

Conceptual and Theoretical Questions

Q1

To minimize the error function, take the derivative with respect to \mathbf{w} and set it to 0:

$$\begin{aligned}
 0 &= -\frac{1}{2} \sum_{n=1}^N 2 \cdot r_n (t_n - \mathbf{w}^T \cdot \phi(x_n)) \phi(x_n)^T \\
 0 &= -\sum_{n=1}^N r_n t_n \phi(x_n)^T - r_n \mathbf{w} \phi(x_n) \phi(x_n)^T \\
 0 &= \sum_{n=1}^N r_n t_n \phi(x_n)^T - \mathbf{w} \sum_{n=1}^N r_n \phi(x_n) \phi(x_n)^T \\
 \mathbf{w} \sum_{n=1}^N r_n \phi(x_n) \phi(x_n)^T &= \sum_{n=1}^N r_n t_n \phi(x_n)^T
 \end{aligned}$$

Take the inverse to isolate \mathbf{w}

$$\mathbf{w} = \left(\sum_{n=1}^N r_n \phi(x_n) \phi(x_n)^T \right)^{-1} \sum_{n=1}^N r_n t_n \phi(x_n)^T$$

Because r_n is defined as the inverse of the variance of the n th data point, a large r_n indicates a small variance, indicating the high confidence in that data. A small weight would then of course indicate low certainty in that data with a high variance.

Alternatively, the weight r_n can simply be seen as how many times the n th data point has been observed, which is in line with the constraint $r_n > 0$, but is more fitting for integral values of r_n .

Q2

Using Bayes' rule, the relationship between the prior, likelihood, and posterior distributions is known to be:

$$p(\mathbf{w}, \beta | \mathbf{t}) \propto p(\mathbf{w}, \beta) \cdot p(\mathbf{t} | \mathbf{X}, w, \beta)$$

Work is shown above

Q3

If a dataset is linearly separable, then its classes can be exactly distinguished with a linear decision boundary. The probability of a class C_1 in a dataset using logistic regression is

$$p(C_1|\phi) = \sigma(w^T \phi)$$

Because the dataset is linearly separable, data of one class x_a may fall one side of the decision boundary with $w^T \phi(x_a) > 0$, and another class denoted by x_b may be found on the other side $w^T \phi(x_b) < 0$.

If $|w| \rightarrow \infty$, then the sigmoid function of the posterior for class a $\sigma(w^T \phi) \rightarrow 1$, which thus maximizes the likelihood. This would also be true for finding class b since $\sigma(x) = 1 - \sigma(-x)$

Q4

We have

$$H = \nabla \nabla E(w) = \sum_{n=1}^N y_n(1 - y_n) \phi_n \phi_n^T = \Phi^T R \Phi \quad (4.97, \text{Bishop})$$

For a symmetric matrix A to be positive definite, then any real valued vector z can be multiplied by A to get all entries being positive:

$$z^T A z = A_{ij} > 0$$

Begin by multiplying the double gradient by z as above to prove it is positive definite:

$$\begin{aligned} z^T \sum_{n=1}^N (y_n(1 - y_n) \phi_n \phi_n^T) z \\ \sum_{n=1}^N y_n(1 - y_n) z^T \phi_n \phi_n^T z \end{aligned}$$

Notice that $z^T \phi_n = (\phi_n^T z)^T$, which is also a dot product (scalar quantity), which allows us to square that term, making it always greater than or equal to zero:

$$\sum_{n=1}^N y_n(1 - y_n) (\phi_n^T z)^2$$

Now, because $1 > y_n > 0$, since that is the range of the sigmoid function in logistic regression, then $y_n(1 - y_n) > 0$ as well, meaning the terms are all positive.

As for proving that the function is concave, it has just been proven that the double gradient is always greater than 0, which ensures $\nabla \nabla E(w) > 0$, which means it is concave by definition.

Q5

Part a

The PMF of the exponential distribution family can be written as

$$p(x|\eta) = h(x)g(\eta) \exp \{ \eta^T y(x) \} \quad (2.194, \text{Bishop})$$

The functions correspond to the following:

$$\begin{aligned} h(x) &= x^{\alpha-1} \\ y(x) &= x \\ \eta^T &= -\beta \\ g(\eta) &= g(\beta) = \beta^\alpha \end{aligned}$$

Part b

In order to find the maximum likelihood, take the log for convenience:

$$\log(L) = \sum_{i=1}^N -\log(T(v)) + v \log\left(\frac{xt_i}{y_i}\right) - \log y_i - \frac{vt_i}{y_i}$$

Expanding and removing constant terms:

$$-(v+1) \sum_{i=1}^N \log y_i - v \sum_{i=1}^N \frac{t_i}{y_i}$$

To find the gradient ascent steps, take the gradient with respect to w_0 and w_1 :

$$\begin{aligned} \frac{\partial \log(L)}{\partial w_0} &= \sum_{i=1}^N -(v+1) + vt_i y_i \\ \frac{\partial \log(L)}{\partial w_1} &= \sum_{i=1}^N x_i (-(v+1) + vt_i y_i) \end{aligned}$$

This leads to the following update rules:

$$\begin{aligned} w_0^{(r+1)} &\leftarrow w_0^r + \alpha \nabla w_0^r \\ w_1^{(r+1)} &\leftarrow w_1^r + \alpha \nabla w_1^r \end{aligned}$$

To determine if the likelihood is concave or not, take the gradient with respect to w_0 and w_1 :

Q6

Part a

The likelihood function would be the probability of observing each t_n , given the parameters w and ϕ , over a normal distribution, so that:

$$L(t_1, t_2, \dots, t_n | w, \phi) = \prod_{i=1}^N N(t_i | w^T \phi(x), \sigma^2)$$

The Laplacian prior is defined as

$$P(w | b) = \frac{1}{2b} \exp\left(-\frac{|w|}{b}\right)$$

Multiply the prior and likelihood:

$$L(w | D, b) \propto L(t_1, t_2, \dots, t_n | w, \phi) \cdot \frac{1}{2b} \exp\left(-\frac{|w|}{2b}\right)$$

Which follows the same general form as the lasso regularization formula:

$$-\frac{1}{2} \sum_i t_i - w^T \phi(x_i) - \lambda \sum_j |w_j|$$

```

In [ ]: # Part b
import numpy as np
import matplotlib.pyplot as plt

mu_ek = 0
sigma_ek = 0.1
ek = np.random.normal(loc=mu_ek, scale=np.sqrt(sigma_ek), size=(100,1))

mu_x = 0
sigma_x = 1
x_i = np.random.normal(loc=mu_x, scale=np.sqrt(sigma_x), size=(100,1))
t_i = 1 + (0.01 * x_i) - 2 * np.multiply(x_i, x_i) + ek

lambda_1 = np.linspace(0, 10, num=100)
eps = 0.0001
phi_x = np.concatenate([np.ones((100, 1)), x_i, np.multiply(x_i, x_i)], axis=
1)
alpha = 1e-4
ws = np.zeros((100, 3))
for l in np.arange(1, 100, step=1):
    t_lambda = lambda_1[l - 1]
    w_r = np.random.randn(3, 1)
    for iter in np.arange(1, 10000, step=1):
        err = t_i - np.matmul(phi_x, w_r) # 100, 1
        dw_0 = sum(err)

        d1 = 1 / max(eps, abs(w_r[1]))
        dw_1 = np.sum(phi_x[:, 1].T * err) - t_lambda * w_r[1] * d1

        d2 = 1 / max(eps, abs(w_r[2]))
        dw_2 = np.sum(phi_x[:, 2].T * err) - t_lambda * w_r[2] * d2

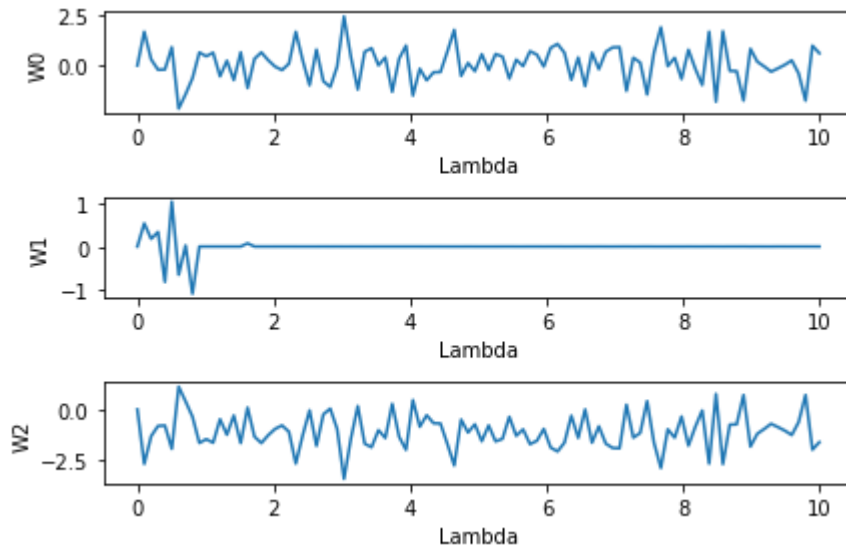
        w_r[0] = w_r[0] + alpha * dw_0
        w_r[1] = w_r[1] + alpha * dw_1
        w_r[2] = w_r[2] + alpha * dw_2
    ws[l,:] = w_r.reshape((3,))

fig_6, ax_6 = plt.subplots(3, 1)
ax_6[0].plot(lambda_1, ws[:, 0])
ax_6[0].set_ylabel("W0")
ax_6[0].set_xlabel("Lambda")

ax_6[1].plot(lambda_1, ws[:, 1])
ax_6[1].set_ylabel("W1")
ax_6[1].set_xlabel("Lambda")

ax_6[2].plot(lambda_1, ws[:, 2])
ax_6[2].set_ylabel("W2")
ax_6[2].set_xlabel("Lambda")
fig_6.tight_layout()

```



Application Questions

```
In [ ]: # Imports

import pandas as pd
import numpy as np
from matplotlib import pyplot as plt
from itertools import combinations
from sklearn import preprocessing
from sklearn.linear_model import LogisticRegression, LinearRegression, BayesianRidge
from sklearn.metrics import mean_squared_error, RocCurveDisplay
from sklearn.naive_bayes import BernoulliNB
from sklearn.model_selection import cross_val_score
from scipy import stats
import pymc3 as pm
```

Linear Regression Problem (25 pts)

