

ECONOMETRICS I

Lecture 4

Understanding $\hat{\beta}$'s standard errors

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Reality (β) vs. estimate ($\hat{\beta}$)

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Tech Adoption: Linear vs Log-Linear Models

Dependent variable:		
	Users Linear (1)	log(Users) Log-Linear (2)
Year	200.094*** (25.205)	0.241*** (0.022)
Constant	-401,968.400*** (50,787.630)	-479.229*** (44.104)
Observations	21	21
R ²	0.768	0.864
Note: *p<0.1; **p<0.05; ***p<0.01		

$\hat{\beta}_2$ estimate

$\hat{\beta}_1$ estimate

R^2

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What is this?

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The information in parentheses corresponds to the respective **standard errors** of each estimated $\hat{\beta}_j$.

Define st. error of $\hat{\beta}_j \equiv s(\hat{\beta}_j) = \sqrt{\text{var}(\hat{\beta}_j)}$

Today's objective: understand what $s(\hat{\beta}_j)$ means and how to estimate it.

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ST. ERROR $s(\hat{\beta})$ & SAMPLE SIZE (n)

Reality (β) vs. estimate ($\hat{\beta}$)

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Consider a data-generator process (a “population”)

$$y = 40 + 2 \cdot x + u, \quad \text{with random error } u \sim N(0, \sigma).$$

That is, the true parameter is $\beta = 2$.

In reality, we don't know the value of β . We only can guess the value through the estimation

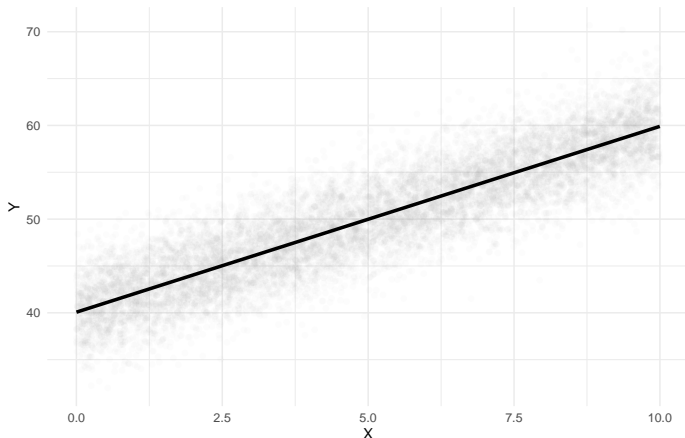
$$\hat{\beta} = \frac{\text{cov}(x, y)}{\text{var}(x)}$$

using n available data points.

Reality (β) vs. estimate ($\hat{\beta}$)

Population Relationship

True slope $\beta = 2$



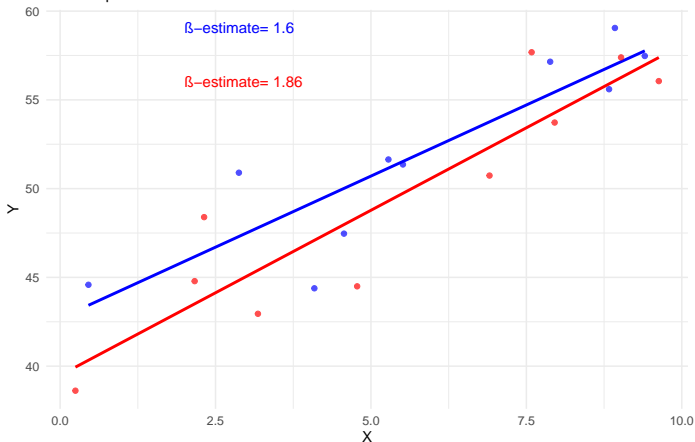
Will the sample size n affect the **precision** of our estimate $\hat{\beta}$?

Sample size n vs. estimation reliability

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Small Samples ($n = 10$)

Two samples with $n = 10$

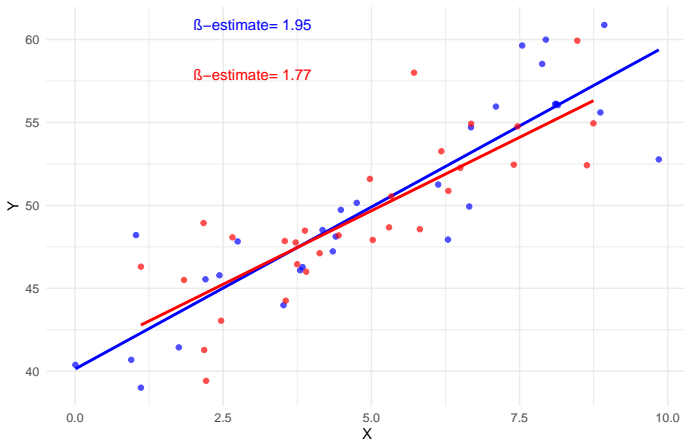


Sample size n vs. estimation reliability

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Medium Samples ($n = 30$)

