{N_0}{1-p^*}$$

$$\boxed{\hat{p}_{\text{MAP}} = \frac{N_1 + k - 1}{n + k - 1}} \quad (2)$$