In []: *# Dataset init*

```
dfr = pd.read_excel("Real estate valuation data set.xlsx")
dfr.rename(columns={
    "X1 transaction date": "X1",
    "X2 house age": "X2",
    "X3 distance to the nearest MRT station": "X3",
    "X4 number of convenience stores": "X4",
    "X5 latitude": "X5",
    "X6 longitude": "X6",
    "Y house price of unit area": "Y"
}, inplace=True)
print(f"Zeroing X1, transaction date, at {min(dfr['X1'])}")
dfr["X1"] = dfr["X1"] - min(dfr["X1"])
dfr.head()
```

Zeroing X1, transaction date, at 2012.6666667

Out[]:

	No	X1	X2	X3	X4	X5	X6	Y
0	1	0.250000	32.0	84.87882	10	24.98298	121.54024	37.9
1	2	0.250000	19.5	306.59470	9	24.98034	121.53951	42.2
2	3	0.916667	13.3	561.98450	5	24.98746	121.54391	47.3
3	4	0.833333	13.3	561.98450	5	24.98746	121.54391	54.8
4	5	0.166667	5.0	390.56840	5	24.97937	121.54245	43.1

In []: *# 1. Visualization*

```
figure_hist, ax_hist = plt.subplots()
dfr.drop("No", axis=1).hist(ax=ax_hist, bins="auto")
figure_hist.suptitle("Predictor histograms")
figure_hist.tight_layout()
figure_hist.set_size_inches(12, 8)

fig_sc_preds, ax_preds = plt.subplots(nrows=5, ncols=3, figsize=(12, 10), constrained_layout=True)
fig_sc_resp, ax_resp = plt.subplots(nrows=2, ncols=3, figsize=(12, 10), constrained_layout=True)
combs = list(combinations(range(1,7), 2))

def lin_to_mat(cols, x):
    return (x // cols, x % cols)

for i in range(1, 7):
    ax_resp[lin_to_mat(3, i - 1)].scatter(dfr[f"X{i}"], dfr["Y"])
    ax_resp[lin_to_mat(3, i - 1)].set_xlabel(f"X{i}")
    ax_resp[lin_to_mat(3, i - 1)].set_ylabel("Y")
fig_sc_resp.suptitle("Scatter plots between predictors and response")

for i in range(len(combs)):
    ax_preds[lin_to_mat(3, i - 1)].scatter(dfr[f"X{combs[i][0]}"], dfr[f"X{combs[i][1]}"])
    ax_preds[lin_to_mat(3, i - 1)].set_ylabel(f"X{combs[i][0]}")
    ax_preds[lin_to_mat(3, i - 1)].set_xlabel(f"X{combs[i][1]}")
fig_sc_preds.suptitle("Scatter plots for predictors")

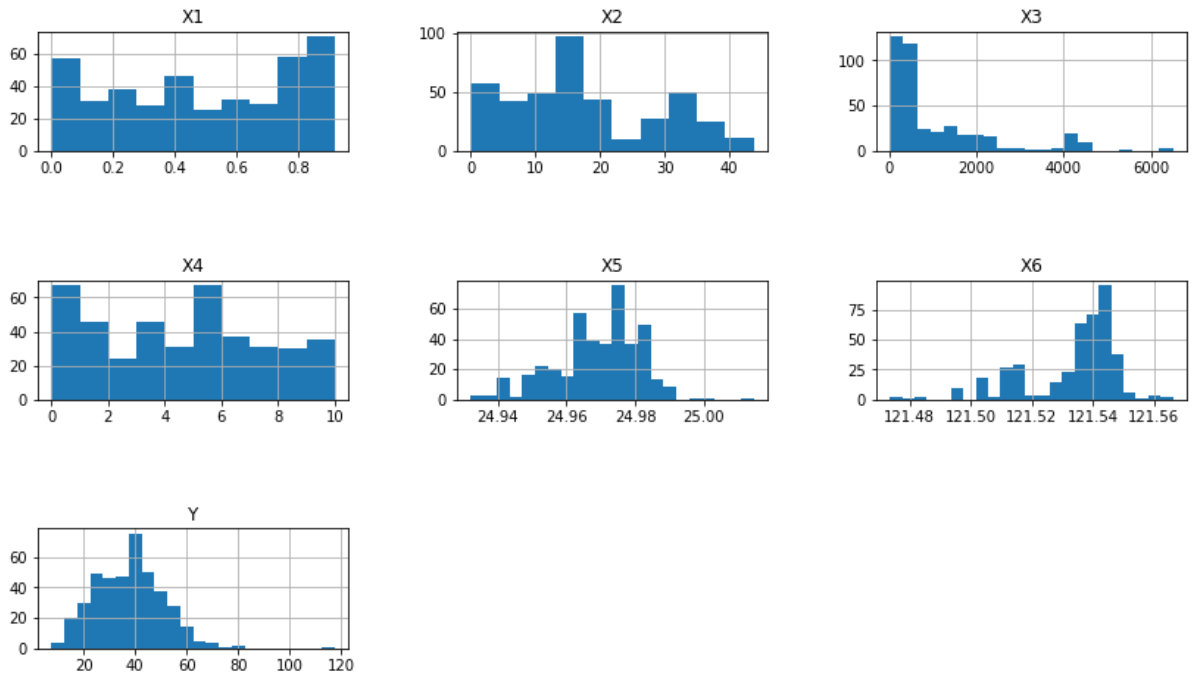
_="""
There appear to be some data points that differ from the rest of the predictor
data, such as in X3 and Y. X3 and X5 appear to have the strongest correlation
with Y.
"""
```

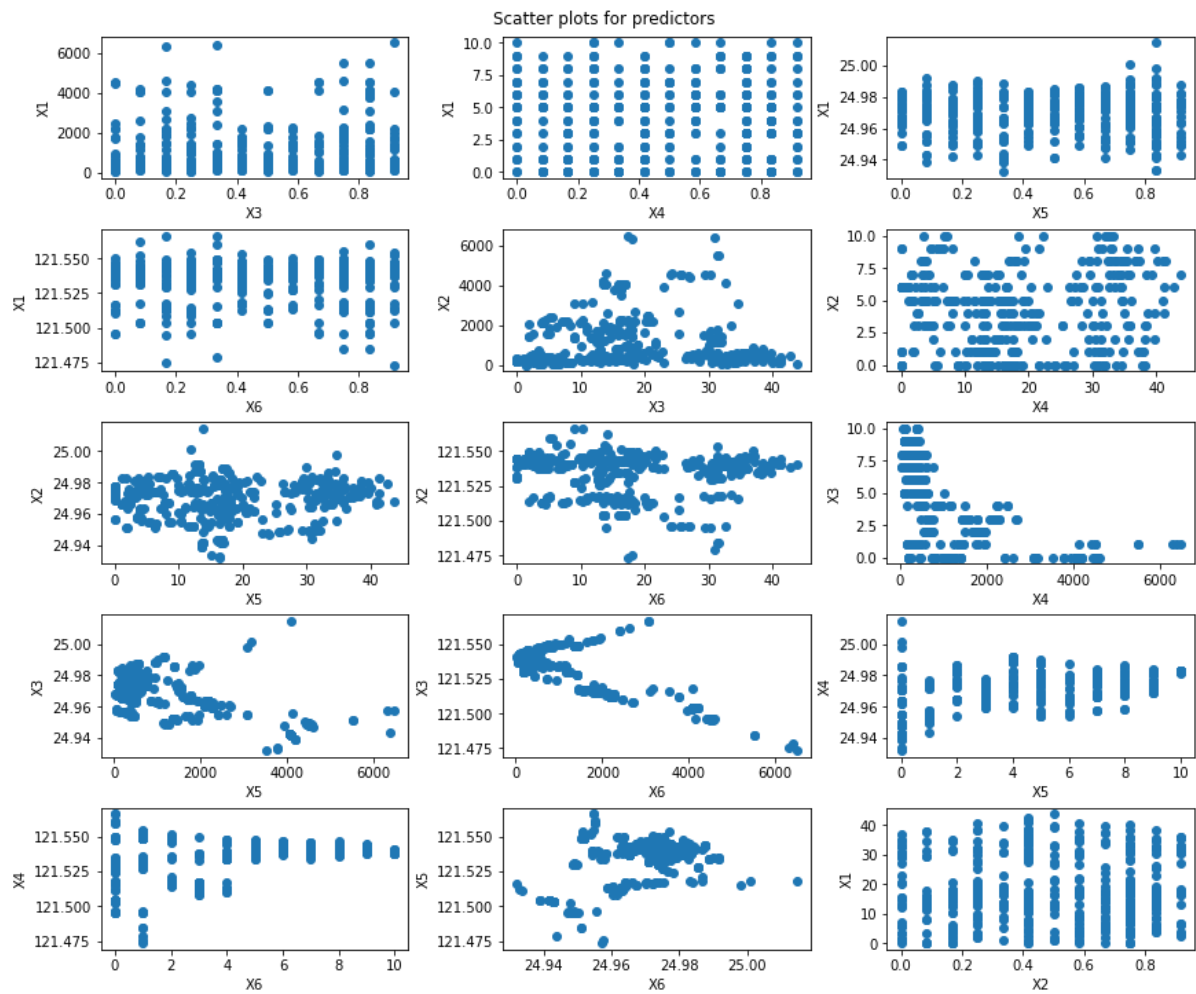
C:\Users\SAADMU~1\AppData\Local\Temp\ipykernel_28320\3992793283.py:4: UserWarning: To output multiple subplots, the figure containing the passed axes is being cleared

```
dfr.drop("No", axis=1).hist(ax=ax_hist, bins="auto")
```