Two samples with $n = 30$

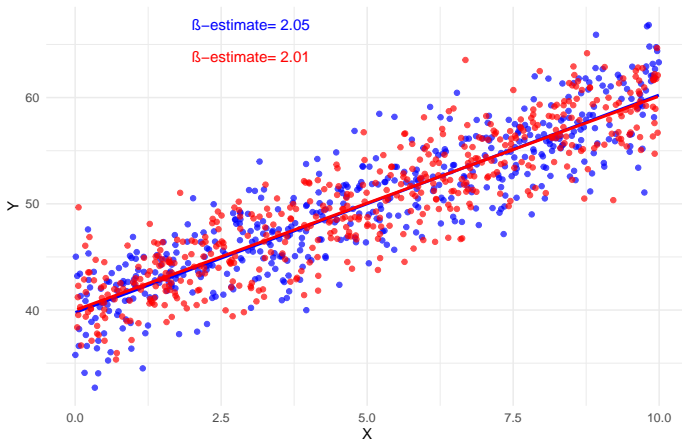


Sample size n vs. estimation reliability

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Large Samples ($n = 500$)

Two samples with $n = 500$



Sample size n vs. estimation reliability

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Takeaway:

- The larger the sample (the higher n) the more stable are the estimates of $\hat{\beta}_j$ – i.e., the lower is $s(\hat{\beta}_j) = \sqrt{\text{var}(\hat{\beta}_j)}$.

ST. ERROR $s(\hat{\beta}_j)$ & ERROR VARIANCE (σ^2)

The importance of the error-variance σ^2

Let us play God and invent some data of the form

$$y_A = 40 + 2 \cdot x + u_A, \quad u_A \sim N(0, 10^2)$$

$$y_B = 40 + 2 \cdot x + u_B, \quad u_B \sim N(0, 40^2)$$

Both populations are identical, except for the variances of the error terms u_A and u_B ($\sigma_B > \sigma_A$).

```
# Create dataset with 100 observations
n <- 100
x <- runif(n, 0, 100) # Uniform distribution between
                        0-100

# Generate y variables with different error variances
yA <- 40 + 2*x + rnorm(n, mean = 0, sd = 10) # Low
                        variance
yB <- 40 + 2*x + rnorm(n, mean = 0, sd = 40) # High
                        variance
```

The importance of the error-variance σ^2

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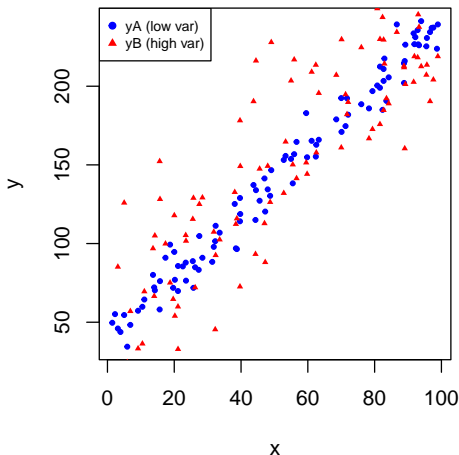
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Different error terms:

- Low-variance u_A vs. high-variance u_B .
- Both normally distributed

$$u_A \sim N(0, 10^2)$$

$$u_B \sim N(0, 40^2)$$

The importance of the error-variance σ^2

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=====		
	Dependent variable:	

	yA	yB
	(sigma = 10)	(sigma = 40)

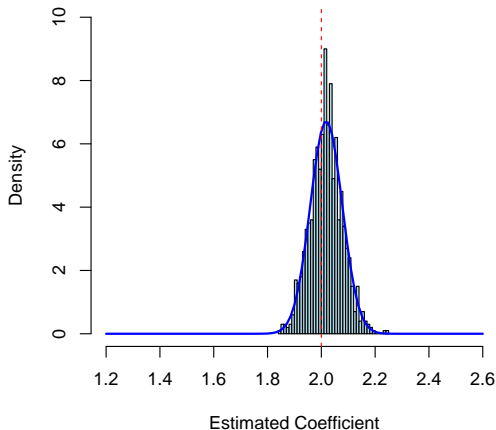
x	2.015*** (0.032)	2.053*** (0.132)
Constant	39.030*** (1.931)	39.310*** (7.924)