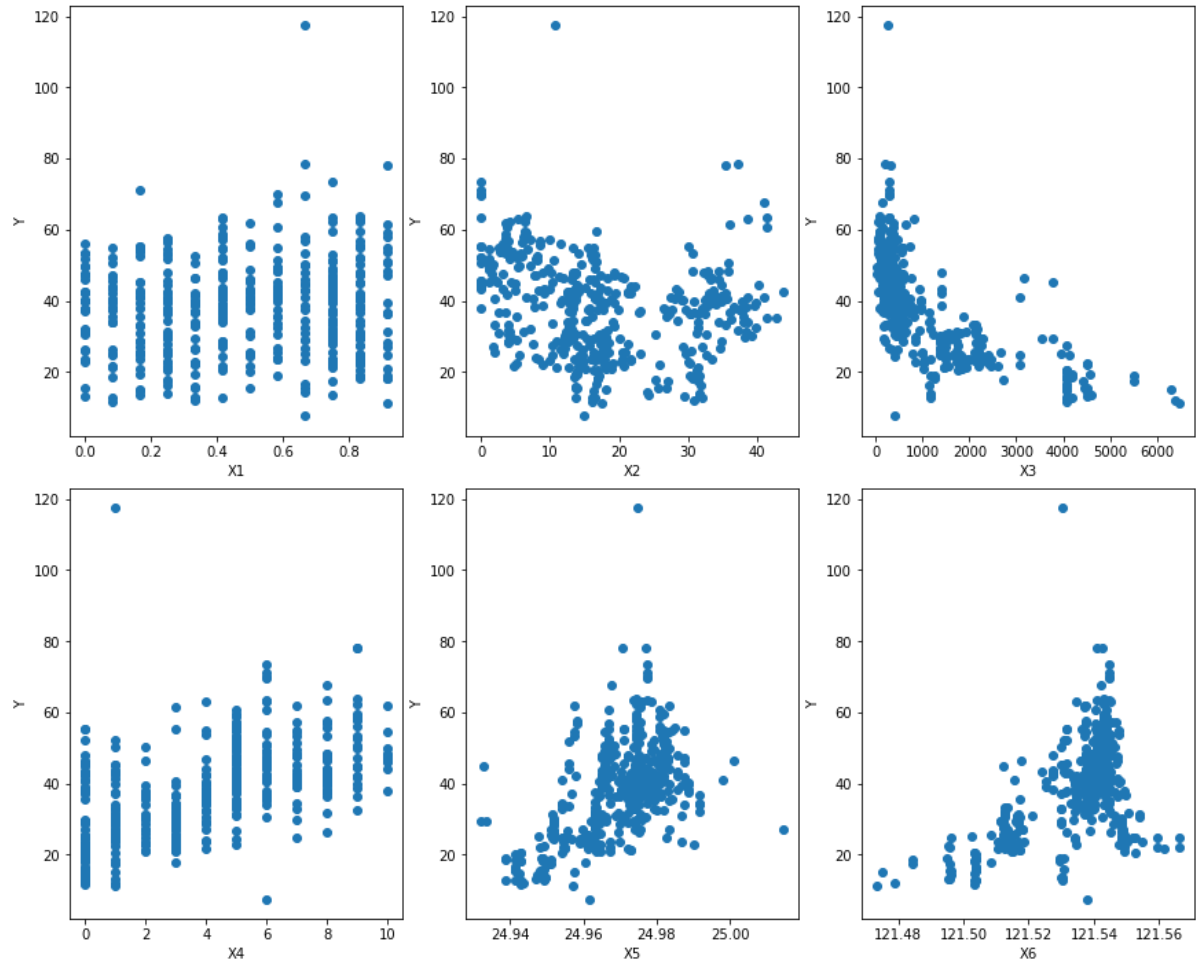
Out[]: Text(0.5, 0.98, 'Scatter plots for predictors')

Predictor histograms





Scatter plots between predictors and response



```

In [ ]: # 2. Linear Regression Model
pred_names = ["X1", "X2", "X3", "X4", "X5", "X6"]
predictors = dfr[pred_names]

lin_model = LinearRegression().fit(X=predictors, y=dfr["Y"])
lin_y_preds = lin_model.predict(predictors)

def show_weights(names, weights):
    print("\n".join("{} weight: {}".format(*i) for i in list(zip(names, weights))))

show_weights(pred_names, lin_model.coef_)
print("\nRMSE: ", mean_squared_error(dfr["Y"], lin_y_preds, squared=False))

fig_time, ax_time = plt.subplots()
tt = np.linspace(0,5, num=414) ## Transaction times

ax_time.plot(tt, np.multiply(tt, lin_model.coef_[0]))
print(len(dfr["Y"]))
ax_time.scatter(tt, dfr["Y"])
ax_time.set_title("House prices with respect to time of transaction")
ax_time.set_ylabel("Price (Y)")
ax_time.set_xlabel("Transaction date (X1, Years since August 2012)")

```

```

X1 weight: 5.149017210936201
X2 weight: -0.2696967345199864
X3 weight: -0.004487508250365868
X4 weight: 1.1333249810148465
X5 weight: 225.47014318107793
X6 weight: -12.42906116553292

```

```

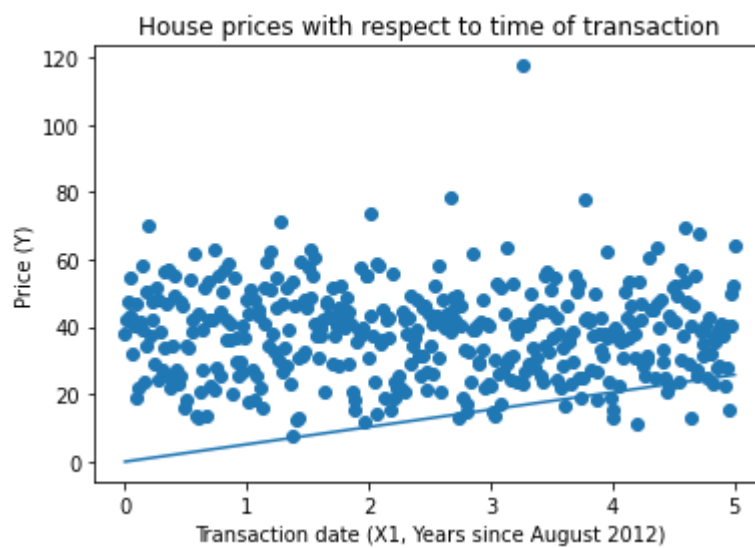
RMSE: 8.782312975361108
414

```

```

Out[ ]: Text(0.5, 0, 'Transaction date (X1, Years since August 2012)')

```



In []: *# 3. Bayesian Linear Regression Model*

```
# x must be a dataframe
def bayesian_reg(x, t, lambda_reg, basis=lambda v: v, basis_params=None):
    alpha = lambda_reg ** 2
    beta = lambda_reg
    def posterior(phi, t):
        # Prior here is defined as  $N(w|\mu=0, cov=a^{-1} * I)$ 
        cov_n = np.linalg.pinv(alpha * np.eye(phi.shape[1]) + beta * phi.T.dot
(phi))
        m_n = beta * cov_n.dot(phi.T).dot(t)
        return m_n, cov_n

    def predict_dist(phi, m_n, cov_n, beta):
        cov = beta ** -1 + np.sum(phi.dot(cov_n) * phi, axis=1)
        mu = phi.dot(m_n)
        return mu, cov

    def design_mat(x, basis):
        if(basis_params == None):
            return np.concatenate([np.ones((x.shape[0], 1)), basis(x)], axis=1
)
        else:
            x.apply(lambda row: basis(row, basis_params[row.name]), axis=1)
            return np.concatenate(
                [
                    np.ones((x.shape[0], 1)),
                    x.to_numpy()
                ], axis=1)

    phi = design_mat(x, basis)
    # print("Design mat", phi)

    m_n, cov_n = posterior(phi, t)
    y, var = predict_dist(phi, m_n, cov_n, beta)

    # Frequentist approach uses penrose pseudo-inverse
    # weights = np.matmul(np.linalg.pinv(phi), t)

    # show_weights(["B0"] + pred_names, m_n)
    # print(np.diag(cov_n))
    _, ax_weights = plt.subplots()
    weight_x = np.linspace(1, len(m_n), len(m_n))
    ax_weights.errorbar(x=weight_x, y=m_n, yerr=2*np.sqrt(np.diag(cov_n)), ls=
"none", capsize=8)
    ax_weights.scatter(x=weight_x, y=m_n)
    ax_weights.set_title(f"Weight Uncertainty,  $\lambda$  = {lambda_reg}")
    ax_weights.set_xlabel("Weight #")
    ax_weights.set_ylabel("Value")

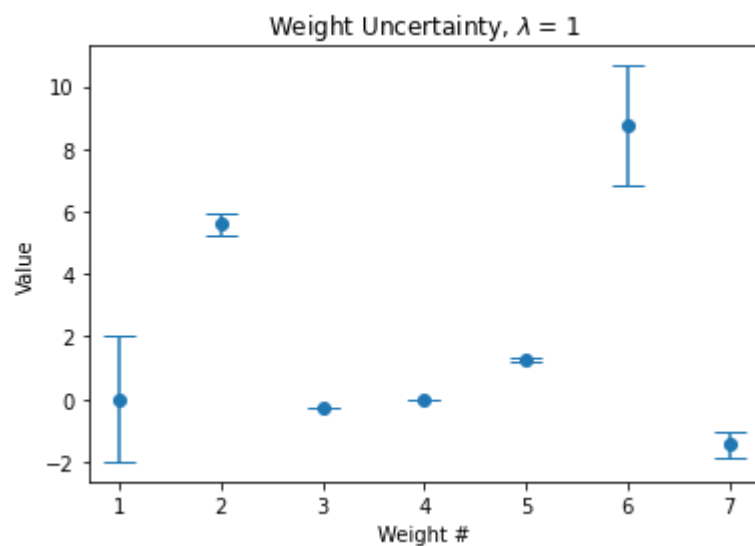
    cols = min(3, x.shape[1])
    rows = x.shape[1] // cols
    fig_bay, ax_bay = plt.subplots(rows, cols)
    for i, ax in enumerate(ax_bay.ravel()):
        names = list(x.keys())
        data = x[names[i]]
```

```

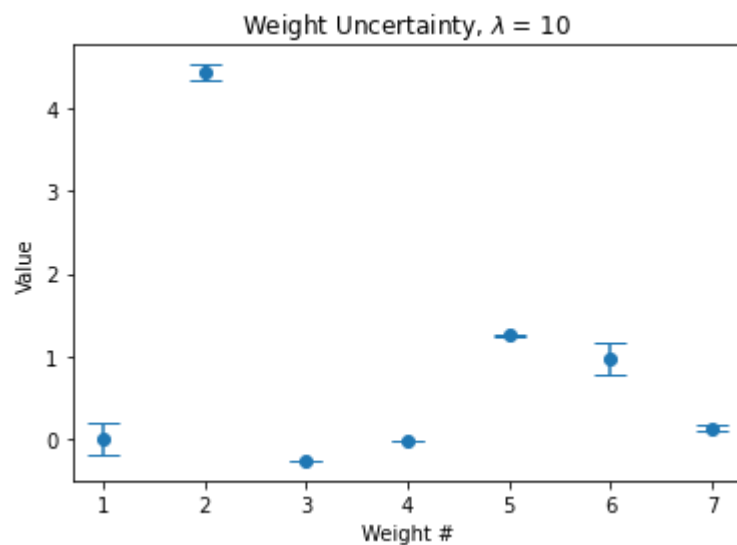
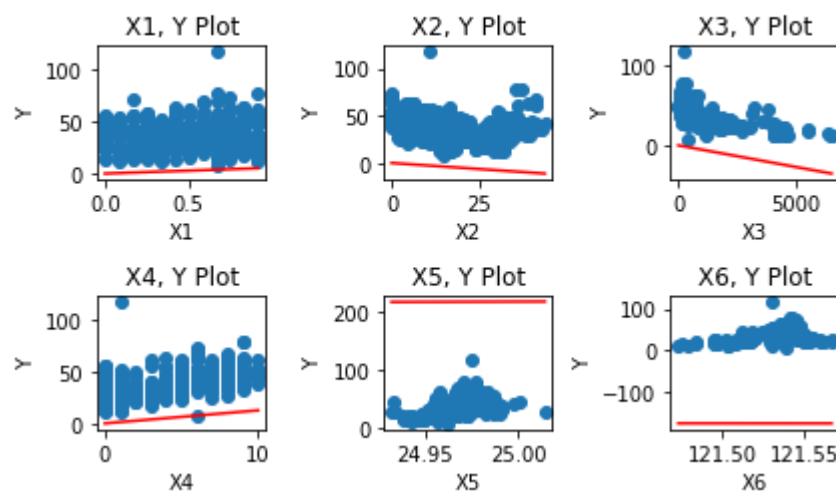
ax.scatter(data, dfr["Y"])
x_test = np.linspace(min(data), max(data), 200)
ax.plot(x_test, (m_n[i + 1] * x_test) + m_n[0], "r")
ax.set_title(f"X{i + 1}, Y Plot")
ax.set_xlabel(f"X{i + 1}")
ax.set_ylabel("Y")
fig_bay.suptitle(f"Predictors + respective regression coeff | $\lambda$={lambda_reg}")
fig_bay.tight_layout()

bayesian_reg(predictors, dfr["Y"], lambda_reg=1)
bayesian_reg(predictors, dfr["Y"], lambda_reg=10)
bayesian_reg(predictors, dfr["Y"], lambda_reg=100)

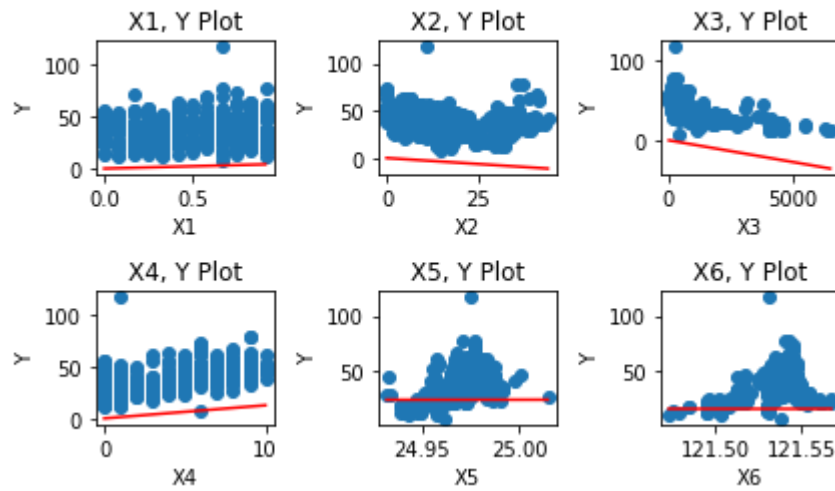
```



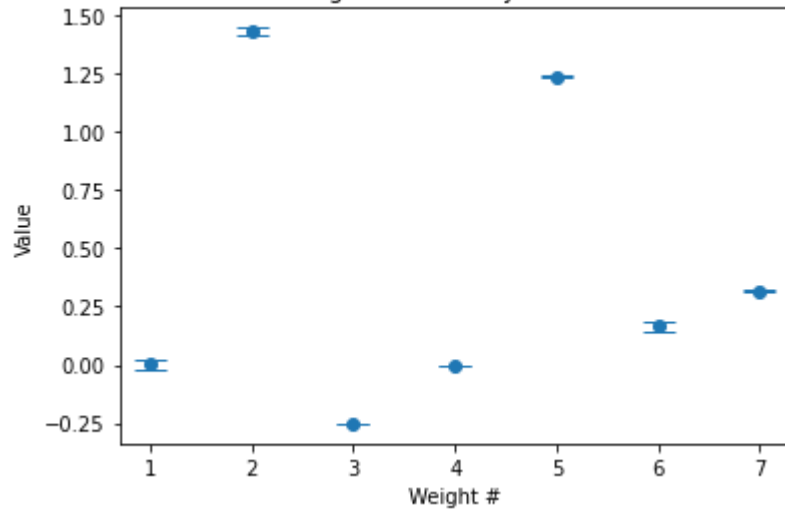
Predictors + respective regression coeff | $\lambda = 1$



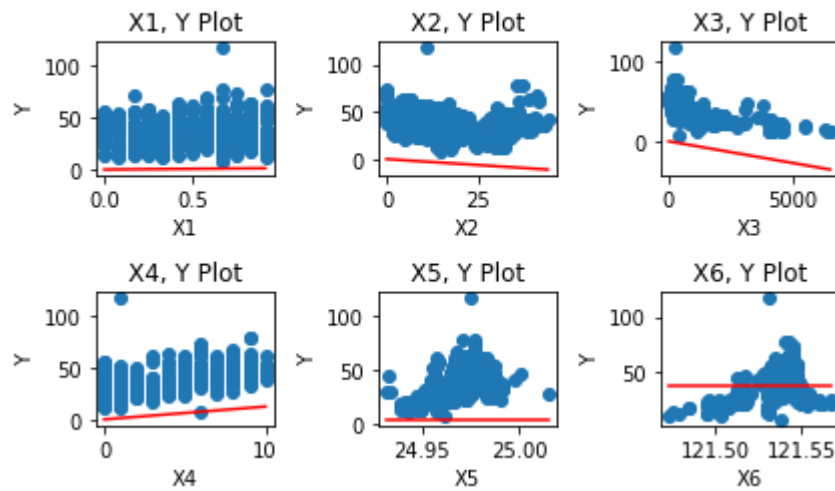
Predictors + respective regression coeff | $\lambda = 10$



Weight Uncertainty, $\lambda = 100$



Predictors + respective regression coeff | $\lambda = 100$



```

In [ ]: # 3. Bayesian Linear Regression, SKLearn
# I attempted to perform this problem with a custom made solution, but the results were poor
def bay_ridge(x, t, lambda_reg, basis=lambda v: v, basis_params=None):
    model = BayesianRidge(lambda_1=lambda_reg)
    design_mat = x.copy()

    if basis_params is not None:
        for c in x.columns:
            for i in basis_params:
                design_mat[f"{c}.mu{i}"] = basis(x[c], i)
    # print(design_mat.shape)
    model.fit(X=design_mat, y=t)
    predicted, std = model.predict(X=design_mat, return_std=True)
    print("Predicted vals", predicted)
    print("Uncertainties", std)
    weights = model.coef_
    weight_std = np.diag(model.sigma_)
    intercept = model.intercept_
    print("Intercept", intercept)

    _, ax_weights = plt.subplots(constrained_layout=True)
    weight_x = np.linspace(1, len(weights), len(weights))
    ax_weights.errorbar(x=weight_x, y=weights, yerr=2*weight_std, ls="none", c='r',
                        apsize=8)
    ax_weights.scatter(x=weight_x, y=weights)
    ax_weights.set_title(f"Weight Uncertainty,  $\lambda = \{\lambda\_reg\}$ ")
    ax_weights.set_xlabel("Weight #")
    ax_weights.set_ylabel("Value")

    cols = min(3, design_mat.shape[1])
    rows = design_mat.shape[1] // cols
    fig_bay, ax_bay = plt.subplots(rows, cols, constrained_layout=True)
    for i, ax in enumerate(ax_bay.ravel()):
        names = list(design_mat.keys())
        data = design_mat[names[i]]
        ax.scatter(data, dfr["Y"])
        x_test = np.linspace(min(data), max(data), 200)
        ax.plot(x_test, (weights[i] * x_test) + intercept, "r")
        ax.set_title(f"{names[i]}, Y Plot")
        ax.set_xlabel(names[i])
        ax.set_ylabel("Y")
    fig_bay.suptitle(f"Predictors + respective regression coeff |  $\lambda = \{\lambda\_reg\}$ ")
    # fig_bay.tight_layout()
    del design_mat
    return model

bay_1 = bay_ridge(x=predictors, t=dfr["Y"], lambda_reg=1)
bay_2 = bay_ridge(x=predictors, t=dfr["Y"], lambda_reg=10)
bay_3 = bay_ridge(x=predictors, t=dfr["Y"], lambda_reg=100)