Observations	100	100
R2	0.976	0.711
=====		
Note:	*p<0.1; **p<0.05; ***p<0.01	

The importance of the error-variance σ^2

1000 different estimates $\hat{\beta}$ using $n = 30$ random samples of y_A

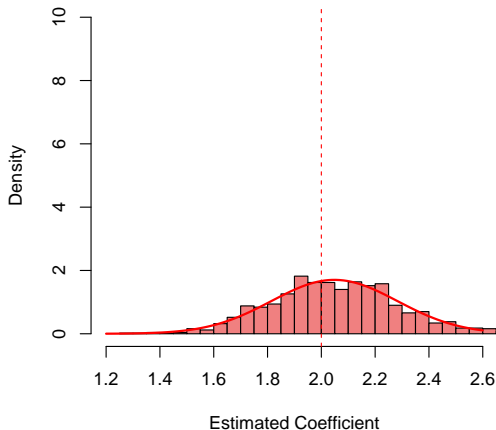
Sampling Distribution of β (y_A)



The importance of the error-variance σ^2

1000 different estimates $\hat{\beta}$ using $n = 30$ random samples of y_B

Sampling Distribution of β (y_B)



The importance of the error-variance σ^2

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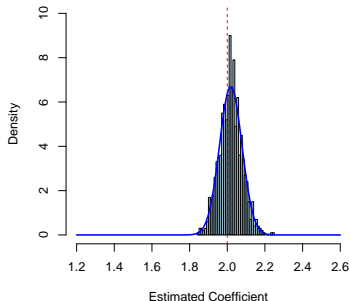
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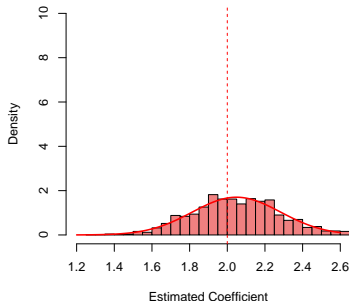
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Comparison of Sampling Distributions

Sampling Distribution of β (yA)



Sampling Distribution of β (yB)



The importance of the error-variance σ^2

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Takeaway:

- The smaller the error variance (the lower σ) the more stable are the estimates of $\hat{\beta}_j$ – i.e., the lower is $s(\hat{\beta}_j) = \sqrt{\text{var}(\hat{\beta}_j)}$.

The importance of the error-variance σ^2

Notice higher standard errors (in parentheses) of the coefficients of y_B :

	Dependent variable:	
	yA (sigma = 10)	yB (sigma = 40)
x	2.015*** (0.032)	2.053*** (0.132)
Constant	39.030*** (1.931)	39.310*** (7.924)
Observations	100	100
R2	0.976	0.711
Note: *p<0.1; **p<0.05; ***p<0.01		

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HOW IS $s(\hat{\beta}_j)$ ESTIMATED?

In practice, we do not run thousands of estimates with different samples (like shown in the previous slides). Instead, we estimate the parameter vector $\hat{\beta} = [\hat{\beta}_1, \hat{\beta}_2, \dots, \hat{\beta}_k]'$ just once.

Q: How do we then obtain an estimate of $s(\hat{\beta}_j) = \sqrt{\text{var}(\hat{\beta}_j)}$?

A: In the simple case of the univariate model ($k = 2$), it is estimated with the analytical solution

$$\widehat{\text{var}}(\hat{\beta}_2|x) = \frac{\hat{\sigma}^2}{\sum (x_i - \bar{x})^2} \quad \text{with } \hat{\sigma}^2 = \frac{\sum_i^n \hat{u}_i^2}{n-2}$$

That is the formula that statistical software use to estimate the st. errors $\hat{s}(\hat{\beta}_2|x) = \sqrt{\widehat{\text{var}}(\hat{\beta}_2|x)}$ shown in parenthesis in the introduction.

Where does the $\widehat{\text{var}}(\hat{\beta}_2|x)$ formula come from?