```

Predicted vals [46.7630763 47.4763383 43.41025225 43.3421758 45.82610157 3
2.89297689
39.38533815 44.22960541 6.53244148 32.97500145 33.74046204 52.3202548
43.32659262 29.11549027 38.83715658 34.14562837 49.14619003 38.05136342
47.00708079 50.97572461 33.59735344 47.91659968 33.33963169 47.74444785
35.61204881 30.29982552 46.20293543 45.44622948 40.70256166 45.16778658
12.3015666 39.8711181 32.2296225 44.8448585 46.55489837 17.75111375
31.33879059 34.21990148 46.31269166 43.85573832 17.39993207 17.09279735
37.80002021 38.90596154 45.0350606 41.13296824 46.46062357 34.94075528
12.1817702 11.98299124 40.25769979 27.11055479 31.46724917 43.38762992
43.81270032 28.93480615 42.80865104 50.49517928 12.43840919 44.25349996
32.69475097 48.20819292 30.56361595 45.06009913 33.81370007 42.55405151
49.24240696 47.07866001 40.72111998 42.99787221 52.44936864 34.63245055
42.92481226 17.48600799 51.43845761 34.41709176 35.07581819 29.61796968
33.12331044 32.27805261 42.68847776 40.75917821 47.67912768 25.58747685
45.91473174 52.04254392 35.97092987 17.13583527 33.35016768 16.19867182
42.97739293 33.59687568 29.26964264 29.03210602 38.14137807 46.72731737
52.3632928 36.53951421 44.00989069 52.3632928 38.87010715 40.37362947
49.28544488 49.20600064 39.92985139 48.80580762 47.65079832 34.30597731
35.0120099 34.57138422 46.83843173 42.80194845 33.20712143 44.95176331
44.2088341 36.63767057 1.41971702 16.76012713 32.29640693 44.02789007
43.45570645 43.65289673 38.20943611 42.75832241 47.65837201 49.35352141
34.1338255 46.50420101 39.60044583 41.25186085 42.76334554 34.50330777
40.17855824 46.22691335 40.89431776 33.04803936 44.15555909 45.79997922
35.85191442 43.14740172 43.92381477 37.35085752 41.20967706 43.65289673
34.17686348 46.28356027 42.21371058 49.21763834 18.77862255 42.93424638
42.04677211 51.15831924 34.35605438 47.05631113 17.06775882 17.48600799
29.22048636 39.84132271 44.30971145 39.5150186 51.71974311 39.12110009
16.68041732 47.07866001 42.28178703 34.27656465 49.28234301 44.38911712
39.68007034 31.98403104 12.59674811 46.4181251 52.44936864 35.69204141
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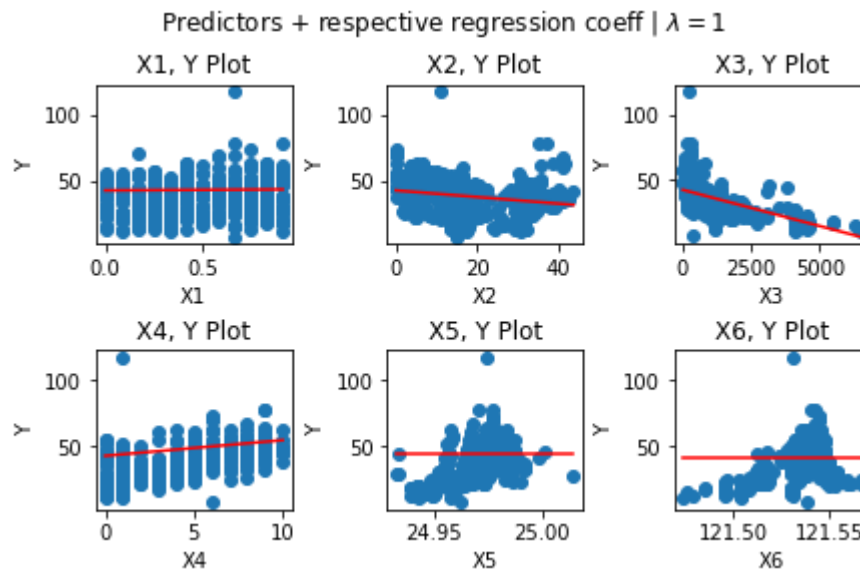
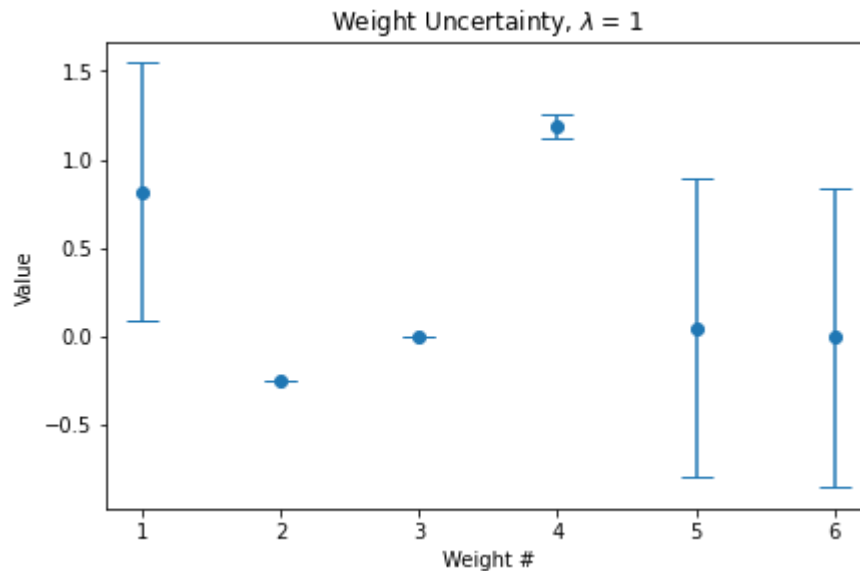
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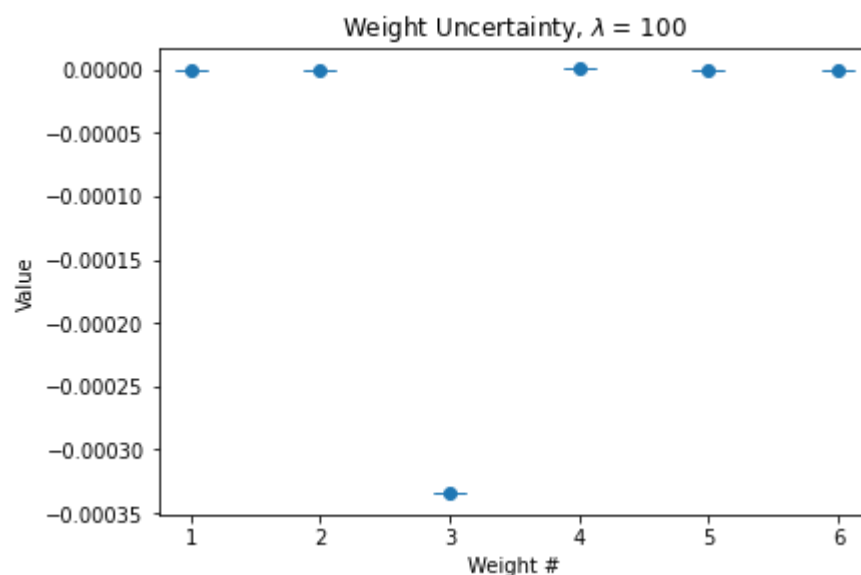
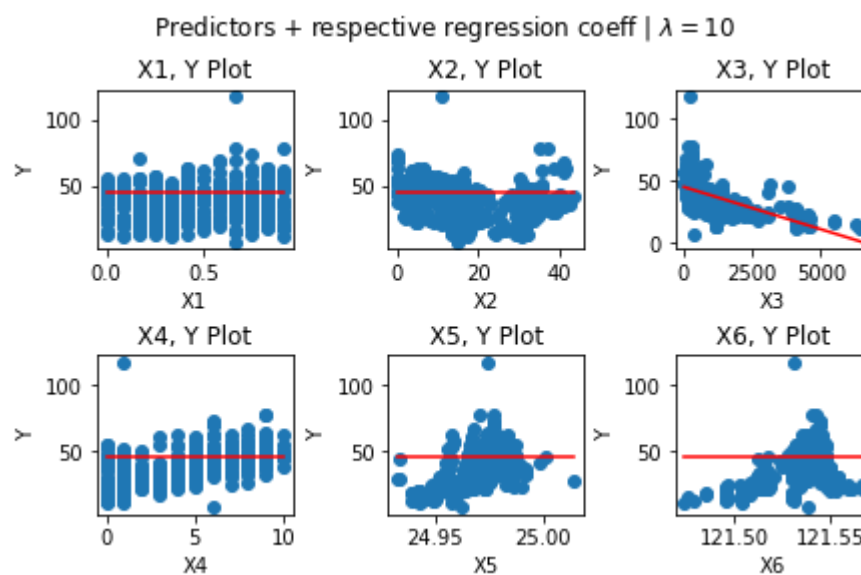
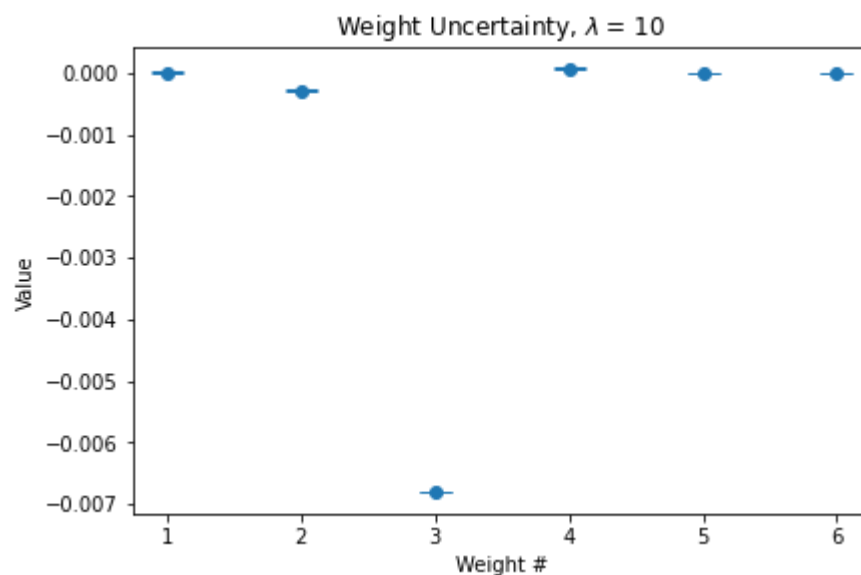
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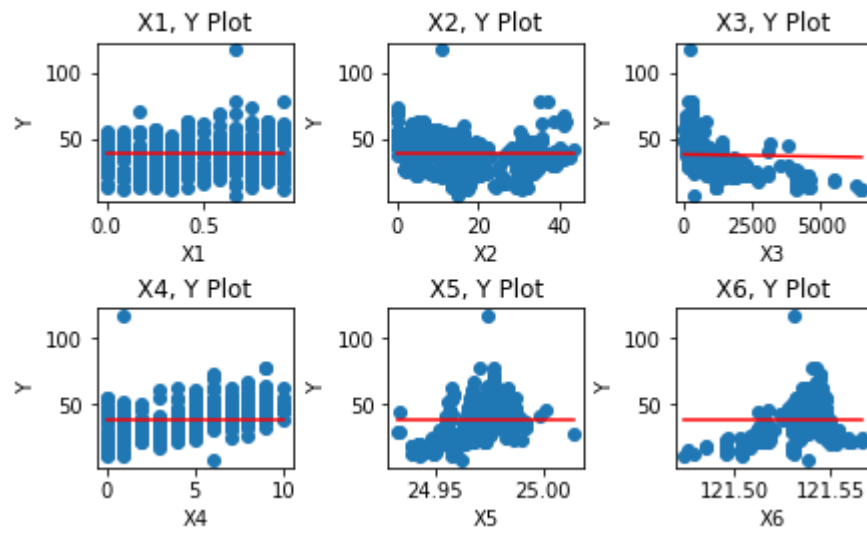
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Intercept 38.34257315585334

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Predictors + respective regression coeff | $\lambda = 100$



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In [ ]: # 4. Standardization + Gaussian Basis
standardized_pred = preprocessing.StandardScaler().fit_transform(dfr[["X2", "X3"]])
standardized_pred = pd.DataFrame(standardized_pred, columns=["X2", "X3"])
gauss = lambda x_i, mu: np.exp(-0.5 * (x_i - mu) ** 2)

mu = [-3, -2, -1, 1, 2, 3]

stan_1 = bay_ridge(standardized_pred, dfr["Y"], lambda_reg=1, basis=gauss, basis_params=mu)
stan_2 = bay_ridge(standardized_pred, dfr["Y"], lambda_reg=10, basis=gauss, basis_params=mu)
stan_3 = bay_ridge(standardized_pred, dfr["Y"], lambda_reg=100, basis=gauss, basis_params=mu)
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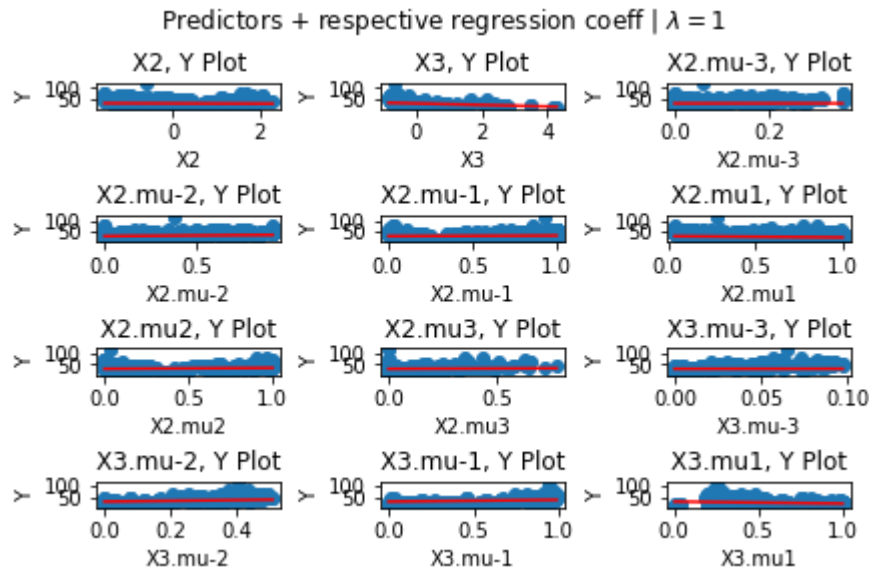
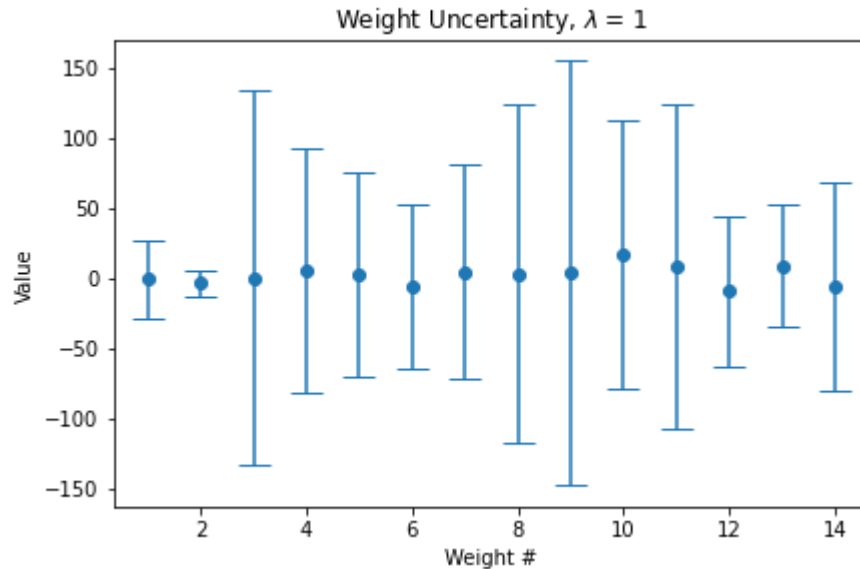
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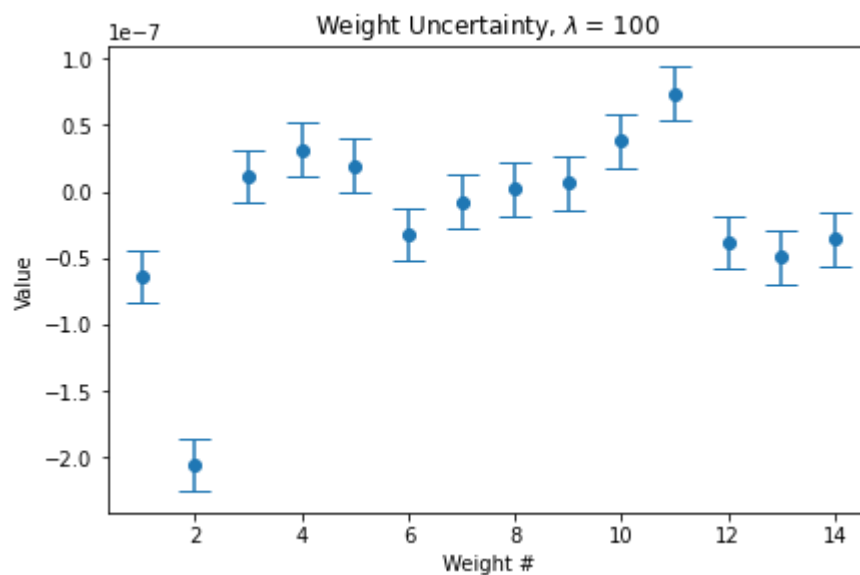
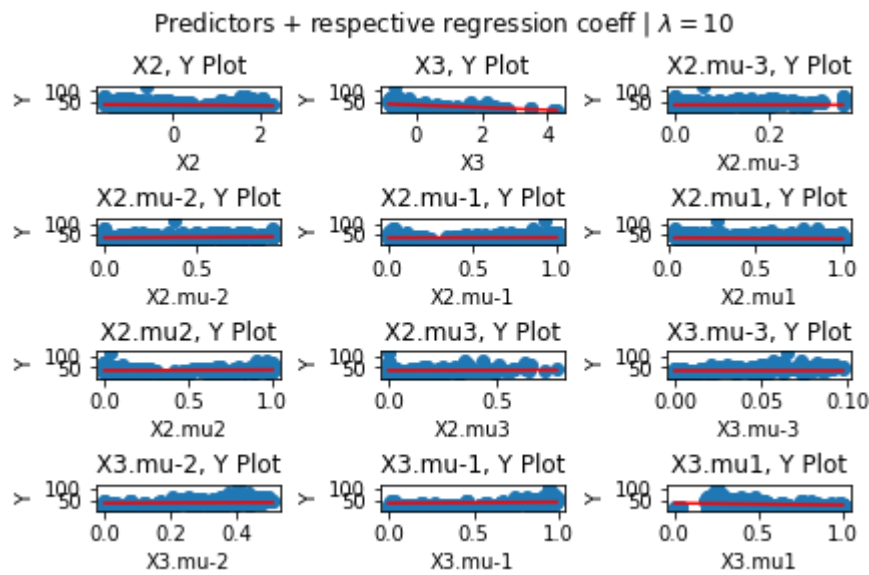
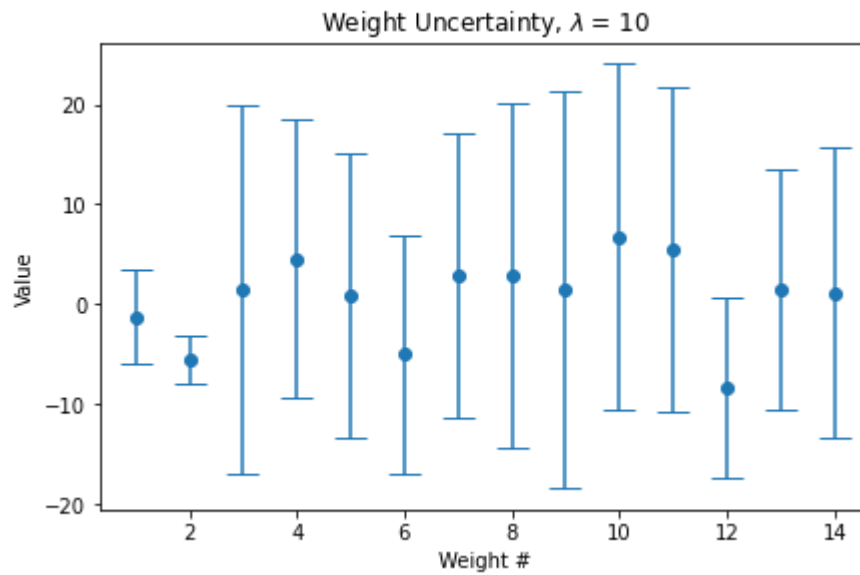
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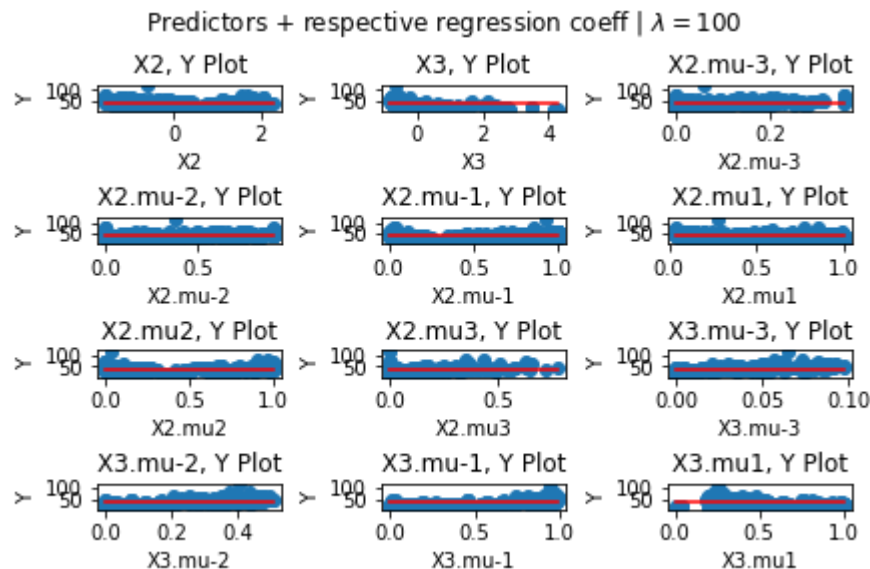
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Intercept 37.980193203288984

```







```
In [ ]: # 5. Compare linear regression model, gaussian basis model
BayesianRidge().fit(X=predictors, y=dfr["Y"])
```