Assume true model is $y = \beta_1 + x\beta_2 + u$. Then

$$\hat{\beta}_2 = \frac{\text{cov}(x, y)}{\text{var}(x)} = \frac{\text{cov}(x, \beta_1 + x\beta_2 + u)}{\text{var}(x)}$$

By the rules described in the appendix:

$$\begin{aligned}\hat{\beta}_2 &= \frac{\text{cov}(x, x\beta_2)}{\text{var}(x)} + \frac{\text{cov}(x, u)}{\text{var}(x)} \\ &= \beta_2 \frac{\text{cov}(x, x)}{\text{var}(x)} + \frac{\text{cov}(x, u)}{\text{var}(x)} \\ &= \beta_2 \frac{\text{var}(x)}{\text{var}(x)} + \frac{\text{cov}(x, u)}{\text{var}(x)} \\ &= \beta_2 + \frac{\text{cov}(x, u)}{\text{var}(x)} = \beta_2 + \frac{\sum_{i=1}^n (x_i - \bar{x})u_i}{\sum_{i=1}^n (x_i - \bar{x})^2}\end{aligned}$$

Where does the $\widehat{\text{var}}(\hat{\beta}_2|x)$ formula come from?

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$$\begin{aligned}\text{var}(\hat{\beta}_2|x) &= \text{var}\left(\beta_2 + \frac{\sum_{i=1}^n (x_i - \bar{x})u_i}{\sum_{i=1}^n (x_i - \bar{x})^2} \middle| x\right) \\ &= \text{var}\left(\frac{\sum_{i=1}^n (x_i - \bar{x})u_i}{\sum_{i=1}^n (x_i - \bar{x})^2} \middle| x\right) \\ &= \frac{1}{(\sum_{i=1}^n (x_i - \bar{x})^2)^2} \text{var}\left(\sum_{i=1}^n (x_i - \bar{x})u_i \middle| x\right)\end{aligned}$$

Let's first expand $\text{var}(\sum_{i=1}^n (x_i - \bar{x})u_i|x)$:

Where does the $\widehat{\text{var}}(\hat{\beta}_2|x)$ formula come from?

Let's first expand $\text{var}(\sum_{i=1}^n (x_i - \bar{x})u_i|x)$ using the appendix rule 3:

$$\text{var}\left(\sum_{i=1}^n (x_i - \bar{x})u_i \mid x\right)$$

$$= \sum_{i=1}^n \sum_{j=1}^n (x_i - \bar{x})(x_j - \bar{x}) \text{cov}(u_i, u_j|x)$$

$$= \sum_{i=1}^n (x_i - \bar{x})^2 \text{var}(u_i|x) + \sum_{i \neq j} (x_i - \bar{x})(x_j - \bar{x}) \text{cov}(u_i, u_j|x)$$

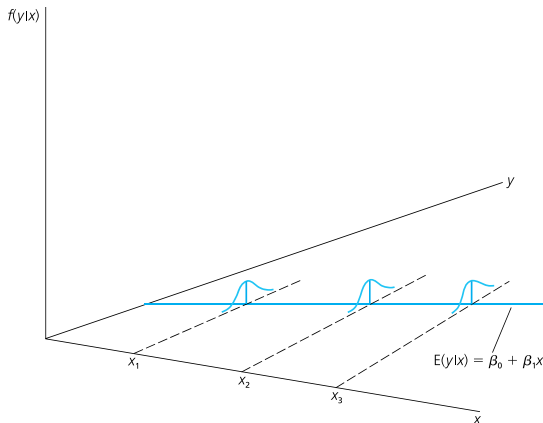
We will now **assume**

- **Homoskedasticity:** $\text{var}(u_i|x) = \sigma^2$ for all i
- **No autocorrelation:** $\text{cov}(u_i, u_j|x) = 0$ for $i \neq j$

Homoskedasticity

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Homoskedastic errors: $\text{var}(u_i|x_i) = \sigma^2$ for all $i = 1 \dots n$

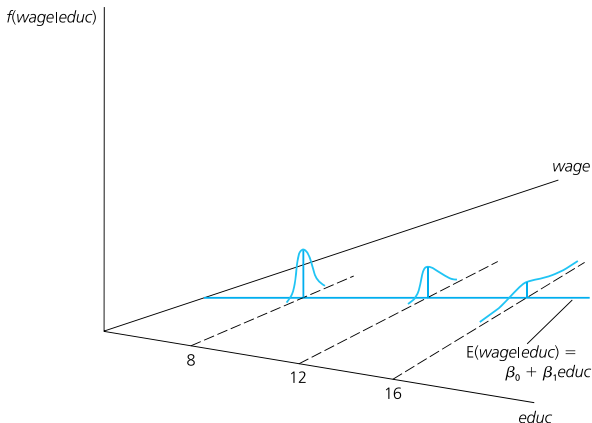


Source: Wooldridge's *Introductory Econometrics*

Homoskedasticity

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Heteroskedastic errors: $\text{var}(u_i|x_i) \neq \text{var}(u_j|x_j)$ for some $i \neq j$



Source: Wooldridge's *Introductory Econometrics*

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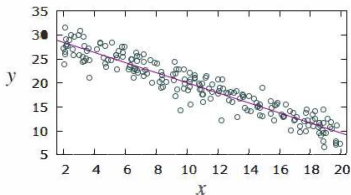
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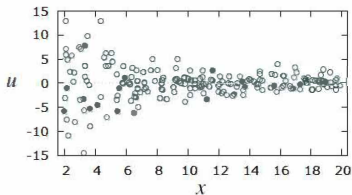
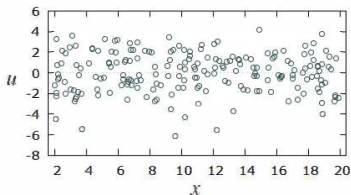
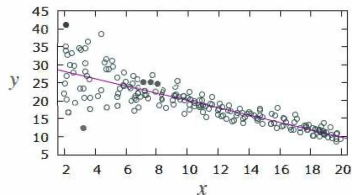
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Variance and covariance rules

(a) Invariant (homoskedastic) variance



(b) Variance decreasing in x



Homoskedastic assumption $\mathbb{E}(u_i^2 | x_i) = \sigma^2$ violated in (b).

Homoskedasticity

Thus:

$$\begin{aligned}\text{var} \left(\sum_{i=1}^n (x_i - \bar{x}) u_i \middle| x \right) &= \sum_{i=1}^n (x_i - \bar{x})^2 \sigma^2 + 0 \\ &= \sigma^2 \sum_{i=1}^n (x_i - \bar{x})^2\end{aligned}$$

Now completing the derivation:

$$\begin{aligned}\text{var}(\hat{\beta}_2 | x) &= \frac{1}{(\sum_{i=1}^n (x_i - \bar{x})^2)^2} \text{var} \left(\sum_{i=1}^n (x_i - \bar{x}) u_i \middle| x \right) \\ &= \frac{1}{(\sum_{i=1}^n (x_i - \bar{x})^2)^2} \cdot \sigma^2 \sum_{i=1}^n (x_i - \bar{x})^2 \\ &= \frac{\sigma^2}{\sum_{i=1}^n (x_i - \bar{x})^2}\end{aligned}$$

The final formula (if $k = 2$)

Thus, if we assume

- **Model = Population:** $y_i = \beta_1 + \beta_2 + u_i$
- **Homoskedasticity:** $\text{var}(u_i|x) = \sigma^2$ for all i
- **No autocorrelation:** $\text{cov}(u_i, u_j|x) = 0$ for $i \neq j$

The variance is

$$\text{var}(\hat{\beta}_2|x) = \frac{\sigma^2}{\sum (x_i - \bar{x})^2}$$

Note that σ^2 is also a population parameter that needs to be estimated.

We can use the SSR for that:

$$\hat{\sigma}^2 = \frac{\sum_i^n \hat{u}_i^2}{n - 2}$$

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Notice that $\sum_{i=1}^n (x_i - \bar{x})^2$ grows with n . This becomes explicit when substituting for $\text{var}(x) = \sum (x_i - \bar{x})^2 / n$:

$$\text{var}(\hat{\beta}_2 | x) = \frac{\sigma^2}{n \text{ var}(x_1)}$$

As we saw before, **precision** will be higher (the variance lower):

- with more data ($\uparrow n$)
- with lower error variance ($\downarrow \sigma^2$)

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You should now know:

- Difference between sample and population
- What standard errors represent
- How $\text{var}(\hat{\beta}_2|x)$ depends on n and σ
- Basic variance and covariance rules
- Homo- vs. heteroskedasticity
- How to derive $\text{var}(\hat{\beta}_2|x)$ in the univariate case

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Basic Variance Rules:

1 $\text{Var}(a) = 0$ (constant a)

2 $\text{Var}(aX + b) = a^2 \text{Var}(X)$

3 $\text{Var}(X \pm Y) = \text{Var}(X) + \text{Var}(Y) \pm 2\text{Cov}(X, Y)$

Covariance Rules:

4 $\text{Cov}(X, X) = \text{Var}(X)$

5 $\text{Cov}(X, Y) = E[(X - \mu_X)(Y - \mu_Y)]$

6 $\text{Cov}(X, a) = 0$ (constant a)

7 $\text{Cov}(aX + b, cY + d) = ac \text{Cov}(X, Y)$

8 $\text{Cov}(X, Y + Z) = \text{Cov}(X, Y) + \text{Cov}(X, Z)$