```
In [ ]: # 6. Cross Validation
from sklearn.metrics import make_scorer

rmse = make_scorer(mean_squared_error)
print(cross_val_score(bay_1, X=predictors, y=dfr["Y"], cv=10, scoring=rmse))
print(cross_val_score(bay_2, X=predictors, y=dfr["Y"], cv=10, scoring=rmse))
print(cross_val_score(bay_3, X=predictors, y=dfr["Y"], cv=10, scoring=rmse))

print(cross_val_score(stan_1, X=standardized_pred, y=dfr["Y"], cv=10, scoring=
rmse))
print(cross_val_score(stan_2, X=standardized_pred, y=dfr["Y"], cv=10, scoring=
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print(cross_val_score(stan_3, X=standardized_pred, y=dfr["Y"], cv=10, scoring=
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Classification Problem (25 pts)

```
In [ ]: # Dataset Init
dfc = pd.read_excel("ENB2012_data.xlsx")
dfc["Y1-Y2"] = np.where(dfc["Y1"] - dfc["Y2"] > 0, 1, 0)
dfc.head()
```

Out[]:

	X1	X2	X3	X4	X5	X6	X7	X8	Y1	Y2	Y1-Y2
0	0.98	514.5	294.0	110.25	7.0	2	0.0	0	15.55	21.33	0
1	0.98	514.5	294.0	110.25	7.0	3	0.0	0	15.55	21.33	0
2	0.98	514.5	294.0	110.25	7.0	4	0.0	0	15.55	21.33	0
3	0.98	514.5	294.0	110.25	7.0	5	0.0	0	15.55	21.33	0
4	0.90	563.5	318.5	122.50	7.0	2	0.0	0	20.84	28.28	0

In []: *# 1. Visualization*

```
excl_hist = ["Y1", "Y2", "Y1-Y2", "X5", "X6", "X7", "X8"]

# Histogram of predictors X1, X2, X3, X4 that have Label 0
fig_hist_0, ax_hist_0 = plt.subplots()
dfc[dfc["Y1-Y2"] == 0].drop(excl_hist, axis=1).hist(bins="auto", ax=ax_hist_0)
fig_hist_0.suptitle("Predictors for y2-y1=0")
fig_hist_0.tight_layout()

# Histogram of predictors X1, X2, X3, X4 that have Label 1
fig_hist_1, ax_hist_1 = plt.subplots()
dfc[dfc["Y1-Y2"] == 1].drop(excl_hist, axis=1).hist(bins="auto", ax=ax_hist_1)
fig_hist_1.suptitle("Predictors for y2-y1=1")
fig_hist_1.tight_layout()

# Construct conditional PMFs
x_discrete = ["X5", "X6", "X7", "X8"]

# P(Xn | Y=0)
fig_pmf_0, ax_pmf_0 = plt.subplots(2, 2)
for i in range(len(x_discrete)):
    data = dfc[dfc["Y1-Y2"] == 0][x_discrete[i]]
    data = data.value_counts().sort_index()
    data = data.divide(sum(data.values))
    # print(data)
    x = data.index
    p = data.values
    ax_pmf_0[lin_to_mat(2, i)].plot(x, p, "bo", ms=8, mec="b")
    ax_pmf_0[lin_to_mat(2, i)].vlines(x, 0, p, linestyle="-", lw=2)
    ax_pmf_0[lin_to_mat(2, i)].set_title(x_discrete[i])
    ax_pmf_0[lin_to_mat(2, i)].set_ylim([0, 0.6])
    ax_pmf_0[lin_to_mat(2, i)].set_xticks(x)
fig_pmf_0.suptitle("Conditional PMFs for Y=0")

fig_pmf_0.tight_layout()

# P(Xn | Y=1)
fig_pmf_1, ax_pmf_1 = plt.subplots(2, 2)
for i in range(len(x_discrete)):
    data = dfc[dfc["Y1-Y2"] == 1][x_discrete[i]]
    data = data.value_counts().sort_index()
    data = data.divide(sum(data.values))
    # print(data)
    x = data.index
    p = data.values
    ax_pmf_1[lin_to_mat(2, i)].plot(x, p, "bo", ms=8, mec="b")
    ax_pmf_1[lin_to_mat(2, i)].vlines(x, 0, p, linestyle="-", lw=2)
    ax_pmf_1[lin_to_mat(2, i)].set_title(x_discrete[i])
    ax_pmf_1[lin_to_mat(2, i)].set_ylim([0, max(0.6, max(p))])
    ax_pmf_1[lin_to_mat(2, i)].set_xticks(x)
fig_pmf_1.suptitle("Conditional PMFs for Y=1")
fig_pmf_1.tight_layout()

_=""
```

```
X5 and X7 appear to be very good predictors of the class labels, looking at th  
eir conditional distributions.  
"""
```

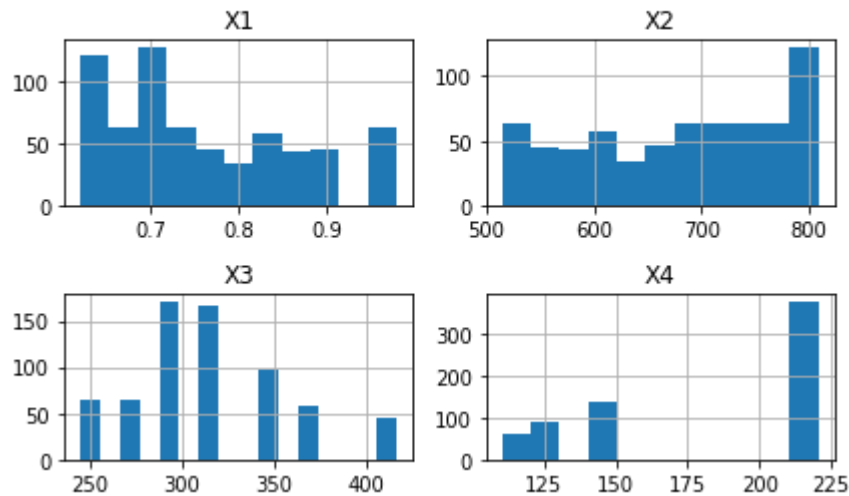
C:\Users\SAADMU~1\AppData\Local\Temp\ipykernel_28320\1939464138.py:7: UserWarning: To output multiple subplots, the figure containing the passed axes is being cleared

```
dfc[dfc["Y1-Y2"] == 0].drop(excl_hist, axis=1).hist(bins="auto", ax=ax_hist_0)
```

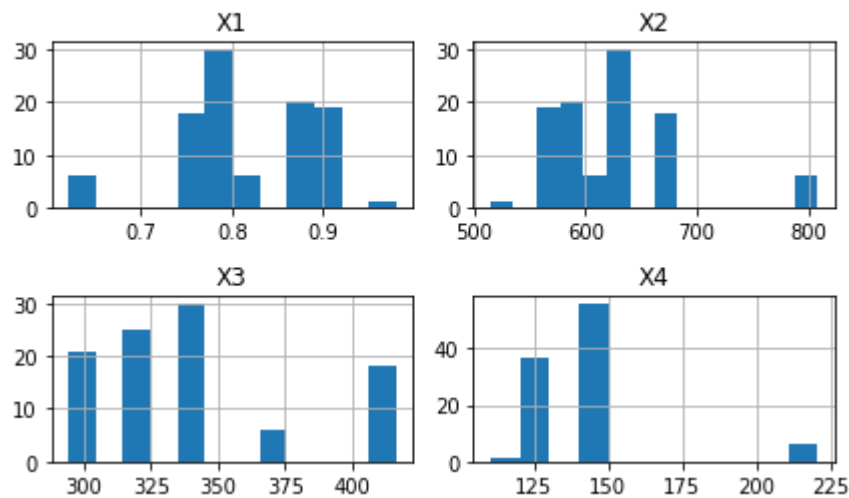
C:\Users\SAADMU~1\AppData\Local\Temp\ipykernel_28320\1939464138.py:13: UserWarning: To output multiple subplots, the figure containing the passed axes is being cleared

```
dfc[dfc["Y1-Y2"] == 1].drop(excl_hist, axis=1).hist(bins="auto", ax=ax_hist_1)
```

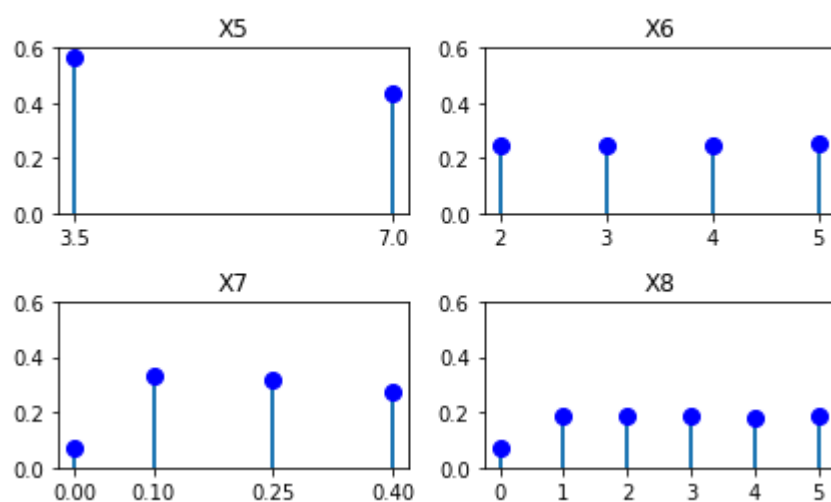
Predictors for $y_2 - y_1 = 0$



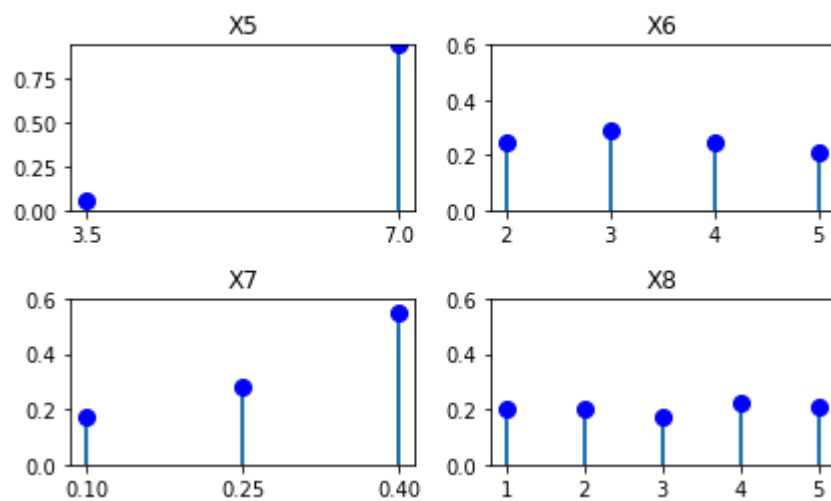
Predictors for $y_2 - y_1 = 1$



Conditional PMFs for $Y=0$



Conditional PMFs for $Y=1$



```

In [ ]: # 2. Logistic Regression
pred_c_names = ["X1", "X2", "X3", "X4", "X5", "X6", "X7", "X8"]
pred_c_data = dfc[pred_c_names]
logistic_model = LogisticRegression(max_iter=500).fit(X=pred_c_data, y=dfc["Y1-Y2"])

prob_preds = logistic_model.predict_proba(X=pred_c_data)[: ,1]
show_weights(pred_c_names, logistic_model.coef_[0])

# Assuming Y=1 is "positive", Y=0 is "negative"
preds = logistic_model.predict(pred_c_data)
preds_df = pd.DataFrame(preds.transpose(), columns=["preds"])
true_pos = len(dfc[(dfc["Y1-Y2"] == 1) & (preds_df["preds"] == 1)])

print("\nAccuracy\nRecall: ", true_pos / len([i for i in preds if i == 1]))
print("Precision: ", true_pos / len(dfc[dfc[pred_c_names] == 1]))
print("Model score:", logistic_model.score(X=pred_c_data, y=dfc["Y1-Y2"]))

X1 weight: 0.1314098373909968
X2 weight: 0.01376272663718662
X3 weight: -0.0019824349199593157
X4 weight: 0.007872580757269867
X5 weight: 1.6519035109073805
X6 weight: -0.08958701981032652
X7 weight: 3.1448815458541177
X8 weight: 0.08805294156710261

Accuracy
Recall: 0.5833333333333334
Precision: 0.009114583333333334
Model score: 0.8723958333333334

```

```

In [ ]: # 3. ROC Curve
fig_roc, ax_roc = plt.subplots()
fig_roc.suptitle("ROC Curve")
RocCurveDisplay.from_predictions(y_true=dfc["Y1-Y2"], y_pred=prob_preds, ax=ax_roc, name="Balanced Model")

fig_thresh, ax_thresh = plt.subplots()

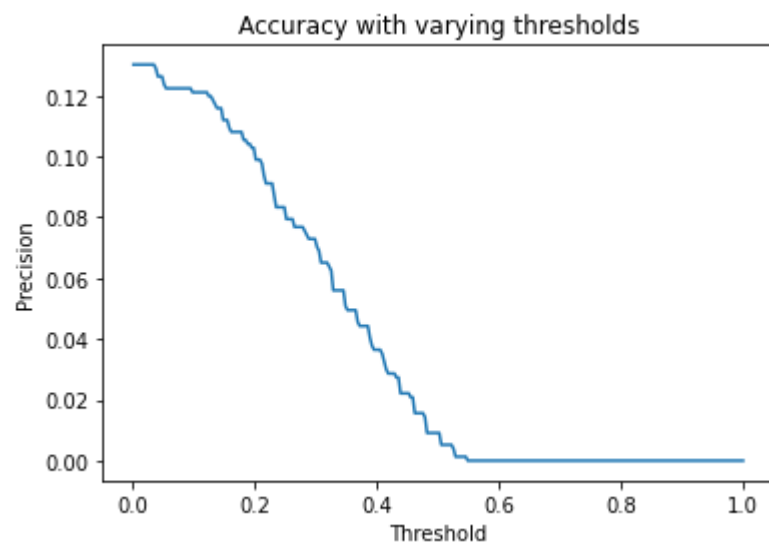
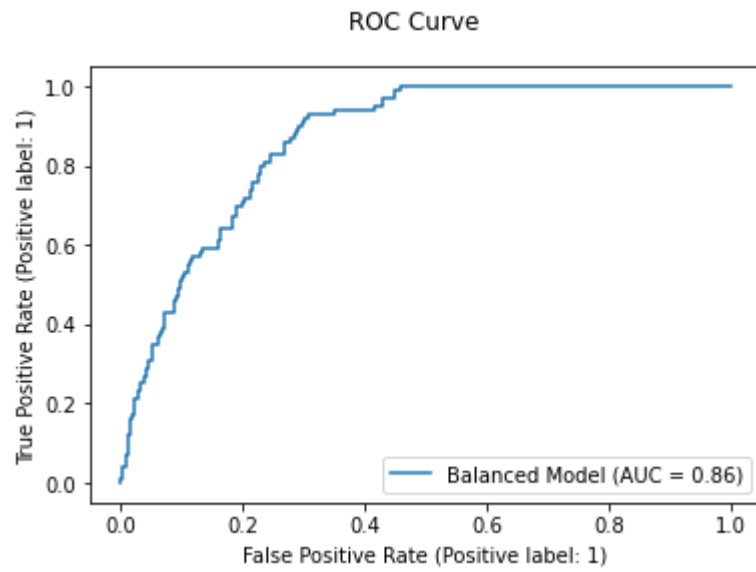
threshs = np.linspace(0.001, 1, 300)
prec_data = []
for thresh in threshs:
    preds_df["thresh"] = [1 if i > thresh else 0 for i in prob_preds]
    true_pos = len(dfc[(dfc["Y1-Y2"] == 1) & (preds_df["thresh"] == 1)])
    prec_data.append(true_pos / len(dfc[dfc[pred_c_names] == 1]))

len(dfc[(dfc["Y1-Y2"] == 1) & (preds_df["preds"] == 1)])
ax_thresh.plot(threshs, prec_data)
ax_thresh.set_title("Accuracy with varying thresholds")
ax_thresh.set_ylabel("Precision")
ax_thresh.set_xlabel("Threshold")

```



```
Out[ ]: Text(0.5, 0, 'Threshold')
```



```

In [ ]: # 4. Bayesian Regression
def bay_logistic(alpha):
    model = BernoulliNB(alpha=alpha, fit_prior=True)
    model.fit(X=pred_c_data, y=dfc["Y1-Y2"])
    model.predict_proba(pred_c_data)
    weights = model.coef_
    print("Weights", weights)
    # print("Weight uncertainties", )
    return model, weights

a_1, w1 = bay_logistic(0.1)
a_2, w2 = bay_logistic(1)
a_3, w3 = bay_logistic(10)
a_4, w4 = bay_logistic(100)

Weights [[-0.0009985 -0.0009985 -0.0009985 -0.0009985 -0.0009985 -0.0009985
          -0.0009985 -0.0009985]]
Weights [[-0.0098523 -0.0098523 -0.0098523 -0.0098523 -0.0098523 -0.0098523
          -0.0098523 -0.0098523]]
Weights [[-0.08701138 -0.08701138 -0.08701138 -0.08701138 -0.08701138 -0.0870
          1138
          -0.08701138 -0.08701138]]
Weights [[-0.40546511 -0.40546511 -0.40546511 -0.40546511 -0.40546511 -0.4054
          6511
          -0.40546511 -0.40546511]]

C:\Python39\lib\site-packages\sklearn\utils\deprecation.py:103: FutureWarnin
g: Attribute `coef_` was deprecated in version 0.24 and will be removed in 1.
1 (renaming of 0.26).
  warnings.warn(msg, category=FutureWarning)
C:\Python39\lib\site-packages\sklearn\utils\deprecation.py:103: FutureWarnin
g: Attribute `coef_` was deprecated in version 0.24 and will be removed in 1.
1 (renaming of 0.26).
  warnings.warn(msg, category=FutureWarning)
C:\Python39\lib\site-packages\sklearn\utils\deprecation.py:103: FutureWarnin
g: Attribute `coef_` was deprecated in version 0.24 and will be removed in 1.
1 (renaming of 0.26).
  warnings.warn(msg, category=FutureWarning)
C:\Python39\lib\site-packages\sklearn\utils\deprecation.py:103: FutureWarnin
g: Attribute `coef_` was deprecated in version 0.24 and will be removed in 1.
1 (renaming of 0.26).
  warnings.warn(msg, category=FutureWarning)

```

```
In [ ]: # 5. Cross validation
print(cross_val_score(estimator=a_1, X=pred_c_data, y=dfc["Y1-Y2"], cv=10))
print(cross_val_score(estimator=a_2, X=pred_c_data, y=dfc["Y1-Y2"], cv=10))
print(cross_val_score(estimator=a_3, X=pred_c_data, y=dfc["Y1-Y2"], cv=10))
print(cross_val_score(estimator=a_4, X=pred_c_data, y=dfc["Y1-Y2"], cv=10))

_=""
The model accuracy does not appear to change with varying values of alpha
""
```

```
[0.24675325 0.87012987 0.87012987 0.87012987 0.87012987 0.87012987
 0.87012987 0.87012987 0.86842105 0.86842105]
[0.24675325 0.87012987 0.87012987 0.87012987 0.87012987 0.87012987
 0.87012987 0.87012987 0.86842105 0.86842105]
[0.24675325 0.87012987 0.87012987 0.87012987 0.87012987 0.87012987
 0.87012987 0.87012987 0.86842105 0.86842105]
[0.87012987 0.87012987 0.87012987 0.87012987 0.87012987 0.87012987
 0.87012987 0.87012987 0.86842105 0.86842105]
```

```
In [ ]: # 6.

x1=np.array([0.8,600.0,286.0,138.1,5,4,0.25,0]).reshape((1, 8))
x2=np.array([0.67,630.0,296.0,238.1,2,6,0.5,3]).reshape((1, 8))

for w in [w1, w2, w3, w4]:
    print((w * x1 + w * x2) / 2)
    # print(m.predict_proba(x2))

[[-7.33899212e-04 -6.14078933e-01 -2.90564178e-01 -1.87818288e-01
 -3.49475815e-03 -4.99251165e-03 -3.74438374e-04 -1.49775349e-03]]
[[-7.24143789e-03 -6.05916231e+00 -2.86701826e+00 -1.85321696e+00
 -3.44830376e-02 -4.92614822e-02 -3.69461117e-03 -1.47784447e-02]]
[[-6.39533621e-02 -5.35119968e+01 -2.53203107e+01 -1.63668400e+01
 -3.04539819e-01 -4.35056885e-01 -3.26292664e-02 -1.30517065e-01]]
[[-2.98016854e-01 -2.49361041e+02 -1.17990346e+02 -7.62679868e+01
 -1.41912788e+00 -2.02732554e+00 -1.52049416e-01 -6.08197662e-01]]
```

```
In [ ]:
```